

## Foreign Ownership, Exporting and Gender Wage Gaps: Evidence from Japanese Linked Employer-Employee Data

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# Foreign Ownership, Exporting and Gender Wage Gaps: Evidence from Japanese Linked Employer-Employee Data

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#### Abstract

We explore potential relationships between international economic activities and gender wage gaps (GWGs) using linked employer-employee data for Japan. We find evidence that exporting and multinational activities are associated with reduced GWGs. Domestic-owned firms that neither export nor invest abroad (i.e., domestic-only firms) report the largest GWG, followed by Japanese-owned multinational enterprises (JMNE), then by locally-owned exporters that do not invest abroad and finally by foreign-owned multinational enterprises (FMNE). We separate FMNE by mode of entry and confirm that FMNE established by greenfield investment deviate more than FMNE established by merger and acquisition from domestic-only firms in terms of wages. Greenfield-born FMNE are associated with the smallest GWG and largest gender-neutral wage premium among the firm types. The estimated GWG among Greenfield-born FMNE is almost 12 percentage-points lower than the 26.8 percent prevailing at domestic-only firms.

*Keywords*: gender wage gap; wage premium; exporters; multinational enterprises *JEL Classification*: F14, F16, J31

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#### 1 Introduction

"Womenomics", a concept first articulated by Kathy Matsui, vice chair of Goldman Sachs Japan, and adopted by Japan's Prime Minister Shinzo Abe, proposes increasing female employment and reducing gender pay disparities in order to achieve economic growth.<sup>1</sup> While Japan has improved its female labor force participation rates and now reports a rate that exceeds the OECD average by over 5 percentage points,<sup>2</sup> it has not been as successful in reducing genderbased earnings inequality. Japan's raw gender earnings gap of 24.5 percent was much higher than the OECD average of 13.2 percent in 2017.3 Only South Korea (34.6 percent) and Estonia (28.3 percent) reported higher gender earnings gaps than Japan.<sup>4</sup>

International activities, both trade and foreign direct investment (FDI), have the potential for impacting a country's gender earnings gap. We seek to explore how Japan's gender pay gap may be affected by exporting and multinational activities. Exposure to foreign markets and foreign gender norms through exporting and FDI outflows and inflows has the potential to impact male and female workers differently. Using linked employer-employee data for the manufacturing sector, we find suggestive evidence that international activities contribute to reducing Japan's gender wage gap (GWG). We find that foreign-owned multinational enterprises (FMNE) report the smallest GWG, followed by locally-owned firms that export but do not invest abroad, then by Japanese-owned MNE (JMNE) and finally by domestic-only firms (i.e., domestic-owned firms that neither export nor invest abroad).

<sup>&</sup>lt;sup>1</sup> Matsui et. al. (2014) estimate that closing the gender employment gap in Japan will produce a nearly 13 percent increase in GDP.

<sup>&</sup>lt;sup>2</sup> Based on 2017 data with Japan's female labor force participation rate at 69.4 percent compared with the OECD average of 64 percent (OECD, 2019).

<sup>&</sup>lt;sup>3</sup> Gender pay gap of median monthly earnings defined as the difference between male and female median monthly earnings divided by male median monthly earnings for full-time employees (OECD, 2017).

<sup>4</sup> Estonia's data reflects 2014, the most recent statistic available in the OECD's online data available at: https://data.oecd.org/earnwage/gender-wage-gap.htm.

After finding that FMNE deviate the most from domestic-only firms in terms of GWGs, we ask whether FMNE mode of entry matters for gender wage outcomes. We propose that FMNE originating from greenfield investment will deviate further from domestic-only firms than FMNE originating from mergers and acquisitions (M&A). A Greenfield-born FMNE involves foreigners in setting up the human resource management (HRM) and pay systems from the firm's establishment, while an M&A-born FMNE involves foreigners influencing and possibly changing existing HRM and pay systems.<sup>5</sup> We find support for our hypothesis that Greenfield-born FMNE deviate the most from the gender wage inequality existing at domestic-only firms. Among majority-owned FMNE, only Greenfield-born ones are associated with significant, large reductions in the GWG while M&A-born FMNE do not differ significantly from domestic-only firms that report a 26.8 percent GWG on average. Quintile regressions indicate that Greenfield-born FMNE deviate the most from domestic-only firms in terms of GWGs in the middle quintiles of the wage distribution.

Our research relates to three strands of literature. The first is extensive research on wage premiums paid by exporters or by foreign-owned firms. Early research on this topic used firmlevel or plant-level data to estimate these wage premia while more recent empirical work has utilized matched employer-employee data in order to control for both firm characteristics and worker characteristics in estimating wage premia.<sup>6</sup> While the research using matched datasets,

<sup>&</sup>lt;sup>5</sup> For example, IBM Japan is a greenfield-born MNE (see: <u>https://www.ibm.com/ibm/jp/ja/history.html</u>) while Nissan Motor Co. is a (minority-foreign-owned) M&A-born FMNE based on Renault's equity links to Nissan (see: <u>https://www.nissan-global.com/EN/COMPANY/PROFILE/HERITAGE/HISTORY/</u>).

<sup>&</sup>lt;sup>6</sup> Early research includes Bernard and Jensen (1995, 1997, 1999) on exporter wage premia, and Aitken et al. (1996) and Lipsey and Sjoholm (2004) on MNE wage premia. More recent work using linked employer-employee data to examine exporter wage premia includes Frias et al. (2018), Munch and Skaksen (2008), Schank et al. (2007), Schank et al. (2010), and Verhoogen (2008), while MNE wage premia are addressed using linked data in Girma and Gorg (2007), Heyman et al. (2007), Hijzen et al. (2013) and Martins (2011).

especially panel matched datasets, has improved our understanding of these globalization wage premia by accounting for firm and worker heterogeneity, differing conclusions have been found. Using quasi-experimental variation in exports due to a currency devaluation, Frías et al. (2018) find a large causal link between exporting and wage premia using Mexican data (i.e., a 1 percent increase in a plant's export share leads to a 3 percent increase in the plant-specific wage premium). Heyman et al. (2007) reject the notion of wage premia paid by foreign-owned firms relative to domestic-owned firms using linked employer-employee data for Sweden, while Hijzen et al. (2013) find foreign-ownership wage premia of 1 to 6 percent using matched data for three developed economies (i.e., Germany, Portugal, the UK) and one developing economy (i.e., Brazil).

A second strand of related literature has focused on the impacts of international activities on gender pay disparities. While several papers have examined the potential links between import competition and GWGs,7 we focus on how exports and multinational activities may impact the GWG. On this topic, the literature is quite recent and relatively sparse but a few hypotheses have been proposed and empirically tested. Bøler et al. (2018) propose that exporting requires more flexible workers who can communicate with trade partners at odd hours and/or travel internationally on short notice, which tends to disfavor female workers who are perceived as being more family-oriented and therefore less flexible. Using matched employer-employee data, they find that entry into exporting by Norwegian firms leads to a three-percentage-point increase in the GWG among college-educated workers. Juhn et al. (2014) use firm-level data for Mexico to show that trade-liberalization-induced exporting increased the relative wage and employment of blue-

<sup>&</sup>lt;sup>7</sup> Black and Brainerd (2004) employ the Becker (1957) model of discrimination to explain their finding that increased product market competition through imports reduces the ability of U.S. manufacturing firms in concentrated industries to gender discriminate in wages. Other studies in a similar vein have used industry-level evidence to examine trade liberalization effects for one or two countries mainly through import competition and impacts on gender wage gaps (e.g., both Seguino (2000) and Berik et al. (2004) study South Korea and Taiwan).

collar women because exporting firms adopted new technologies that reduced the need for physically demanding skills. Blue-collar men were replaced by blue-collar women and new production technologies. These two papers suggest that the relationship between exporting and GWGs may differ substantially based on the skill or education level of workers.

Researchers also have explored the link between MNE and gender wage gaps. Oostendorp (2009) finds that the within-occupation GWG tends to decrease with trade and FDI in richer countries but no significant effects are found in poorer countries. Valter and Masso (2019) use Estonian matched employer-employee data to conclude that foreign-owned firms have larger GWGs than locally-owned firms. They argue that the larger GWG is due to men capturing a larger share of MNE wage premiums. One previous study has used matched employer-employee data for Japan to study this issue. Ono and Odaki (2011) use matched data for 1998 to document MNE wage premiums and smaller GWGs within foreign-owned MNE. They suggest that FMNE may have an advantage in hiring career-minded Japanese women, who are under-valued by Japanese firms due to the seniority-wage and lifetime employment systems.8 They find a GWG of 32 percent in locally-owned firms, but only 21 percent in FMNE defined using a 10 percent foreign equity minimum to define FMNE. Our study updates and extends the work of Ono and Odaki (2011) by using more recent data (i.e., for 2012), by considering two different capital ownership criteria for defining foreign ownership (i.e., 50 percent and 10 percent) and by separating firms in our dataset into five different types, instead of only two types (i.e., foreign and domestic).

A third strand of related literature examines the transmission of culture across international borders through economic activity. Of particular interest is the transmission of more gender-neutral corporate culture into a country with less gender-neutral norms. Tang and Zhang (2017) find that

<sup>8</sup> See Ono (2010) and Kambayashi and Kato (2010) for more details on Japan's seniority wage and lifetime employment systems.

foreign affiliates in China from more gender-equal countries tend to have higher female shares of employment compared with foreign affiliates from less gender-equal countries. Siegal et al. (2019) find that FMNE are more likely to hire and/or promote female managers in Korea, while Choi and Greaney (2020) find that FMNE acquisitions of Korean firms result in workforce reorganization that favors proportionately more women in the most elite postings within a firm. They also show that this workforce reorganization produces productivity gains for the acquired firms. 9 Kodama et al. (2018) find that FMNE tend to employ proportionately more women and more female managers than locally-owned firms in Japan. They also find that the FMNE tend to have more family-friendly workplace policies, such as flextime, telecommuting, used vacation days and childcare subsidies. Our results complement the Kodama et al. (2018) results in that we find wage-based evidence that foreign-owned MNE have more gender-equal pay structures than locally-owned firms in Japan.

We contribute to the literature in three distinctive ways. Our study is the first, to our knowledge, to use linked employer-employee data for Japan to examine gender-based differences in employee returns to both exporting and multinational enterprise activities. <sup>10</sup> These two international activities typically have been examined by separate strands of literature that seek to identify wage premiums, whereas we are able to directly compare and disentangle the two effects on wages and GWGs.<sup>11</sup>

Second, we distinguish between domestic-owned MNE and foreign-owned MNE and find substantial differences in worker wages, especially for women in Japan. We find that FMNE pay much larger gender-neutral wage premiums than JMNE operating in Japan, and female workers

<sup>9</sup> This result relates to another strand of research that examines how culture impacts firms' international economic performance through differences in management practices. See, for example, Guiso et al. (2006), Bloom and Van Reenan (2007) and Bloom et al. (2012).

<sup>&</sup>lt;sup>10</sup> Ono and Odaki (2011) use Japanese linked employer-employee data for 1998 to examine the GWG between foreign-owned and locally-owned firms, but they do not address export activities nor distinguish between locally-owned and foreign-owned MNE or between foreign-owned MNE based on entry mode.

<sup>11</sup> Greaney and Li (2016) also disentangle exporter and MNE wage premiums using firm-level data for China.

receive larger added wage premiums by working at FMNE so these firms feature the smallest GWGs. This implies that the "multinationality" of a firm matters in a quantitatively small way among Japanese-owned firms, but foreign ownership matters in a quantitatively large way in Japan's labor market.

Our third and perhaps most unique contribution is to separate FMNE based on their mode of entry which has not to our knowledge been previously done in the literature on FMNE wage premiums. This distinction allows us to test the hypothesis that mode of entry matters for the transmission of corporate gender norms. We propose that Greenfield-born FMNE are the most likely to deviate from the existing corporate gender norms in Japan,12 while other international firm types (i.e., M&A-born FMNE, JMNE and exporters) may be influenced by foreign gender norms to a lesser degree, using domestic-only firms as the reference group. Since our dataset does not distinguish between Greenfield-born FMNE and M&A-born FMNE, we make a unique contribution to the data by examining the history of each FMNE to determine whether its origin involved a foreign partner from the start. The advantage of separating FMNE based on mode of entry is it allows us to address the endogeneity problem for a specific group of foreign-owned firms using cross-sectional data. Smaller GWGs associated with FMNE may be due to foreign owners instituting a more gender-neutral pay system or due to foreign owners preferring to acquire local firms with more gender-neutral pay systems already in place. In order to conclude that foreign ownership drives gender wage outcomes, rather than vice versa, we separately identify FMNE established through greenfield investment. For these firms, foreign influence occurs from firm establishment so we can reasonably interpret the direction of causality as flowing from foreign ownership to gender wage outcomes.

<sup>&</sup>lt;sup>12</sup> This argument assumes that most FMNE come from countries more gender equal than Japan. Statistics to support this assumption are presented in the next paragraph.

Japan is an ideal country to use for this study for several reasons. First, its recent growth promotion policies include encouraging better utilization of its female labor force and enticing more FDI inflows so the relationship between the two should interest policymakers.<sup>13</sup> Second, Japan has some unique labor market characteristics such as the lifetime employment system and seniority wage system that have been identified as posing barriers to women in the labor market.14 Third, Japan has a large GWG but a relatively low level of male wage inequality, making it an interesting outlier among the 22 developed countries studied in Blau and Kahn (2003).15 Fourth, even though we cannot identify the countries of origin of the FMNE in our dataset, we can reasonably assume that most come from countries that are more gender-equal than Japan in terms of employment practices. In the World Economic Forum's Global Gender Gap Report 2012, Japan was ranked 102 out of 135 countries for Economic Participation and Opportunity.16 In 2012, the largest foreign investors in Japan based on FDI stock ranked well above Japan in terms of gender equality using the same gender gap index, with the US (ranked 8) followed by the Netherlands (24), France (62) and the UK (33).17 Focusing specifically on FMNE in manufacturing, European affiliates accounted for 61 percent of total FMNE employment, followed by North American affiliates with 31 percent, while Asian affiliates accounted for only 5 percent of total FMNE employment in 2012.18

<sup>&</sup>lt;sup>13</sup> See Japan Revitalization Strategy: Japan is Back, 2013, for details on Prime Minister Abe's economic growth strategies involving "Promoting active participation by women" (p. 44-48) and "Activating inward direct investment (p. 137-138). See Hoshi and Kiyota (2019) for details and analysis of Japan's promotion of inward FDI since 2013. <sup>14</sup> See Kawaguchi (2015) and Hara (2018) on this topic.

<sup>&</sup>lt;sup>15</sup> They find that compressed returns to skills (i.e., low measured male wage inequality) tend to correspond with low gender wage gaps among developed economies, with Japan being the exception.

<sup>&</sup>lt;sup>16</sup> Japan's ranking in the most recent WEF Gender Gap Report 2020 was 115 out of 153 countries for the same subindex, but we present 2012 data above to match the 2012 wage data used in our study.

<sup>17</sup> Japan's inward FDI stock data obtained from the OECD (2019) database.

<sup>18</sup> Data from the Ministry of Economy, Trade and Industry, Survey of Trends in Business Activities of Foreign Affiliates, 47th Survey, for fiscal year 2012, with a response rate of 60 percent.

We present our methodology in the next section, followed by a description of our data sources and summary statistics. Section 4 presents our empirical results, followed by a section on robustness checks. We conduct robustness checks by changing our definition of foreign ownership from majority-foreign-ownership to a more expansive definition that includes all firms with 10 percent or more foreign capital ownership. Our conclusions are presented in Section 6.

#### 2 Methodology

This study employs both ordinary least squares (OLS) and quantile regression (QR) to estimate the Mincer (1974) wage equations.<sup>19</sup> After investigating the wage premium using OLS, we conduct QR to examine wage premiums on a particular percentile of the wage distribution to investigate whether these premiums vary across the range of wages. For example, we examine whether high-wage workers obtain higher multinational wage premiums than low-wage workers and whether GWGs vary between high-wage and low-wage workers. QR has several attractive features, as explained in Koenker and Hallock (2001) and Cameron and Trivedi (2010). First, QR enables us to investigate the effects of a covariate on the full distribution or any particular percentile of the distribution, whereas OLS reveals the average relationship between the wage and explanatory variables. Second, QR is robust to the presence of outliers, whereas OLS regression is sensitive to such presence.

First, we employ OLS and the following Mincer wage equation to examine whether exporters or multinationals pay higher wages than our reference group that we call "domestic-only" establishments (i.e., locally-owned plants that neither export nor own subsidiaries abroad):

<sup>&</sup>lt;sup>19</sup> Ideally we would use the difference-in-differences technique to examine our research questions with panel data, but our linked employer-employee data is not available as a panel, as we explain in the Data section.

$$lnWage_{ip} = \alpha + \beta Female_i + \gamma' \sum_p EstType_p + \delta' \sum_p Female * EstType_p + \theta' EstChar_p + \mu' WorkerChar_i + \varepsilon_{ip}.$$
(1)

InWage<sub>*ip*</sub> is the log of hourly wage for worker *i* in plant *p*, Female<sub>*i*</sub> is an indicator variable for female workers, and EstType<sub>*p*</sub> is a vector of dummy variables for our various international establishment types of interest (i.e., initially, Exporter, JMNE and FMNE). In some specifications we replace the FMNE dummy with separate dummy variables for Greenfield-born FMNE and M&A-born FMNE. We control for the following establishment characteristics (EstChar<sub>*p*</sub>): size (measured by the log of establishment employment), a headquarters dummy variable, average worker productivity (measured by the log of plant sales per worker), and industry and region fixed effects.<sup>20</sup> We also control for individual worker characteristics (WorkerChar<sub>*i*</sub>) that are typically included in a Mincer wage regression (i.e., education,<sup>21</sup> experience, experience squared, tenure, and tenure squared), along with dummy variables for white-collar workers and workers hired for an indefinite period. The latter is used to proxy for long-term employees of Japanese firms. The descriptive statistics of all variables for male and female workers are presented in Appendix Table A1.22

In an alternate OLS specification, we measure the international activities variables of interest using continuous variables. For exporting, we use the export share of total sales which allows us to control for exporting by local exporters and by JMNE and FMNE.23 To measure JMNE activities with a continuous variable, we use the log of the number of foreign subsidiaries reported by each firm. More foreign subsidiaries indicates greater engagement in foreign markets and

<sup>20</sup> We have 24 manufacturing industries at the 2-digit level and 47 prefectures.

<sup>&</sup>lt;sup>21</sup> We include three dummy variables to indicate the highest educational attainment achieved: high school graduate, junior college graduate, or college graduate.

<sup>22</sup> Appendix tables appear online here: <u>https://sites.google.com/site/greaneyecon/research</u>.

<sup>&</sup>lt;sup>23</sup> Note that this differs from our treatment with dummy variables in which the establishment types are mutually exclusive (i.e., an exporting firm is classified as a locally-owned exporter, a JMNE or a FMNE).

stronger potential for the transmission of more gender-equal workplace norms to filter back to Japan. For FMNE, the foreign ownership ratio is used as the continuous variable. The larger the foreign ownership ratio, the stronger the foreign influence on the workplace norms.

Second, this study employs QR to estimate the wage premium on any particular percentile of the wage distribution. Using the linear programming method, we obtain the qth estimator  $\hat{\gamma}_q$ , which minimizes over  $\gamma_q$  the objective function:

$$Q(\gamma_q) = \sum_{i: y_i \ge x_i' \gamma}^N q |y_i - x_i' \gamma_q| + \sum_{i: y_i \le x_i \gamma}^N (1 - q) |y_i - x_i' \gamma_q|$$
(2)

where 0 < q < 1,  $y_i$  is the log of hourly wage and  $x_i$  is a vector of explanatory variables. We divide the wage distribution into quintiles and use the same explanatory variables as were used in the OLS regression. The QR estimates have the advantage of being more robust to non-normal errors and outliers.

#### **3** Data and summary statistics

#### 3.1 Data

To construct the matched employer–employee data, this study uses confidential data from three official surveys: (i) *The Basic Survey on Wage Structure* (2012), (ii) *The Economic Census for Business Frame* (2009), and (iii) *The Economic Census for Business Activity* (2012).<sub>24</sub> We merge the data from these three surveys at the plant level using the common plant-level census ID.

First, we obtain the data on most variables used in the analysis, such as worker-level wage and other worker-, plant-, and firm-level variables from *the Basic Survey on Wage Structure*, and construct the worker-level cross-sectional data. *The Basic Survey on Wage Structure* is conducted

<sup>24</sup> All three surveys are conducted as Fundamental Statistics according to Japan's Statistics Act.

every year from July 1st to 31st by the Ministry of Health, Labour and Welfare (MHLW).25 The survey covers plants with five or more regular employees in major industries in Japan. Plants are selected using a uniform sampling method. Employees are also selected using a uniform sampling method from among the plants selected for the survey.

Second, we merge the worker-level data with the firm-level data on foreign ownership from *The Economic Census for Business Frame*. This is a newly-created census to identify the basic structure of establishments and enterprises in Japan and is conducted by the Ministry of Internal Affairs and Communications (MIC).<sub>26</sub> The 2009 census was the first one and it was conducted as of July 1, 2009.

Third, we also merge the worker-level data with data on firms' export status from *The Economic Census for Business Activity*. This is another newly-created census to investigate the economic activity of establishments and enterprises in all industries. The purpose of the census is to obtain basic information for conducting various statistical surveys.<sup>27</sup> Using the results of the *2009 Economic Census for Business Frame*, the *2012 Economic Census for Business Activity* was conducted by MIC and the Ministry of Economy, Trade and Industry (METI) in February 2012 to investigate the activities of establishments and enterprises during 2011.

Full details on the matching procedures appear in Appendix Figure A1.28 As indicated, we use worker-level data and plant-level data from 2012 except for export data, which is reported in the 2012 Economic Census for Business Activity for 2011, and JMNE number of foreign subsidiaries and FMNE foreign equity shares, which are from *The Economic Census for Business* 

<sup>25</sup> See the MHLW website (http://www.mhlw.go.jp/english/database/db-l/wage-structure.html) for more details.

<sup>26</sup> See the MIC website (http://www.stat.go.jp/english/data/e-census/index.htm) for more details.

<sup>27</sup> See the MIC website (http://www.stat.go.jp/english/data/e-census/2012/index.htm) for more details.

<sup>28</sup> Appendix Figure A1 is available online here: https://sites.google.com/site/greaneyecon/research.

*Frame* for 2009. The matching procedure to link employees and employers is successful only for the 2012 data so we are restricted to using single-year cross-sectional data.<sup>29</sup>

After constructing the linked employer–employee data through these steps, we develop a dataset to estimate the Mincer wage equation. This study's procedure follows that of Kawaguchi (2011). In particular, we restrict our analysis to the sample of full-time workers under the age of 60 years who work at private manufacturing firms. The reasons for these restrictions are as follows.<sub>30</sub> First, we restrict our analysis to full-time workers because information on the education level of part-time workers is unavailable in the survey.<sub>31</sub> Second, we drop the data on workers over 60 years of age to address the fact that workers in Japan at the age of 60 years tend to face large declines in wages due mainly to mandatory retirement policies at almost all Japanese firms.<sub>32</sub> Lastly, we restrict our sample to manufacturing firms because we do not have export data for non-manufacturing firms that can be matched with the employee data.

Following Kawaguchi (2011), the monthly total wage is calculated as the sum of monthly wages plus one-twelfth of yearly bonuses. Then, the hourly wage used in our analysis is calculated as the monthly total wage divided by hours worked expressed in hundreds of Japanese yen (JPY). To control for the potential influence of outliers, we exclude hourly wages that fall below the minimum wage at the prefectural level and hourly wages that exceed the mean wage plus 10 standard deviations.

<sup>&</sup>lt;sup>29</sup> Kawaguchi (2011) noted the same single-year matching outcome. Note that *The Basic Survey on Wage Structure* is not a longitudinal survey so even studies that use only this data for multiple years cannot use worker-fixed-effects models or establishment-fixed-effects models.

<sup>30</sup> See Kawaguchi (2011) for more details.

<sup>&</sup>lt;sup>31</sup> Full-time workers include regular (*sei shain*) and non-regular (*hisei shain*) workers, with both groups further divided into staff hired for an indefinite period (i.e., long-term employees) or staff hired for a definite period. Regular workers make up approximately 90 percent of our sample.

<sup>&</sup>lt;sup>32</sup> Mandatory retirement policies were used at 93.3 percent of Japanese firms overall, and 99.8 percent of firms with 1000 or more employees, according to the 2013 General Survey on Working Conditions conducted by the MHLW. Workers at such firms are expected to formally retire at age 60 but often continue to work at the same firm at a lower wage.

Utilizing the data from the two Economic Censuses, we initially classify firms into four types: domestic-only firms, exporters, Japan-owned MNE (JMNE), and foreign-owned MNE (FMNE). Domestic-only firms are locally-owned non-exporters without foreign subsidiaries, whereas exporters are locally-owned exporters without foreign subsidiaries. JMNE are locallyowned firms with foreign subsidiaries which are defined as foreign affiliates that are 50 percent or more equity-owned by the JMNE and its other subsidiaries. Finally, FMNE are defined as firms for which more than 50 percent of the equity is foreign owned. We conduct robustness tests using a 10 percent foreign equity cutoff to check our results. We also separate FMNE based on their mode of entry into Greenfield-born FMNE and M&A-born FMNE for further tests. This separation is based on the authors' research into the corporate history of each FMNE to determine whether foreign ownership was involved at the time of the firm's establishment. It should be noted that we use the existence of foreign subsidiaries to identify JMNE and the presence of foreign equity ownership to identify FMNE, but our wage data is based solely on each establishment's employees in Japan. Therefore, foreigners working in Japan will be included but all employees of JMNE and FMNE working outside Japan are excluded.

#### **3.2 Summary statistics**

Table 1 presents the number of firms, plants, workers and wages in the sample by firm type using the 50 percent criteria for FMNE.<sub>33</sub> Among 7,932 firms, 6,485 firms (82 percent) are domestic-only firms. The number of JMNE is 971 (12.2 percent), which is the second largest group. The third largest group is exporters, at 437 firms (5.5 percent). The number of majority-owned

<sup>&</sup>lt;sup>33</sup> Table 1 also shows summary statistics for observations using a 10 percent foreign-ownership criteria for defining FMNE. This alternate definition will be used later for robustness checks.

FMNE is only 39 (0.5 percent), with 20 Greenfield-born and 19 M&A-born FMNE.<sub>34</sub> The shares of plants by firm type follow the same ranking.

The tendency for exporters and MNE to be larger than domestic-only firms is confirmed in the employment statistics. Domestic-only firms account for 82 percent of firms but only 61 percent of employees in our data, while JMNE employ 31 percent, exporters employ 7 percent and FMNE employ 1 percent of the 188,787 workers in our dataset. Of the 43,260 female workers in our dataset, 70 percent work in domestic-only firms, 23 percent in JMNE, 6 percent in exporters and 0.7 percent in FMNE.

<Table 1>

Table 1 also provides worker-level descriptive statistics for mean hourly wage by firm type and indicates that FMNE tend to pay the highest wages (3170 JPY), followed by JMNE (2760 JPY), exporters (2150 JPY), and domestic-only firms (1770 JPY) in descending order of wage.

#### **4** Estimation results

Table 2 shows regression results for the Mincer wage regression in equation (1) using dummy variables for firm types and majority foreign-ownership as the criteria for defining FMNE. Industry and region fixed effects are included in all regressions and standard errors are clustered at the firm-level. The results in column (1) indicate a large "raw" GWG of 34.0 percent when we only control for industry and region fixed effects. 35 In column (2) we reduce the GWG to 27.4

<sup>&</sup>lt;sup>34</sup> It should be noted that the small number of majority-owned FMNE in our employer-employee linked dataset is indicative of their scarcity in Japan. Majority-owned FMNE account for only 0.2 percent (or 498) of the 277,066 total number of manufacturing enterprises included in the *The Economic Census for Business Frame* in 2009. <sup>35</sup> Wage premiums for all dummy variables are calculated as exponent(coefficient)-1.

percent by adding worker characteristics to the estimation. We next add other plant characteristics in column (3) to reduce the GWG to 26.2 percent. The estimated coefficients on the other control variables used in Table 2 are significant and have the expected signs with positive values for plant size, plant sales per worker, all education dummy variables, experience, tenure, lifetime employee and white collar.<sub>36</sub>

Column (4) shows the effect of our establishment-type variables on wages and on the GWG. The coefficient on Female now represents the average GWG among domestic-only firms, which is estimated to be 26.8 percent. This result implies that women earn 26.8 percent less than men working in domestic-only firms after controlling for observable plant and worker characteristics. Exporters, JMNE and FMNE all pay wage premia (WP) relative to domestic-only firms and women receive significantly higher WP than men from these international firm types. Compared with domestic-only firms, exporters and JMNE pay modest gender-neutral WP of 2.8 percent and 7.4 percent, respectively, while majority-owned FMNE pay a large gender-neutral WP of 24.7 percent. Female workers at these types of firms receive these gender-neutral WP plus added ones of 4.2 percent, 2.6 percent and 10.2 percent, respectively. These added WP for female workers imply smaller GWGs among these international firms compared with domestic-only firms. The GWG is 19.3 percent at FMNE, a 7.5 percentage-point improvement compared with the 26.8 percent gender wage gap at domestic-only firms. 37 Due to the large gender-neutral WP plus female-specific WP associated with working at FMNE, women working at these firms earn 0.6 percent more than men working at domestic-only firms, after controlling for other plant and worker

<sup>&</sup>lt;sup>36</sup> We find negative but quantitatively very small coefficients for the squares of experience and tenure, and a negative small coefficient on the headquarters (HQ) dummy variable. The latter result may reflect that HQ workers receive larger non-wage compensation through transportation and housing allowances that are not captured in our wage data. Full regression results are suppressed in our tables due to space constraints but are available online here: https://sites.google.com/site/greaneyecon/research.

<sup>37</sup> GWG for each international firm type is calculated as exp(Female + Female\*EstType) – 1.

characteristics. This implies approximate wage parity between a female worker and a male worker with similar observable worker, plant, industry and region characteristics except that the woman works for an FMNE while the man works at a domestic-only firm.

#### <Table 2>

In column (5) of Table 2 we separate FMNE based on their mode of market entry and we confirm strong differences both in WP and GWG based on entry mode. Greenfield-born FMNE pay a much larger gender-neutral WP of 35.5 percent relative to domestic-only firms while M&A-born FMNE pay a corresponding WP of only 14.7 percent. Of even greater interest, only Greenfield-born FMNE are associated with a significant reduction in the GWG, with an estimated GWG of 14.9 percent. This implies that the GWG among Greenfield-born FMNE is 11.9 percentage-points lower than it is among domestic-only firms or among M&A-born FMNE. These results support our hypothesis that foreign owners have a greater influence on corporate HRM and pay structure when they are involved from the firm's birth rather than taking control at a later stage through M&A.38

Based on these results, we cannot determine whether Greenfield-born FMNE hire very similar workers to other firm types but gender discriminate less in paying them or hire more able female workers who only appear similar to those at other firms based on observable characteristics that cannot capture ability or ambition. Either interpretation (i.e., pay similar workers more or

<sup>&</sup>lt;sup>38</sup> An alternative hypothesis regarding the wage differences found between Greenfield-born and M&A-born FMNE may be that the former tend to be 100 percent foreign-owned while the latter are not. However, the six percentage point difference in average foreign equity shares across the two firm types (i.e., 87.8 percent for Greenfield-born FMNE versus 81.5 percent for M&A-born FMNE) seems much too small to explain such large differences in WP and GWG.

employ better workers) supports our hypothesis of the transmission of more gender-equal pay structures through Greenfield-born FMNE. In the latter interpretation, the employee-employer matching that pairs more able female workers with Greenfield-born FMNE is not random but arguably reflects the more gender-equal corporate culture of the foreign-owned firms that deviate the most from the local corporate norms of gender inequality.<sup>39</sup> Greenfield-born FMNE also may be the most likely to deviate from the lifetime employment plus seniority wage system that fails to adequately reward merit,<sup>40</sup> which implies that more ambitious workers who want a more marketbased, competitive pay structure seek out employment with these foreign-owned firms. The large gender-neutral WP at FMNE, and particularly at Greenfield-born ones, supports this hypothesis.<sup>41</sup> In addition, Ono (2007) concludes that workers at foreign firms in Japan have higher levels of general human capital and are more likely to have changed jobs in the past, indicating a higher mobility of these workers compared with workers in Japanese-owned firms who are less mobile in part due to their firm-specific human capital.<sup>42</sup>

In Table 3, we use continuous variables rather than dummy variables for each of our international firm activities. In column (1) we control for exporting using the export share of sales for each establishment. This allows us to control for any WP associated with export-related activities by "Exporters" and by MNE, but it changes the reference group to now include both

<sup>39</sup> Kodama et al. (2018) cite anecdotal evidence that Japanese women are likely to have better career opportunities working in FMNE than in Japanese-owned firms, using evidence from IBM Japan and Nissan Motor Co. following its acquisition by Renault. Both firms made substantial increases in their numbers of female managers.
<sup>40</sup> Survey results from 2005 indicate that 67.2 percent of foreign-owned firms reported adopting meritocracy in their labor management (i.e., compensation, job assignments and promotions) while only 55.9 percent of Japanese-owned firms responded similarly. (The Japan Institute for Labour Policy and Training, "Report on the Survey on Foreign Firms Industrial Relations", *Gaishi Kei Kigyo no Roshi Kankei Tou Jittai Chousa Kekka Houkokusyo.*)
<sup>41</sup> A reasonable question is whether Japanese job applicants can identify Greenfield-born FMNE relative to other types of firms, including M&A-born FMNE, but the corporate name is often a clear indication since foreign names are written in "*katakana*", a separate Japanese syllabary used specifically for foreign-language words.
<sup>42</sup> Firm-specific human capital is developed through intra-firm training that is typical of the Japanese HRM system involving lifetime employment and seniority wages. See Ono (2007) for more details.

domestic-only firms and locally-owned exporters without foreign subsidiaries. We find that a 10 percentage-point increase in export share of sales is associated with a 0.45 percent gender-neutral WP and an added 0.68 percent female WP. Greater export intensity corresponds with higher wages and lower GWGs.

<Table 3>

The regression results in columns (2) through (4) of Table 3 change our measurement of MNE activities to continuous variables and establish that being "more international" as an MNE is associated with higher gender-neutral WP for workers but not necessarily smaller gender wage gaps. Column (2) results include the log of the number of foreign subsidiaries held by each JMNE as a continuous indicator of their international activities. We estimate that a 10 percent increase in the number of foreign affiliates for a JMNE corresponds with 0.3 percent gender-neutral WP and an additional 0.11 percent WP for women at these firms. In column (3), FMNE are measured using the foreign capital ownership ratio, which can vary from 50 percent to 100 percent in this specification. A 10 percentage-point increase in the foreign capital ownership ratio of majorityowned FMNE is associated with a 3 percent gender-neutral WP and an additional 1 percent WP for women.43 Column (4) shows our results using the foreign capital ownership ratio again but dividing FMNE into those that are Greenfield-born versus M&A-born. The estimated coefficients confirm our previous results that Greenfield-born FMNE deviate more than M&A-born FMNE from local norms in worker compensation, and we find the deviation increases in the level of foreign ownership. Greenfield-born FMNE pay larger gender-neutral WP than M&A-born FMNE

<sup>&</sup>lt;sup>43</sup> Ono and Odaki (2011) also find a 3 percent wage premium associated with a similar increase in foreign equity share using 1998 wage data, despite using fewer worker and firm characteristics to control for worker productivity.

(i.e., 3 percent versus 2 percent WP, respectively, from a 10 percentage-point increase in the foreign capital ownership ratio) and Greenfield-born FMNE are the only ones that pay a significant added WP to women (i.e., 2 percent WP from a 10 percentage-point increase in the foreign capital ownership ratio).

In column (5) of Table 3, we include continuous variables for all of our international firmtype activities. More foreign subsidiaries for JMNE and higher foreign ownership ratios for Greenfield-born and M&A-born FMNE generate significant gender-neutral WP while higher export shares do not. Higher export shares, more foreign subsidiaries for JMNE and higher foreign ownership ratios for Greenfield-born FMNE all correspond with lower GWG. Our results using continuous variables show that more international activities tend to generate larger gender-neutral and female-specific WP. Our results using both indicator and continuous variables support our hypotheses that mode of entry matters for gender wage outcomes and that larger deviations from the host-country gender wage outcomes are associated with Greenfield-born rather than M&Aborn FMNE.

Quantile regression results using the majority-foreign-ownership criteria and controlling for both plant and worker characteristics are shown in Table 4. Having established significant differences between Greenfield-born and M&A-born FMNE in previous results, we now include these two types of FMNE separately in each estimation. The strongest changes across the income distribution for our establishment-type variables of interest are for FMNE WP, which increase strongly in wage quintiles, particular for Greenfield-born FMNE. The gender-neutral WP associated with Greenfield-born FMNE relative to domestic-only firms is 27.0 percent for workers in the lowest wage quintile while it is a whopping 41.6 percent for workers in the highest quintile. <Table 4>

The Table 4 results indicate that the additional WP for female workers at Greenfield-born FMNE are largest in the middle three quintiles of the wage distribution, while M&A-born FMNE pay no significant added WP to women in these wage quintiles. These WP can be used to calculate GWGs by firm type, which are presented in Figure 1 across the five wage quintiles. The figure clearly shows that Greenfield-born FMNE deviate the most from domestic-only firms in terms of having much lower GWGs. Greenfield-born FMNE are associated with 13.2, 16.1 and 12.2 percentage-point reductions in the GWGs of 26.1 to 26.4 percent prevailing at domestic-only firms at the 25th, 50th and 75th wage quintiles, respectively.44 The next highest reductions in GWG at these quintiles are only 3.4 to 3.5 percentage-points for exporters.

<Figure 1>

Greenfield-born FMNE are associated with the smallest reduction in GWG at the highest income quintile, with a 4.8 percentage-point reduction, but this is still larger than the next highest GWG reduction of 3.2 percentage-points for M&A-born FMNE. This may indicate that women still face barriers to achieving the highest-paying corporate jobs in Japan, even at FMNE. In this context, it is interesting to note that the 30 percent Club Japan was launched on May 1, 2019, with the goal of increasing female representation on corporate boards of TOPIX100 companies to 10 percent by 2020 and to 30 percent by 2030. Of the group's 36 members thus far, it is noteworthy

<sup>&</sup>lt;sup>44</sup> See Appendix Table A2 to see the gender wage gap differentials for all firm types that correspond with the coefficients in Tables 2, 4, 5 and 7.

that 14 (i.e., almost 39 percent) are from foreign-owned firms or organizations.<sup>45</sup> This suggests that gender-based differences between FMNE and Japanese-owned firms in the highest wage quintile may continue and even grow in the future as the latter are under-represented in this public commitment to corporate change. Overall, the QR results show strong differences between Greenfield-born FMNE and M&A-born FMNE with the former contributing much more in terms of gender-neutral WP and reductions in the GWG across income quintiles.

#### 5 Robustness Check: Foreign ownership criteria

Our results thus far are based on majority-owned FMNE, but the typical cut-off for classifying FDI is 10 percent or more foreign ownership of a firm. Therefore, we check the robustness of our results by considering a less stringent criteria for defining foreign ownership. That is, we lower our criteria from 50 percent to 10 percent minimum foreign equity ownership to classify FMNE. This reclassification moves minority-foreign-owned firms (i.e., those with foreign-equity ownership of at least 10 percent but less than 50 percent) into the FMNE group and out of one of the three domestic-owned firm groups. Summary statistics in Table 1 show that FMNE gain a total of 153 firms, or 19,140 workers, with the change from a 50 percent to a 10 percent cutoff, most of which come from JMNE (135 firms with 18,206 workers), while only a handful come from domestic-only firms (10 firms with 646 workers) and exporters (8 firms with 288 workers). Note that this change in our definition of FMNE to include minority-owned FMNE means 11 percent of our sample is now classified as working for FMNE whereas we previously had just under 1 percent classified as working for majority-owned FMNE. This expanded

<sup>45</sup> Based on 30 percent Club Japan website accessed on July 24, 2019:

https://30percentclub.org/about/chapters/japan . The 30 percent Club originated in the United Kingdom in 2010, and the Japan Chapter was the club's 14th chapter.

definition of FMNE allows us to check our results for the influence of outliers and/or measurement errors due to the relatively small number of workers at majority-owned FMNE.

Table 5 shows the Mincer wage regression results using dummy variables to indicate firm type and the 10 percent criteria for defining FMNE and other firm types. The qualitative results of interest hold, while quantitatively the WP for FMNE are smaller as anticipated due to the reclassification of so many previously-classified JMNE workers who work in minority-owned FMNE. Column (4) results show the same firm-type ranking of significant positive WP for women that was previously established, going from FMNE (4.5 percent), to exporters (4.0 percent), and then JMNE (2.5 percent) all paying women more than they earn in domestic-only firms. Column (5) of Table 5 shows results after grouping FMNE by mode of entry. Similar to our previous findings, Greenfield-born FMNE are associated with a larger gender-neutral WP of 25.1 percent versus only 10.2 percent for M&A-born FMNE. With the 10 percent foreign ownership criteria we find that both modes of entry are associated with significant reductions in GWGs. Greenfield-born FMNE are associated with a larger reduction in the GWG than M&A-born FMNE which is consistent with our prior findings (i.e., 6.5 percentage-points for Greenfield-born FMNE versus 2.6 percentage points for M&A-born FMNE).

#### <Table 5>

Next we report the estimation results using continuous variables to measure international activities and the 10 percent criteria for defining FMNE in Table 6. The elasticity estimates using the various continuous variables are very similar to previous estimates presented in Table 3 using the 50 percent criteria for FMNE. As we found previously, the GWG is reduced by increases in

export share of sales, JMNE number of foreign subsidiaries and the foreign ownership ratio of Greenfield-born, but not M&A-born, FMNE.

<Table 6>

Table 7 reports QR results using the 10 percent foreign-ownership criteria for identifying FMNE and the other firm types. As noted previously, QR analysis is more robust than OLS analysis to non-normal errors and outliers. The inclusion of minority-owned FMNE makes the previously-observed strong upward trend in the Greenfield-born FMNE gender-neutral WP disappear, but we again see an upward trend until the 50th or 75th percentiles of the wage distribution and then a decline in this firm type's additional WP for women. Greenfield-born FMNE are associated with declines in the GWG of 6.2, 12.6 and 8.7 percentage-point at the 50th, 75th, and 90th percentiles of the wage distribution, but this firm type is no longer a significant determinant of the GWG at the lowest two wage quintiles. M&A-born FMNE are now associated with reductions in the GWG at all wage quintiles, with an increasing trend from 1.5 percentage-points at the lowest wage quintile to 3.4 percentage-points at the highest wage quintile. Overall, we find that our main results are robust in qualitative terms to relaxing the definition of FMNE to include minority-foreign-owned firms. This implies that our results are not driven by measurement errors or outliers among the relatively small sample of majority-owned FMNE.

<Table 7>

#### 6 Conclusions

Using linked employer-employee data for 2012, we find that FMNE pay the largest genderneutral wage premia and they feature the smallest gender wage gaps in Japan compared with domestic-only firms, exporters and JMNE. We separate FMNE by their mode of market entry— Greenfield-born versus M&A-born—and find that Greenfield-born FMNE deviate the most from domestic-only firms in term of wages and gender wage gaps. Separating FMNE by mode of entry allows us to address the endogeneity problem, at least for Greenfield-born FMNE, in using crosssectional, rather than panel, data for this study. We confirm that the influence of foreign workplace gender norms is most apparent at Greenfield-born FMNE, which report a much lower GWG than any of our other firm types. Greenfield-born FMNE are associated with a GWG that is almost 12 percentage-points lower than the 26.8 percent prevailing among domestic-only firms.

The smaller gender-based pay discrepancy is likely due to both lower levels of gender discrimination in pay and unobservable worker quality differences that match more ambitious and able female workers with these foreign-owned firms. Greenfield-born FMNE also report the highest gender-neutral wage premia (i.e., 35.5 percent) which suggests more ambitious and able male workers also may seek employment with these FMNE that are less likely to practice the seniority wage plus lifetime employment system of Japanese-owned firms. We interpret the lower GWG and higher gender-neutral wage premia associated with Greenfield-born FMNE as being caused by foreign influences on a firm's corporate culture and human resource management. 46

Admittedly our study would benefit from the availability of panel data that links employer and employee observations. That would allow us to control for unobservable worker and plant characteristics so that we could more precisely examine the causal links between foreign

<sup>&</sup>lt;sup>46</sup> See Kato and Kodama (2018) for an extensive review of the literature on management practices and gender-based outcomes in the workforce, including suggested directions for further research.

ownership, multinationality, exporting and worker compensation. Since such data is currently unavailable for Japan, our unique approach with cross-sectional data addresses the endogeneity issue for a subset of FMNE by separating them by mode of entry. The strong differences found between Greenfield-born FMNE and M&A-born FMNE provide support for this approach.

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#### References

- Aitken, B., A. Harrison and R. E. Lipsey (1996), Wages and Foreign Ownership: A Comparative Study of Mexico, Venezuela, and the United States, *Journal of International Economics* 40, 3-4,345–71.
- Becker, G. S. (1957), (1971). The Economics of Discrimination, Chicago: University of Chicago Press.
- Berik, G., Y. van der Meulen Rodgers, and J.E. Zveglich, Jr. (2004). International Trade and Gender Wage Discrimination: Evidence from East Asia, *Review of Development Economics* 8(2): 237-254.
- Bernard, A. B. and B. Jensen (1995), Exporters, Jobs, and Wages in U.S. Manufacturing: 1976– 87, *Brookings Papers on Economic Activity: Microeconomics*, 1995, 67–112.
- Bernard, A. B. and B. Jensen (1997), Exporters, Skill Upgrading and the Wage Gap, *Journal of International Economics*, 42, 1–2, 3–31.
- Bernard, A. B. and Jensen, J. B. (1999). Exceptional Exporter Performance: Cause, Effect, or Both?, *Journal of International Economics* 47: 1-25.

- Black, S. E. and E. Brainerd (2004). "Importing equality? The impact of globalization on gender discrimination, *Industrial and Labor Relations Review* 57: 540-559.
- Blinder, A. (1973). Wage Discrimination: Reduced Form and Structural Estimates, *Journal of Human Resources* 8 (4): 436-455.
- Blau, F. D. and L. M. Kahn, (2003). Understanding International Differences in the Gender Pay Gap, *Journal of Labor Economics* 21(1): 106-144.
- Bloom, N., R. Sadun, and J. Van Reenen (2012). The Organization of Firms Across Countries, Quarterly Journal of Economics 127: 1663-1705.
- Bloom, N. and J. Van Reenen (2007). Measuring and Explaining Management Practices across Firms and Countries, *Quarterly Journal of Economics* 122: 1351–1408.
- Bøler, E. A., B. Javorcik, and K. H. Ulltveit-Moe (2018), "Working across time zones: Exporters and the gender wage gap," *Journal of International Economics* 111: 122-33.
- Cameron, A. C. and P. K. Trivedi (2010). *Microeconometrics using Stata*, College Station, TX: Stata Press.
- Choi, J. and T. M. Greaney (2020). Global Influences on Gender Inequality: Evidence from Female Employment in Korea, University of Hawai'i Economics Department Working Paper No. 20-3R.
- Frías, J. A., D. S. Kaplan, E. A. Verhoogen and D. Alfaro-Serrano (2018). Exports and Wage Premia: Evidence from Mexican Employer-Employee Data, unpublished, Columbia University.
- Girma, S. and H. Gorg (2007). "Evaluating the Foreign Ownership Wage Premium using a Difference-in-Differences Matching Approach," *Journal of International Economics*, 72, 97–112.
- Greaney, T. M. and Y. Li (2016). Examining Determinants of Foreign Wage Premiums in China, *The World Economy* 40: 2056-2077.
- Guiso, L., P. Sapienza, and L. Zingales (2006). Does Culture Affect Economic Outcomes?, *Journal* of Economic Perspectives 20: 23–48.
- Hara, H. (2018). The gender wage gap across the wage distribution in Japan: Within- and betweenestablishment effects, *Labour Economics* 53: 213-229.
- Heyman, F., F. Sjoholm, and P. G. Tingvall (2007). Is There Really a Foreign Ownership Wage Premium? Evidence from Matched Employer-Employee Data, *Journal of International Economics*, 73, 355–376.

- Hijzen, A., P. S. Martins, T. Schank, and R. Upward (2013). Foreign-Owned Firms around the World: A Comparative Analysis of Wages and Employment at the Micro-Level, *European Economic Review*, 60, 170–188.
- Hoshi, T. and K. Kiyota (2019). Potential for inward foreign direct investment in Japan, *Journal* of the Japanese and International Economies: 52: 32-52.
- Japan Revitalization Strategy: Japan is Back, June 14, 2013, (accessed on July 26, 2019), available from URL: <u>http://www.kantei.go.jp/jp/singi/keizaisaisei/pdf/en\_saikou\_jpn\_hon.pdf</u>.
- Juhn, Chinhui, Gergely Ujhelyi, and Carolina Villegas-Sanchez (2014). Men, women, and machines: How trade impacts gender inequality, *Journal of Development Economics 106*: 179–193.
- Kambayashi, R. and T. Kato (2010). The Japanese Employment System after the Bubble Burst: New Evidence, in K. Hamada, A.K. Kashyap, and D.E. Weinstein (eds.), Japan's Bubble, Deflation, and Long-Term Stagnation, The MIT Press, pp. 217-262.
- Kato, T. and N. Kodama (2018). Women in the Workplace and Management Practices: Theory and Evidence, in S.L. Averett, L.M. Argys, and S.D. Hoffman (eds.), Oxford Handbook of the Economics of Women, Oxford University Press, New York.
- Kawaguchi, A. (2015). Internal labor markets and gender inequality: Evidence from Japanese micro data, 1990-2009, *Journal of The Japanese and International Economies* 38: 193-213.
- Kawaguchi, D. (2007). A Market Test for Sex Discrimination: Evidence from Japanese Panel Data, International Journal of Industrial Organization 25(3): 441-460.
  - (2011). Applying the Mincer Wage Equation to Japanese Data (in Japanese), *RIETI Discussion Paper Series*, 11-J-026.
- Kodama, N., B. S. Javorcik, and Y. Abe (2018). Transplanting corporate culture across international borders: Foreign direct investment and female employment in Japan, *The World Economy 41* (5): 1148–1165.
- Koenker, R. and K. Hallock (2001). Quantile Regression: An Introduction, *Journal of Economic Perspectives*, 15, 43–56.
- Lipsey, R. and Sjoholm, F. (2004). Foreign direct investment, education and wages in Indonesian manufacturing, *Journal of Development Economics* 73: 415-422.
- Martins, P. S. (2011). Paying More to Hire the Best? Foreign Firms, Wages, and Worker Mobility, *Economic Inquiry*, 49, 349–363.
- Matsui, K., H. Suzuki, K. Tatebe, and T. Akiba (2014). Womenomics 4.0: Time to Walk the Talk, Japan Portfolio Strategy, May 30, 2014, The Goldman Sachs Group, Inc.

Mincer, J. (1974). Schooling, Experience, and Earnings, New York: NBER Press.

- Munch, J. R. and J. R. Skaksen (2008). Human Capital and Wages in Exporting Firms, *Journal of International Economics*, 75, 363–372.
- Oaxaca, R. (1973). Male-Female Wage Differentials in Urban Labor Markets, *International Economic Review* 14 (3): 693-709.
- OECD (2019). International Direct Investment Statistics, data downloaded from OECD.Stat on May 20, 2019. https://stats.oecd.org/index.aspx?lang=en
- OECD (2017). The Pursuit of Gender Equality: An Uphill Battle, OECD Publishing, Paris. http://dx.doi.org/10.1787/9789264281318-en
- Oostendorp, R. H. (2009). Globalization and the Gender Wage Gap, *World Bank Economic Review* 23 (1): 141–161.
- Ono, H. (2007). Careers in Foreign-Owned Firms in Japan, American Sociological Review 72: 267-290.
  - (2010). Lifetime employment in Japan: Concepts and measurements, *Journal of the Japanese and International Economies* 24: 1-27.
- Ono, H. and K. Odaki (2011). Foreign Ownership, Human Capital, and the Structure of Wages in Japan, *The International Journal of Human Resource Management* 22(15): 3036-3050.
- Schank, T., C. Schnabel, and J. Wagner (2007). Do Exporters Really Pay Higher Wages? First Evidence from German Linked Employer-Employee data, *Journal of International Economics* 72: 52–74.
- (2010). Higher Wages in Exporting Firms: Self-Selection, Export Effect, or Both? First Evidence from Linked Employer-Employee Data, *Review of World Economics*, 146, 303– 322.
- Seguino, S. (2000). The effects of structural change and economic liberalization on gender wage differentials in South Korea and Taiwan, *Cambridge Journal of Economics* 24: 437-459.
- Siegel, J., L. Pyun, and B. Y. Cheon (2019). Multinational Firms, Labor Market Dis- crimination, and the Capture of Outsider's Advantage by Exploiting the Social Divide, *Administrative Science Quarterly* 64 (2): 370–397.
- Tang, H. and Y. Zhang (2017). Do Multinationals Transfer Culture? Evidence on Female Employment in China, Technical Report, CESifo Group Munich.
- Vahter, P. and J. Masso (2019). "The contribution of multinationals to wage inequality: foreign ownership and the gender pay gap", *Review of World Economics* 155: 105-148.

Verhoogen, E. A. (2008). Trade, Quality Upgrading, and Wage Inequality in the Mexican Manufacturing Sector, *The Quarterly Journal of Economics*, 123, 489–530.

World Economic Forum, The Global Gender Gap Report 2020, 2012.

Table 1: Summary statistics by firm type

50% criteria														
	N. of Firms	l.	N. of Plant	s	N. of Emp	oloyees	N. Female	Empl	Hourly Wa	ge (100 JPY	Y)			
Firm type	Freq.	Percent	Freq.	Percent	Freq.	Percent	Freq.	Percent	Mean	Median	St Dev 1	Min	М	lax
Domestic-only	6,485	81.8	6,706	76.8	115,931	61.4	30,413	70.3	17.7	15.7	8.6		6.5	171.6
Exporter	437	5.5	478	5.5	12,322	6.5	2,656	6.1	21.5	19.4	10.1		6.6	140.0
JMNE	971	12.2	1,504	17.2	58,797	31.1	9,885	22.9	27.6	25.0	13.4		6.6	162.3
FMNE	39	0.5	45	0.5	1,737	0.9	306	0.7	31.7	27.7	16.5		7.7	144.1
Greenfield-born	20	0.3	21	0.2	868	0.5	156	0.4	32.6	28.5	16.7		8.3	128.2
M&A-born	19	0.2	24	0.3	869	0.5	150	0.4	30.8	26.8	16.2		7.7	144.1
Total	7,932	100	8,733	100	188,787	100	43,260	100	21.1	18.3	11.5		6.5	171.6
10% criteria														
	N. of Firms		N. of Plant	s	N. of Emp	oloyees	N. Female	Empl	Hourly Wa	ge (100 JPY	Y)			
Firm type	Freq.	Percent	Freq.	Percent	Freq.	Percent	Freq.	Percent	Mean	Median	St Dev 1	Min	М	lax
Domestic-only	6,475	81.6	6,692	76.6	115,285	61.1	30,327	70.1	17.6	15.7	8.6		6.5	171.6
Exporter	429	5.4	469	5.4	12,034	6.4	2,595	6.0	21.5	19.3	10.1		6.6	140.0
JMNE	836	10.5	1,142	13.1	40,591	21.5	7,470	17.3	25.8	23.2	12.5		6.6	162.3
FMNE	192	2.4	430	4.9	20,877	11.1	2,868	6.6	31.3	28.8	14.5		7.4	161.7
Greenfield-born	30	0.4	43	0.5	2,070	1.1	338	0.8	29.8	26.7	14.5		7.4	128.2
M&A-born	162	2.0	387	4.4	18,807	10.0	2,530	5.9	31.5	29.0	14.5		7.4	161.7
Total	7,932	100	8,733	100	188,787	100	43,260	100	21.1	18.3	11.5		6.5	171.6

Notes: 50% criteria and 10% criteria refer to the minimum foreign equity shares used for defining FMNEs relative to other firm types.

	(1)	(2)	(3)	(4)	(5)
			In Hourly Wa	ige	
Female	-0.416***	-0.320***	-0.304***	-0.312***	-0.312***
	[0.005]	[0.004]	[0.004]	[0.004]	[0.004]
Japanese-Owned Exporters (50%)				0.028**	0.028**
without Foreign Subsidiaries				[0.011]	[0.011]
Female * Japanese-Owned Exporters (50%)				0.041***	0.041***
				[0.011]	[0.011]
JMNE (50%)				0.071***	0.071***
				[0.009]	[0.009]
Female * JMNE (50%)				0.026***	0.026***
				[0.008]	[0.008]
FMNE (50%)				0.221***	
				[0.038]	
Female * FMNE (50%)				0.097**	
				[0.040]	
Greenfield-born FMNE (50%)					0.304***
					[0.046]
Female * Greenfield-born FMNE (50%)					0.151***
					[0.035]
M&A-born FMNE (50%)					0.137**
					[0.053]
Female * M&A-born FMNE (50%)					0.037
					[0.046]
Worker characteristics included	No	Yes	Yes	Yes	Yes
Other plant-level characteristics included	No	No	Yes	Yes	Yes
Observations	200,230	192,917	192,891	182,094	182,094
R-squared	0.299	0.646	0.707	0.712	0.712

Table 2: Wage regressions using 50% foreign ownership criteria and firm-type dummy variables

Notes: The "50%" indicators are a reminder that all firm-types are defined using that criteria for identifying FMNE. Standard errors, clustered at the firm level, appear in brackets. Industry and region fixed effects are suppressed. \*\*\*, \*\*, and \* indicate statistical significance at the 1%, 5% and 10% levels, respectively.

Table 3: Wage regressions using 50% foreign ownership criteria and firm-type continuous variables $(1)$ $(2)$ $(2)$ $(4)$ $(5)$									
	(1)	(2)	(3) In Hourly Wa	(4)	(5)				
Female	-0.310***	-0.311***	-0.312***	-0.312***	-0.309***				
	[0.004]	[0.004]	[0.004]	[0.004]	[0.004]				
Japanese-Owned Exporters (50%)		0.028**	0.029**	0.028**					
without Foreign Subsidiaries		[0.011]	[0.011]	[0.011]					
Formala * Jananasa Ownad Exportana (50%)		0.039***	0.040***	0.041***					
Female * Japanese-Owned Exporters (50%)		[0.011]	[0.011]	[0.011]					
IMANIE (500/)	0.062***		0.060***	0.070***					
JMNE (50%)	[0.009]		[0.009]	[0.009]					
			0.005444	0.00 (****					
Female * JMNE (50%)	0.016* [0.009]		0.025*** [0.009]	0.026*** [0.008]					
			[0.007]	[0.000]					
FMNE (50%)	0.210*** [0.038]	0.213*** [0.037]							
	[0.038]	[0.037]							
Female * FMNE (50%)	0.090**	0.086**							
	[0.040]	[0.043]							
Export Share	0.045*				0.013				
	[0.027]				[0.026]				
Female * Export Share	0.068**				0.060**				
	[0.029]				[0.028]				
In N. of Foreign Subsidiaries		0.030***			0.028***				
		[0.003]			[0.003]				
Female * InN_Foreign_Subsidiaries		0.011***			0.008***				
Female Iniv_Foldgn_Subsidiants		[0.003]			[0.003]				
			0.002***						
Foreign Ownership Ratio (FOR)			0.003*** [0.000]						
Female * Foreign Ownership Ratio			0.001* [0.000]						
			[0.000]						
Greenfield-born FMNE FOR				0.003***	0.003***				
				[0.001]	[0.001]				
Female * Greenfield-born FMNE FOR				0.002***	0.001***				
				[0.000]	[0.000]				
M&A-born FMNE FOR				0.002**	0.001**				
				[0.001]	[0.001]				
Female * M&A-born FMNE FOR				0.000	0.000				
				[0.001]	[0.001]				
Worker characteristics included	Yes	Yes	Yes	Yes	Yes				
Other plant-level characteristics included Observations	Yes 182,094	Yes 182,094	Yes 182,094	Yes 182,094	Yes 182,094				
R-squared	0.712	0.713	0.713	0.712	0.713				

Notes: The "50%" indicators are a reminder that all firm-types are defined using that criteria for identifying FMNE. Standard errors, clustered at the firm level, appear in brackets. Industry and region fixed effects are suppressed. \*\*\*, \*\*, and \* indicate statistical significance at the 1%, 5% and 10% levels, respectively.

Table 3: Wage regressions using 50% foreign ownership criteria and firm-type continuous variables

Dependent variable = Ln Hourly Wage					
	q10	q25	q50	q75	q90
Female	-0.289***	-0.303***	-0.306***	-0.307***	-0.304***
	[0.005]	[0.005]	[0.004]	[0.003]	[0.003]
Japanese-Owned Exporters (50%)	0.043***	0.028***	0.026***	0.029***	0.036***
without Foreign Subsidiaries	[0.004]	[0.001]	[0.002]	[0.001]	[0.005]
Female * J-Owned Exporters (50%)	0.041***	0.046***	0.045***	0.046***	0.025**
	[0.003]	[0.009]	[0.005]	[0.007]	[0.010]
JMNE (50%)	0.065***	0.063***	0.067***	0.075***	0.079***
	[0.002]	[0.004]	[0.005]	[0.003]	[0.002]
Female * JMNE (50%)	0.011***	0.024***	0.023***	0.027***	0.033***
	[0.004]	[0.005]	[0.006]	[0.003]	[0.011]
Greenfield-born FMNE (50%)	0.239***	0.281***	0.277***	0.289***	0.348***
	[0.006]	[0.004]	[0.005]	[0.009]	[0.035]
Female * Greenfield-born FMNE (50%)	0.128**	0.165***	0.198***	0.153***	0.063**
	[0.057]	[0.057]	[0.033]	[0.019]	[0.025]
M&A-born FMNE (50%)	0.114***	0.115***	0.113***	0.148***	0.186***
	[0.011]	[0.031]	[0.022]	[0.027]	[0.032]
Female * M&A-born FMNE (50%)	0.036***	0.028	0.038	0.019	0.043***
	[0.013]	[0.041]	[0.043]	[0.047]	[0.006]
Worker characteristics included	Yes	Yes	Yes	Yes	Yes
Other plant-level characteristics included	Yes	Yes	Yes	Yes	Yes
Pseudo R-squared	0.4434	0.4737	0.4844	0.4849	0.473
Observations					182,094

Table 4: Wage regressions using 50% foreign ownership criteria and firm-type dummy variables, QR results Dependent variable = Ln Hourly Wage

Notes: The "50%" indicators are a reminder that all firm-types are defined using that criteria for identifying FMNE. Standard errors, clustered at the firm level, appear in brackets. Industry and region fixed effects are suppressed. \*\*\*, \*\*, and \* indicate statistical significance at the 1%, 5% and 10% levels, respectively.

	(1)	(2)	(3)	(4)	(5)
			In Hourly Wa	age	<u> </u>
Female	-0.416***	-0.320***	-0.304***	-0.312***	-0.311***
	[0.005]	[0.004]	[0.004]	[0.005]	[0.004]
Japanese-Owned Exporters (10%)				0.029**	0.028**
without Foreign Subsidiaries				[0.012]	[0.012]
Female * Japanese-Owned Exporters (10%)				0.039***	0.039***
				[0.011]	[0.011]
JMNE (10%)				0.068***	0.066***
				[0.009]	[0.009]
Female * JMNE (10%)				0.025**	0.025***
				[0.010]	[0.010]
FMNE (10%)				0.111***	
				[0.016]	
Female * FMNE (10%)				0.044***	
				[0.012]	
Greenfield-born FMNE (10%)					0.224***
					[0.035]
Female * Greenfield-born FMNE (10%)					0.086***
					[0.031]
M&A-born FMNE (10%)					0.097***
					[0.017]
Female * M&A-born FMNE (10%)					0.035***
					[0.013]
Worker characteristics included	No	Yes	Yes	Yes	Yes
Other plant-level characteristics included	No	No	Yes	Yes	Yes
Observations	200,230	192,917	192,891	182,094	182,094
R-squared	0.299	0.646	0.707	0.711	0.712

Table 5: Wage regressions using 10% foreign ownership criteria and firm-type dummy variables

Notes: The "10%" indicators are a reminder that all firm-types are defined using that criteria for identifying FMNE. Standard errors, clustered at the firm level, appear in brackets. Industry and region fixed effects are suppressed. \*\*\*, \*\*, and \* indicate statistical significance at the 1%, 5% and 10% levels, respectively.

Table 6: Wage regressions using 10% foreign			(3)		
	(1)		(5) In Hourly Wa	(4)	(5)
Female	-0.310***	-0.310***	-0.311***	-0.311***	-0.308***
remaie	[0.004]	[0.004]	[0.004]	[0.004]	[0.004]
Japanese-Owned Exporters (10%)		0.023**	0.024**	0.023**	
without Foreign Subsidiaries		[0.012]	[0.012]	[0.012]	
Female * Japanese-Owned Exporters (10%)		0.037*** [0.011]	0.039*** [0.011]	0.040*** [0.011]	
JMNE (10%)	0.060*** [0.009]		0.057*** [0.009]	0.057*** [0.009]	
Female * JMNE (10%)	0.017 [0.011]		0.026*** [0.010]	0.026*** [0.010]	
FMNE (10%)	0.101*** [0.017]	0.029* [0.017]			
Female * FMNE (10%)	0.033*** [0.012]	0.022 [0.016]			
Export Share	0.038 [0.027]				0.014 [0.026]
Female*Export Share	0.060** [0.028]				0.062** [0.027]
In N. of Foreign Subsidiaries		0.024*** [0.004]			0.025*** [0.004]
Female*InN_Foreign_Subsidiaries		0.009*** [0.004]			0.007** [0.003]
Foreign Ownership Ratio (FOR)			0.003*** [0.000]		
Female*Foreign Ownership Ratio			0.001** [0.000]		
Greenfield-born FMNE FOR				0.003*** [0.001]	0.003*** [0.001]
Female*Greenfield-born FMNE FOR				0.001*** [0.000]	0.001*** [0.000]
M&A-bom FMNE FOR				0.003*** [0.000]	0.001** [0.000]
Female*M&A-born FMNE FOR				0.001 [0.000]	-0.000 [0.000]
Worker characteristics included	Yes	Yes	Yes	Yes	Yes
Other plant-level characteristics included	Yes	Yes	Yes	Yes	Yes
Observations	182,094	182,094	182,094	182,094	182,094
R-squared	0.711	0.711	0.713	0.712	0.713

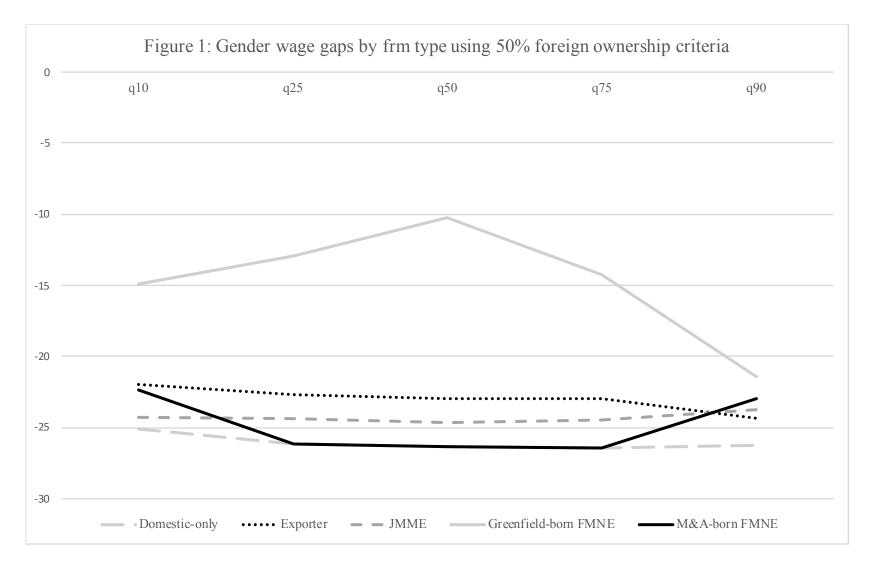
Notes: The "10%" indicators are a reminder that all firm-types are defined using that criteria for identifying FMNE. Standard errors, clustered at the firm level, appear in brackets. Industry and region fixed effects are suppressed. \*\*\*, \*\*, and \* indicate statistical significance at the 1%, 5% and 10% levels, respectively.

Dependent variable = Ln Hourly Wage					
	q10	q25	q50	q75	q90
Female	-0.289***	-0.302***	-0.305***	-0.306***	-0.305***
	[0.002]	[0.004]	[0.003]	[0.002]	[0.004]
Japanese-Owned Exporters (10%)	0.041***	0.027***	0.027***	0.030***	0.038***
without Foreign Subsidiaries	[0.007]	[0.004]	[0.001]	[0.003]	[0.004]
Female * J-Owned Exporters (10%)	0.041***	0.046***	0.043***	0.041***	0.022
	[0.001]	[0.012]	[0.006]	[0.016]	[0.017]
JMNE (10%)	0.063***	0.061***	0.062***	0.067***	0.074***
	[0.002]	[0.001]	[0.000]	[0.000]	[0.003]
Female * JMNE (10%)	0.010***	0.019***	0.024***	0.032***	0.032***
	[0.003]	[0.003]	[0.005]	[0.005]	[0.009]
Greenfield-born FMNE (10%)	0.206***	0.204***	0.199***	0.191***	0.198***
	[0.011]	[0.010]	[0.007]	[0.004]	[0.008]
Female * Greenfield-born FMNE (10%)	0.017	0.007	0.081*	0.158***	0.111***
	[0.011]	[0.036]	[0.044]	[0.031]	[0.012]
M&A-born FMNE (10%)	0.096***	0.089***	0.095***	0.100***	0.104***
	[0.004]	[0.003]	[0.005]	[0.007]	[0.004]
Female * M&A-born FMNE (10%)	0.020***	0.032***	0.029***	0.035**	0.045**
· · ·	[0.007]	[0.009]	[0.009]	[0.016]	[0.022]
Worker characteristics included	Yes	Yes	Yes	Yes	Yes
Other plant-level characteristics included	Yes	Yes	Yes	Yes	Yes
Pseudo R-squared	0.4439	0.474	0.4845	0.4848	0.4725
Observations					182,094

Table 7: Wage regressions using 10% foreign ownership ratio criteria and firm-type dummy variables, QR results

Dependent variable = Ln Hourly Wage

Notes: The "10%" indicators are a reminder that all firm-types are defined using that criteria for identifying FMNE. Standard errors, clustered at the firm level, appear in brackets. Industry and region fixed effects are suppressed. \*\*\*, \*\*, and \* indicate statistical significance at the 1%, 5% and 10% levels, respectively.



Notes: Based on authors' QR results in Table 4.

#### APPENDIX TABLES AND FIGURE NOT FOR PUBLICATION AVAILABLE ONLINE AS INDICATED IN THE MANUSCRIPT

## Appendix Figure A1: Data flow chart

Basic Survey on Wage Structure (2012)

https://www.mhlw.go.jp/english/database/db-l/wage-structure.html

- Annual sample survey. Repeated cross-sectional data.
- All industries (1,276,813 workers in 57,078 establishments). Manufacturing (272,624 workers in 10,604 establishments)
- We have 2009—2012 data but matching with the two censuses is successful for 2012 data only.

## Matching using the establishment ID of the 2009 Economic Census for Business Frame

Perfect match For 2012 data	<ul> <li>2009 Economic Census for Business Frame</li> <li>https://www.stat.go.jp/english/data/e-census/index.html</li> <li>First complete census, followed by the 2014 one.</li> <li>No. of foreign subsidiaries. Foreign ownership ratio.</li> <li>Establishment-Firm concordances. 7,271,801 establishments of 6,104,514 firms in all industries.</li> </ul>
	2012 Economic Census for Business Activity
9,564 of 10,604	https://www.stat.go.jp/english/data/e-census/2012/index.html
establishments	First complete census, followed by the 2016 one.
and 243,173 of	• The response rate is over 80%. It covers establishments' and firms' activites during 2011.
272,624 workers	Questionnaires vary across industries. We use manufacturing industries to merge the data on
are matched.	exports.
(89% of workers,	<ul> <li>440,562 establishments of 398,608 firms in manufacturing.</li> </ul>
90% of plants.)	

## The merged data (2012, manufacturing)

 Restricted to full-time workers under 60 years old in private manufacturing plants only. Exclude outliers whose wage exceeds average plus 10 standard deviations. Exclude outliers whose wage is below minimum wage at prefecture level (1,079 workers).

→ 200,230 workers (154,067 male + 46,163 female workers) in 9,420 establishments of 8,614 firms in manufacturing.
 188,787 workers (145,527 male + 43,260 female workers) in firms with data on no. of foreign subsidiaries and exports.

#### Appendix Table A1: Summary statistics by worker gender

		Male Wor	rkers			Female We	orkers	
_	Mean	St Dev	Min	Max	Mean	St Dev	Min	Max
Ln Wage	3.030	0.458	1.878	5.145	2.564	0.402	1.875	5.019
Domestic-only	0.555	0.497	0	1	0.659	0.474	0	1
Exporter	0.063	0.242	0	1	0.058	0.233	0	1
JMNE	0.335	0.472	0	1	0.228	0.420	0	1
FMNE (50% criteria)	0.010	0.099	0	1	0.007	0.084	0	1
Jr High School	0.039	0.194	0	1	0.050	0.217	0	1
High School	0.635	0.481	0	1	0.709	0.454	0	1
Jr College	0.083	0.275	0	1	0.145	0.352	0	1
College	0.244	0.429	0	1	0.096	0.295	0	1
Experience	20.477	10.983	0	44	21.673	11.668	0	44
Tenure	13.474	10.819	0	44	10.699	9.331	0	44
Employee for an Indefinite Period	0.943	0.232	0	1	0.814	0.389	0	1
White Collar	0.363	0.481	0	1	0.395	0.489	0	1
Blue Collar	0.613	0.487	0	1	0.582	0.493	0	1
In Plant Size (In N. of Employees)	4.917	1.558	1.609	9.468	4.586	1.379	1.609	9.468
Headquarters	0.252	0.434	0	1	0.266	0.442	0	1
In Plant Sales per Worker	7.213	2.398	0	12.407	6.798	2.324	0	12.407
Observations	154,067				46,163			

Notes: Employee for an Indefinite Period indicates a long-term or "lifetime" employee of the firm.

Baseline GWG Significant differentials from baseline (percentage points)								
					Greent	field-born	M&A-born	
FMNE criteria	Quintile	Domestic-only	Exporter	JMME	FMNE	,	FMNE	
50% criteria	All	-26.8		3.1	1.9	11.9		0.0
50% criteria	q10	-25.1		3.1	0.8	10.2		2.7
50% criteria	q25	-26.1		3.5	1.8	13.2		0.0
50% criteria	q50	-26.4		3.4	1.7	16.1		0.0
50% criteria	q75	-26.4		3.5	2.0	12.2		0.0
50% criteria	q90	-26.2		1.9	2.5	4.8		3.2
10% criteria	All	-26.8		2.9	1.9	6.5		2.6
10% criteria	q10	-25.1		3.1	0.8	0.0		1.5
10% criteria	q25	-26.1		3.5	1.4	0.0		2.4
10% criteria	q50	-26.3		3.2	1.8	6.2		2.2
10% criteria	q75	-26.4		3.1	2.4	12.6		2.6
10% criteria	q90	-26.3		0.0	2.4	8.7		3.4

Appendix Table A2: Gender wage gaps and differentials by firm type and wage quintile

Notes: Based on estimated coefficients in Tables 2, 5 and 7; zero indicates no significant difference.