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Trade Sustainability in the Forestry Domain: Evidence from Malaysia using Johansen and Bound Test Method

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Abstract

This study was undertaken to explore the relationship between export and import, in the category of Forestry domain for Malaysia, which includes sub domain (1) industrial roundwood; (2) wood pulp; (3) wood fuel; (4) paper and paper board; (5) sawn wood; (6) recovered paper and (7) wood base panel. To test the robustness, Johansen (1991) cointegration method and ARDL approach were employed and the period of the study covers annual data from 1961 to 2007. The results clearly show that the export and import of forestry domain is highly cointegrated. This shows trade is indeed sustainable in the domain of forestry and a sign of good trade policies adapted and implemented by the relevant authorities. Bi-directional granger causality could also be detected based on VECM (vector-error correction model) method. Imports seem to positively and significantly affect exports, both in the long run and short run, *vice versa*. The results are robust when we employed the Bounds test as well.

Key Words: *Johansen cointegration test, forestry trade, VECM, ARDL*

Introduction

The forest products sector is estimated to contribute about one percent of world gross domestic product and account for three percent of international merchandise trade. As for the case of Malaysia, we produced an estimated 26.4 million cu m (932 million cu ft) of round wood from a forest area of 19.3 million ha (47.7 million acres) in 2000. About 32% of the forest area is located in Peninsular Malaysia, 22% in Sabah, and 46% in Sarawak. Exports of timber products in 2000 amounted to \$3.8 billion, or 4.2% of total exports. Malaysia is the world's third leading producer (after Brazil and Slovakia) of veneer sheets, accounting for 7% of global production in 2000. While, forest products imports by Malaysia are dominated by paper products. In 1994 paper and paperboard comprised 83 percent of Malaysia's forest products imports. Malaysia, likewise Indonesia, is a major producer of hardwood forest products.

Trade sustainability studies are gaining momentum all over the world, due to the intrinsic information obtained from the empirical studies are able to imply the policy suitability. Though numerous studies have flourished on the external imbalances in the current account, most of these studies concentrated on aggregated trade imbalances, seldom have researcher's concentrated on the trade imbalances of a particular domain as what we are trying to attempt in this study, whereby we would like to focus on the trade imbalances in the category of Forestry

domain. More precisely we intend to investigate the existence of any co-integrating factors between the export and import of the above category, if any. In the study of sustainability of external imbalances for 22 Least Developed Countries (LDC), Tang and Smyth (2008) found that the macroeconomic policies have been effective in causing the exports and imports to converge in the long run. It is a well known fact that much of the discussion in the empirical literature on external imbalances has focused on their sustainability.

The empirical literature that has examined the sustainability of external imbalances has adopted one of the two major empirical approaches. The first approach is to examine whether exports and imports are cointegrated, the same approach that we intend to employ in this study. The second approach to examine the sustainability of external imbalances is to test for a unit root in the current account. Using this method, if the current account or trade balance is found to be stationary, this implies that external imbalances are sustainable and that the international budget constraint is satisfied. If, however, the trade balance is non-stationary this is either a reflection of bad policies or indicative of the productivity gap hypothesis. To this point, the results in the empirical literature on the sustainability of external imbalances are mixed. Among the notable studies that followed the first approach are as those of Husted (1992), Bahmani-Oskooee (1994), Fountas and Wu (1999) and Irandoust and Ericsson (2004).

Husted (1992) and Fountas and Wu (1999) used quarterly data for the periods 1967-1989 and 1967-1994 respectively, to examine whether exports and imports are cointegrated in the United States. While Fountas and Wu (1999) found no long-run relationship, Husted (1992) found that exports and imports were cointegrated. Employing quarterly data from 1971-1997, Irandoust and Ericsson (2004) found that exports and imports were cointegrated in Germany, Sweden and the United States. Bahmani-Oskooee (1994) found that Australian exports and imports are cointegrated, and further concluded that the cointegrating coefficient was close to one, implying that Australia's external account is sustainable. Using quarterly data, Bahmani-Oskooee and Rhee (1997) found that South Korea's exports and imports are cointegrated and the coefficient on exports was positive. This result implies that South Korea does not violate its international budget constraint.

Employing the Gregory and Hansen (1996) cointegration test to examine the sustainability of external imbalances in Indonesia, Malaysia, the Philippines and Thailand, using annual data for the period, 1961-1999, Baharumshah et al. (2003) found that exports and imports are cointegrated in Indonesia, the Philippines, and Thailand, but not in Malaysia. In contrary, Tang (2002), employing bounds test for cointegration developed by Pesaran et al. (2001) to annual data for the period 1968-1998 (1974-1998 for Singapore) found cointegration between exports and imports for Malaysia and Singapore, but not for the Philippines, Indonesia and Thailand. Narayan and Narayan (2004) similarly employed the bounds test to examine long-run relationship between exports and imports in Fiji and Papua New Guinea and found that while exports and imports were cointegrated in both the countries, the coefficient on exports was equal to one, for only Fiji.

The purpose of this paper is to investigate the relationship between the export and import in the category of forestry domain which includes subdomains (1) industrial roundwood; (2) wood pulp; (3) wood fuel; (4) paper and paper board; (5) sawn wood; (6) recovered paper and (7) wood base panel. Johansen (1991) cointegration method was chosen to explore the probable cointegration. Vector error correction model (VECM) would be employed in the later stage to investigate the

dynamic and long run relationship between these variables. Granger causality test based on VECM will also be conducted. Then the policy implication will be proposed which will benefit the sector itself and also the country.

Data and Methodology

The data used in this analysis are gathered from the Food and Agriculture Organization of The United Nations Statistical Database. The data was downloaded from the website faostat.fao.org. This study only focused on trade data for forestry domain namely, (1) industrial roundwood; (2) wood pulp; (3) wood fuel; (4) paper and paper board; (5) sawn wood; (6) recovered paper and (7) wood base panel in Malaysia over the period 1961 to 2007.

In empirical economics macroeconomic variables comprises of non stationary series. Treating non stationary variables in empirical analysis is important so that the results of spurious regression can be avoided. According to the concept of cointegration, two or more non-stationary time series share a common trend, then they are said to be cointegrated. The theoretical framework highlighted are expressed as follows: the component of the vector $Y_t = (y_{1t}, y_{2t}, \dots, y_{nt})'$ are considered to be cointegrated of order d, b , denoted $Y_t \sim CI(d, b)$ if (i) all the component Y_t are stationary after n difference, or integrated of order d and noted as $Y_t \sim I(d)$. (ii) presence of a vector $\hat{a} = (\hat{a}_1, \hat{a}_2, \dots, \hat{a}_n)$ in such that linear combination $\hat{a}Y_t = \hat{a}_1 y_{1t} + \hat{a}_2 y_{2t} + \dots + \hat{a}_n y_{nt}$ whereby the vector \hat{a} is named the cointegrating vector. A few major characteristics of this model are that the cointegration relationship obtained indicates a linear combination of non-stationary variables, in which all variables must be integrated of the same order and lastly if there are n series of variables, there may be as many as $n-1$ linearly independent cointegrating vectors.

Johansen's (1991) cointegration test is adopted to determine whether the linear combination of the series possesses a long-run equilibrium relationship. The numbers of significant cointegrating vectors in non-stationary time series are tested by using the maximum likelihood based $\tilde{\epsilon}$ trace and $\tilde{\epsilon}$ max statistics introduced by Johansen and Juselius (1990). The advantage of this test is that it utilizes test statistic that can be used to evaluate cointegration relationship among a group of two or more variables. Therefore, it is a superior test as it can deal with two or more variables that may be more than one cointegrating vector in the system.

Prior to testing for the number of significant cointegrating vectors, the likelihood ratio (LR) tests are performed to determine the lag length of the vector autoregressive system. In the Johansen procedure, following a vector autoregressive (VAR) model, it involves the identification of rank of the $n \times n$ matrix Φ in the specification given by:

$$\Delta Y_t = \delta + \sum_{i=1}^{k-1} \Gamma_i \Delta Y_{t-i} + \Pi Y_{t-k} + \varepsilon_t \quad (1)$$

whereby Y_t is a column vector of the n variables, Δ is the difference operator, \tilde{A} and Φ are the coefficient matrices, k denotes the lag length and \tilde{a} is a constant. In the absence of cointegrating vector, Φ is a singular matrix, which means that the cointegrating vector rank is equal to zero. On the other hand, in a cointegrated scenario, the rank of Φ could be anywhere

between zero. In other words, the Johansen cointegration test can determine the number of cointegrating equation and this number is named the *cointegrating rank*.

The Johansen Maximum likelihood test provides a test for the rank of \mathfrak{D} , namely the trace test ($\hat{\epsilon}$ trace) and the maximum eigenvalue test ($\hat{\epsilon}$ max). Firstly, the $\hat{\epsilon}$ trace statistic test whether the number of cointegrating vector is zero or one. Then, the $\hat{\epsilon}$ max statistic test whether a single cointegration equation is sufficient. Both test statistics are given as follows:

$$\lambda_{\text{trace}}(r) = -T \sum_{i=r+1}^p \ln(1 - \hat{\lambda}_i) \quad (2)$$

$$\lambda_{\text{trace}}(r, r+1) = -T \sum_{i=r+1}^p \ln(1 - \hat{\lambda}_{r+1}) \quad (3)$$

where p is the number of separate series to be analysed, T is the number of usable observations and $\hat{\epsilon}$ is the estimated eigenvalues.

Now the newly proposed autoregressive distributed lag (ARDL) approach (Pesaran and Shin, 1995, 1998; Pesaran *et al.*, 1996; Pesaran *et al.*, 2001). More recent studies have indicated that the ARDL approach to cointegration is preferable to other conventional cointegration approaches such as Engle and Granger (1987), and Gregory and Hansen (1996). One of the reasons for preferring the ARDL is that it is applicable irrespective of whether the underlying regressors are purely $I(0)$, purely $I(1)$ or mutually co-integrated. The statistic underlying this procedure is the familiar Wald or F-statistic in a generalized Dickey-Fuller type regression, which is used to test the significance of lagged levels of the variables under consideration in a conditional unrestricted equilibrium error correction model (ECM) (Pesaran, *et al.*, 2001). Another reason for using the ARDL approach is that it is more robust and performs better for small sample sizes (such as in this study) than other cointegration techniques.

The ARDL approach to cointegration (Pesaran *et al.*, 2001) involves estimating the conditional error correction (EC) version of the ARDL model to estimate the hypothesis is:

$$\Delta \ln y_t = \delta_0 + \lambda_1 \ln y_{t-1} + \lambda_2 \ln x_{t-1} + \sum_{i=1}^{\rho} \theta_i \ln y_{t-i} + \sum_{i=0}^{\rho} \phi_i \ln x_{t-i} + \varepsilon_t \quad (1)$$

whereby λ_1 and λ_2 are long-run multipliers and δ_0 is the drift. Lagged values of Δy_t and current and lagged values of Δx_t are used to model the short-run dynamic structure. All variables in natural logarithm and ρ is the optimal lag length.

The F -test is conducted to test the existing of the long-run relationship. The alternative hypothesis for a cointegration among variables in equation (1) is $H_1 : \lambda_1 \neq \lambda_2 \neq 0$ against the null hypothesis $H_0 : \lambda_1 = \lambda_2 = 0$. The F -test has a non-standard distribution which depends on (i) the number of regressors, (ii) whether variables included in the model are $I(1)$ or $I(0)$ and (iii) whether the model consists and intercept and/or a trend. Given a relatively small sample size in this study of 40 observations, the critical values used are as reported by Narayan (2005) which are based on a small sample size between 30 and 80.

If the F -test statistic exceeds their respective upper critical values, we can conclude that there is enough evidence of a long-run relationship between the variables. If the F -test statistic is below the upper critical values, we failed to reject the null hypothesis of no cointegration and if it lies between the bounds, a conclusive inference cannot be made without knowing the order of integration of the underlying regressors. If there is a cointegration, the following long-run model is estimated:

$$\ln y_t = \beta_1 + \delta \ln y_{t-1} + \phi \ln x_{t-1} + v_t \quad (2)$$

The orders of the lags in the ARDL model are selected by either the Schwarz Bayesian criterion (SBC) or the Akaike information criterion (AIC), before the selected model is estimated by ordinary least squares (OLS).

The ARDL specification of the short-run dynamics can be derived by constructing an error correction model (ECM) of the following form:

$$\Delta \ln y_t = \alpha_1 + \sum_{i=1}^p \lambda_{2i} \Delta \ln y_{t-i} + \sum_{i=0}^p \omega_{2i} \Delta \ln x_{t-i} + \psi ECM_{t-1} + \xi_t \quad (3)$$

whereby ECM_{t-1} is the error correction term, defined as:

$$ECM_t = \ln y_t - \beta_1 - \delta \ln y_{t-1} - \phi \ln x_{t-1} \quad (4)$$

All coefficients of the short-run equation are coefficients relating to the short-run dynamics of the model's convergence to equilibrium and ψ represent the speed of adjustment.

Vector Error Correction Model (VECM)

VECM is a restricted VAR designed for use with non stationary variables that are known to be co-integrated. VECM specification restricts the long run behaviour of the endogenous variables to converge to their co-integrating relationships while allowing for short-run adjustment dynamics. Engle and Granger (1987) showed that if the variables, say X_t and Y_t is found to be cointegrated, there will be an error representatives which is linked to the said equation, which gives the implication that changes in dependent variable is a function of the imbalance in cointegration relation (represented by the error correction term) and by other explanatory variables. Intuitively, if X_t and Y_t have the same stochastic trend, current variables in X_t (dependent variable) is in part, the result of X_t moving in line with trend value of Y_t (independent variable). Through error correction term, VECM allows the discovery of Granger Causality relation which has been abandoned by Granger (1968) and Sims (1972). The VAR constraint model may derive a VECM model as shown below:

$$\Delta y_t = \delta_0 + \sum_{i=1}^p \theta_i \Delta y_{t-i} + \sum_{i=1}^p \phi_i \Delta x_{t-i} + \varepsilon_t$$

where y_t – in the form of $n \times 1$ vector

$\hat{\theta}_i$ and $\hat{\phi}_i$ – estimated parameters

Δ – difference operator

$\hat{a}t$ – reactional vector which explains unanticipated movements in

Y_t and x (error correction term)

In the Granger causality test, the degree of exogeneity can be identified through the t test for the lagged error correction term ($\hat{\delta}_i$), or F test applied to the lags of the coefficients of each variable separately of the non dependent variable ($\hat{\epsilon}_i$). In addition to the above, VECM method allows the differentiation of the short term and long term relationship. Error term with lagged parameter (ECT ($\epsilon_{1, t-1}$)) is an adaptive parameter where it measures the short term dispersal from long term equilibrium. In the short term, the variables may disperse from one another which will cause in-equilibrium in the system. Hence, the statistical significance of the coefficients associated with ECT provides evidence of an error correction mechanism that drives the variables back to their long-run relationship.

Empirical Result

Table 1 reports the results of Augmented Dickey-Fuller (ADF) and Phillips-Perron (PP) unit root test that describes the stationary properties of the variables (export and import) in Malaysia. Schwartz Information Criterion (SIC) is used to select the optimal truncation lag length to ensure the errors are white noise in ADF. The both results clearly show that the null hypothesis of unit root test for all variables fail to be rejected in the level form but only appear stationary in the first different. In other words, all variables are said to be integrated of order one, which is $I(1)$.

Table 1: Results of the Unit Root Tests

Variables	ADF		PP	
	Level (Trend and Intercept)	First Different (Trend and Intercept)	Level (Trend and Intercept)	First Different (Trend and Intercept)
Export	-1.5571 (0)	-6.9264*** (1)	-1.4166	-8.4279***
Import	-2.6017 (0)	-7.8283*** (0)	-2.5829	-7.8325***

Note: *** denotes significant at 1% level. Figures in parenthesis () refer to the selected length. The lag length was arbitrarily selected using SIC.

Since all the variables are integrated of same order, the cointegration test is conducted to determine the presence of a long run equilibrium relationship between export and import in the categories of forestry domain. The results of Johansen cointegration test is reported in Table 2. The results indicate that null hypothesis of no cointegration can be rejected in the period. Therefore, it can be concluded that there is positive and significant long run relationship at one percent level, implying that there is common trend exists within import and export.

Table 2: Results of Cointegration Test (Based on Johansen test)

Vectors	r=0	r≤1
Trace test	20.38	4.37
1% level	20.04	6.65
Max-Eigen test	16.01	4.37
1% level	18.63	6.65

Note: r indicates the number of cointegrating vectors. Trace and Max-Eigen denote the trace statistic and maximum eigenvalue statistic. The critical values obtained from Osterwald-Lenum (1992). Lag selection (k) is based on Schwert (1987) formula.

Table 2a shows the result for cointegration test based on ARDL approach. There is significant cointegrating relationship between export and import at 5 per cent level.

Table 2a: Results of Cointegration Test (Based on ARDL Approach)

Long-run specification (<i>F-test</i>)	Long-run specification (<i>F-test</i>)
F (Im Ex) = 2.59	F (Ex Im) = 5.28**

Note: The 10% critical values for I(0) and I(1) are 3.190 and 3.730. The corresponding 5% and 1% critical values are 3.877, 4.460, and 5.607 and 6.193 respectively. The critical values are obtained from Narayan (2005).

Table 3 and 3a report the results of long-run elasticities among import and export based on Johansen and ARDL approach. The results clearly show that significant long-run relationship could be detected in both variables. In terms of dynamic relationship, the result shows that there is significant short run impact both ways. Based on VECM, bidirectional causality could be detected as shown in Table 4 and 4a whereby the ECT is significant.

Table 3: Long-run elasticities (Based on Johansen test)

Dependent Variables	Independent Variables	
	In Im	In Ex
In Im		3.31*** (-6.22)
In Ex	0.30** (-2.45)	

Note: ** and *** denote significant at 5% and 1% levels respectively. t-stats in parentheses.

Table 3a: Long-run elasticities (Based on ARDL Approach)

Dependent Variables	Independent Variables	
	In Im	In Ex
In Im		1.43 *** (7.36)
In Ex	0.38* (1.84)	

Note: * and *** denote significant at 10% and 1% levels respectively. t-stats in parentheses.

Table 4: Short-run elasticities (Based on Johansen test)

Dependent Variables	Independent Variables		
	$\Delta \ln \text{Im}$	$\Delta \ln \text{Ex}$	ECM_{t-1}
$\Delta \ln \text{Im}$		0.04** (2.97)	-0.02** (-2.28)
$\Delta \ln \text{Ex}$	0.07* (2.28)		-0.14** (-2.93)

Note: * and ** denote significant at 10% and 5% levels respectively. t-stats in parentheses.

Table 4a: Short-run elasticities (Based on ARDL Approach)

Dependent Variables	Independent Variables		
	$\Delta \ln \text{Im}$	$\Delta \ln \text{Ex}$	ECM_{t-1}
$\Delta \ln \text{Im}$		0.22** (3.16)	-0.15** (-3.06)
$\Delta \ln \text{Ex}$	0.35** (2.64)		-0.14** (-2.09)

Note: ** denotes significant at 5% level. t-stats in parentheses.

Conclusion

In this study, the Johansen cointegration test is employed to test the relationship between export and import in categories of forestry domain. The reason is to uncover whether trade is indeed sustainable in the Forestry domain of Malaysia, and also to unearth grappling questions of whether Malaysia is indeed adopting suitable trade policy in the forestry domain. The sample period of study was 1961 to 2007 using annual data. All the data went through log-log transformation so that the estimates will be less sensitive to outliers or influential observations and also order to reduce the data range. From the analysis, we found that the import and export in the forestry domain is highly cointegrated, a finding concurrent with a number of past studies such Husted(1992), Irandoust and Ericsson (2004), Bahmani-Oskooee (1994), Baharumshah et al. (2003) and Narayan and Narayan (2004) to name a few. The results proves that the trade in the forestry domain is sustainable. The results also imply intrinsically that the relevant authorities in Malaysia are practicing good trade policies in the forestry domain, and any deficits are short term phenomenon and indeed will be corrected in the long run. There also exist significant and positive, both dynamic and long-run relationship between both variables. Based on VECM, bidirectional granger causality could be detected among the variables. The same sets of results are also obtained when bounds test (ARDL) is employed. The signs are the asame, even the coefficients are generally identical. From a policy point of view, restricting imports in the forestry

domain could backfire and could turn out to be harmful for the Malaysian economy. Policy makers should be cautious in implementing any trade inflow policies in the domain of forestry. Trade also seems to be sustainable in the area forestry domain, it can be considered as a signal of good governance and appropriate policy selection in this area.

References

- Baharumshah A Z, Lau E and Fountas S (2003), "On the Sustainability of Current Account Deficits: Evidence from Four ASEAN Countries", *Journal of Asian Economics*, Vol. 14, No.3, pp. 465-487.
- Bahmani-Oskooee M (1994), "Are Imports and Exports of Australia Cointegrated?", *Journal of Economic Integration*, Vol. 9, No. 4, pp. 525-533.
- Bahmani-Oskooee M and Rhee H J (1997), "Are Imports and Exports of Korea Cointegrated?", *International Economic Journal*, Vol. 11, No. 1, pp. 109-114.
- Engle, R.F. and Granger, C.W.J. (1987) Cointegration and error correction: representation, estimation, and testing. *Econometrica*, 55, 251 – 276.
- Fountas S and Wu J L (1999), "Are the US Current Account Deficits Really Sustainable?", *International Economic Journal*, Vol. 13, No. 3, pp. 51-58.
- Gregory, A. and Hansen, B. (1996). Residual-based tests for cointegration in models with regime shifts, *Journal of Econometrics*, 70, 99-126.
- Iranoust M and Ericsson J (2004), "Are Import and Exports Cointegrated? An International Comparison", *Metroeconomica*, Vol. 55, No. 1, pp. 49-64.
- Johansen, S. (1991). "Estimation and hypothesis testing of cointegration vectors in Gaussian vector autoregressive models". *Econometrica* 59, 1551–1580.
- Johansen, S. and Juselius, K. (1990). "The full information maximum likelihood procedure for inference on cointegration—with applications to the demand for money". *Oxford Bulletin of Economics and Statistics*, 52, 169–210.
- Narayan P K and Narayan S (2004), "Are Exports and Imports Cointegrated? Evidence from Two Pacific Island Countries", *Economic Papers*, Vol. 23, No. 2, pp. 152-164.
- Puah, C.H., Sze-Wei, Y., Mansor, S.A and Lau, E (2008) Exchange rate and trade balance nexus in Asean-5. *Labuan Bulletin of International Business and Finance*, 6, 19-37.
- Tang T C (2002), "Are Imports and Exports of the Five ASEAN Economies Co-integrated? An Empirical Study", *International Journal of Management*, Vol. 20, pp. 88-91.
- Tang, T.C and Smyth, R (2008) Sustainability of External Imbalances for 22 Least Developed Countries: An Empirical Investigation. *ICFAI Journal of Applied Economics*, Vol. 7 Issue 3, p50-60