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Bogliacino, F and Rampa, G

Universidad EAFIT, università di Pavia

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## Monopolistic Competition and New Products: A Conjectural Equilibrium Approach

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## Francesco Bogliacino and Giorgio Rampa

Universidad EAFIT, Medellin, francesco.bogliacino@gmail.com University of Pavia, giorgio.rampa@unipv.it

## Abstract

In this paper we generalize the heterogeneous risk adverse agents model of diffusion of new products in a multi-firm, heterogeneous and interacting agents environment. We use a model of choice under uncertainty based on Bayesian theory. We discuss the possibility of product failures, the set of equilibria, their stability and some welfare properties.

*Keywords*: Product diffusion, Risk aversion, Lock-in, Monopolistic competition, Multiple equilibria.

### JEL Classification System: L15, D81, O33

Corresponding author: Giorgio Rampa Dipartimento di Economia Politica e Metodi Quantitativi, Via S. Felice, 5 – 27100 Pavia – Italy, tel. +39 0382 986214: giorgio.rampa@unipv.it

## Introduction

The analysis of diffusion processes is interesting under at least two different perspectives. First of all, scholars usually concentrate on new *products*, but it is possible to generalize many conclusions to the adoption of new technologies, behaviours, fashions and strategies (in the game-theoretic sense), so enlarging the focus significantly. Second, diffusion is in essence a multi-disciplinary matter: the literature that has studied the problem spans from management to sociology, from psychology to physics including, obviously, alternative economic approaches<sup>1</sup>.

The literature has discussed both the conditions that favour or hamper diffusion –bringing eventually to failure or success– and the speed of diffusion, looking at the factors giving rise to different possible patterns, and in particular to an epidemiologic-like *S*-shaped curve.

A satisfactory picture should be grounded on some essential building blocks. The first one is *uncertainty*: the very novelty of goods (ideas, technologies, behaviours etc.) implies that agents must act using conjectures over some unknown feature, as in standard Bayesian approaches (Jensen 1982, Feder and O'Mara 1982, Tsur et al. 1990, Chatterjee and Eliashberg 1990, Young 2006). The second block is *heterogeneity*: individual models are necessarily different at the outset, since they summarize personal conjectures, previous learning and *a priori* ideas (Abrahamson and Rosenkopf, 1993; Cowan and Jonard 2003 and 2004; Lopez Pintado and Watts, 2006). The third block is *interaction*: the learning activity on the part of agents exploits past observations, stemming mainly from other agents' choices. Interaction thus shapes the overall process, making it path dependent. Coupling all this with some degree of *non-linearity* might finally allow for multiple equilibria, and hence non-uniqueness of outcomes (lock-in: see Arthur 1994, Amable 1992, Agliardi 1998, Aoki and Yoshikawa 2002, Young 2007).

In Bogliacino and Rampa (2008) we developed a setup, exploiting Bayesian tools, which includes risk aversion and the interaction between demand for and supply of a single new product. Risk aversion is relevant, because during the learning process the emergence of information shapes the confidence of agents (as captured by individual precisions), so altering their willingness to pay. Demand-supply interaction allows one to free the analysis from the single-sided approach prevailing in the literature<sup>2</sup>; in addition, this allows to model explicitly firms' uncertainty over demand.

<sup>&</sup>lt;sup>1</sup> The milestone for the literature on diffusion is the Bass model of epidemiologic diffusion pattern (Bass, 1969). There is a sociological strand of literature focussed on heterogeneity and social effects, e.g. Granovetter (1978), Macy (1991), Abrahamson and Rosenkopf (1993), Valente (1996), Lopez-Pintado and Watts (2006). The orthodox Economics literature is more interested in grounding the choice process on robust roots, using Bayesian theory (Jensen 1982, Feder and O'Mara 1982, Birkhchandani et al. 1992, Bergemann and Välimäki, 1997 Vettas, 1998), but some discussion on more general behaviour rules can be found in Nelson et al. (2002) and Geroski (2000). An excellent review is Hall (2005); an overall discussion of the properties of diffusion curves under alternative setups is Young (2007). On the physics side, one should consider percolation theory as a model of diffusion of ideas and innovations in networks: see e.g. Grimmett (1999) and, as an economic application, Duffie and Manso (2006); an econophysics example is offered by Yanagita and Onozaki (2008).

<sup>&</sup>lt;sup>2</sup> Some noteworthy exceptions are Bergemann and Välimäki (1997) and Vettas (1998).

In the present paper we generalize the previous results analyzing a multiple good case, abandoning monopoly and moving to monopolistic competition. As in the first paper, we provide purely analytical results, characterizing the full set of equilibria of the diffusion process together with their stability properties, without relying on simple simulations exercises which in the end give only a partial understanding of the overall process.

The paper proceeds as follows: Section 2 describes the demand side, Section 3 the supply one, Section 4 presents and discusses the main results, Section 5 concludes. All proofs are collected in the Appendix.

#### Consumers

The individual consumer j (j=1,...M) maximizes her utility choosing the level of consumption of each new good i (i=1,...n), over whose qualities she is uncertain. Qualities are independent normal variables, with known precision and unknown mean. Following a standard Bayesian setting, we assume consumers to be endowed with a prior over the unknown mean quality of each good, defined by two hyper-parameters  $\mu_{j,i,t}$  and  $\tau_{j,i,t}$ , respectively the mean and the precision (the inverse of the variance, see DeGroot, 1970) that evolve through time being updated using Bayes' rule. We assume additively separable preferences. From now on t denotes time, ranging discretely from zero onwards.

We represent the consumer problem in the following way:

$$\max_{\{x_{j,i,t}\}_{i=1,\dots,n}} E[U(x_{j,i,t},\lambda_i) | \mu_{j,i,t-1}, \tau_{j,i,t-1}] = E[\sum_{i=1}^n u(x_{j,i,t}) f(\lambda_i) | \mu_{j,i,t-1}, \tau_{j,i,t-1}]$$
(1)

such that  $\sum_{i=1}^{n} p_{i,t} x_{j,i,t} \le w$ 

where w is the income endowment, for simplicity equal in time and through all consumers<sup>3</sup>.

The function  $f(\cdot)$  is the way the quality of each good is incorporated into agents' preferences. In particular, as in the single good framework of Bogliacino and Rampa

(2008), we assume that U satisfies (i)  $\frac{\partial^2 U}{\partial x_{j,i,t} \partial \lambda_i} > 0$ , meaning that the consumer wishes to

purchase more if quality is higher, for given price; and (*ii*)  $\frac{\partial^3 U}{\partial x_{j,i,t}} \langle \partial \lambda_i \rangle^2 \langle 0, i.e.$  consumers

are risk averse in quality<sup>4</sup>: this suggests that a higher variance of quality tends to depress (expected) marginal utility and hence consumption, for given price.

<sup>&</sup>lt;sup>3</sup> In (1) agents take expectations with respect to all the available information at time t, which obviously includes the information revealed by the market in the previous period, thus we use the time subscript t-1. The reason will become clear in a while.

<sup>&</sup>lt;sup>4</sup> In the standard choice theory, risk aversion is deemed as negativity of the *second* derivative. In our setup, this property obviously holds for quality, since  $u(\cdot)$  is strictly increasing and the utility function U is multiplicatively separable in quality and quantity. However, we preferred to present this characteristic in terms of *third* cross-derivative, because we want to stress the implication for the *quantity* purchased.

As in Bogliacino and Rampa (2008), we posit  $u(\cdot) = (\cdot)^{\delta}$ , and  $f(\lambda_i) = A - \exp(-\lambda_i)$ ; we assume in addition  $\lambda_i \sim N(\mu_i, r)$ , due to random production and/or delivery factors, where the true mean  $\mu_i$  is unknown, and r is known, to consumers; the different qualities are statistically independent. The individual prior, defined over the mean of each quality, is also assumed normal, which allows us to use the properties of the conjugate family. The advantage of these assumptions is threefold: first, they satisfy the two conditions (*i-ii*) above; second, they allow us to "pass through" the expected value operator using the fact that, owing to normality and to the exponential,  $f(\lambda_i)$  is log-normal; finally, they imply, as we shall see, that consumers are not bound to buy a positive quantity of each good. This last property is useful to study the effects of noisy quality signals on consumers' choices, addressing the possibility of lock-in, *i.e.* the failure of a diffusion of a "good" product<sup>5</sup>.

As regards the timing of events, the consumer makes her choice at time t using all information available at that time, which is captured through her posterior, and before knowing the others' choices at t. All the new information refers then to choices made at t-1, hence the hyper-parameters relevant for the choice at t are  $\mu_{i,i,t-1}$  and  $\tau_{i,i,t-1}$ .

Standard maximization implies the following individual demand curve:

$$x_{j,i,t} = \frac{p_{i,t}^{1/(\delta-1)} \left[ \delta \left( A - \exp\left(-\mu_{j,i,t-1} + \frac{\tau_{j,i,t-1} + r}{2}\right) \right) \right]^{1/(1-\delta)}}{\sum_{i=1}^{n} p_{i,t}^{-\delta/(\delta-1)} \left[ \delta \left( A - \exp\left(-\mu_{j,i,t-1} + \frac{\tau_{j,i,t-1} + r}{2}\right) \right) \right]^{1/(1-\delta)} w$$
(2)

where one must intend  $x_{j,i,t} = 0$  whenever  $A - \exp(\cdot) \le 0^6$ . If  $A - \exp(\cdot) > 0$ , we say that consumer *j* is *active* on market *i* at time *t*.

The interpretation is straightforward: each consumer spends a share of its total income on good *i*, depending on the ratio of its price-quality term to that of the whole bundle of goods. Total actual market demand for good *i*,  $Q_{i,i}^D$ , is simply the summation over the *j* index.

After buying the chosen quantity, each active consumer receives a quality signal that she publicly announces to *all* consumers: these signals are used by each of them to update her conjecture. Using the properties of conjugate families (DeGroot, 1970), the posterior parameters for the normal-normal couple (respectively, the likelihood and the prior) are calculated simply as:

$$\mu_{j,i,t} = \frac{\tau_{j,i,t-1} \mu_{j,i,t-1} + r M_{i,t} \lambda_{i,t}}{\tau_{j,i,t-1} + r M_{i,t}}, \quad \tau_{j,i,t} = \tau_{j,i,t-1} + r M_{i,t}$$
(3)

<sup>&</sup>lt;sup>5</sup> A utility function similar to that used in the present setup was proposed also by Roberts and Urban (1988), who however did not explore analytically the dynamic implications of learning and of demand-supply interaction, limiting themselves to simulations exercises.

<sup>&</sup>lt;sup>6</sup> In fact, although the sub-utility  $u(\cdot) = (\cdot)^{\delta}$  satisfies Inada conditions, when this condition holds, the perperiod utility becomes negative, except if the quantity is zero: thus not buying becomes the rational choice.

where  $\overline{\lambda}_{i,t}$  is the quality sample mean, computed from the announced perceived qualities, and  $M_{i,t} \leq M$  is the number of active buyers at date *t*. Notice that consumers treat qualities as independent, and update their conjectures accordingly (that is, separately for each good).

The above equation simply tells us that consumers average their own prior opinions and the sample mean of quality from the new observations, the weight being the relative precisions of the two measures. Moreover, through time individual precisions grow linearly: as one can imagine, given the assumptions of quality-risk-aversion, this fact tends to raise demand in time, due to a simple informational effect.

## Firms

Firms interact in monopolistic competition, each producing a new good at a constant marginal cost  $c_i$ : since each firm corresponds to a different product, as in standard monopolistic competition, we use the *i* index to define a firm. Every firm is uncertain over its own demand. To make things as simple as possible, we assume that it conjectures a linear demand defined by two parameters: more precisely, given the price  $p_{i,t}$ , firm *i* believes that its demand is a random normal variable with mean  $Q_{i,t} = a_i - b_i p_{i,t}$  and precision equal to 1. In addition, firm *i* does not know  $a_i$  and  $b_i$ , and maintains the hypothesis that the distribution of the two parameters is a normal bivariate: the mean and the precision hyper-parameters of this distribution at date *t* are as follows<sup>7</sup>:

$$\mathbf{m}_{i,t} = \begin{bmatrix} \alpha_{i,t} \\ \beta_{i,t} \end{bmatrix}, \quad \mathbf{\Gamma}_{i,t} = \begin{bmatrix} \gamma_{i,1,t} & \gamma_{i,12,t} \\ \gamma_{i,21,t} & \gamma_{i,2,t} \end{bmatrix}$$
(4)

where  $\gamma_{i,1,i}$  and  $\gamma_{i,2,i}$  are positive. Since the firm has surely no reason to conjecture any particular initial value for the correlation among the two mean hyper-parameters, we assume  $\gamma_{i,12,0} = \gamma_{i,21,0} = 0$ . Define also  $\gamma_{i,k} \equiv \gamma_{i,k,0}$ , k = 1,2 as the firm's initial precisions of the mean parameters.

As in the consumer case, the timing is as follows: the firm announces the price before observing demand, hence it uses its (t-1)-conjecture, formed observing demand at time t-1. The firm chooses the price so as to maximize expected profit. Therefore, from standard First Order Condition in monopoly, the price announced at date t is:

$$p_{i,t} = \frac{\alpha_{i,t-1}}{2\beta_{i,t-1}} + \frac{c_i}{2}$$
(5)

and expects the following demand:

<sup>&</sup>lt;sup>7</sup> This derives from our assumption that the conditional distribution of  $Q_{i,t}$  has known precision equal to 1; if this precision were different from 1, the precision matrix  $\Gamma_{i,t}$  would be multiplied by its value. Things could be generalized, but this would be immaterial for our results, since firm's expected profit does not depend on precisions, given risk neutrality.

$$Q_{i,t}^{e}(p_{i,t}) = \frac{\alpha_{i,t-1} - c_i \beta_{i,t-1}}{2}$$
(6)

We neglect any capacity constraint, and assume that the firm can meet all demand<sup>8</sup>.

The updating process on the part of firm *i* follows, again, standard Bayesian rules: using primes to denote transposed vectors, define the row vector  $\mathbf{x}'_{i,i} \equiv \begin{bmatrix} 1 & -p_{i,i} \end{bmatrix}$ . Given our assumptions, one has (DeGroot, 1970, Chapter 11):

$$\mathbf{m}_{i,t} = [\Gamma_{i,t-1} + \mathbf{x}_{i,t-1} \mathbf{x}_{i,t-1}']^{-1} [\Gamma_{i,t-1} \mathbf{m}_{i,t-1} + \mathbf{x}_{i,t-1} Q_{i,t}^{D}]$$
(7)

and

$$\boldsymbol{\Gamma}_{i,t} = [\boldsymbol{\Gamma}_{i,t-1} + \boldsymbol{\mathbf{x}}_{i,t-1} \boldsymbol{\mathbf{x}}_{i,t-1}']$$
(8)

By simple algebra, (7) can be rewritten as:

$$\mathbf{m}_{i,t} = [\Gamma_{i,t-1} + \mathbf{x}_{i,t-1}\mathbf{x}'_{i,t-1}]^{-1}[(\Gamma_{i,t-1} + \mathbf{x}_{i,t-1}\mathbf{x}'_{i,t-1} - \mathbf{x}_{i,t-1}\mathbf{x}'_{i,t-1})\mathbf{m}_{i,t-1} + \mathbf{x}_{i,t-1}Q_{i,t}^{D}] = = [\Gamma_{i,t-1} + \mathbf{x}_{i,t-1}\mathbf{x}'_{i,t-1}]^{-1}[(\Gamma_{i,t-1} + \mathbf{x}_{i,t-1}\mathbf{x}'_{i,t-1})\mathbf{m}_{i,t-1} + \mathbf{x}_{i,t-1}(Q_{i,t}^{D} - \mathbf{x}'_{i,t-1}\mathbf{m}_{i,t-1})] = = \mathbf{m}_{i,t-1} + [\Gamma_{i,t-1} + \mathbf{x}_{i,t-1}\mathbf{x}'_{i,t-1}]^{-1}[\mathbf{x}_{i,t-1}(Q_{i,t}^{D} - \mathbf{x}'_{i,t-1}\mathbf{m}_{i,t-1})] =$$
(9)

In a nutshell, the above expression tells us that the new mean parameters are equal to the previous period's ones, plus a correction term depending the prediction  $error^9$  and adjusted for the new precision matrix.

#### **Equilibria: Main Results**

The system can be fully characterized in terms of firms' and consumers' hyperparameters.

Define  $\boldsymbol{\mu}_{j,t} = [\boldsymbol{\mu}_{j,1,t} \dots \boldsymbol{\mu}_{j,n,t}]'$  and  $\boldsymbol{\tau}_{j,t} = [\boldsymbol{\tau}_{j,1,t} \dots \boldsymbol{\tau}_{j,n,t}]'$  as the vectors of consumer *j*'s hyperparameters at time *t*. Then define  $\boldsymbol{\mu}_t = [\boldsymbol{\mu}_{1,t}' \dots \boldsymbol{\mu}_{M,t}']'$  and  $\boldsymbol{\tau}_t = [\boldsymbol{\tau}_{1,t}' \dots \boldsymbol{\tau}_{M,t}']'$  for all consumers. As regards firms, call  $\gamma_{i,t} = [\gamma_{i,1,t}, \gamma_{i,12,t}, \gamma_{i,21,t}, \gamma_{i,2,t}]'$  the vectorization of the precision matrix of firm *i*'s conjecture at time *t*; posit finally  $\gamma_t = [\gamma_{1,t}' \dots \gamma_{n,t}']'$ , and  $\mathbf{m}_t = [\mathbf{m}_{1,t}' \dots \mathbf{m}_{n,t}']'$ .

Defining  $\mathbf{y}_t = [\mathbf{\mu}'_t \ \mathbf{\tau}'_t \ \mathbf{m}'_t \ \mathbf{\gamma}'_t]'$ , we compact all the updating equations<sup>10</sup> in the following system of 2nM + 6n first order difference equations:

$$\mathbf{y}_t = F(\mathbf{y}_{t-1}) \tag{10}$$

<sup>&</sup>lt;sup>8</sup> An interesting aspect of our setup is the possibility of analysing disequilibrium processes leaving its main features unaltered. In fact, assuming for instance production lags, i.e. the need for the firms to decide quantity *and* price, then equilibrium, as defined short below, is also a *market* equilibrium in the standard sense. From (9) it is clear that all that is needed to discuss the disequilibrium path and the convergence to a market equilibrium is the possibility for firms to observe the true demand for given price in each period and to adjust supply accordingly, e.g. through the use of inventories. However this full characterization is beyond the scope of the present paper, being a question more related to discussion of the features of a general equilibrium.

<sup>&</sup>lt;sup>9</sup> Notice in fact that  $\mathbf{x}'_{i,t-1}\mathbf{m}_{i,t-1}$  is expected demand, given the prior.

<sup>&</sup>lt;sup>10</sup> Taking account of (2) and (5).

which completely describes the learning and diffusion dynamics.

Risk aversion on the part of consumers makes them sensitive to all piece of information available: as time goes by, new information can increase precisions and raise their demand, ceteris paribus. For this reason the system shows path dependence and irreversibility. The relevant equilibrium concept is thus a steady state one, meaning the agents' conjectures remain fixed in time. We use in fact a *conjectural equilibrium* notion: a conjectural equilibrium is a fixed point of (10).

One might think that a conjectural equilibrium requires that all consumers have necessarily learnt the true qualities of the goods. In fact, *if* new information keeps arriving, the Law of Large Numbers implies that consumers are bound to learn the true qualities. It is also possible, however, that consumers are endowed initially with pessimistic conjectures about one of the goods, so demanding a null quantity of it: a null demand, in turn, implies that no signal will arrive at next date, and conjectures remain unchanged (*lock-in*). More importantly, it might happen that, even starting from a positive demand at date t, a highly biased signal switches demand off at date t+1: we term "failure" this phenomenon.

As regards this last point, we recall one of the results of Bogliacino and Rampa (2008).

**Proposition 1.** Suppose that demand for good i is positive at time t. Then, there exists positive probability of failure of the i-th product at time t + 1.

Proof. See Bogliacino and Rampa (2008), Proposition 1.

The argument runs as follows: at every time *t* we can build a complete ordering over the set of consumers in terms of a function of their mean and precision hyper-parameters: the higher its value, the higher a consumer's 'optimism'. If a signal is such biased as to drive the most optimistic consumer below a certain threshold (recall that  $A < \exp(-\lambda_i)$  implies no purchase), then all demand is driven to zero. But then no information is made available to update conjectures, and consumers are locked-in at zero demand<sup>11</sup>.

We come now to a different set of results, assuming that failure does not occur. In this case a conjectural equilibrium is a situation in which consumers' conjectured means have converged to the true mean qualities, *and* in addition firms' conjectures are confirmed by the true demands, so that prediction errors are zero and firms' conjectures remain unchanged at subsequent dates<sup>12</sup>. We can fix the ideas taking  $\mu_{i,j,t} = \mu_i, \forall i, j$ , and studying the dynamics in *expected value* terms<sup>13</sup>, i.e. with the signals always equal to the true qualities, so that demands stay constant for given prices (and consumers' precisions are free to diverge as in the standard Bayesian setting).

This given, define

 $g_i(\mathbf{m}_{i,t-1}) = Q_{i,t}^D [p_{i,t}(\mathbf{m}_{i,t-1}), \mu_i] - \mathbf{x'}_{i,t} \mathbf{m}_{i,t-1}$ 

as the excess of actual demand over expected one for good i; thus the equilibrium condition can be written as follows:

<sup>&</sup>lt;sup>11</sup> One can also, as in Bogliacino and Rampa (2008), study the diffusion dynamics and microfound logistic or concave diffusion patterns depending on initial consumer conjectures, an aim which is beyond the scope of this paper.

<sup>&</sup>lt;sup>12</sup> See (9).

<sup>&</sup>lt;sup>13</sup> With respect to the true distribution of  $\mu_i$ .

$$g(\mathbf{m}_{t}) = [g_{1}(\mathbf{m}_{1t}) \dots g_{n}(\mathbf{m}_{nt})]' = \mathbf{0}, \qquad (11)$$

a set of *n* equations. Then the following Proposition holds.

**Proposition 2.** There exists a n-dimensional equilibrium manifold in the space of firms' parameters.

*Proof.* Trivial: (11) is a system of *n* equations in 2*n* variables.

Conjectural equilibria, then, form a *continuum*: there is not a unique steady state that can be attained by the system.

A natural questions is now the *stability* of equilibria along the manifold. Given the continuum, we must speak of *Lyapunov* stability: that is, stable equilibria are not asymptotically (locally) stable, since a small displacement from one stable equilibrium to another does not cause convergence back to the former. In addition, in the case of stability, different initial conditions lead to different final states.

We study stability of equilibria at any finite time, recalling that we are assuming  $\mu_{i,j,t} = \mu_i, \forall i, j$  and are working in expected values. Hence the stability of equilibria depends entirely on the firms' parameters: indeed we can prove the following Proposition.

**Proposition 3.** *The equilibria where conjectured demand is more elastic than the true one are locally unstable.* 

*Proof.* See Appendix A.1.



Figure 1. An equilibrium where conjectured demand is more elastic than the true one.

The intuition for this result can be seen using Figure 1. A possible equilibrium position is A, where the firm maximizes profits, given its conjecture, and there is no prediction error. From the definition of equilibrium, price and quantity are common to both the true and the conjectured demand, so the condition of Proposition 3 implies that the derivative of the conjectured demand is higher (in absolute value) than that of true demand. On the

contrary, a B-like equilibrium is one where the true demand is less rigid than the conjectured one.

Look at expression (9), and at how it can be rewritten according to Appendix A.2: in the presence of excess demand a firm updates its parameters in such a way that the  $\alpha$  parameter grows and the  $\beta$  parameter decreases<sup>14</sup>. Hence, using (5), it follows that firm will raise its price at the subsequent date. The opposite holds in the presence of excess supply.

Consider now what is happening in a neighbourhood of A; a higher (resp. smaller) price, such as  $p^{h}$  (resp.  $p^{l}$ ) generates excess supply (resp. demand), thus inducing the firm to raising (resp. lowering) further its price. It is then apparent that the system moves away from the A equilibrium. A similar reasoning for a B-like equilibrium shows that in this case there can exist a basin of attraction (unless there is overshooting, a possibility shown by Proposition 4 below). This type of instability is obviously local, since we can only study linear approximations.

In a B-like equilibrium we could still observe local instability at some finite time, instability being of the oscillatory type. This property, however, is smoothed by the passing of time and the instability is rapidly reabsorbed. In fact we have the following Proposition.

**Proposition 4.** In an equilibrium where the true elasticity is high and the demand conjectured by a single firm is more rigid than the true one, there can exist oscillatory instability as long as t is small, and provided that the firms' initial precisions are low.

Proof. See Appendix A.2.

Under the condition of this Proposition, if the system starts in a neighbourhood of some equilibrium the variables will be pushed away from it, and, given the continuum of equilibria, the location of the steady state depends on initial conditions. Observe however that the same unstable equilibria are turned into stable ones by the passing of time, that has the effect of increasing firms' precisions, as apparent from the proof of Proposition 4.

We can finally add some further results in terms of welfare. In Bogliacino and Rampa (2008), studying a single firm, we analyzed the relation between welfare and stability along the equilibrium manifold. In the present context the higher dimensionality makes things more complex: it is not so easy to identify how individual parameters change together along the manifold; and we cannot block n-1 firms, trying to concentrate on a single one, since changing one price implies obviously changes in all expenditure shares. We leave this point for further research.

Our multiple-good setup, however, allows us to analyze the degree of diversification of the decentralized economy and its welfare properties, although under some stricter

<sup>14</sup> In A.1 it is shown that (9) is equivalent to 
$$\mathbf{m}_{i,t} = \mathbf{m}_{i,t-1} + \mathbf{C}_{i,t}\mathbf{g}_i(\cdot)$$
, where  $\mathbf{C}_{i,t} = d_{i,t} \begin{bmatrix} \frac{\gamma_{i,2}}{p_{i,t}} + tp_{i,t} & tp_{i,t} \\ t & \gamma_{i,1} + t \end{bmatrix}$ ,

 $d_{i,t} > 0$ ,  $\mathbf{g}_i(\cdot) \equiv [g_i(\cdot) - g_i(\cdot)]'$ , and  $g_i(\cdot)$  was defined before expression (11) above: see (16) and (23) in that Appendix. As a consequence, one can easily check that if  $g_i(\cdot) > 0$ , that is, if true demand exceeds conjectured demand, then the first element of  $\mathbf{C}_{i,t}\mathbf{g}_i(\cdot)$  is positive, while the second is negative.

assumptions. This is a fairly standard procedure in Monopolistic Competition literature: we need to endogenize the number of firms (i.e. the number of varieties) by means of a fixed cost of entry (see Dixit and Stiglitz, 1977; Tirole, 1988; Bertoletti et al. 2008), then free entry implies a zero profit condition, which closes the model. Indeed, the following Proposition holds.

**Proposition 5.** In equilibrium with endogenous number of firms (assuming a positive fixed cost of entry) and identical marginal cost and qualities of goods, there is over (resp. under) diversification, if for the marginal firms –defined as that who fix the price at the lowest level in equilibrium– the true elasticity is greater (resp. lower) than the conjectured one.

*Proof.* See Appendix A.3.

The interpretation is fairly obvious. Define  $\varepsilon_T$  and  $\varepsilon_C$  to be the true elasticity and that conjectured by firms: when the full information case is characterised by efficiency,  $\frac{\varepsilon_T}{\varepsilon_C} > 1$  in equilibrium makes firms less able to appropriate surplus, pushing entry. The opposite holds for  $\frac{\varepsilon_T}{\varepsilon_C} < 1$ . Thus, interestingly, not only the case  $\frac{\varepsilon_T}{\varepsilon_C}$  is stable, as in Bogliacino and Rampa (2008) and in the present case: it is also efficient in terms of diversification.

A caveat about this result: it is partly dependent on the particular form of the utility function. In general, the relation between the optimal degree of diversification and that prevailing under perfect information depends on how consumers' preferences affect the mark-up, since the latter is related to the ability of firms to appropriate the surplus (see Dixit and Stiglitz 1977). In our case, the iso-elastic assumption guarantees efficiency. However in the general case the ratio among the true elasticity and the conjectured one still allows us to characterize over and under diversification with respect to the perfect information case; of course one cannot say any longer that the degree of diversification under perfect information is also optimal.

## Conclusions

This work studies a monopolistic competitive market, where firms innovate introducing new products and are uncertain about demand; at the same time, consumers are heterogeneous in their expectations on quality, which they are uncertain about. There is interaction in time among and between the market sides: this interaction shapes the learning process and the final pattern observed. This setup is fruitful, in that it allows for the analysis of diffusion, failure, dynamic stability and welfare. Heterogeneity, interaction and non-linearity can coexist with analytical tractability: indeed, we are able to characterize analytically the set of equilibria and their stability properties.

Further research includes the use of more sophisticated firms (oligopoly or conjectural variations models) and the characterization of the welfare properties along the manifold. Of course the model could be simulated to study different diffusion curves and how final outcomes depend on initial conditions.

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## Appendix

#### A.1. Proof of Proposition 3

The system is highly non-linear, so we should limit ourselves to discuss local stability, using a linear approximation in a neighbourhood of one equilibrium. The Jacobian matrix of  $\mathbf{y}_t = F(\mathbf{y}_{t-1})$  is easily checked to be the following one:

$$\mathbf{J} = \begin{bmatrix} \begin{bmatrix} \frac{\partial \boldsymbol{\mu}_{t}}{\partial \boldsymbol{\mu}_{t-1}} \end{bmatrix} & \begin{bmatrix} \frac{\partial \boldsymbol{\mu}_{t}}{\partial \boldsymbol{\tau}_{t-1}} \end{bmatrix} & \mathbf{0}_{nM,2n} & \mathbf{0}_{nM,4n} \\ \mathbf{0}_{nM,nM} & \mathbf{I}_{nM} & \mathbf{0}_{nM,2n} & \mathbf{0}_{nM,4n} \\ \begin{bmatrix} \frac{\partial \mathbf{m}_{t}}{\partial \boldsymbol{\mu}_{t-1}} \end{bmatrix} & \begin{bmatrix} \frac{\partial \mathbf{m}_{t}}{\partial \boldsymbol{\tau}_{t-1}} \end{bmatrix} & \begin{bmatrix} \frac{\partial \mathbf{m}_{t}}{\partial \mathbf{m}_{t-1}} \end{bmatrix} & \begin{bmatrix} \frac{\partial \mathbf{m}_{t}}{\partial \boldsymbol{\gamma}_{t-1}} \end{bmatrix} \\ \mathbf{0}_{4n,nM} & \mathbf{0}_{4n,nM} & \begin{bmatrix} \frac{\partial \boldsymbol{\gamma}_{t}}{\partial \mathbf{m}_{t-1}} \end{bmatrix} & \mathbf{I}_{4n} \end{bmatrix}$$
(12)

where  $\mathbf{I}_k$  is the *k*-identity matrix and  $\mathbf{0}_{k,p}$  is a *k*-by-*p* null matrix.

The stability condition is that all the eigenvalues of J, evaluated at an equilibrium, do not lie outside the unit circle. We need some preliminary results.

Claim 1. At an equilibrium, the eigenvalues of J are those of the four blocks along its main diagonal.

*Proof.* We need simply to prove that  $\frac{\partial \mathbf{m}_{t}}{\partial \gamma_{t-1}} = \mathbf{0}$ . Define first

$$\begin{bmatrix} \boldsymbol{\Gamma}_{i,t-1} + \boldsymbol{\mathbf{x}}_{i,t} \boldsymbol{\mathbf{x}}_{i,t}' \end{bmatrix} = \begin{bmatrix} \boldsymbol{\Gamma}_{i,t-1} + \begin{bmatrix} 1 & -p_{i,t} \\ -p_{i,t} & p^{2}_{i,t} \end{bmatrix} \end{bmatrix} \equiv \mathbf{A}_{i,t}(\boldsymbol{\Gamma}_{i,t-1}, \mathbf{m}_{i,t-1})$$
(13)

and

$$\begin{bmatrix} \mathbf{x}_{i,t} (Q_{i,t}^{D} - \mathbf{x}_{i,t}' \mathbf{m}_{i,t-1}) \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & p_t \end{bmatrix} \begin{bmatrix} (Q_{i,t}^{D} - \mathbf{x}_{i,t}' \mathbf{m}_{i,t-1}) \\ - (Q_{i,t}^{D} - \mathbf{x}_{i,t}' \mathbf{m}_{i,t-1}) \end{bmatrix}$$

$$\equiv \mathbf{B}_{i,t} (\mathbf{m}_{i,t-1}) \mathbf{g}_i (\mathbf{m}_{i,t-1}, \mathbf{\mu}_{t-1}, \mathbf{\tau}_{t-1})$$
(14)

where  $\mathbf{g}_i(\cdot) \equiv [g_i(\cdot) - g_i(\cdot)]'$  and  $g_i(\cdot)$  was defined before expression (11). Define finally

$$\mathbf{C}_{i,t}(\mathbf{\Gamma}_{i,t-1},\mathbf{m}_{i,t-1}) \equiv [\mathbf{A}_{i,t}(\mathbf{\Gamma}_{i,t-1},\mathbf{m}_{i,t-1})]^{-1} \mathbf{B}_{i,t}(\mathbf{m}_{i,t-1})$$
(15)

Summing up, firm *i*'s updating formula can be written as

$$\mathbf{m}_{i,t} = \mathbf{m}_{i,t-1} + \mathbf{C}_{i,t}(\Gamma_{i,t-1}, \mathbf{m}_{i,t-1})\mathbf{g}_i(\mathbf{m}_{i,t-1})$$
(16)

and the block which interests us now is:

$$\begin{bmatrix} \frac{\partial \mathbf{m}_{t}}{\partial \gamma_{t-1}} \end{bmatrix} = \begin{bmatrix} \begin{bmatrix} \frac{\partial \mathbf{m}_{1,t}}{\partial \gamma_{t-1}} \end{bmatrix} \\ \vdots \\ \begin{bmatrix} \frac{\partial \mathbf{m}_{n,t}}{\partial \gamma_{t-1}} \end{bmatrix} \end{bmatrix} = \begin{bmatrix} \begin{bmatrix} \frac{\partial \mathbf{C}_{1,t}}{\partial \gamma_{t-1}} \end{bmatrix} \mathbf{g}_{1} + \mathbf{C}_{1,t} \begin{bmatrix} \frac{\partial \mathbf{g}_{1}}{\partial \gamma_{t-1}} \end{bmatrix} \\ \vdots \\ \begin{bmatrix} \frac{\partial \mathbf{C}_{n,t}}{\partial \gamma_{t-1}} \end{bmatrix} \mathbf{g}_{n} + \mathbf{C}_{n,t} \begin{bmatrix} \frac{\partial \mathbf{g}_{n}}{\partial \gamma_{t-1}} \end{bmatrix} \end{bmatrix}$$
(17)

which is clearly equal to zero, since from (11)  $\mathbf{g}_i(\cdot) = 0$  in equilibrium, and the  $\mathbf{g}_i$ 's themselves do not depend on firms' precisions. *QED* 

We can thus concentrate on the four principal blocks of **J**. The NW block has eigenvalues lower than one, and tending to one as time goes to infinity: they are the weights attached to consumers' prior means in the updating formulae: see (3) above. The second and fourth blocks give rise to respectively nM and 4n eigenvalues equal to one: they relate to the updating of consumers' and firms' precisions, and are immaterial for stability. In fact changes in the precisions do not affect the equilibrium itself, being more important in the initial, rather than the final, phases of the learning process (Rampa, 1989).

We are thus left with the 2*n* eigenvalues of the block 
$$\left[\frac{\partial \mathbf{m}_{t}}{\partial \mathbf{m}_{t-1}}\right]$$
  
Claim 2. The eigenvalues of  $\left[\frac{\partial \mathbf{m}_{t}}{\partial \mathbf{m}_{t-1}}\right]$  are as follows:

(i) n eigenvalues are equal to one, implied by the continuum of equilibria;

(ii) the other n eigenvalues are equal to  $\rho(\mathbf{D}_t\mathbf{G}_t)+1$ , where  $\rho(\cdot)$  is the column vector of the eigenvalues of the argument, 1 is a column vector of ones,  $\mathbf{D}_t$  is a diagonal matrix with positive diagonal elements, and  $\mathbf{G}_t$  is a matrix with positive extra-diagonal elements.

#### Proof

(*i*) Define the following matrix:

$$\mathbf{C}_{t} = \begin{bmatrix} \mathbf{C}_{1,t} & \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \dots & \mathbf{0} \\ \mathbf{0} & \mathbf{0} & \mathbf{C}_{n,t} \end{bmatrix}$$
(18)

and

$$\mathbf{g} = [\mathbf{g}_1(\mathbf{m}_{t-1})' \quad \dots \quad \mathbf{g}_n(\mathbf{m}_{t-1})']'.$$
(19)

Given (16), and given that  $\mathbf{g}_i(\cdot) = 0$  in equilibrium, one deduces:

$$\left[\frac{\partial \mathbf{m}_{t}}{\partial \mathbf{m}_{t-1}}\right] = \mathbf{I}_{2n} + \mathbf{C}_{t} \left[\frac{\partial \mathbf{g}}{\partial \mathbf{m}_{t-1}}\right]$$
(20)

This matrix has 2n eigenvalues equal to 1 plus those of the second term. Since by construction  $\mathbf{g}(\cdot)$  is formed by 2n terms, n of which are the opposite of the remaining n,  $\left[\frac{\partial \mathbf{g}}{\partial \mathbf{m}_{t-1}}\right]$  has rank n,

and the same is generically true for  $\mathbf{C}_{t}\left[\frac{\partial \mathbf{g}}{\partial \mathbf{m}_{t-1}}\right]$ : hence the latter has *n* eigenvalues equal to zero.

Thus, we can conclude that *n* eigenvalues of  $\left[\frac{\partial \mathbf{m}_{t}}{\partial \mathbf{m}_{t-1}}\right]$  are unitary. These *n* unitary eigenvalues correspond precisely to the very existence of the *n*-dimensional continuum of equilibria: a move along this continuum is followed neither by divergence nor by convergence to the previous point. This completes the proof of part (*i*) of Claim 2.

(*ii*) In order to study the remaining *n* eigenvalues of  $\mathbf{C}_t \begin{bmatrix} \frac{\partial \mathbf{g}}{\partial \mathbf{m}_{t-1}} \end{bmatrix}$ , we can write:

$$\begin{bmatrix} \frac{\partial \mathbf{g}}{\partial \mathbf{m}_{t-1}} \end{bmatrix} = \begin{bmatrix} \frac{\partial \mathbf{g}}{\partial \mathbf{p}_t} \end{bmatrix} \begin{bmatrix} \frac{\partial \mathbf{p}_t}{\partial \mathbf{m}_{t-1}} \end{bmatrix}$$
where  $\begin{bmatrix} \frac{\partial \mathbf{g}}{\partial \mathbf{p}_t} \end{bmatrix}$  is 2*n* by *n*, and  $\begin{bmatrix} \frac{\partial \mathbf{p}_t}{\partial \mathbf{m}_{t-1}} \end{bmatrix}$  is *n* by 2*n*, and  $\mathbf{p}_t = [p_{1,t} \dots p_{n,t}]'$ .

It is well known that the non-zero eigenvalues of  $\mathbf{C}_t \left[ \frac{\partial \mathbf{g}}{\partial \mathbf{p}_t} \right] \left[ \frac{\partial \mathbf{p}_t}{\partial \mathbf{m}_{t-1}} \right]$  are the same as those of

$$\left\lfloor \frac{\partial \mathbf{p}_t}{\partial \mathbf{m}_{t-1}} \right\rfloor \mathbf{C}_t \left\lfloor \frac{\partial \mathbf{g}}{\partial \mathbf{p}_t} \right\rfloor$$
. Exploiting (5),  $\frac{\partial \mathbf{p}_t}{\partial \mathbf{m}_{t-1}}$  has the following expression:

$$\begin{bmatrix} \frac{\partial \mathbf{p}_{i}}{\partial \mathbf{m}_{i-1}} \end{bmatrix} = \begin{bmatrix} \frac{\partial p_{1}}{\partial \alpha_{1}} & \frac{\partial p_{1}}{\partial \beta_{1}} & 0 & 0 & \cdots & \cdots & 0 & 0 \\ 0 & 0 & \frac{\partial p_{2}}{\partial \alpha_{2}} & \frac{\partial p_{2}}{\partial \beta_{2}} & \cdots & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & 0 & \cdots & \cdots & \frac{\partial p_{n}}{\partial \alpha_{n}} & \frac{\partial p_{n}}{\partial \beta_{n}} \end{bmatrix}$$

$$= \frac{1}{2} \begin{bmatrix} \frac{1}{\beta_{1}} & \frac{-\alpha_{1}}{\beta_{1}^{2}} & 0 & 0 & \cdots & \cdots & 0 & 0 \\ 0 & 0 & \frac{1}{\beta_{2}} & \frac{-\alpha_{2}}{\beta_{2}^{2}} & \cdots & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & 0 & \cdots & \cdots & \frac{1}{\beta_{n}} & \frac{-\alpha_{n}}{\beta_{n}^{2}} \end{bmatrix}$$

$$(22)$$

Each of the diagonal blocks of matrix  $\mathbf{C}_{t}$ , in turn, can be written<sup>15</sup> as:

$$\mathbf{C}_{i,t} = d_{i,t} \begin{bmatrix} \frac{\gamma_{i,2}}{p_{i,t}} + tp_{i,t} & tp_{i,t} \\ t & \gamma_{i,1} + t \end{bmatrix}$$
(23)

where  $d_{i,t} = \frac{p_{i,t}}{\gamma_{i,1}\gamma_{i,2} + t(\gamma_{i,2} + \gamma_{i,1}p_{i,t}^2)} > 0$  for  $t < \infty$ , and recalling that  $\gamma_{i,k} \equiv \gamma_{i,k,0}, k = 1, 2$ .

<sup>&</sup>lt;sup>15</sup> See Bogliacino and Rampa (2008), expression C.2 of the Appendix.

Hence, the product  $\mathbf{C}_{t} \begin{bmatrix} \frac{\partial \mathbf{g}}{\partial \mathbf{m}_{t-1}} \end{bmatrix}$  is easily seen to be a 2*n*-by-*n* matrix composed by *n*-by-*n* column vectors of the following form:

$$d_{i,t} \begin{bmatrix} \frac{\gamma_{i,2}}{p_{i,t}} \\ -\gamma_{i,1} \end{bmatrix} \frac{\partial g_i}{\partial p_{k,t}}$$
(24)

As a result of (22)-(24), the product  $\left[\frac{\partial \mathbf{p}_t}{\partial \mathbf{m}_{t-1}}\right] \mathbf{C}_t \left[\frac{\partial \mathbf{g}}{\partial \mathbf{p}_t}\right]$  is a *n*-by-*n* matrix with elements:

$$\frac{1}{2}d_{i,t}\left(\frac{\gamma_{i2}}{\beta_{i1,t}p_{i,t}} + \frac{\alpha_{i,t}}{\beta_{i,t}^2}\gamma_{i1}\right)\frac{\partial g_i}{\partial p_{k,t}} = \frac{d_{i,t}}{2\beta_{i1,t}}\left(\frac{\gamma_{i2}}{p_{i,t}} + \frac{\alpha_{i,t}}{\beta_{i,t}}\gamma_{i1}\right)\frac{\partial g_i}{\partial p_{k,t}}$$
(25)

We can thus write the expression:

$$\begin{bmatrix} \frac{\partial \mathbf{p}_{t}}{\partial \mathbf{m}_{t-1}} \end{bmatrix} \mathbf{C}_{t} \begin{bmatrix} \frac{\partial \mathbf{g}}{\partial \mathbf{p}_{t}} \end{bmatrix} = \begin{bmatrix} \frac{d_{1,t}}{2\beta_{1,1,t}} \begin{pmatrix} \frac{\gamma_{1,2}}{p_{1,t}} + \frac{\alpha_{1,t}}{\beta_{1,t}} \gamma_{1,1} \end{pmatrix} & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \frac{d_{n,t}}{2\beta_{n1,t}} \begin{pmatrix} \frac{\gamma_{n,2}}{p_{n,t}} + \frac{\alpha_{n,t}}{\beta_{n,t}} \gamma_{n,1} \end{pmatrix} \end{bmatrix} \begin{bmatrix} \frac{\partial g_{1}}{\partial p_{1,t}} & \cdots & \frac{\partial g_{1}}{\partial p_{n,t}} \\ \vdots & \ddots & \vdots \\ \frac{\partial g_{n}}{\partial p_{1,t}} & \cdots & \frac{\partial g_{n}}{\partial p_{n,t}} \end{bmatrix} = \mathbf{D}_{t} \mathbf{G}_{t}$$
(26)

where  $\mathbf{D}_t$  is a diagonal matrix with positive diagonal elements.

We need now to prove that the extra-diagonal elements of the matrix  $\mathbf{G}_t = \begin{bmatrix} \frac{\partial g_i}{\partial p_{k,t}} \end{bmatrix}$  are positive. Using the definition of  $\mathbf{g}_i(\cdot)$  after expression (14), one can see that the elements outside the main diagonal of  $\mathbf{G}_t$  are the derivatives of the demand w.r.t. prices of the other goods  $(\frac{\partial x_i}{\partial p_k}, i \neq k)$ . Remind that the consumer's problem is:

$$\max\sum_{i=1}^{n} u(x_i) f(\lambda_i) s.t. \sum_{i=1}^{n} p_i x_i \le w$$
(27)

whose first order condition is (calling *z* the Lagrange multiplier):

$$u'(x_i)f(\lambda_i) - zp_i = 0.$$
<sup>(28)</sup>

If we calculate  $\frac{\partial x_i}{\partial p_k}$  using the implicit function theorem, we get  $\frac{\partial x_i}{\partial p_k} = \frac{p_k}{u''(x_i)f(\lambda_i)}\frac{\partial z}{\partial p_k}$ , where  $\frac{\partial x_i}{\partial p_k}$ 

 $\frac{\partial z}{\partial p_k}$  is defined by the boundary condition:

$$\sum_{i=1}^{n} p_i \left[ u^{i-1} \left( \frac{p_i z}{f(\lambda_i)} \right) \right] - w = 0$$
<sup>(29)</sup>

Thus we have:

$$\frac{\partial z}{\partial p_k} = -\frac{u^{\prime-1} \left(\frac{p_i z}{f(\lambda_i)}\right) + \frac{p_i z}{f(\lambda_k)} \left[ \left[ u^{\prime-1} \right] \left(\frac{p_k z}{f(\lambda_k)}\right) \right]}{\sum_{i=1}^n \frac{p_i z}{f(\lambda_k)} \left[ \left[ u^{\prime-1} \right] \left(\frac{p_i z}{f(\lambda_i)}\right) \right]}$$
(30)

Simple manipulation of the numerator above, in particular using the Inverse Function Theorem, shows that the numerator itself is negative as long as:

$$\left|\frac{u'(x_i^*)}{u''(x_i^*)x_i^*}\right| > 1$$
(31)

where  $x_i^*$  is a solution to (28). With additively separable preferences the LHS is nothing else than the elasticity of demand (Bertoletti et al., 2008), so (31) is certainly satisfied with our formulation<sup>16</sup>  $u(x) = x^{\delta}$ . So (30) is negative, and we can conclude that  $\frac{\partial x_i}{\partial p_k} > 0, i \neq k$ . This completes the proof of Claim 2. *QED* 

**Claim 3.** If at an equilibrium the elasticity of the conjectured demand is greater than the elasticity of the true demand,  $\mathbf{D}_t \mathbf{G}_t$  has at least one positive eigenvalue.

#### Proof

. . .

As we said, the elements of  $\mathbf{D}_i$  are positive for  $t < \infty$ . From the definition of  $g_i(\cdot)$  and from (26) it follows that the *i*-th element along the main diagonal of  $\mathbf{G}_i$  can be written as

$$\frac{\partial g_i}{\partial p_i} = \frac{\partial x_i}{\partial p_i} \bigg|_{TrueDemand} - \frac{\partial x_i}{\partial p_i} \bigg|_{ConjecturedDemand}$$
(32)

By definition of elasticity, using the fact that at an equilibrium the price-quantity couple is the same for the true and the conjectured demand, the assumption of Claim 3 is equivalent to

$$\left|\frac{\partial x_i}{\partial p_i}\right|_{TrueDemand} - \left|\frac{\partial x_i}{\partial p_i}\right|_{ConjecturedDemand} < 0.$$
(33)

Since both true and conjectured demand are negatively sloped, (33) implies that (32) is positive.

Using the results of Claim 2, part (*ii*), the fact that (29) is positive, and finally the fact that  $\mathbf{D}_t$  is a positive diagonal matrix, we conclude that all elements of  $\mathbf{D}_t \mathbf{G}_t$  are positive. Claim 3 then follows from the Perron-Frobenius Theorem<sup>17</sup>. *QED* 

We can finally complete the proof of **Proposition 2**. From Claim 3 and Claim 2, part (*ii*), it follows that  $\left[\frac{\partial \mathbf{m}_{t}}{\partial \mathbf{m}_{t-1}}\right]$  has an eigenvalue greater than one. Claim 1 says that the eigenvalues of  $\left[\frac{\partial \mathbf{m}_{t}}{\partial \mathbf{m}_{t-1}}\right]$  are also eigenvalues of **J**; hence **J** has an eigenvalue greater than one. *QED* 

<sup>&</sup>lt;sup>16</sup> Indeed one can argue that the property is completely general, since a firm will never find optimal to fix a price where the elasticity of demand is lower than one. However, this condition is true only for conjectured demand, and not for the true one.

<sup>&</sup>lt;sup>17</sup> See Lancaster-Tismenetsky (1985), Theorem 1 on page 536.

#### A.2 Proof of Proposition 4

Using Claim 2, part (*ii*), we need to prove that  $\mathbf{D}_t \mathbf{G}_t$  can have a real eigenvalue lower than -1. In what follows we will drop the time subscripts for easiness of notation, writing **DG** instead of  $\mathbf{D}_t \mathbf{G}_t$ : in fact we are evaluating the jacobian matrix **J** at an equilibrium (all consumers have converged to the true quality value and prices are fixed at the equilibrium values). We start once more from some preliminary results.

Claim 4. The following statements hold:

- i) **G** is symmetric, and one has  $\mathbf{G} = \mathbf{G}_1 + \mathbf{G}_2$ , where  $\mathbf{G}_1$  is diagonal and  $\mathbf{G}_2$  has rank 1;
- *ii)* **DG** has the same eigenvalues as  $\mathbf{D}^{1/2}\mathbf{G}\mathbf{D}^{1/2}$ ;
- *iii*)  $\mathbf{D}^{1/2}\mathbf{G}\mathbf{D}^{1/2} = \mathbf{D}\mathbf{G}_1 + \mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2}$ .

#### Proof

(*i*) To find the elements of matrix  $\mathbf{G}$ , we differentiate the equilibrium conditions (11) w.r.t. prices, using the definition of demand (3) and imposing equilibrium condition. We get:

$$\begin{cases}
\frac{\partial g_i}{\partial p_i} = \frac{1}{p_i(\delta - 1)}Q_i^D + \beta_i + \frac{1}{(1 - \delta)}\frac{\delta Q_i^D}{M_W}Q_i^D & i = k \\
\frac{\partial g_i}{\partial p_k} = \frac{1}{(1 - \delta)}\frac{\delta Q_i^D}{M_W}Q_k^D > 0 & i \neq k
\end{cases}$$
(34)

which implies that **G** is symmetric<sup>18</sup>.

Hence,  $\mathbf{G} = \mathbf{G}_1 + \mathbf{G}_2$ : the first matrix is a diagonal matrix with elements  $\frac{1}{p_i(\delta - 1)}Q_i^D + \beta_i$ , and

 $\mathbf{G}_2 = \frac{\delta}{1-\delta} \frac{1}{Mw} \mathbf{Q} \mathbf{Q}'$ , where  $\mathbf{Q}$  is the vector of equilibrium quantities. But the non-zero  $\delta$  1

eigenvalues of  $\mathbf{G}_2$  are the same as those of  $\frac{\delta}{1-\delta} \frac{1}{Mw} \mathbf{Q}' \mathbf{Q}$ , so  $\mathbf{G}_2$  has rank 1.

(*ii*) We know from Claim 2, part *ii*, that **D** is diagonal and non singular, thus it admits  $\mathbf{D}^{1/2}$ . But **DG** is similar to  $\mathbf{D}^{-1/2}(\mathbf{DG})\mathbf{D}^{1/2} = \mathbf{D}^{1/2}\mathbf{GD}^{1/2}$ , and the two matrices have the same eigenvalues.

(iii) Given part (i) above, and since diagonal matrices commute, one can write

$$\mathbf{D}^{1/2}\mathbf{G}\mathbf{D}^{1/2} = \mathbf{D}^{1/2}\mathbf{G}_1\mathbf{D}^{1/2} + \mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2} = \mathbf{D}\mathbf{G}_1 + \mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2}$$
(35)

where the first term is a diagonal matrix, while the second term is  $\frac{\delta}{1-\delta}\frac{1}{Mw}$  times the external product of  $\mathbf{D}^{1/2}\mathbf{Q}$  and itself, hence is symmetric. *QED* 

<sup>&</sup>lt;sup>18</sup> The  $\frac{\partial g_i}{\partial p_k}$ ,  $i \neq k$ , are equal to the extra-diagonal terms of the Jacobian of the demand functions, thus in

the general case symmetry holds only for compensated demands (see Theorem 1: McKenzie, 2002, p. 10), and not for the Marshallian ones, because of income effects (McKenzie, 2002, p. 12). However the condition holds under our present assumptions.

Define now the following terms:  $\phi_i$ , as the *i*-th element of the main diagonal of **D**;  $v_i \equiv \frac{\gamma_{i1}}{\gamma_{i2}}$ , as the ratio between firm *i*'s initial precisions; and  $s_i \equiv \frac{\varepsilon_T}{\varepsilon_i}$ , as the ratio between the true elasticity and the conjectured elasticity in the market for product *i*. We proceed with the following

Claim 5. The following statements hold:

(i) the eigenvalues of 
$$\mathbf{DG}_1$$
 are  $(1-s_i)\frac{1}{2}\frac{1+v_ip_i^2\left(1+\frac{1}{\varepsilon_i}\right)}{\gamma_{i1}+t\left(1+v_ip_i^2\right)}, \qquad i=1,\ldots,n$  (36)

(ii) the non zero eigenvalue of 
$$\mathbf{D}^{1/2}\mathbf{G}_{2}\mathbf{D}^{1/2}$$
 is  $\frac{1}{2}\delta\sum_{k}\left[\frac{p_{k}Q_{k}}{\sum_{k}p_{k}Q_{k}}s_{k}\frac{1+v_{k}p_{k}^{2}\left(1+\frac{1}{\varepsilon_{k}}\right)}{\gamma_{k1}+t\left(1+v_{k}p_{k}^{2}\right)}\right]$  (37)

Proof

(*i*) We concentrate first on the *i*-diagonal element of the diagonal matrix **D**,  $\phi_i$ . From the definitions after (23) we have  $\phi_i = \frac{d_i}{2\beta_i} \left( \frac{\gamma_{i2}}{p_i} + \frac{\alpha_i}{\beta_i} \gamma_{i1} \right) = \frac{1}{2\beta_i} \frac{\gamma_{i2} + \gamma_{i1}p_i \frac{\alpha_i}{\beta_i}}{\gamma_{i1}\gamma_{i2} + t(\gamma_{i2} + \gamma_{i1}p_i^2)}$ . From (5) one deduces in addition that  $\frac{\alpha_i}{\beta_i} = 2p_i - c_i$ , hence one can write  $p_i \frac{\alpha_i}{\beta_i} = p_i^2 + (p_i - c_i)p_i$ . But the monopoly pricing rule (Tirole, 1988) is that the *Lerner Index* is equal to the inverse of the

conjectured elasticity, whence  $(p_i - c_i) = \frac{p_i}{\varepsilon_i}$ ; then we have  $p_i \frac{\alpha_i}{\beta_i} = p_i^2 + \frac{p_i^2}{\varepsilon_i}$ . Given the

definition 
$$v_i \equiv \frac{\gamma_{i1}}{\gamma_{i2}}$$
, we get thus  $\phi_i = \frac{1}{2\beta_i} \frac{1 + v_i p_i^2 \left(1 + \frac{1}{\varepsilon_i}\right)}{\gamma_{i1} + t \left(1 + v_i p_i^2\right)}$ 

Now, recall from the proof of Claim 4, part (*i*), that the *i*-diagonal element of the diagonal matrix  $\mathbf{G}_1$  is  $\frac{1}{p_i(\delta-1)}Q_i + \beta_i$ . In equilibrium one has  $Q_i = (p_i - c_i)\beta_i$ , which is equal to  $\frac{p_i}{\varepsilon_i}\beta_i$  by the monopoly pricing rule. Using the true elasticity  $\varepsilon_T = \frac{1}{1-\delta}$  and the definition  $s_i \equiv \frac{\varepsilon_T}{\varepsilon_i}$ , we obtain

$$\frac{1}{p_i(\delta-1)}Q_i + \beta_i = \beta_i(1-s_i).$$

The *i*-th diagonal element of the diagonal matrix  $\mathbf{DG}_1$ , that is its *i*-th eigenvalue, is then equal to  $\beta_i (1 - s_i) \phi_i$ . Substituting the value of  $\phi_i$  found above, we get (36).

(*ii*) Using Claim 4, we write  $\mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2} = \frac{\delta}{1-\delta}\frac{1}{Mw}\mathbf{D}^{1/2}\mathbf{Q}\mathbf{Q}'\mathbf{D}^{1/2}$ , a matrix that has rank 1 and thus a single non-zero eigenvalue. It is easily checked that this matrix has the same non-zero eigenvalue as,  $\frac{\delta}{1-\delta}\frac{1}{Mw}\mathbf{Q}'\mathbf{D}\mathbf{Q} = \frac{\delta}{1-\delta}\frac{1}{Mw}\sum_{k}\phi_kQ_k^2$ . Exploit again the equilibrium fact

 $Q_k = (p_k - c_k)\beta_k$ , the monopoly pricing rule  $(p_k - c_k) = \frac{p_k}{\varepsilon_k}$ , and the definitions  $s_k \equiv \frac{\varepsilon_T}{\varepsilon_k}$  and

 $\varepsilon_T = \frac{1}{1-\delta}$ , which together imply  $\frac{1}{1-\delta}Q_k^2 = Q_k p_k s_k \beta_k$ . Substituting finally the value of  $\phi_i$  found in part (*i*) above, and using the budget constraint  $Mw = \sum_k p_k Q_k$ , one gets (37). *QED* 

We will use the following result, which we call Claim 6.

**Claim 6.** If **A** and **B** are two symmetric matrices and if rank(**B**) = 1, then the i-th eigenvalue of **A+B**, say  $\rho_i(\mathbf{A} + \mathbf{B})$ , is equal to  $\rho_i(\mathbf{A}) + m_i \cdot \rho(\mathbf{B})$ , where  $m_i \in [0,1]$ , and  $\rho(\mathbf{B})$  is the only non-zero eigenvalue of **B**.

Proof. See Wilkinson (1965), pp. 97-98.

QED

We are finally ready to complete the proof of **Proposition 4**, stating that an eigenvalue of **J**, say  $\rho_i(\mathbf{J})$ , can be lower than -1. By Claim 2, part (*ii*), this means  $\rho_i(\mathbf{DG}) < -2$ . We will posit sufficient conditions for this result.

We know from Claim 4 that **DG** has the same eigenvalues as  $\mathbf{D}^{1/2}\mathbf{G}\mathbf{D}^{1/2}$ , and that  $\mathbf{D}^{1/2}\mathbf{G}\mathbf{D}^{1/2} = \mathbf{D}\mathbf{G}_1 + \mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2}$ , the sum of two symmetric matrices. Claim 5, in turn, gives expressions for the eigenvalues of  $\mathbf{D}\mathbf{G}_1$  and  $\mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2}$ . Claim 6 asserts finally that  $\rho_i(\mathbf{D}\mathbf{G}) = \rho_i(\mathbf{D}\mathbf{G}_1) + m_i \cdot \rho(\mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2})$ , with  $m_i \in [0,1]$ , implying  $\rho_i(\mathbf{D}\mathbf{G}) \le \rho_i(\mathbf{D}\mathbf{G}_1) + \rho(\mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2})$ .

Suppose now that the true elasticity  $\varepsilon_T$  is very high, implying  $\delta \approx 1$  (recall that  $\delta < 1$  anyway), and that all firms but the *i*-th one conjecture an elasticity  $\varepsilon_k$  very near to the true one, while the *i*-th firm conjectures a low elasticity  $\varepsilon_i$ . This implies  $s_k \approx 1$  and  $\frac{1}{\varepsilon_k} \approx 0$  for all  $k \neq i$ ; at the same time, the *i*-th firm will price very high, so that (2) implies a low share of consumers' expenditure on good *i*; in addition,  $s_i$  is very high. Suppose further that all firms have low initial precisions of their  $\alpha$  parameters, so that  $\gamma_{k1}$  is near to zero,  $\forall k$ . Finally, consider the system at the very start of the learning process, meaning t = 1. Looking carefully at (37), all this implies that  $\rho(\mathbf{D}^{1/2}\mathbf{G}_2\mathbf{D}^{1/2}) \approx \frac{1}{2}$ .

This given, Claim 6 ca be written as  $\rho_i(\mathbf{DG}) \leq (1-s_i)\frac{1}{2}\frac{1+v_ip_i^2\left(1+\frac{1}{\varepsilon_i}\right)}{1+v_ip_i^2} + \frac{1}{2}$  for the *i*-th eigenvalue of **DG**.

We need to have 
$$(1-s_i)\frac{1}{2}\frac{1+v_ip_i^2\left(1+\frac{1}{\varepsilon_i}\right)}{1+v_ip_i^2} + \frac{1}{2} < -2$$
, meaning  $(1-s_i)\frac{1+v_ip_i^2\left(1+\frac{1}{\varepsilon_i}\right)}{1+v_ip_i^2} < -5$ . This

might well be the case, given our current assumptions of a low  $\varepsilon_i$  and a high  $s_i$ , and if in addition one assumes that  $v_i$  is high, *i.e.* firm *i* is initially more uncertain on the  $\beta$  parameter than on the  $\alpha$ parameter. Notice that, as time passes (t > 1) and hence  $\gamma_{k1}$  grows above zero, the result does not hold any longer.

This completes the proof of Proposition 4.

QED

## A.3 Proof of Proposition 5

Let introduce a fixed cost of entry equal to F. In order to calculate the optimal degree of diversification, we need to fix price equal to marginal cost, introduce lump sum taxation for an amount  $\frac{nF}{M}$  for each consumer (reducing her income to  $w - \frac{nF}{M}$ ) and maximize the indirect utility function in n. The demand for goods of identical quality is:

$$Q^{D} = M(w - nF/M) \frac{p^{1/(1-\delta)} \left[\delta f(\lambda)\right]^{1/(1-\delta)}}{np^{\delta/(1-\delta)} \left[\delta f(\lambda)\right]^{1/(1-\delta)}} = \frac{M(w - nF/M)}{np}$$
(38)

Replacing the price equal to marginal cost, the indirect utility function is given by  $n^{1-\delta}c^{-\delta}M^{\delta}(w-nF/M)^{\delta}$ , which must be maximized in *n*, considered as a real variable for simplicity. The first order condition is also sufficient, due to the strict concavity of the indirect utility function, and is the following:

$$n^{e} = (1 - \delta)Mw/F = \frac{Mw}{F\varepsilon_{T}}$$
(39)

The equilibrium condition with endogenous number of firms is a zero profit condition for the marginal firm, defined by the price  $\overline{p} = \min\{p_i \mid p_i = p_i^*\}$  (where  $p_i^*$  are equilibrium prices), given the equality of marginal cost and quality through firms (and convergence of consumers' conjectures in equilibrium). Hence:

$$(\overline{p} - c)Q = F \frac{\alpha - c\beta}{2\beta} \frac{Mw}{np} = F$$
(40)

By simple algebra we get

$$n^* = \frac{\alpha - c\beta}{\alpha + c\beta} \frac{Mw}{F} = \frac{Mw}{F\varepsilon_c}$$
(41)

Over (respectively under) diversification is the case  $n^* > n^e$  (respectively  $n^* < n^e$ ). Replacing with (39) and (41) completes the proof. *QED*