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**A GENERALIZED NEURAL LOGIT MODEL
FOR AIRPORT AND ACCESS MODE CHOICE IN GERMANY**

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Abstract

The purpose of this paper is to present an airport and access mode choice model based on a new discrete choice approach called “Generalized Neural Logit-Model”. This approach employs artificial neural networks to model the utility function and correlations within the choice set and genetic algorithms to optimize the network structure. To evaluate the new approach, a nested logit approach is chosen as a benchmark. The concept of alternative groups is employed for estimating a market segment-specific airport and access mode choice model and therefore it is generally applicable to any number of airports and combinations of airports and access modes. Hence it is possible to analyze future scenarios in terms of new airport constellations and new airport access modes. To achieve this, Kohonen’s Self-Organizing-Maps are used to identify different airport clusters from a demand-oriented point of view and assign every airport to the appropriate cluster.

Keywords: Airport and access mode choice model, Artificial neural networks, Concept of alternative groups, Discrete choice model, Generalized Neural Logit-Model, Kohonen’s Self Organizing Maps

1. Introduction

This paper presents a novel approach in discrete choice modeling based on artificial neural networks and is an excerpt of the doctoral thesis of the author. Research is mainly focused on the distribution assumptions of the random component of the utility function to model correlations among alternatives within the choice set due to unobserved alternative attributes. Only a few works deal with the subject of nonlinear utility functions partly because of the difficulties arising in determining a priori the form of nonlinearity of the utility function. Box-Cox and Box-Tukey transformations (see i.e. Maier and Weiss 1990, pp. 126ff.) enable to model some limited forms of nonlinear utility functions.

Some research has been done in combining discrete choice models with artificial neural networks to model a nonparametric nonlinear utility function. Bentz and Merunka (Bentz and Merunka 2000) and Gelhausen (Gelhausen 2003) describe two different ways to represent a logit-model as an artificial neural network. This approach shows significant better empirical results than a standard logit-model with a linear utility function (Bentz and Merunka 2000; Hruschka et al. 2002; Probst 2002). An implementation of a nested logit-model with an arbitrary nesting structure is possible (Wilken and Gelhausen 2005, pp. 25ff.), but because of the complexity of the resulting network severe estimation and performance problems occur.

The Generalized Neural Logit-Model enables to model a nonparametric nonlinear utility function and arbitrary correlations among alternatives in the choice set due to unobserved attributes. Correlations among alternatives are modeled similar to the dogit-model. An efficient implementation of the Generalized Neural Logit-Model is possible as it is close to standard artificial neural networks.

The outline of this paper is as follows:

Chapter two explains the concept of alternative groups in discrete choice models as already introduced in Gelhausen (2006) and Gelhausen and Wilken (2006). This concept facilitates complexity reduction and the development of a model, which is applicable to alternatives outside the estimation data set.

Chapter three describes the theory of the Generalized Neural Logit-Model and its implementation as artificial neural network.

The Generalized Neural Logit-Model is applied to the case study of airport and access mode choice of air travelers in Germany in chapter four. A nested logit approach serves as a benchmark to evaluate the new model empirically (Gelhausen and Wilken 2006).

The paper ends with a summary and conclusion.

2. Grouping of Alternatives in Discrete Choice Models

The fundamental hypothesis of discrete choice models is the assumption of individual utility maximization. Alternatives are evaluated by means of a utility function and the one with the highest utility is supposed to be chosen. From an external point of view the utility of an alternative for a specific individual is a random variable, so that the utility U_i for alternative i is composed of a deterministic component V_i and a random component ε_i (Maier and Weiss 1990, p. 100):

$$(2.01) \quad U_i = V_i + \varepsilon_i$$

The random component of the utility function is introduced for various reasons, i.e. a lack of observability of the relevant alternative attributes or their incomplete measurability (Maier and Weiss, pp. 98f.).

As a result of the random component in the utility function only evidence in terms of the probability of an alternative being the one with the highest utility can be given from an external point of view. Specific discrete choice models differ in terms of their assumptions regarding the random component. The most prominent member of this class of models is the logit-model with independently and identically distributed random components. The choice probability of an alternative i is computed as (Train 2003, p. 40):

$$(2.02) \quad P(a_i = a_{\text{opt}}) = \frac{e^{\mu V_i}}{\sum_j e^{\mu V_j}}$$

As a consequence of the independently and identically distributed random components of the utility functions the ratio of two choice probabilities is only dependent on the utility of those two alternatives (Ben-Akiva and Lerman 1985, p. 108):

$$(2.03) \quad \frac{P(a_i = a_{\text{opt}})}{P(a_j = a_{\text{opt}})} = \frac{\frac{e^{\mu V_i}}{\sum_k e^{\mu V_k}}}{\frac{e^{\mu V_j}}{\sum_k e^{\mu V_k}}} = \frac{e^{\mu V_i}}{e^{\mu V_j}}$$

This property of the logit-model is called “Independence from Irrelevant Alternatives” (IIA) and it is both a weakness and a strength of the model. Due to the distribution assumptions of the random component of the utility function it is not possible to model correlations among alternatives owing to unobserved factors. A major advantage of the IIA-property is the possibility to estimate model parameters, excluding alternative-specific variables, on a subset of the alternatives (McFadden 1974, p. 113; McFadden 1978, pp. 87ff.; Ortuzar and Willumsen 2001, pp. 227f.; Train 2003, pp. 52f.) and the possibility of an evaluation of new alternatives without the need to re-estimate alternative-unspecific model parameters (Domencich and McFadden 1975, pp. 69f.). The problem of estimating alternative-specific variables from a subset of alternatives is discussed below.

The nested logit-model relaxes the IIA-restriction to some extent without losing the closed-form expression of the choice probabilities. For this purpose the random component in (2.01) is split up into a part ε_i^a , which varies over all alternatives i , and a part ε_k^c , which is identical for all alternatives of a nest k (Maier and Weiss 1990, pp. 154f.):

$$(2.04) \quad U_i = V_i + \varepsilon_i^a + \varepsilon_k^c$$

The nested logit approach enables to model correlations due to unobserved factors among subsets of the alternatives, so that the choice set is partitioned into clusters with highly correlated alternatives. (2.05) is an example of a covariance matrix consisting of four alternatives partitioned into two clusters with the first two belonging to cluster one and the last two assigned to cluster two.

$$(2.05) \quad \Omega = \begin{bmatrix} \sigma_{11}^2(\mu_1^c) & \sigma_{12}^2(\varepsilon_1^c) & 0 & 0 \\ \sigma_{21}^2(\varepsilon_1^c) & \sigma_{22}^2(\mu_1^c) & 0 & 0 \\ 0 & 0 & \sigma_{33}^2(\mu_2^c) & \sigma_{34}^2(\varepsilon_2^c) \\ 0 & 0 & \sigma_{43}^2(\varepsilon_2^c) & \sigma_{44}^2(\mu_2^c) \end{bmatrix}$$

Each cluster k is characterized by an individual scale parameter μ_k^c and an identical non-negative covariance for all alternatives i within a cluster k . Alternatives of different clusters are assumed not to be correlated.

For technical reasons the choice probabilities $P(a_i = a_{opt})$ are decomposed into an unconditional choice probability $P(c_k = c_{opt})$, that cluster k is chosen, and a conditional choice probability $P(a_i = a_{opt} \mid a_i \in c_k)$, that alternative i from cluster k is chosen (Maier and Weiss 1990, p. 156):

$$(2.06) \quad P(a_i = a_{opt}) = P(a_i = a_{opt} \mid a_i \in c_k) * P(c_k = c_{opt})$$

The conditional choice probabilities comply with the logit-model and the choice set is restricted to the alternatives of the appropriate nest. The choice probability of a nest k is determined by its maximum utility V_k^c (Maier and Weiss 1990, p. 157):

$$(2.07) \quad V_k^c = \frac{1}{\mu} \ln \sum_{i \in k} e^{\mu V_i}$$

The choice probability of an alternative i in nest k can be written as (Maier and Weiss 1990, p. 158):

$$(2.08) \quad P(a_i = a_{\text{opt}}) = \frac{e^{\mu V_i}}{\sum_{j \in k} e^{\mu V_j}} * \frac{e^{\mu_k^c V_k^c}}{\sum_l e^{\mu_l^c V_l^c}}$$

The hierarchical structure of (2.08) does not imply a sequential decision process. An extension to more than two levels is possible (see i.e. Ben-Akiva and Lerman 1985, pp. 291ff.).

In the nested logit-model the IIA-property holds only for two alternatives of the same cluster:

$$(2.09) \quad \begin{aligned} & \frac{P(a_1 = a_{\text{opt}} | a_1 \in c_1) * P(c_1 = c_{\text{opt}})}{P(a_2 = a_{\text{opt}} | a_2 \in c_1) * P(c_1 = c_{\text{opt}})} \\ &= \frac{\frac{e^{\mu V_1}}{\sum_{j \in c_1} e^{\mu V_j}} * \frac{e^{\mu_1^c V_1^c}}{\sum_l e^{\mu_l^c V_l^c}}}{\frac{e^{\mu V_2}}{\sum_{j \in c_1} e^{\mu V_j}} * \frac{e^{\mu_1^c V_1^c}}{\sum_l e^{\mu_l^c V_l^c}}} \\ &= \frac{e^{\mu V_1}}{e^{\mu V_2}} \end{aligned}$$

The ratio of the choice probabilities for two alternatives of different clusters depends on the characteristics of all alternatives of those two clusters:

$$\begin{aligned}
 & P(a_1 = a_{opt} | a_1 \in c_1) * P(c_1 = c_{opt}) / P(a_2 = a_{opt} | a_2 \in c_2) * P(c_2 = c_{opt}) \\
 &= \frac{e^{\mu V_1}}{\sum_{j \in c_1} e^{\mu V_j}} * \frac{e^{\mu_1^c V_1^c}}{\sum_I e^{\mu_1^c V_1^c}} \quad / \quad \frac{e^{\mu V_2}}{\sum_{j \in c_2} e^{\mu V_j}} * \frac{e^{\mu_2^c V_2^c}}{\sum_I e^{\mu_1^c V_1^c}} \\
 &= \frac{e^{\mu V_1}}{\sum_{j \in c_1} e^{\mu V_j}} \quad / \quad \frac{e^{\mu V_2}}{\sum_{j \in c_2} e^{\mu V_j}} * \frac{e^{\mu_1^c V_1^c}}{e^{\mu_2^c V_2^c}}
 \end{aligned}
 \tag{2.10}$$

As the nested logit-model lacks the IIA-property for some pairs of alternatives, model estimation on a subset of the choice set equal to the logit-model is not possible.

If it is feasible to form groups of at least approximately similar clusters and to assign an identical covariance matrix for all clusters of the same group, an estimation of alternative-unspecific model-parameters equal to the logit-model on a subset of alternatives is possible. Each group of clusters must be represented by at least one member in this subset to enable the estimation of all cluster-specific scale parameters. (2.11) shows a covariance-matrix for six alternatives belonging to three groups, with two alternatives per group. Figure 2.01 shows the relationship between a group and a cluster for this example.

$$\Omega = \begin{bmatrix} A & 0 & 0 & 0 & 0 & 0 \\ 0 & B & 0 & 0 & 0 & 0 \\ 0 & 0 & B & 0 & 0 & 0 \\ 0 & 0 & 0 & C & 0 & 0 \\ 0 & 0 & 0 & 0 & A & 0 \\ 0 & 0 & 0 & 0 & 0 & C \end{bmatrix}
 \tag{2.11}$$

The letters A, B and C represent the covariance structure of a cluster. Same letters indicate an equal covariance structure for different clusters. Figure 2.01 illustrates the assignment of clusters to groups.

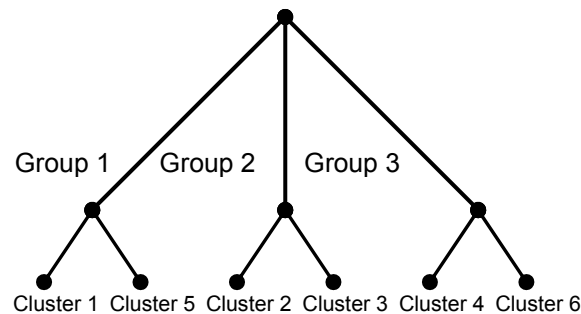


Figure 2.01: Dependence between Clusters and Groups

If identical alternative-specific model-parameters, especially alternative-specific constants, can be assumed reasonably well for different clusters of the same group, an estimation of all model-parameters is feasible on a subset of all alternatives as described above.

Applying the concept of grouping to the logit-model is possible, however, serves only to estimating alternative-specific variables, as there are no different scale parameters due to independently and identically distributed random components in the utility function.

The main advantage of this approach does not only lie in the reduction of computational costs for very large choice sets, as many econometric software packages limit the maximum number of clusters and alternatives for nested logit estimations, but primarily in a better way of developing a more generally applicable choice model beyond the alternatives of the estimation data set, i.e. in the context of scenario analysis.

A less popular member of discrete choice models is the dogit-model. Correlations among alternatives in the choice set are modeled by means of a functional

combination of the utility functions of each alternative with an alternative-specific parameter θ_i (Gaudry and Dagenais 1979, p. 105):

$$(2.12) \quad P(a_i = a_{opt}) = \frac{e^{\mu V_i} + \theta_i \sum_j e^{\mu V_j}}{\left(1 + \sum_j \theta_j\right) \sum_j e^{\mu V_j}}$$

Dogit- and logit-model are equal for all θ_i being zero so that the IIA-property holds for arbitrary pairs of alternatives. The vector of parameters θ describes to what extend the IIA-property does not hold.

In some empirical cases the dogit-model is superior to a logit approach in terms of model fit (see i.e. McCarthy 1997). However, the IIA-property does not hold in a systematic way in a genuine dogit-model with a nonzero vector θ , so that the aforementioned concept of alternative groups is not applicable.

3. The Generalized Neural Logit-Model

3.1 Theory of the Generalized Neural Logit-Model

The distribution assumptions regarding the random component of the utility function of the Generalized Neural Logit-Model are equal to the logit-model. Correlations among alternatives due to unobserved attributes are modeled by means of a combination of the utility functions of each alternative. This approach shows some similarities to the dogit-model, however, it offers more flexibility in terms of modeling correlations among alternatives.

An essential part of the Generalized Neural Logit-Model is a linear combination of the utility functions of each alternative:

$$(3.01) \quad V_i^{LK} = \sum_{j \in A_p^{LK}} \gamma_{ij} * V_j \quad i, j \in A_p^{LK}$$

with

γ_{ij} : Coefficient of the linear combination of the alternatives i and j

Alternatives within a subset A_p^{LK} of the complete choice set A^l are correlated. The correlation structure among alternatives is modeled by means of a hierarchy of utility functions. This approach shows some similarities to the nested logit-model. Due to the linear dependence between utility functions of different levels a two-stage hierarchy is sufficient. The choice probabilities are computed in the same way as in the logit-model with (3.01) being the utility function:

$$(3.02) \quad P(a_i = a_{opt}) = \frac{e^{\mu V_i^{LK}}}{\sum_j e^{\mu V_j^{LK}}}$$

The Generalized Neural Logit-Model belongs to the class of General Extreme Value-Models, so that utility maximizing behavior is modeled. The derivation of the model is identical to the logit-model with (3.01) being the utility function (Train 2003, p. 97f):

The function G is defined as:

$$(3.03) \quad G = \sum_j Y_j$$

with

$$(3.04) \quad Y_j = e^{\mu V_j^{LK}}$$

Inserting G and its first derivation G_i in

$$(3.05) \quad P(a_i = a_{opt}) = \frac{Y_i G_i}{G}$$

produces

$$(3.06) \quad P(a_i = a_{\text{opt}}) = \frac{Y_i}{\sum_j Y_j}$$

(3.06) equals (3.02).

The definition of the subsets A_p^{LK} depends on the correlations among the alternatives to be modeled. Four cases are distinguished:

- No correlations (logit-model)
- Correlations among all alternatives in the choice set
- Correlations among alternatives in disjoint clusters
- Limited correlations among all alternatives in the choice set

No correlations (logit-model)

Each subset A_p^{LK} equals exact one alternative and all coefficients γ_{ij} are set to a value of one. The IIA-property holds for arbitrary combinations of alternatives.

Correlations among all alternatives in the choice set

There is only one subset A_p^{LK} , which equals the complete choice set. The coefficients can take any values. The IIA-property does not hold for any combination of alternatives.

Correlations among alternatives in disjoint clusters

In this case alternatives are grouped in disjoint clusters A_p^{LK} similar to the nested logit approach to model arbitrary correlations among alternatives in each subset. The IIA-property does only hold on the cluster level. The aforementioned concept of alternative groups can be applied. Clusters of the same group have an identical

matrix of coefficients of linear combination instead of an identical covariance matrix:

$$(3.07) \quad \Omega^g = \begin{bmatrix} \gamma_{11}^g & \cdots & \gamma_{1j}^g \\ \vdots & \ddots & \vdots \\ \gamma_{i1}^g & \cdots & \gamma_{ij}^g \end{bmatrix}$$

with

γ_{ij}^g : Coefficient of linear combination of the alternatives i and j of a cluster of group g

Limited correlations among all alternatives in the choice set

In this case all alternatives may be correlated, so that the IIA-property does not hold for any pair of alternatives, but the coefficients of linear combination underlie a systematic structure, which enables model estimation on a subset of the complete choice set. It is possible to identify structural groups of alternatives and clusters according to the logit- and nested logit-model. Their definition is problem-dependent. In this study correlations among alternatives of the same cluster and between alternatives of different cluster groups are considered. Therefore this approach is a medium between case two and case three. The alternative subsets A_p^{LK} are composed of the cluster of the considered alternative and all clusters of different groups.

A group-dependent coefficient of linear combination γ_{im}^{klg} is assigned to every alternative i of the cluster k and alternative m of the cluster l . This coefficient is identical for two pairs of alternatives (a, b) and (c, d) , if (a, c) and (b, d) belong to different clusters of the same group.

The number of coefficients $\gamma_{im}^{kl\ g}$ equals the square of the number of alternatives on the lowest level of the cluster structure. Because of the identity of certain coefficients every cluster of group g has an identical matrix of structural coefficients of linear combination:

$$(3.08) \quad \Omega^g = \begin{bmatrix} o_{11}^{g1} & \dots & o_{1j}^{g1} & \dots & o_{11}^{gG} & \dots & o_{1M}^{gG} \\ \vdots & \ddots & \vdots & \dots & \vdots & \ddots & \vdots \\ o_{l1}^{g1} & \dots & o_{lj}^{g1} & \dots & o_{l1}^{gG} & \dots & o_{lM}^{gG} \end{bmatrix}$$

with

o_{im}^{gG} : Coefficient of linear combination of alternative i of a cluster of group g and alternative m of a cluster of group G

The assignment of elements of matrix (3.08) to the coefficients $\gamma_{im}^{kl\ g}$ results from the cluster groups:

$$(3.09) \quad o_{im}^{gG} = \gamma_{im}^{kl\ g} \Rightarrow k \in \text{group } g \wedge l \in \text{group } G$$

The coefficients o_{im}^{gG} and $\gamma_{im}^{kl\ g}$ receive the value zero in the case of an assignment of two different clusters to the same group, as this type of correlations among alternatives is not considered in this study. This fact is pointed out above in the definition of the subsets A_p^{LK} .

Matrix (3.07) is a special case of (3.08).

In the cases one, two and three formula (3.01) contains only non-equivalent alternatives. Two alternatives are equivalent, if they are on the same position in the cluster structure and their clusters belong to the same group. They have the same index m in (3.09).

In case four the above mentioned does not necessarily have to hold. By a normalization of the coefficients the value of (3.01) is according to the other three cases only dependent on the quality of non-equivalent linear combined alternatives and independent from the number of equivalent alternatives. For this purpose the coefficients O_{im}^{gG} and γ_{im}^{klg} have to be divided by the number of equivalent alternatives in (3.01) for model estimation and model application:

$$(3.10) \quad O_{im}^{gG\text{Norm}} = \frac{O_{im}^{gG}}{N_m^G}$$

with

N_m^G : Number of equivalent alternatives m of different cluster of the same group G in (3.01)

Model estimation on a subset of the total choice set by means of the IIA-property equal to the logit- and nested logit-model is not possible, as the IIA-property does hold neither on cluster- nor on alternative-level. However, identical coefficients being independent from the number of summands in (3.01) are assigned to certain alternatives as a result of the grouping of clusters and the normalization of coefficients. According to a random sample model estimation is possible in the case of every cluster group being represented at least with one member.

3.2 The Generalized Neural Logit-Model as Artificial Neural Network

The Generalized Neural Logit-Model represented as artificial neural network is composed of different modules, which have to be configured and put together problem-specific:

- Utility functions
- Linear combinations

- Logit-function

Utility function

For the problem considered a three-layer multilayer perceptron is sufficient for universal function approximation (Hornik et al. 1989, pp. 359ff.; Fausett 1994, p. 329).

The activation function of the input and output neurons constitutes the identical function. Hidden neurons have a tangens hyperbolicus function as activation function. The linear part of the utility function is described by means of the direct connections from the input to the output neurons marked in blue. This represents a linear perceptron in itself. The nonlinear part of the utility function is modeled by means of the connections between the neurons marked in black. The input neurons correspond to alternative attributes and the output neurons match the utility of an alternative. Figure 3.01 displays a nonlinear utility function as artificial neural network in an abstract way. The box in the upper part of the figure shows the type of activation function for the appropriate layer of the artificial neural network.

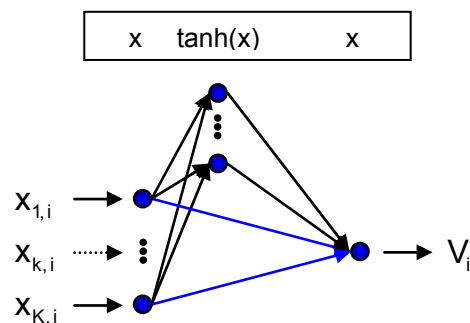


Figure 3.01: Nonlinear Utility Function as Multilayer Perceptron

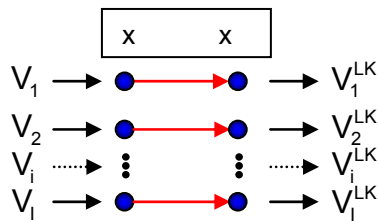
Linear Combinations

Linear combinations are modeled by means of a two-layer linear perceptron. The input and output neurons possess the identical function as activation function. The

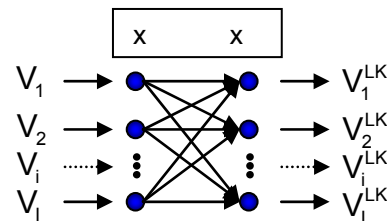
input neurons correspond to utility values of an alternative and the output neurons match the alternative-specific linear combinations of those utility values. The connection weights correspond to the coefficients of linear combination. Connections marked in red are constrained to a value of one.

Figure 3.02 displays the aforementioned four cases of correlation among the alternatives of the choice set as artificial neural network. A possible grouping structure of clusters is indicated by an appropriate highlighting of the connections in black and blue. As only one cluster per group is displayed this structure is easily identifiable. Therefore an additional indexing of the utility functions relating to the clusters is omitted for reasons of clearness.

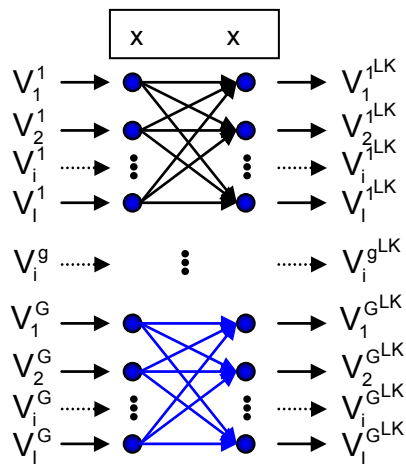
(a) Case 1



(b) Case 2



(c) Case 3



(d) Case 4

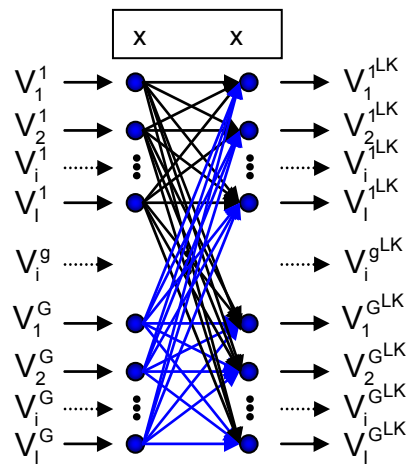


Figure 3.02: Linear Combinations as Artificial Neural Network

Logit-Function

The logit-function is modeled by means of a three-layer multilayer perceptron. The activation function is $f(x)=e^x$ for the input neurons and $f(x)=1/x$ for the hidden neurons. The output neurons possess the identical function as activation function with (3.11) being the net input function. Instead of the usual summation the inputs into a neuron are multiplied.

$$(3.11) \quad \text{net}_j = \prod_i w_{ij} * o_i$$

with

w_{ij} : Connection weight between a neuron of layer i and a neuron of layer j

o_i : Output of a neuron of layer i

This type of neuron is called “combiner neuron” (NeuroDimension 2005, pp. 276f.). The input neurons represent the linear combinations of the utility values and the output neurons the choice probabilities of the alternatives in the choice set. Figure 3.03 shows the logit-function represented as artificial neural network. Connections marked in red are constrained to a value of one.

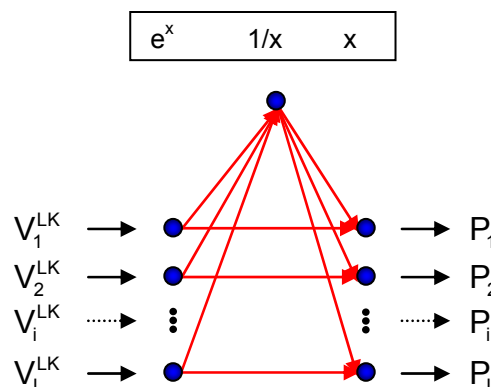


Figure 3.03: Logit-Function as Artificial Neural Network

Generalized Neural Logit-Model

Figures 3.04 to 3.07 display the Generalized Neural Logit-Model for all four aforementioned cases. For reasons of a concise implementation alternatives and clusters are grouped although this is not necessary for the cases one and two. The number of utility functions and output neurons equals the number of alternatives on the lowest level of the cluster structure. Identical utility functions relating to the connection structure and weights can be achieved by weight sharing (Bishop 2003, p. 349; LeCun et al. 1989, pp. 542ff.; Rumelhart et al. 1986, p. 349) in the case of all alternatives being evaluated by means of the same utility function. If endogeneity of exogenous factors due to unobserved alternative attributes is present (Bhat 2003, pp. 16f.), a dependence of the utility function on the considered alternative is possible as alternative attributes are evaluated differently dependent on the relevant alternative.

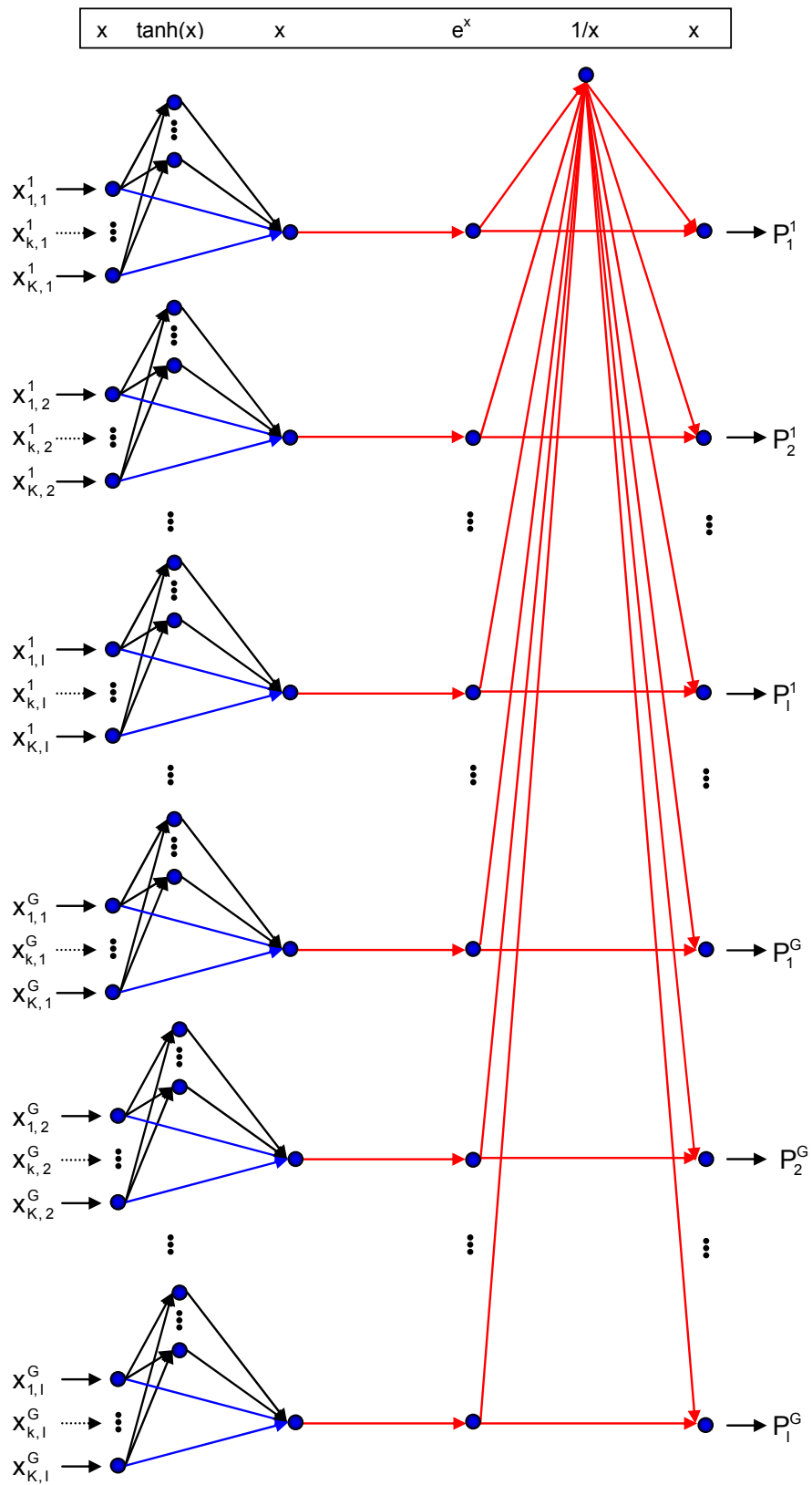


Figure 3.04: Generalized Neural Logit-Model for Case 1

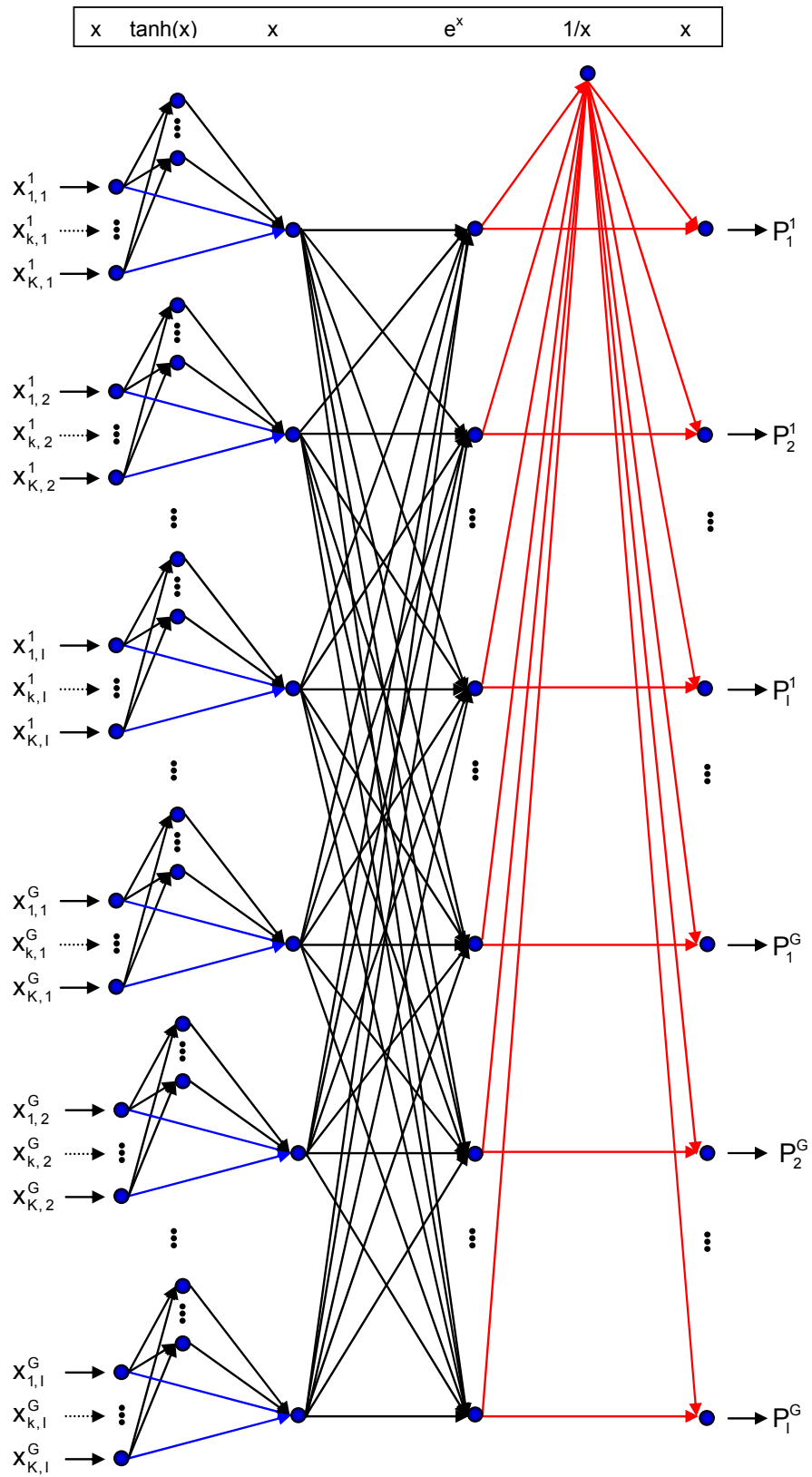


Figure 3.05: Generalized Neural Logit-Model for Case 2

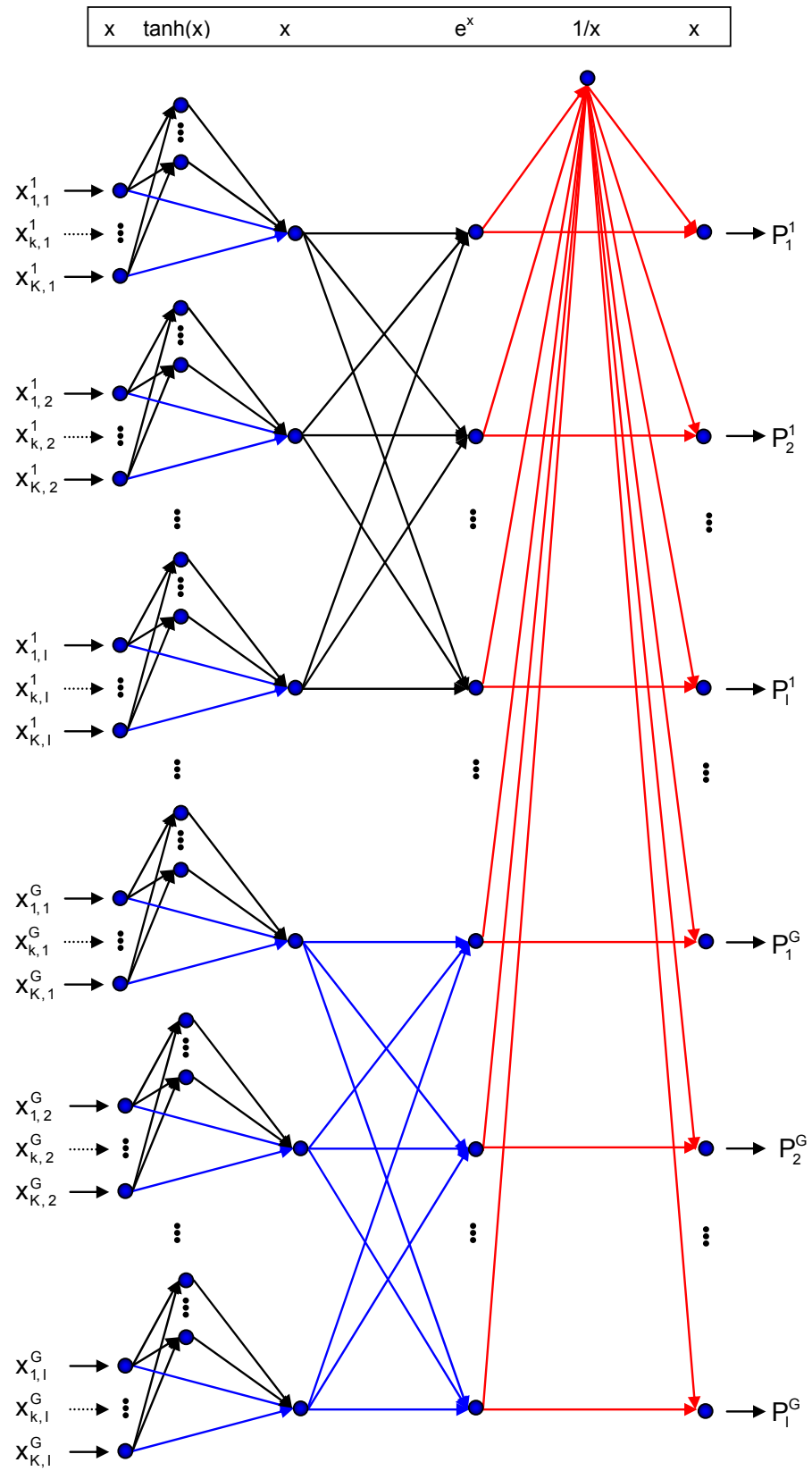


Figure 3.06: Generalized Neural Logit-Model for Case 3

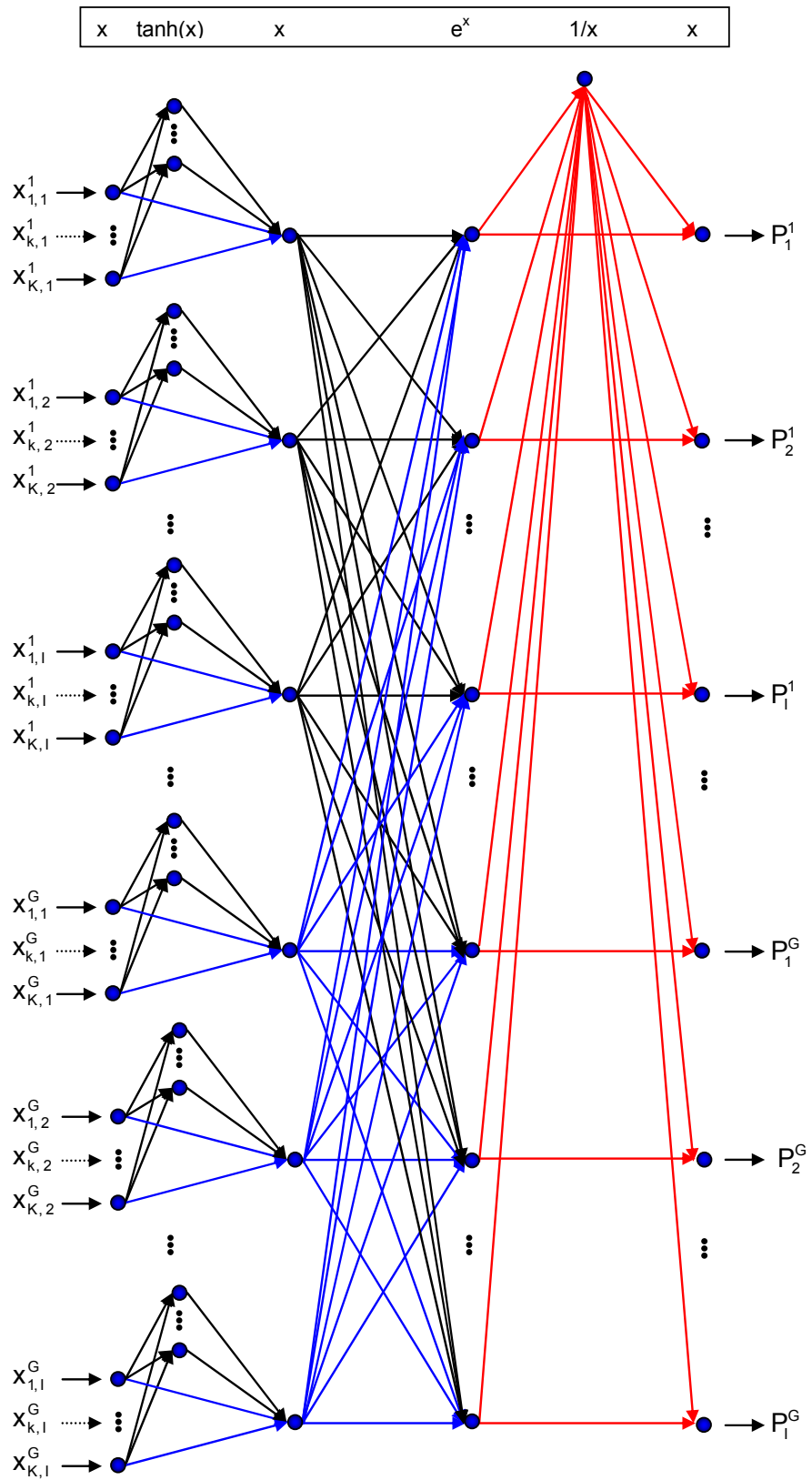


Figure 3.07: Generalized Neural Logit-Model for Case 4

4. Case Study: Airport and Access Mode Choice in Germany

4.1 Introduction

The Generalized Neural Logit-Model is applied to the case study of airport and access mode choice of air travelers in Germany. A nested logit approach with a linear utility function serves as a benchmark to evaluate the new model empirically in terms of model fit. The available database, airport categories, model definition and model estimation of the nested logit-model is discussed in great detail in Gelhausen and Wilken (2006, pp. 10ff.). Only some fundamental facts are introduced briefly below. A full discussion of these issues would be beyond the scope of this paper. Purpose of this chapter is to discuss both approaches concerning model quality and to present some conclusions relating to air traveler's choice behavior in airport and access mode choice.

Table 4.01 shows the full alternative set of the database (Gelhausen and Wilken 2006, p. 11). Only the access mode "car" includes parking at the airport for the duration of the journey. For "kiss and ride" the number of trips is doubled compared to all other access modes as the car is parked at the trip origin. The "taxi" alternative includes taxis and private bus services operating on demand only. The access mode "bus" contains scheduled public-transit buses. "urban railway" and "train" are distinguished in terms of the tariff paid. If the tariff of the Deutsche Bahn applies, it is a train; otherwise it is an urban railway (Gelhausen and Wilken 2006, p. 11).

	Car	Kiss and Ride	Rental Car	Taxi	Bus	Urban Railway	Train
Berlin	x	x	x	x	x	x	
Bremen	x	x	x	x		x	
Dortmund	x	x	x	x	x		
Dresden	x	x	x	x	x	x	
Düsseldorf	x	x	x	x	x	x	x
Erfurt	x	x	x	x	x		
Frankfurt a. M.	x	x	x	x	x	x	x
Frankfurt Hahn	x	x	x	x	x		
Friedrichshafen	x	x	x	x	x	x	
Hamburg	x	x	x	x	x		
Hannover	x	x	x	x	x	x	
Karlsruhe-Baden	x	x	x	x	x		
Köln/Bonn	x	x	x	x	x		
Leipzig/Halle	x	x	x	x	x		x
Lübeck	x	x	x	x	x		
München	x	x	x	x	x	x	
Münster/Osnabrück	x	x	x	x	x		
Niederrhein	x	x	x	x	x		
Nürnberg	x	x	x	x	x	x	
Paderborn/Lippstadt	x	x	x	x	x		
Saarbrücken	x	x	x	x	x		
Stuttgart	x	x	x	x	x	x	

Table 4.01: Airports and Available Access Modes

According to the length and purpose of a journey different market segments are defined (Gelhausen and Wilken 2006, pp. 10f.):

- Journeys to domestic destinations, subdivided into private (BRD P) and business (BRD B) trip purpose
- Journeys to European destinations for business trip purpose (EUR B)
- Journeys to European destinations for private short stay reasons up to four days (EUR S)
- Journeys to European destinations for holiday reasons for five days or longer (EUR H)
- Journeys to intercontinental destinations, subdivided into private (INT P) and business (INT B) trip purpose

Table 4.02 displays the chosen alternative attributes and their definitions (Gelhausen and Wilken 2006, p. 12).

Variable (Abbreviation)	Definition
Access Cost (COST)	Cost in € per Person incl. Parking Fees, Double Trip Length
Access Time (TIME)	Time in Minutes, Double Trip Length
Waiting Time (WAIT)	Inverse of the Daily Frequency
Inverse of the Population Density (INVPD)	Inverse of Residents per km ²
Inverse of the Competition on a Direct Flight Connection (COMP)	Inverse of the Number of Alliances and Independent Airlines
Quality of Terminal Access (AAS)	binary (good/bad)
Existence of a Direct Flight Connection (DIRECT)	binary (good/bad)
Frequency of a Direct Flight Connection (DFREQ)	Number Flights per week
Existence of a Low-Cost Connection (LC)	binary (yes/no)
Frequency of a Low-Cost Connection (LCFREQ)	Number Low-Cost Flights per week
Existence of a Charter Flight Connection (CC)	binary (yes/no)
Frequency of a Charter Flight Connection (CCFREQ)	Number Charter Flights per week

Table 4.02: Definition of Alternative Attributes

Airports are categorized from a demand-oriented point of view by means of a clustering technique based on artificial neural networks called “Kohonen’s Self-Organizing Maps” (see i.e. Kohonen 2001, pp. 109ff.) to form groups of clusters composed of one airport category and all access modes (Gelhausen and Wilken 2006, pp. 14ff.). Table 4.03 shows the relevant attributes for distinguishing airport categories (Gelhausen and Wilken 2006, p. 12).

Attributes (Abbreviation)	Definition
Number of Domestic Low-Cost Flights (LCBRD)	Flights per Week
Number of Domestic Charter Flights (CCBRD)	Flights per Week
Number of Domestic Full Service Flights (FSBRD)	Flights per Week
Number of European Low-Cost Flights (LCEUR)	Flights per Week
Number of European Charter Flights (CCEUR)	Flights per Week
Number of European Full Service Flights (FSEUR)	Flights per Week
Number of Intercontinental Low-Cost Flights (LCINT)	Flights per Week
Number of Intercontinental Charter Flights (CCINT)	Flights per Week
Number of Intercontinental Full Service Flights (FSINT)	Flights per Week
Number of Domestic Destinations (NUMBRD)	Number of Destinations
Number of European Destinations (NUMEUR)	Number of Destinations
Number of Intercontinental Destinations (NUMINT)	Number of Destinations

Table 4.03: Attributes for Airport Categorization

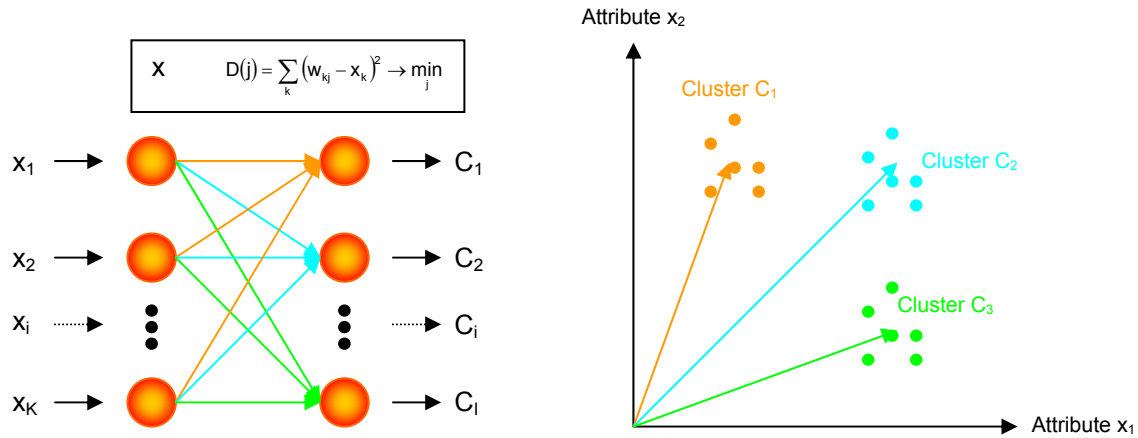


Figure 4.01: Self-Organizing Map

Figure 4.01 is a schematic illustration of a Self-Organizing Map. The neurons are simple computational units connected by weighted edges. Computations in a neuron are performed according to a simple transfer function. Input neurons correspond to clustering attributes and output neurons represent the clusters. The transfer function of the input neurons is the identical function $f(x) = x$. The output neurons have a “winner-takes-all” transfer function. The neuron with the smallest distance between the input vector and its synaptic weight vector wins the competition and is activated. During learning of the self-organizing map the synaptic weight vector of the output neurons approach the corresponding cluster centroid as the right part of figure 4.01 illustrates.

Table 4.04 shows the parameters for optimal cluster identification. The self-organizing map was not highly sensitive with regard to parameter variations.

Parameter	Value
Topology of output neurons	Linear
Measure of distance	Euclidean
Neighborhood function	linear: $2 - 0.002 * \text{Iteration}$
Learning rate	0.01
Number of iterations	10 000
Data normalization	yes, [-1; 1]
Number of input neurons	12
Number of output neurons	3

Table 4.04: Parameters of a Self-Organizing Map for Airport Categorization

Three airport categories have been identified. The output neurons are arranged in a linear grid and distances between input vectors and output neurons are measured Euclidean. A linear neighborhood function is used and the neighborhood contains all output neurons at the beginning of the learning process. It shrinks to 0 within 1 000 iterations. The number of learning iterations is 10 000 and the learning rate is chosen rather small with 0.01. Each element of the input vector is normalized to the interval $[-1; 1]$.

Table 4.05 shows the synaptic weights for the trained self-organizing map. The color of the columns equals the color of the synaptic weights in figure 5.2.

Attributes	Airport		
	Category 1	Category 2	Category 3
LCBRD	0.054281	0.026181	-0.973566
CCBRD	0.63343	-0.23698	-0.902359
FSBRD	0.820399	-0.16164	-0.810737
LCEUR	-0.814996	-0.248973	-0.717447
CCEUR	0.673964	0.145995	-0.811895
FSEUR	0.767974	-0.596754	-0.967617
LCINT	-0.999997	-0.507511	-0.862715
CCINT	0.459986	-0.679604	-0.986041
FSINT	0.128171	-0.975403	-0.999997
NUMBRD	0.810002	0.570222	-0.409338
NUMEUR	0.791409	-0.012681	-0.737397
NUMINT	0.314031	-0.817745	-0.991489

Table 4.05: Cluster Centroids of Airport Categories

Table 4.06 displays the airports of the German Air Traveler Survey (Berster et al. 2005), which was used as a database for model estimation, and the appropriate category for each of those airports.

Category	Airport (IATA-Code)
AP 1	Frankfurt a. M. (FRA)
AP 1	München (MUC)
AP 2	Düsseldorf (DUS)
AP 2	Hamburg (HAM)
AP 2	Köln/Bonn (CGN)
AP 2	Stuttgart (STR)
AP 3	Bremen (BRE)
AP 3	Dortmund (DTM)
AP 3	Dresden (DRS)
AP 3	Erfurt (ERF)
AP 3	Frankfurt Hahn (HHN)
AP 3	Friedrichshafen (FDH)
AP 3	Hannover (HAJ)
AP 3	Karlsruhe/Baden (FKB)
AP 3	Leipzig/Halle (LEJ)
AP 3	Lübeck (LBC)
AP 3	Münster/Osnabrück (FMO)
AP 3	Niederrhein (NRN)
AP 3	Nürnberg (NUE)
AP 3	Paderborn/Lippstadt (PAD)
AP 3	Saarbrücken (SCN)

Table 4.06: Assignment of Airports to Categories

Tables 4.07 and 4.08 show the properties of the identified three airport categories both in percentages and in absolute values (Gelhausen and Wilken 2006, p. 17).

	LCBRD	CCBRD	FSBRD	LCEUR	CCEUR	FSEUR	LCINT	CCINT	FSINT	NUMBRD	NUMEUR	NUMINT
AP 1	3.18	0.43	20.39	0.87	5.83	55.81	0.00	1.24	12.25	8.31	60.27	31.42
AP 2	8.97	0.58	28.27	11.65	11.76	37.24	0.02	0.71	0.79	16.23	74.62	9.16
AP 3	1.29	0.86	39.22	32.57	15.57	10.05	0.02	0.42	0.00	19.94	78.90	1.16

Table 4.07: Properties of Airport Categories (in %)

	LCBRD	CCBRD	FSBRD	LCEUR	CCEUR	FSEUR	LCINT	CCINT	FSINT	NUMBRD	NUMEUR	NUMINT
AP 1	106	16	756	32	225	2138	0	49	517	19	144	83
AP 2	104	7	348	129	153	487	0	11	11	17	80	12
AP 3	3	1	80	47	25	39	0	0	0	6	22	1

Table 4.08: Properties of Airport Categories (absolute)

Figure 4.02 illustrates the nesting structure of each cluster group composed of one airport category and all access modes (Gelhausen and Wilken 2006, p. 20).

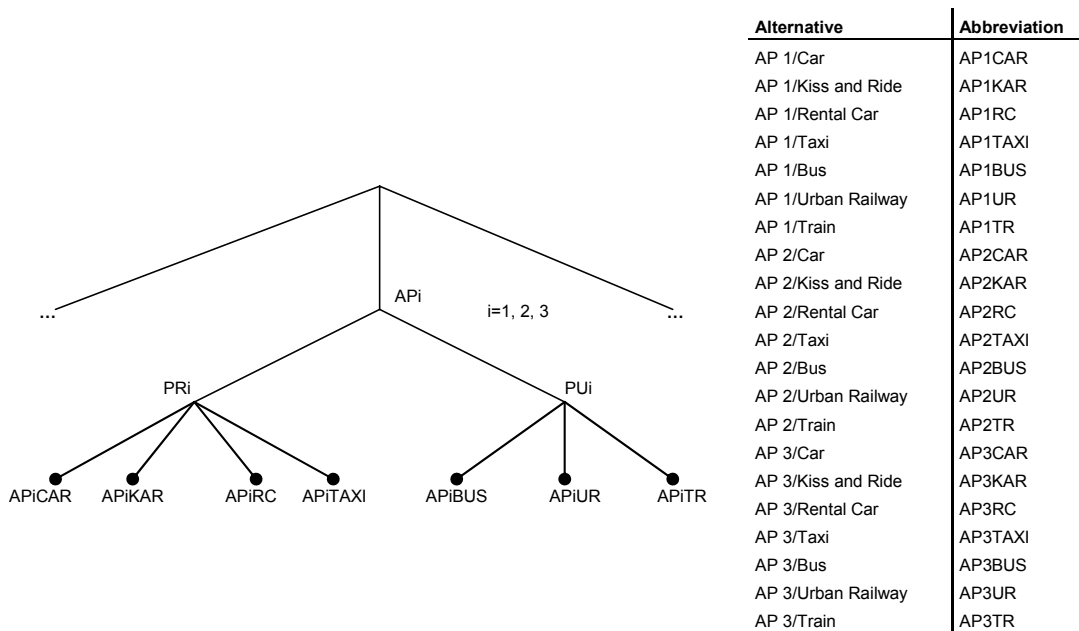


Figure 4.02: Nesting Structure

For model estimation the data set is partitioned into several disjoint data subsets. Each data subset contains only a subset of the full set of airport-access mode alternatives, namely one airport of each category and its access modes. Each data subset includes observations of individuals, who have chosen one of the alternatives of the reduced alternative set. By a suitable definition of data subsets, it is possible to estimate a model with the full set of seven access modes for all three airport categories. For this purpose, the inclusion of the airports Frankfurt a. M., Düsseldorf and Leipzig/Halle is necessary, as these are the only airports of their category with an airport access via train in 2003. The individual data subsets are merged into a single new estimation data set. The number of alternatives is reduced from 122 to 21. By weighting each observation the estimation data set is statistically representative. Figure 4.03 shows the definition of the data subsets. The nearest airport of each category is assigned to each data set marked in different colors. Every subset is named according to its airport of the third category (Gelhausen and Wilken 2006, p. 18).

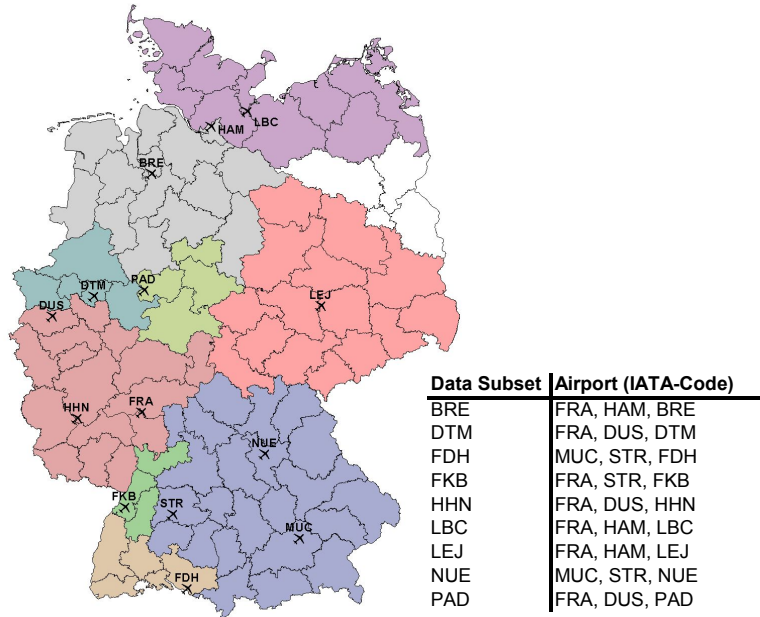


Figure 4.03: Data Subsets and Assignment of Airports

After selecting the airports and access modes for the specific application case, they are assigned to categories with the appropriate model parameters. Model application is possible to any number of airports and arbitrary airport/access mode combinations as a result of the clusters groups. Figure 4.04 summarizes the general process of model estimation and its application (Gelhausen and Wilken 2006, pp. 18f.).

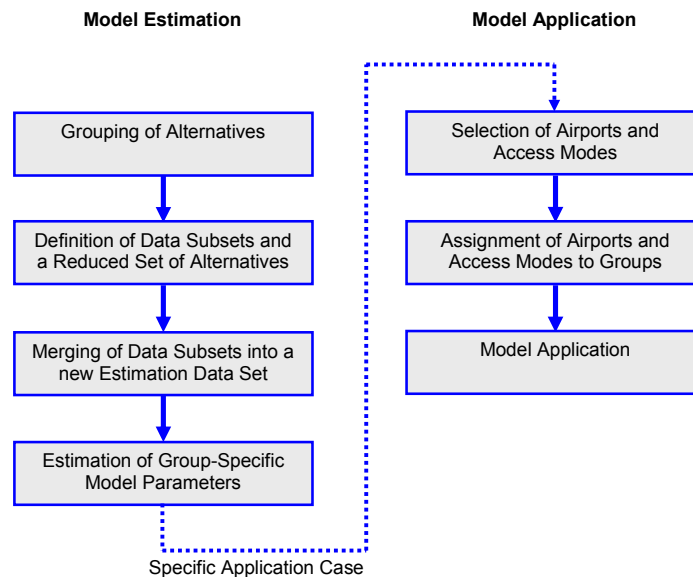


Figure 4.04: Estimation and Application of Airport and Access Mode Choice Model

The deterministic part of the utility function of the nested logit-model is of linear form (Gelhausen and Wilken 2006, p. 19):

$$(4.01) \quad V_i = \text{alt}_i + \sum_k b_k * x_{k,i}$$

with

alt_i : Alternative-specific constant of alternative i

b_k : Coefficient of attribute k

$x_{k,i}$: Value of attribute k for alternative i

Table 4.09 displays the estimated coefficients of the alternative attributes, scale parameters, goodness-of-fit measures and the likelihood-ratio test statistics for all seven market segments (Gelhausen and Wilken 2006, p. 28).

Variable	BRD P	BRD B	EUR S	EUR H	EUR B	INT P	INT B
COST	-0.0263035	-0.0204609	-0.0199987	-0.0173617	-0.0216885	-0.0138527	-0.00936472
TIME	-0.0081889	-0.0152572	-0.0061063	-0.00857067	-0.00795957	-0.00541014	-0.00535887
WAIT	-28.8061	-18.935	-8.33078	-4.40982	-9.94709	-18.7546	-35.7591
INVPD	-187.86	-21.8829	-215.876	-235.641	x	-25.6109	-32.2589
COMP	-0.158635	x	-1.22176	-1.13258	-0.182127	x	x
AAS	0.920627	1.12781	0.20336	0.46823	0.504623	0.840462	0.382595
DIRECT	2.29637	3.64119	3.63327	3.31697	1.43564	1.85847	0.439344
DFREQ	0.00682913	0.00601159	0.0104684	0.0153856	0.0177437	x	x
LC	x	x	0.0863075	0.563633	0.275153	x	x
LCFREQ	x	x	0.0631856	x	0.0761092	x	x
PR1	1.07092	1.02375	0.764486	0.61189	0.808397	1.13266	1.03073
PU1	0.745385	0.978059	0.593257	0.3847	0.386155	0.983045	0.32899
PR2	0.492518	1.00829	0.767123	0.570138	0.783306	1.06067	1.3532
PU2	0.390636	0.992109	0.543582	0.437515	0.708662	0.927296	0.832438
PR3	0.817955	1.00988	0.821821	0.610065	0.937914	0.813943	0.91783
PU3	0.428619	0.999286	0.395656	0.551239	0.805435	0.137029	0.718249
AP1	1.81029	1.01119	1.80601	1.65075	1.61072	1.10489	2.10553
AP2	2.10174	1.00887	1.76862	1.92646	1.67197	1.19742	1.16102
AP3	2.35248	1.01164	1.74828	1.99236	1.77295	1.23031	1.73837
pseudo-R ² (null) in %	57.41	54.10	52.40	52.29	48.58	48.89	47.46
pseudo-R ² (const) in %	43.82	40.47	41.94	38.22	35.96	32.86	28.30
LR (MNL)	82414	8740	43774	349740	311756	599974	131576
$\alpha=0.5\%$	25.19	23.59	23.59	23.59	23.59	23.59	23.59

Table 4.09: Overview Estimation Results per Market Segment

Scale parameters are normalized to a value of one on the lowest level of the nesting structure. For the alternative-specific constants, p- and t-values and the

standard deviation of the estimated coefficients see Gelhausen and Wilken (Gelhausen and Wilken 2006, pp. 21ff.).

4.2 Generalized Neural Logit-Model

For model estimation, the same data set and cluster group structure is used as for the nested logit-model. The subdivision of access modes into private and public modes of travel is omitted as a two-stage hierarchy is sufficient. Case four is chosen in terms of correlation among alternatives as this approach is more flexible than a nested logit-model yet it enables the development of a model, which is applicable to alternatives outside the estimation data set. The selection of explanatory variables is based on the nested logit-model because of the possibility of statistical significance tests and simple plausibility checks. To consider endogeneity of exogenous factors no weight sharing is applied. Figure 4.05 exemplifies the structure of the Generalized Neural Logit-Model for the case of private journeys to domestic destinations (BRD P).

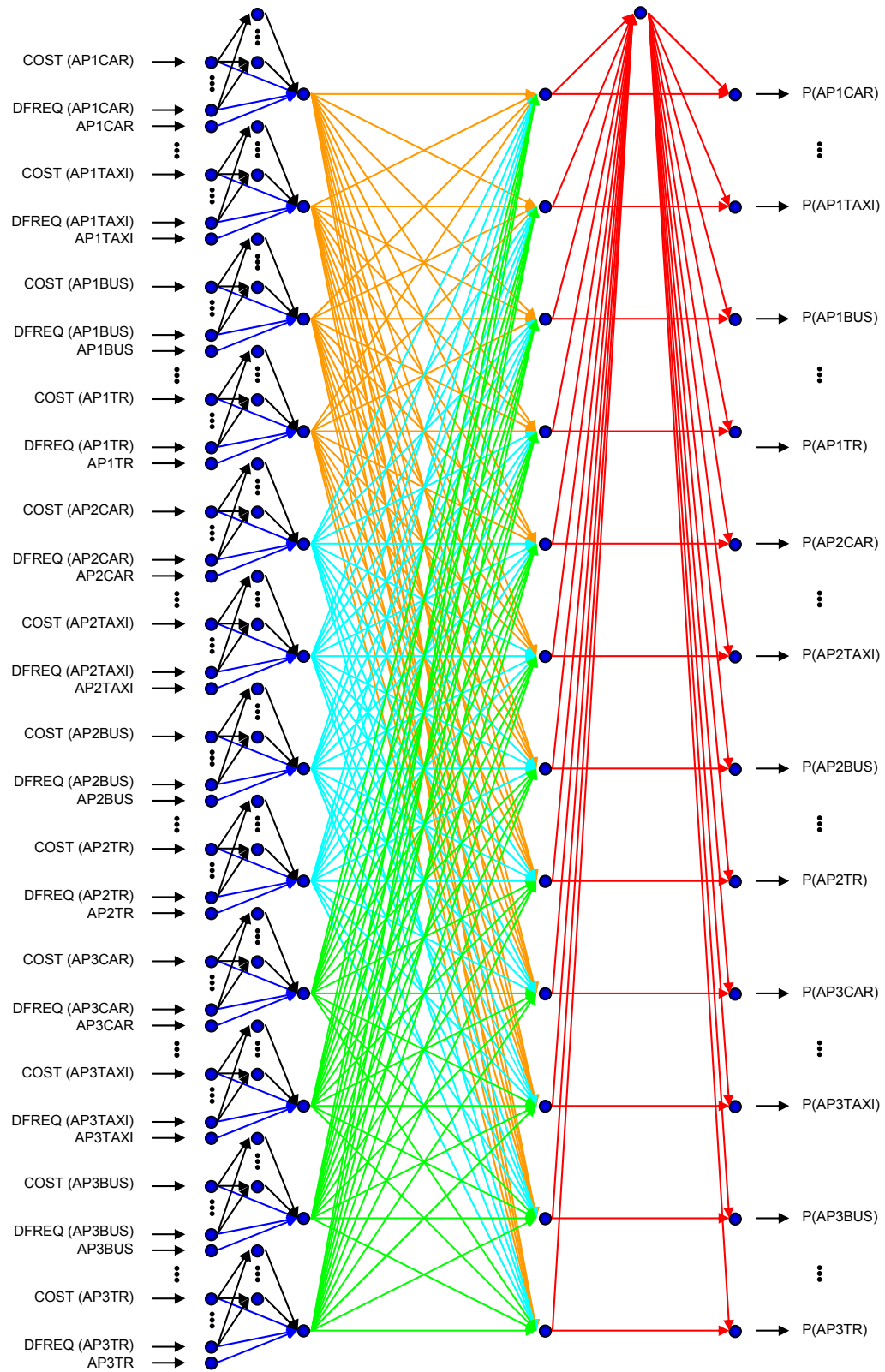


Figure 4.05: Structure of the Generalized Neural Logit-Model for the Market Segment BRD P

The method of network structure specification follows the idea of Miller, Todd and Hedge (Miller et al. 1989). The estimation data set is split up into a training set and a cross-validation set with a share of 85% and 15% respectively. Different network topologies, which are generated by a genetic search, are trained on the training set and evaluated on the cross-validation set in terms of their ability to generalize. The best network topology in terms of a minimal cross-validation error the network is trained on the entire estimation data set without an early stopping of training, so that a maximum of information is available for the final estimation of the connection weights. This ensures a maximum of statistical efficiency while the ability of an artificial neural network to generalize does not decline in the case of an appropriate network structure (Anders 1997, pp. 116ff.). The method of least squares is employed for the estimation of connection weights with conjugate gradients being the numerical optimization method. Input variables are scaled on the interval [-1; 1]. Table 4.10 summarizes the training parameters. Because of computational costs the population size is chosen small, but an optimal network topology is found within ten generations.

Estimation Method	Least Squares
Optimization Method	Conjugate Gradients
Scaling	yes, [-1; 1]
Genetic Search	
Share of Cross-Validation	15%
Population Size	10
Selection Rule	Roulette
Cross-over	Multi-Point
P(Cross-over)	0.9
P(Mutation)	0.01
Coding	Direct Encoding

Table 4.10: Training Parameters

Model quality is measured in terms of model fit and assessed by means of the pseudo- R^2 . Benchmark is a model without any variables (R^2_{null}) and a market share model (R^2_{const}). Table 4.11 illustrates the pseudo- R^2 by market segment for the Generalized Neural Logit-Model (GNL) and the nested logit approach (NL).

Market Segment	R2(null) in %			R2(const) in %		
	NL	GNL	Diff. to NL	NL	GNL	Diff. to NL
BRD P	57.41	61.35	3.94	43.82	49.74	5.92
BRD B	54.10	58.13	4.03	40.47	47.16	6.69
EUR S	52.40	58.09	5.69	41.94	49.99	8.05
EUR H	52.29	56.51	4.22	38.22	45.10	6.88
EUR B	48.58	51.96	3.38	35.96	41.79	5.83
INT P	48.89	55.10	6.21	32.86	42.01	9.15
INT B	47.46	56.01	8.55	28.30	41.26	12.96

Table 4.11: Comparison of Model Fit

The Generalized Neural Logit-Model shows especially for the market segments of intercontinental journeys for both private and business purpose a clear increase in model fit compared to the nested logit approach. For example, the increase of R2(const) is about 45% for the market segment INT B compared to the nested logit-model. The pseudo-R² is more evenly distributed over the market segments and lies between 41% and 49% in the case of R2(const). This corresponds to an R² of linear regression of 82% and 92% (Domencich et al. 1975, p. 124). Table 4.12 contrasts relative alternative frequencies and computed choice probabilities for the market segment EUR S.

Alternative	Relative	NL-Model		GNL-Model	
	Frequency		abs. Diff.		abs. Diff.
AP1CAR	4.7068	5.4012	0.69	5.2468	0.54
AP1KR	9.1049	10.0309	0.93	9.4534	0.35
AP1RC	0.1929	0.1543	0.04	0.2043	0.01
AP1TAXI	2.7006	2.7392	0.04	2.7109	0.01
AP1BUS	0.5401	0.5015	0.04	0.4558	0.08
AP1UR	5.5170	6.4429	0.93	6.2622	0.75
AP1TR	1.3503	1.7747	0.42	2.3327	0.98
AP2CAR	9.9537	9.3364	0.62	9.9918	0.04
AP2KR	16.4738	14.1975	2.28	15.1639	1.31
AP2RC	0.3472	0.1929	0.15	0.2491	0.10
AP2TAXI	6.5201	6.5972	0.08	6.6302	0.11
AP2BUS	2.1219	2.6620	0.54	2.3235	0.20
AP2UR	4.2438	5.3627	1.12	4.2061	0.04
AP2TR	4.4753	3.7809	0.69	3.8147	0.66
AP3CAR	16.3580	16.2423	0.12	17.5289	1.17
AP3KR	11.0725	9.7994	1.27	9.6736	1.40
AP3RC	0.1157	0.0772	0.04	0.1048	0.01
AP3TAXI	2.4306	2.7392	0.31	2.1798	0.25
AP3BUS	1.2731	1.3889	0.12	0.9809	0.29
AP3UR	0.4244	0.5401	0.12	0.4718	0.05
AP3TR	0.0772	0.0386	0.04	0.0149	0.06
E(abs. Diff.)			0.50		0.40
σ (abs. Diff.)			0.55		0.45

Table 4.12: Relative Alternative Frequencies and Computed Choice Probabilities for EUR S

Airport and access mode choice behavior of air travelers is governed by a complex nonlinear utility function and correlations among alternatives beyond the capabilities of a nested logit approach with a linear utility function as the clear increase in model fit demonstrates. Figures 4.06 and 4.07 illustrate the dependencies between two selected alternative attributes and the choice probability of a specific alternative. There is a clear nonlinear relationship between access time, access cost and the choice probability in domestic air travel for business purpose. These travelers are very access time-sensitive. This relationship is of more linear form with a greater importance of access cost in the market segment of intercontinental air travel for private reasons.

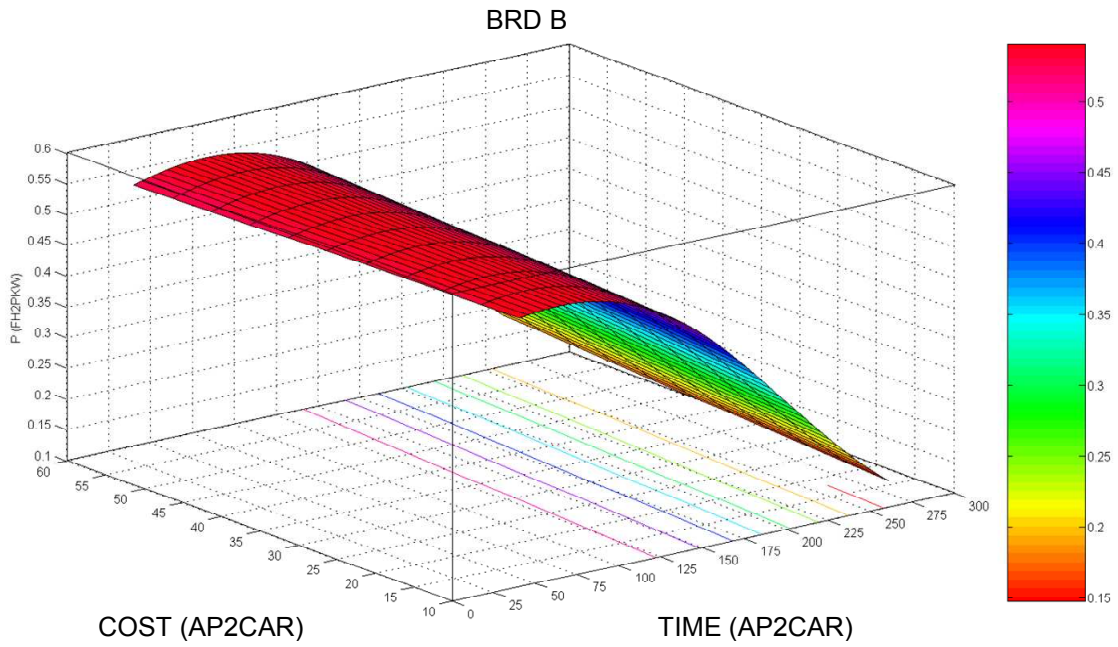


Figure 4.06: Analysis of the Utility Function for BRD B

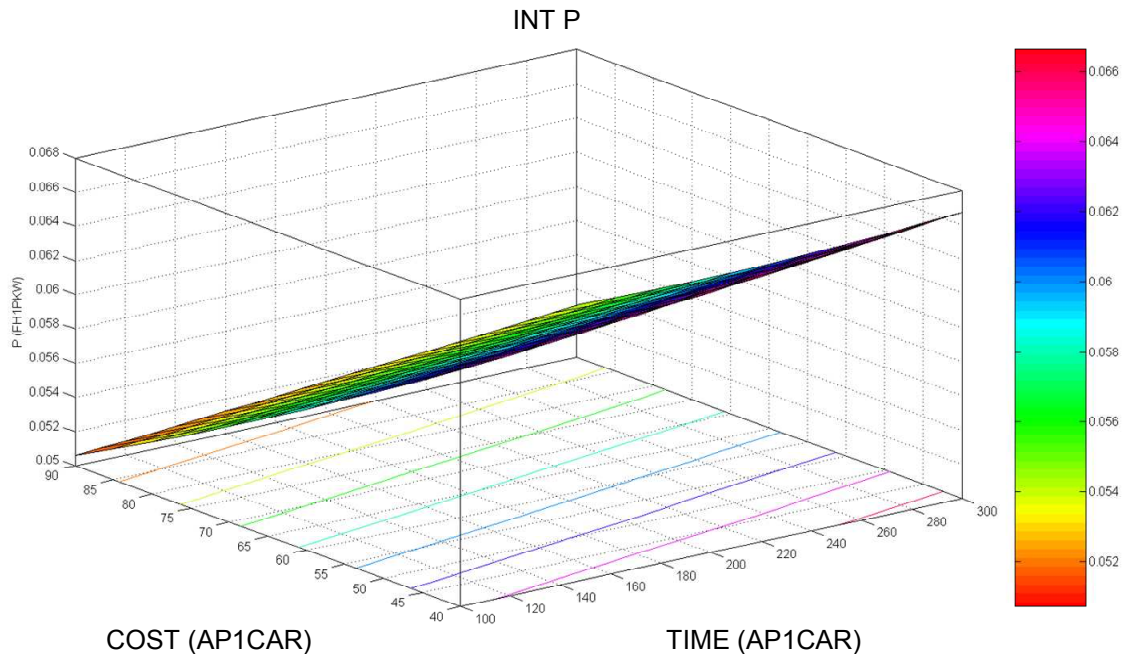


Figure 4.07: Analysis of the Utility Function for INT P

5. Summary and Conclusions

This paper presents a novel approach in discrete choice modeling called “Generalized Neural Logit-Model”. This approach is based upon the General Extreme Value-framework and is implemented as artificial neural network. Its main advantages lie in a nonparametric nonlinear utility function and the capability to model arbitrary correlations among alternatives in the choice set.

The first part of this paper deals with the concept of alternatives and cluster groups. It enables the development of discrete choice models applicable to alternatives outside the estimation data set.

The next chapter introduces the Generalized Neural Logit-Model. The first part is about the theoretical framework followed by an implementation as artificial neural network.

The Generalized Neural Logit-Model is evaluated empirically by means of a case study. The case of airport and access mode choice in Germany is chosen and a nested logit approach based on the concept of cluster groups serves as a benchmark. To form cluster groups airports are categorized from a demand-oriented point of view by means of a clustering technique based on artificial neural networks called “Kohonen’s Self-Organizing Maps”.

The Generalized Neural Logit-Model is superior to the nested logit approach in terms of model fit as the considered problem is governed by a complex nonlinear utility function and correlations among alternatives beyond the capabilities of a nested logit approach with a linear utility function. The pseudo- R^2 based on a market share model as a benchmark lies within the range of 41% to 49% and is up to 45% above the nested logit approach depending on the market segment. This corresponds to an R^2 of linear regression of 82% to 92%, so that a model of very good quality can be obtained for all market segments.

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