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2. November 2010

Online at <https://mpra.ub.uni-muenchen.de/26609/>

MPRA Paper No. 26609, posted 12. November 2010 08:59 UTC

Interpreting Markups in Spanish Manufacturing: The Exponential Model*

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First Draft: May 10nd, 2008. This draft, November 2th, 2010.

Abstract

In this paper we attempt to rationalize markups in a sample of Spanish manufacturing by assuming a representative consumer, profit-maximizing firms and constant returns to scale. We find that the standard forms of demand (CES and linear) do not provide a good explanation of markups. In contrast, a model where the representative consumer has an exponential utility function yields results that match data more closely.

*We thank Carmen Beviá, Luis Cabral, Natalia Fabra, José María Sastre, Galina Zudenkova and the participants in a seminar at Carlos III University, especially Cesar Alonso and Paco Marhuenda, for very helpful comments, and Diego Rodríguez for programming support and data. The first author acknowledges financial support from ECO2008-02738/ECON. The second author acknowledges financial support from SEJ2007-66520/ECON.

1 Introduction

This paper attempts to rationalize markups -price-marginal cost margins- by assuming a textbook model of industrial organization with a representative consumer, constant returns to scale, profit-maximizing firms and quantity competition. Assuming certain functional forms of demand -explained below- we derive first order conditions of profit maximization and transform them such that markups depend on observable variables. We regress the markups following the methodology used in Moreno and Rodriguez (2010b) with the corresponding observable variables. The data correspond to a panel of Spanish manufacturing firms for the period 1990-2005 obtained from the Encuesta Sobre Estrategias Empresariales (Survey on Business Strategies).

We first consider the most popular forms of demand, namely, CES, a combination of CES and linear, and CES with a subsistence level. We find that all these forms perform poorly when trying to explain markups in our sample: R^2 's seldom go over .1, the explanatory variables are barely significant and when they are significant, they often have values that are inconsistent with the theory. The details can be found in the following tables and summaries provided below them: Tables 1a and 1b (for the CES model), Table 2 (for the CES-linear model) and Table 3 (for the CES with a subsistence level).

We present a functional form for utility which, to the best of our knowledge, is new in the field of imperfect competition, namely the exponential form.¹ We assume that the representative consumer derives utility from a bundle of goods by means of a function which aggregates this bundle into a real number, call it m , and derives utility from m in the form $-\frac{A}{B}e^{-Bm}$ where A and B are positive numbers. We prove the existence and the uniqueness of equilibrium when firms face inverse demands that follow from these preferences, see Proposition 1.

From the empirical point of view, the exponential model implies that markups are an increasing function of output. When the product is homogeneous, this function is linear. We present empirical evidence that this model explains markups better than previous forms: the R^2 's of all sectors are larger than in any estimation of previous forms and in most cases explanatory variables are significant -with values consistent with those predicted by theory- at 99% level, see Tables 4a, 4b and 5 and the subsequent comments.

¹This form has been used extensively in macroeconomics (see, e.g. Lucas, 1985) and welfare economics (see, e.g. Ray and Ueda, 1996),

Next, we present a variation of the previous model which we call the exponential model Mark II. Now utility is derived as an exponential of a sum of quadratic terms involving outputs. In this case, markups are a quadratic function of output. From the theory point of view, this model presents two problems: The representative consumer has an atypical utility function and the second order condition of profit maximization (SOC) does not hold globally.

From the empirical point of view, this model fits data better than the previous exponential model: all R^2 's are larger than in the exponential model and, again, most variables are significant at 99% with values consistent with those predicted by the theory. We also show that in 95% of the cases, SOC hold in the vicinity of output data for the estimated parameters.

A criticism of our procedure is that we perform a test of some functional forms by running regressions which involve, on the left hand side of the equations, markups, and on the right hand side, variables like outputs or prices. But both variables are determined jointly by profit maximization so our estimation procedure might be plagued with problems like endogeneity, etc.

A response to the previous criticism is that the models presented in this paper do not admit a closed form solution, except in the extreme case in which the product is homogeneous. This means that it is not possible to perform a test like the one explained in the previous paragraphs. All we can do is test if FOCs, which involve in a natural way the equalization between markups and the corresponding elasticities of inverse demand, are satisfied. Thus, our procedure can be seen as a preliminary test of the necessary conditions of equilibrium.

Our procedure is related to that of the so-called New Empirical Industrial Organization, Gollop and Roberts (1979), see Kim and Knittel (2006) for a recent entry and references on this literature. A difference with these papers is that we do not consider the degree of collusion but focus on independent profit maximization, see our comments at the end of the paper.

The rest of the paper goes as follows. In Section 2 we explain our approach from both the theory and the empirical side. In Section 3 we test standard demand forms (CES and variations). In Section 4 we present the two exponential models and perform the relevant tests. Section 5 concludes and outlines some research paths. An Appendix explains the estimation of markups.

2 The Model

In this section we present the common ingredients of the models used in this paper. We divide this section in two subsections devoted to explaining the theoretical model and the empirical implementation of the theoretical model respectively.

2.1 The Theoretical Model

There are k sectors in the economy. A generic sector is composed of n firms producing a single output each denoted by x_1, x_2, \dots, x_n .

There is a representative consumer consuming goods $0, 1, 2, \dots, n$ in quantities M, x_1, x_2, \dots, x_n at prices $1, p_1, p_2, \dots, p_n$. Her preferences are representable by a quasi-linear utility function

$$U = V(x_1, x_2, \dots, x_n) + M, \quad (1)$$

where $V(\cdot)$ is strictly concave and three times continuously differentiable. The budget constraint of the representative consumer is $\sum_{i=1}^n p_i x_i + M = I$ where I is her income, assumed to be exogenous. Substituting the value of M in (1) and dropping I , which is just a constant, we obtain

$$U = V(x_1, x_2, \dots, x_n) - \sum_{i=1}^n p_i x_i. \quad (2)$$

We will call (2) the consumer surplus. First order conditions of consumer surplus maximization yields inverse demand functions

$$p_i = \frac{\partial V(x_1, x_2, \dots, x_n)}{\partial x_i}, \quad i = 1, 2, \dots, n. \quad (3)$$

All firms have constant marginal costs denoted by c_i . Profits for firm i denoted by π_i are

$$\pi_i = \frac{\partial V(x_1, \dots, x_n)}{\partial x_i} x_i - c_i x_i, \quad i = 1, 2, \dots, n. \quad (4)$$

The markup for firm i , denoted by μ_i , is defined as

$$\mu_i = \frac{p_i - c_i}{p_i} \quad (5)$$

We look for a Nash equilibrium in quantities where each firm maximizes profits taking the output of the other firms as given. Denoting equilibrium values with

a star superindex, first order conditions of profit maximization are

$$\frac{\partial V(x_1^*, \dots, x_n^*)}{\partial x_i} + \frac{\partial^2 V(x_1^*, \dots, x_n^*)}{\partial x_i^2} x_i^* - c_i = 0, \quad i = 1, 2, \dots, n. \quad (6)$$

From (3), (5) and (6) we obtain our fundamental equation:

$$\mu_i = - \frac{\frac{\partial^2 V(x_1^*, \dots, x_n^*)}{\partial x_i^2} x_i^*}{\frac{\partial V(x_1^*, \dots, x_n^*)}{\partial x_i}}, \quad i = 1, 2, \dots, n.. \quad (7)$$

Our research strategy is as follows. 1) We assume a parametric form of $V(\cdot)$ and find the corresponding Nash equilibrium. 2) We obtain an estimate for μ_i which does not depend on the characteristics of demand. 3) We estimate equation (7) and discuss the validity of this functional form.

2.2 The Empirical Implementation

In the Appendix we provide details of our estimation procedure. Here we only comment on the aspects that are essential for the understanding of the main text.

We estimate markups using the methodology of Roeger (1995). This approach allows us to obtain markups –by industry or by firm- estimating only one parameter and controlling for potential endogeneity of the productivity shocks.

The estimations are carried out using a panel of Spanish manufacturing firms for the period 1990-2005 obtained from the Encuesta Sobre Estrategias Empresariales (ESEE). Due to the fact that a long time period is available, we estimate individual markups for firms with at least nine observations. In this sense, we have one estimated markup for each firm. Specifically, we have information for 947 firms belonging to 20 manufacturing sectors.

We estimate equation (7) by cross-section data methodology. The independent variable is the previous estimated markup (or transformations of this variable according to the different specifications of $V(\cdot)$). The dependent variables that also vary with the theoretical specification are outputs, prices, market shares. Temporal averages of these variables are included as repressors.

Equation (7) is estimated for 15 manufacturing sectors. Although the original classification in the Appendix considers twenty sectors, in order to have a sufficient number of observations some sectors have been grouped together.

3 The CES Model and its Variants

In this section we describe the pure CES model and two extensions of it. The case in which preferences can be represented by a utility function which has a CES part and a linear part (notice that this case subsumes the linear case as a special case), and the case of CES preferences with a minimum subsistence level.

3.1 CES Preferences

In this case preferences of the representative consumer are representable by a CES utility function

$$U = F\left(\sum_{i=1}^n q_i x_i^\rho\right) + M, \quad q_i > 0, \quad \rho \in (0, 1), \quad (8)$$

where $F(\cdot)$ is strictly concave and three times continuously differentiable and q_i is interpreted as the quality of good i . In order to simplify the notation let $m \equiv \sum_{i=1}^n q_i x_i^\rho$. The inverse demand function for firm $i = 1, 2, \dots, n$ is:

$$p_i = \frac{dF(m)}{dm} q_i \rho x_i^{\rho-1}. \quad (9)$$

Profits for firm i are

$$\pi_i = \frac{dF(m)}{dm} q_i \rho x_i^\rho - c_i x_i. \quad (10)$$

The first order condition of profit maximization for firm i is

$$\frac{dF(m)}{dm} q_i \rho^2 x_i^{\rho-1} + \frac{d^2F(m)}{dm^2} q_i^2 \rho^2 x_i^{2\rho-1} = c_i. \quad (11)$$

Taking into account (5), (11) can be written as

$$\frac{d^2F(m)}{dm^2} q_i \rho x_i^\rho = \frac{dF(m)}{dm} (1 - \mu_i - \rho). \quad (12)$$

Define the elasticity of $\frac{dF(m)}{dm}$ with respect to m as $\varepsilon = \frac{d^2F(m)}{dm^2} \frac{m}{\frac{dF(m)}{dm}}$. Taking this notation into account and rearranging (12), we obtain that

$$\mu_i = 1 - \rho - \varepsilon \rho \frac{q_i x_i^\rho}{\sum_{i=1}^n q_i x_i^\rho} \quad (13)$$

Equation (13) is difficult to estimate because of the non linear terms x_i^ρ . In order to get rid of these terms , we write (9) as

$$p_i x_i = \frac{dF(m)}{dm} q_i \rho x_i^\rho. \quad (14)$$

Adding over goods (14) we obtain,

$$\sum_{i=1}^n p_i x_i = \rho \frac{dF(m)}{dm} \sum_{i=1}^n q_i x_i^\rho. \quad (15)$$

and dividing (14) by (15) we obtain that

$$\frac{p_i x_i}{\sum_{i=1}^n p_i x_i} = \frac{q_i x_i^\rho}{\sum_{i=1}^n q_i x_i^\rho}. \quad (16)$$

Thus, equation (13) can now be written as

$$\mu_i = 1 - \rho - \varepsilon \rho \frac{p_i x_i}{\sum_{i=1}^n p_i x_i}. \quad (17)$$

When $p_i x_i / \sum_{i=1}^n p_i x_i$ is small (17) yields $\mu_i \simeq 1 - \rho$ so we obtain the usual equation in which markups are constant.

Equation (17) can be estimated by making some assumptions ensuring that ε is constant and running a linear regression between markups and $p_i x_i / \sum_{i=1}^n p_i x_i$. We have two subcases, each corresponding to a different assumption on ε .²

The first case arises if $F(m) = \log(m)$, $\varepsilon = -1$ and equation (17) is

$$\mu_i = 1 - \rho + \rho \frac{p_i x_i}{\sum_{i=1}^n p_i x_i}. \quad (18)$$

The second case arises if $F(m) = (m)^{\frac{r}{\rho}}$ with $r > 0$ but $r \leq \rho$. In this case, $\varepsilon = (r - \rho)/\rho$ and equation (17) is

$$\mu_i = 1 - \rho + (\rho - r) \frac{p_i x_i}{\sum_{i=1}^n p_i x_i}. \quad (19)$$

²We do not consider that ρ 's might be different for different firms because in this case it is not possible to transform $x_i^{\rho_i} / \sum_{j=1}^n x_j^{\rho_j}$ into $p_i x_i / \sum_{i=1}^n p_i x_i$. Indeed $p_i x_i / \sum_{i=1}^n p_i x_i = \rho_i x_i^{\rho_i} / \sum_{j=1}^n \rho_j x_j^{\rho_j} \neq \rho_i x_i^{\rho_i} / \rho \sum_{j=1}^n x_j^{\rho_j}$.

In order to estimate equations (18) and (19) we write them as follows

$$\mu_i = a + b \frac{p_i x_i}{\sum_{i=1}^n p_i x_i}, \quad (20)$$

with $a = 1 - \rho$ and $b = \rho$ in the case (18) and $b = \rho - r$ in the case of (19). Notice that if $\rho = r$, the expected value for the estimate of b is zero. We present our results in Tables 1 and 1b. Table 1 offers our results for positive and negative markups. Since the latter are not compatible with the theoretical model, in Table 1b we consider only those firms with positive markups.

INSERT TABLES 1a and 1b

Consider the estimation for positive and negative markups. Firstly, the values of a are, as predicted by the theoretical model, between 0.071 and 0.276 and they are significant at 100% in all sectors. But the market share is not significant for most sectors.³ The only two sectors which present a significant effect of this variable are *Timber and Paper* and *Chemical*. However the values of b for both sectors –larger than one for *Timber and Paper* and negative for *Chemicals*- are not in accordance with the values predicted by the theoretical model (i.e. between zero and one). Therefore, our estimations yield an estimated ρ which is approximately 1 minus the estimated average markups. Notice that the values of R^2 are small (even if we take into account that our regression is cross section) being larger than .1 only in *Timber and Paper*.

When we consider positive markups only, we find similar results. Firstly, the values of a are between 0.178 and 0.305 and they are significant at 100% in all sectors. Secondly, as in the previous estimates, market shares are not significant in most sectors. For sectors where the market share is significant - *Timber and Paper*, *Chemicals*, *Fabricated Metal Products* and *Furniture* - the estimated values for b are not in accordance with the theoretical model. Thirdly, as expected, the values of R^2 are larger than in Table 1 when negative markups are considered but still smaller than 0.1 in most sectors. Finally, for the whole industry - the last rows of both Tables- we find the same results.

Summing up, the CES model with constant ε estimated by OLS does not perform well in any sector of our sample. In the next two sections we will consider two extensions of the CES model.

³This result is compatible with ρ equal to r .

3.2 The Linear-CES Model

In this case, the utility function of the representative consumer is a sum of a CES and a linear utility function, namely

$$U = F\left(\sum_{i=1}^n q_i x_i^\rho\right) + \alpha \sum_{i=1}^n x_i - \frac{\beta}{2} \sum_{i=1}^n x_i^2 - \frac{\gamma}{2} \sum_{i=1}^n x_i \sum_{j \neq i} x_j + M, \quad \beta > \gamma \geq 0, \alpha \geq 0. \quad (21)$$

Under our assumptions

$$p_i = \frac{dF(m)}{dm} q_i \rho x_i^{\rho-1} + \alpha - \beta x_i - \gamma \sum_{j \neq i} x_j, \quad i = 1, 2, \dots, n. \quad (22)$$

Thus, profits of firm i are

$$\pi_i = \frac{dF(m)}{dm} q_i \rho x_i^\rho + \alpha x_i - \beta x_i^2 - \gamma x_i \sum_{j \neq i} x_j - c_i x_i. \quad (23)$$

To simplify the model we assume, as did Spence (1976), that when a firm maximizes profits it takes m as given. The first order condition of profit maximization is

$$\frac{dF(m)}{dm} q_i \rho^2 x_i^{\rho-1} + \alpha - 2\beta x_i - \gamma \sum_{j \neq i} x_j = c_i. \quad (24)$$

Multiplying both sides of (22) by ρ , using this equation and that $c_i = p_i(1 - \mu_i)$, (24) is now

$$\rho p_i - \rho \alpha + \rho \beta x_i + \rho \gamma \sum_{j \neq i} x_j + \alpha - 2\beta x_i - \gamma \sum_{j \neq i} x_j = p_i(1 - \mu_i)$$

or

$$\mu_i p_i = -(1 - \rho)\alpha + (1 - \rho)p_i + \beta(2 - \rho)x_i + (1 - \rho)\gamma \sum_{j \neq i} x_j \quad (25)$$

In order to estimate equation (25) we write it as follows

$$\mu_i p_i = a + b p_i + d x_i + e \sum_{j \neq i} x_j \quad (26)$$

with $a = -(1 - \rho)\alpha$, $b = (1 - \rho)$, $d = \beta(2 - \rho)$ and $e = (1 - \rho)\gamma$.⁴ We present our results for positive and negative markups in Table 2. Similar results (available upon request) are obtained when only positive markups are considered.

⁴Throughout the paper, abusing notation, we use lower case latin letters a and b to denote estimated parameters. Of course, the value and the meaning of these parameters change from equation to equation.

INSERT TABLE 2

Firstly, the value of b is consistent with the predictions of the theoretical model. The implicit values of ρ tend to be smaller than the values of ρ found in the previous estimation. Nevertheless, in this case the number of sectors for which this variable is significant is smaller than in the previous estimation. With respect to the rest of explanatory variables, the coefficients are non significant in almost all sectors and when they are significant the sign is not the one predicted by the theory. For the whole sample the sign of the coefficient associated with the output of rival firms is not the one predicted by the theory.⁵ The R^2 are low but tend to be larger than those obtained in the previous estimation.

Summing up, the Linear-CES model estimated by OLS does not perform well in any sector of our sample.

3.3 The CES Model with a Subsistence Level

We now consider another variation of a CES utility function namely

$$U = F\left(\sum_{i=1}^n q_i(x_i - \bar{x})^\rho\right) + M \quad (27)$$

\bar{x} is the minimum consumption of a brand (Brown and Heien, 1972).⁶ Defining $o \equiv \sum_{i=1}^n q_i(x_i - \bar{x})^\rho$, the inverse demand function of good i in this case is:

$$p_i = \frac{dF(o)}{do} q_i \rho (x_i - \bar{x})^{\rho-1}, \quad i = 1, 2, \dots, n. \quad (28)$$

Profits in this case are

$$\pi_i = \frac{dF(o)}{do} q_i \rho (x_i - \bar{x})^{\rho-1} x_i - c_i x_i.$$

The first order condition of profit maximization yields the following equation

$$\frac{dF(o)}{do} q_i \rho (x_i - \bar{x})^{\rho-1} \left[1 + \frac{\rho - 1}{x_i - \bar{x}} x_i + \xi \rho x_i \frac{q_i (x_i - \bar{x})^{\rho-1}}{\sum_{i=1}^n q_i (x_i - \bar{x})^\rho}\right] = c_i, \quad (29)$$

⁵In any case, the sign predicted by the theory is counterintuitive because an increase in the output of rival firms should decrease the monopoly power and thus the markup of a firm.

⁶ \bar{x} could be considered to be different for different goods but this will cause problems when we estimate the relevant equation, see (32) below.

where ξ is the elasticity of $\frac{dF(\cdot)}{do}$ with respect to o . Taking into account (28) and the definition of μ_i , (29) can be written as

$$\mu_i = \frac{1 - \rho}{x_i - \bar{x}} x_i - \xi \rho \frac{p_i x_i}{\sum_{i=1}^n (p_i x_i - p_i \bar{x})}. \quad (30)$$

Notice that when $\bar{x} = 0$, (30) reduces to (17). The equation (30) is non-linear and difficult to estimate so, as we did in the Linear-CES model, we will assume when a firm maximizes profits it takes o as given. Thus $\xi = 0$ and the last term in (30) disappears. Manipulating (30) we obtain that

$$\frac{1}{\mu_i} = \frac{1}{1 - \rho} - \frac{\bar{x}}{1 - \rho} \frac{1}{x_i}. \quad (31)$$

Equation (31) relates the inverse of markups with the inverse of output. In order to estimate this equation we write it as follows:

$$\frac{1}{\mu_i} = a + b \frac{1}{x_i}. \quad (32)$$

with $a = \frac{1}{1 - \rho}$ and $b = \frac{-\bar{x}}{1 - \rho}$. We present our results for positive markups in Table 3. Due to the non-linearity of the specification, the results obtained by considering negative markups are not very relevant and are not presented.

INSERT TABLE 3

Notice that the coefficient associated with a is significant in all sectors. The implicit value of ρ lies between 0.8 and 0.963 so it agrees with the predictions of the theory. But the coefficient associated with b is not significant in 12 out of 15 sectors. Moreover, R^2 are small, never exceeding .04. Summing up, the CES model with a subsistence level does not perform any better than previous models in explaining markups.

4 The Exponential Model

In this section we present two models in which demand is an exponential function of outputs. In these models markups are a simple function of outputs.

4.1 The Basic Exponential Model

In this subsection we assume that the utility function takes the following form

$$U = -\frac{A}{B}e^{-Bm} + M, \quad A, B > 0, \quad (33)$$

where m is the variable defined in subsection 3.1. Since, to the best of our knowledge, this is the first time that this model has been used in industrial organization, we provide a complete analysis of it.

Firstly, consider the maximization of consumer surplus. We readily see that second order conditions are satisfied. Thus, first order conditions of surplus maximization yield an optimum for the consumer, namely

$$p_i = Aq_i\rho e^{-Bm}x_i^{\rho-1}, \quad i = 1, \dots, n. \quad (34)$$

Secondly, consider profit maximization. Profits for a typical firm, say i , are

$$\pi_i = Aq_i\rho e^{-Bm}x_i^\rho - c_i x_i. \quad (35)$$

First order condition of profit maximization when $x_i > 0$ are

$$\frac{\partial \pi_i}{\partial x_i} = Aq_i\rho^2 e^{-Bm}x_i^{\rho-1} - ABq_i\rho^2 x_i^\rho e^{-Bm}q_i x_i^{\rho-1} - c_i = 0, \quad (36)$$

which can be written

$$Aq_i\rho^2 e^{-Bm}x_i^{\rho-1}(1 - Bx_i^\rho q_i) - c_i = 0. \quad (37)$$

The second order condition of profit maximization holds because the left hand side of (37) is decreasing in x_i . Also, when $x_i \rightarrow 0$, $\frac{\partial \pi_i}{\partial x_i} \rightarrow \infty$ so all firms produce a positive output. We now investigate the basic properties of equilibrium.

Proposition 1. *The exponential model has a unique Nash equilibrium*

Proof. By defining a new variable $y_i \equiv q_i x_i^\rho$, the game becomes aggregative, i.e. payoffs and first order conditions for each firm, say i , can be written in terms

of y_i and $\sum_{j=1}^n y_j$, indeed equations (35) and (36) become

$$\pi_i = A\rho e^{-B\sum_{j=1}^n y_j} y_i - c_i \left(\frac{y_i}{q_i}\right)^{\frac{1}{\rho}}. \quad (38)$$

$$\frac{\partial \pi_i}{\partial x_i} = A\rho e^{-B\sum_{j=1}^n y_j} (1 - By_i) - \frac{c_i}{\rho q_i^{\frac{1-\rho}{\rho}}} y_i^{\frac{1-\rho}{\rho}}. \quad (39)$$

We see that $\frac{\partial \pi_i}{\partial x_i}$ is decreasing in $\sum y_j$ given y_i and decreasing in y_i given $\sum y_j$. Thus, assumptions A1 and A2 in Corchón (1994) are satisfied. The compactness requirement also holds because y_i can be taken such that $1 \geq By_i$. Thus, as observed in Corchón op. cit. p. 156, the model has a unique equilibrium.⁷

Proposition 1 shows that the exponential model is amenable to analysis even though a closed form solution is, in general, hard to obtain since equation (36) is difficult to solve. In order to simplify our problem, in the rest of this section we set $q_i = 1$ for all i . From the definition of a markup in (5) and (34), (36) can be written as

$$\mu_i = 1 - \rho + B\rho x_i^{\rho}. \quad (40)$$

First, let us estimate ρ . Consider the equation (34). Dividing by p_n we get

$$\frac{p_i}{p_n} = \left(\frac{x_i}{x_n}\right)^{\rho-1}, \quad i = 1, \dots, n-1. \quad (41)$$

Taking logs, the previous equation can be written as

$$\log \frac{p_i}{p_n} = (\rho - 1) \log \frac{x_i}{x_n}, \quad i = 1, \dots, n-1. \quad (42)$$

Thus, by running a linear regression between $\log \frac{p_i}{p_n}$ and $\log \frac{x_i}{x_n}$ $i = 1, \dots, n-1$ we can estimate the value of ρ . With this value of ρ , say $\bar{\rho}$, we estimate equation (40) which we write as follows

$$\mu_i - 1 + \bar{\rho} = ax_i^{\bar{\rho}}. \quad (43)$$

We present our estimation of ρ in the first column of Tables 4a (for positive and negative markups) and 4b (positive markups only).

INSERT TABLES 4a AND 4b

We see that the estimated values of ρ are, in all sectors, very close to one. The estimated value of a is shown in the first column. Notice the following:

⁷These properties also imply well-behaved comparative statics, see Corchón (1994).

Firstly, the signs of the values of a agree with those predicted by the theoretical model (i.e. positive). Secondly, the values of a are highly significant except in *Industrial and Agricultural Equipment* (with positive markups). Thirdly the values of R^2 are relatively high (taking into account that our regression is cross section). The estimation appears particularly promising in the following four sectors: *Meat Related Products*, *Timber and Paper*, *Basic Metal Products*, *Furniture and Leather*, *Fur and Footwear* and *Miscellaneous*.

The previous result suggests that taking $\rho = 1$, i. e. that goods in each sector are perfect substitutes, may be a good approximation, at least around equilibrium.⁸ In this case (40) becomes

$$\mu_i = Bx_i, \tag{44}$$

i.e. markups are a linear function of outputs. We now check the robustness of the previous estimation by testing the model imposing $\rho = 1$ for all sectors. We write equation (44) as follows:

$$\mu_i = ax_i. \tag{45}$$

with $B = a$. We present our results in Table 5 for both the case when positive and negative markups are considered and for the case when only positive markups are considered.

INSERT TABLE 5

Results here are very close to those presented before. Again, the explanatory variable is significant in all cases and in the four sectors mentioned before, the average of the R^2 in both estimations is larger than .175 and the explanatory variable is significant at 100%. To further check the robustness of this result we estimate the model for $\rho = .95$ with identical results (available upon request).

4.2 The Exponential Model, Mark II

As we will see in a moment, the estimation improves by introducing a quadratic term in (45) so $\mu_i = Bx_i - Dx_i^2$, say. This can be rationalized as follows. Suppose that the inverse demand function for firm i is

$$p_i = A_i(x_{-i})e^{-Bx_i + \frac{D}{2}x_i^2}, \quad A_i, B, D > 0, \quad i = 1, \dots, n, \tag{46}$$

⁸The homogeneity assumption allows us to find an explicit solution of equilibrium. A proof of this assertion is available upon request.

where x_{-i} is a list of all outputs minus i and $A_i(\cdot)$ is a function of x_{-i} . Under (46), profits for firm i are

$$\pi_i = A_i(x_{-i})e^{-Bx_i + \frac{D}{2}x_i^2}x_i - c_i x_i. \quad (47)$$

We will call this model, the exponential model Mark II. First order condition of profit maximization (assuming interiority) is:

$$\frac{\partial \pi_i}{\partial x_i} = A_i(x_{-i})e^{-Bx_i + \frac{D}{2}x_i^2} - A_i(x_{-i})e^{-Bx_i + \frac{D}{2}x_i^2}(B - Dx_i)x_i - c_i = 0, \quad (48)$$

which can be written as

$$A_i(x_{-i})e^{-Bx_i + \frac{D}{2}x_i^2}(1 - Bx_i + Dx_i^2) - c_i = 0. \quad (49)$$

From (46) and (49) we have that

$$p_i(1 - Bx_i + Dx_i^2) - c_i = 0. \quad (50)$$

Now, from the definition of markup we obtain that

$$\mu_i = Bx_i - Dx_i^2, \quad (51)$$

as desired. The reader can easily verify if (51) holds for all x_i , (46) is the unique inverse demand function that yields (51).

The first problem of the exponential model Mark II is to find a representative consumer yielding (46). Our next result addresses this question.

Proposition 2. *If there is a representative consumer yielding inverse demand functions as in (46) her preferences can be represented as*

$$U = \frac{e^{-\frac{1}{2}\frac{B^2}{D}}\sqrt{\pi}}{\sqrt{-2D}} \prod_{i=1}^n \operatorname{erf}\left(\sqrt{\frac{-1}{2D}}(-B + Dx_i)\right) + M \quad (52)$$

where $\operatorname{erf}(\cdot)$ is the so-called error function and

$$A_i(x_{-i}) = \prod_{j \neq i} \operatorname{erf}\left(\sqrt{\frac{-1}{2D}}(-B + Dx_j)\right), \quad i = 1, 2, \dots, n. \quad (53)$$

Proof. Consider x_{-i} as given. Integrating (46) with respect to x_i we obtain

$$\begin{aligned} V(x_1, x_2, \dots, x_n) &= \int A_i(x_{-i})e^{-Bx_i + \frac{D}{2}x_i^2} dx_i \\ &= \frac{e^{-\frac{1}{2}\frac{B^2}{D}}\sqrt{\pi}}{\sqrt{-2D}} A_i(x_{-i}) \operatorname{erf}\left(\sqrt{\frac{-1}{2D}}(-B + Dx_i)\right). \end{aligned} \quad (54)$$

where $V(\cdot)$ has been defined in (1). In order to simplify notation set

$$\frac{e^{-\frac{1}{2}\frac{B^2}{D}}\sqrt{\pi}}{\sqrt{-2D}} \equiv \lambda \text{ and } \operatorname{erf}\left(\sqrt{\frac{-1}{2D}}(-B + Dx_i)\right) \equiv f(x_i).$$

Since the previous argument works for any $i = 1, \dots, n$ we have a system of functional equations where the functions $V(\cdot)$ and $A_i(\cdot)$ are the unknowns, namely

$$\begin{aligned} V(x_1, x_2, \dots, x_n) &= \lambda A_1(x_{-1})f(x_1) \\ V(x_1, x_2, \dots, x_n) &= \lambda A_2(x_{-2})f(x_2) \\ &\dots\dots \\ V(x_1, x_2, \dots, x_n) &= \lambda A_n(x_{-n})f(x_n), \end{aligned}$$

A solution for the above system is

$$A_i(x_{-i}) = \prod_{j \neq i}^n f(x_j) \quad i = 1, 2, \dots, n. \quad (55)$$

Thus, the utility function in (52) is validated. Now since

$$\begin{aligned} \frac{\partial V(x_1, x_2, \dots, x_n)}{\partial x_i} &= \lambda \frac{\partial f(x_i)}{\partial x_i} \prod_{j \neq i}^n f(x_j) = e^{-Bx_i + \frac{D}{2}x_i^2} \prod_{j \neq i}^n f(x_j) = \\ p_i &= A_i(x_{-i})e^{-Bx_i + \frac{D}{2}x_i^2} \end{aligned}$$

and the inverse demand function is independent of the particular representation of preferences, it must be that

$$A_i(x_{-i}) = \prod_{j \neq i}^n f(x_j)$$

as desired.

Proposition 2 is a negative result because (52) is an atypical utility function and (53) yields atypical demand functions. Moreover, it can be shown that (52) is not concave (an example is available upon request). Thus, first order conditions might reflect just a local maximum. Finally, it is also not clear how to invert inverse demand functions to obtain demand functions except in very special cases.⁹

⁹For instance when $n = 1$ and $2D \log A_1 \leq B^2$, the demand function exists.

The second problem of the exponential model Mark II, is that the second order condition fails to be satisfied globally. Indeed, computing,

$$\frac{\partial p_i}{\partial x_i} = (-B + Dx_i)p_i \quad (56)$$

$$\frac{\partial^2 p_i}{\partial x_i^2} = Dp_i + (-B + Dx_i)\frac{\partial p_i}{\partial x_i} \quad (57)$$

Taking into account (56), (57) can be written as

$$\frac{\partial^2 p_i}{\partial x_i^2} = Dp_i + (-B + Dx_i)^2 p_i. \quad (58)$$

Second order conditions of profit maximization are

$$2\frac{\partial p_i}{\partial x_i} + \frac{\partial^2 p_i}{\partial x_i^2} x_i < 0. \quad (59)$$

Taking into account (56) and (58), (59) can be written as

$$2(-B + Dx_i) + Dx_i + (-B + Dx_i)^2 x_i < 0. \quad (60)$$

Now using (51), (60) can be written as

$$\mu_i(\mu_i - 2) + Dx_i^2 < 0. \quad (61)$$

which is the equation to be checked. Clearly (61) cannot hold for all x_i because for output close to infinity, the expression on the left hand side is positive.

We test equation (51). In the notation used in previous sections,

$$\mu_i = ax_i + bx_i^2, \quad (62)$$

where $a = B$ and $b = -D$. Results are given in Tables 6a and 6b below for the cases of only positive markups and positive and negative markups respectively. We also test if second order conditions hold for the estimated values of μ_i and D and for the data on x_i . Thus, we test if firms are maximizing profits locally.¹⁰

INSERT TABLES 6a AND 6b HERE

¹⁰If firms maximize profits locally we have at least three interpretations of the behavior of firms. 1) Firms are just local maximizers, Silvestre (1979). 2) Firms maximize globally but with capacity constraints. 3) Firms maximize globally, but the form of the demand function is locally but not globally the exponential.

We see that the estimated values of B are significant at 99% in twelve sectors and the aggregate. They are significant at 95% in three sectors. The estimated values of D are significant at 99% in seven sectors and the aggregate. They are significant at 95% in five sectors. All R^2 's are higher than the previous estimation; SOC never hold when $\mu_i < 0$. When $\mu_i > 0$ SOC hold locally in 94.6% of cases when $\mu_i > 0$. Summing up, the exponential model estimated by OLS performs quite well in a relatively large number of sectors in our sample. This model seems to be a reasonable starting point for understanding markups in several sectors of our sample.

5 CONCLUSIONS

In this paper we have presented a new functional form, namely the exponential one, which explains FOC better than standard forms like CES, linear, etc. for a sample of Spanish manufacturing data. Our model relies on assumptions like quantity competition, quasi-linear utility and stable preferences that must be subject to further inquiry before definitive conclusions are drawn. We also do not consider factors that are likely to play an important role in the determination of markups such as Collusion (Bresnahan, (1989)), International Trade (Melitz and Ottaviano (2008), Feenstra and Kee (2008), De Loecker and Warzynski, 2009, Moreno and Rodriguez (2010a) and non-linear pricing (Miravete and Röller (2004)). Our model is admittedly simple but might serve as a useful starting point for future endeavours.

Summing up, the exponential models presented in this paper are workable and provide us with a new tool to analyze markets characterized by imperfect competition.

We emphasize that our paper cannot be construed as a criticism of linear or CES models. On the contrary, the exponential model presented in this paper must be considered as a new addition to the toolbox of Industrial Economics that might be useful in some instances and should not be regarded as a competitor, but rather as a complement to other models.

6 Appendix: Estimation of Markups

6.1 Methodology

To estimate the firm's markups we use the methodology proposed by Roeger (1995). Some papers have used this approach to analyze the effect of trade liberalization on markups (see, for example Konings et al. (2005a & 2005b)). This approach starts with a production function which is linear homogeneous in the inputs:

$$x_{it} = A_{it}F(L_{it}, M_{it}, K_{it}) \quad (63)$$

where x_{it} is the output, L_{it} , M_{it} and K_{it} are the labour, materials and the capital inputs and A_{it} is the productivity of firm i at time t . The constant returns to scale assumption could bias upwardly (downwardly) the estimated levels (changes) in the markup. For a discussion, see Konings et al. (2005a).

Under imperfect competition (see Hall (1988)), the output growth rate of the relevant variables can be expressed as:

$$y_{it} = R_{it}(s_{it}^L \cdot l_{it} + s_{it}^M \cdot m_{it} + s_{it}^K \cdot k_{it}) + \theta_{it} \quad (64)$$

where R_{it} is the ratio between price and marginal cost of firm i and time t . y_{it} , l_{it} , m_{it} and k_{it} are the growth rate of output, labour, materials and capital, respectively. s_{it}^J is the input cost share of factor J in sales (or total value of production), namely

$$s_{it}^J = \frac{P_{it}^J \cdot J_{it}}{P_{it} \cdot x_{it}}, \quad J = L, M, K. \quad (65)$$

where P_{it}^J and P_{it} are the prices of inputs, labour, material and physical capital, and output, respectively. Finally, θ_{it} is the growth of productivity.

Equation (64) can be rewritten after some algebra to decompose the Solow residual (SR_{it}) into two terms:

$$SR_{it} = y_{it} - s_{it}^L \cdot l_{it} - s_{it}^M \cdot m_{it} - (1 - s_{it}^L - s_{it}^M) \cdot k_{it} = \mu_{it}(y_{it} - k_{it}) + (1 - \mu_{it})\theta_{it} \quad (66)$$

where μ_{it} is the price marginal cost markup as defined in the main text.

The problem with the estimation of μ_{it} is the correlation with the inputs and the unobservable growth of the productivity. Previous empirical research has used Equations (64) or (66) to estimate markups by using instrumental variables but some times is difficult to find exogenous instruments (see, for example, Olley

and Pakes (1996) and Levison and Petrin (2003)). Roeger (1995) proposes using the price-based (or dual) Solow residual or the dual Solow residual to deal with this problem. Specifically, using the cost minimization problem and also imposing constant returns to scale, Roeger proposes decomposing the dual Solow residual (DSR) as in equation (66):

$$DSR_{it} = s_{it}^L \cdot p_{it}^L + s_{it}^M \cdot p_{it}^M + (1 - s_{it}^L - s_{it}^M) p_{it}^K - p_{it} = -\mu_{it}(p_{it} - p_{it}^K) + (1 - \mu_{it})\theta_{it} \quad (67)$$

where $p_{it}^L, p_{it}^M, p_{it}^K$ and p_{it} are the growth rate of the wage, price of intermediates inputs (or materials), the rental price of capital and the output price, respectively.

Subtracting equation (67) from equation (66), we obtain that

$$\begin{aligned} SR_{it} - DSR_{it} &= (y_{it} + p_{it}) - s_{it}^L(l_{it} + p_{it}^L) - s_{it}^M(m_{it} + p_{it}^M) - (1 - s_{it}^L - s_{it}^M)(k_{it} + p_{it}^K) \\ &= \mu_{it}(y_{it} + p_{it} - k_{it} - p_{it}^K) \end{aligned} \quad (68)$$

In equation (68) the term which refers to the growth of productivity is eliminated and the problem of simultaneity disappears. In this sense, the Lerner index can be estimated consistently. To simplify notation, we denote the left-hand side in equation (68) as dx_{it} which can be interpreted as the difference between the growth rate of sales (nominal production) sales and a weighted average of the growth rate of the factors cost weighted by their respective share in sales. We denote the term in brackets on the right-hand side of the equation (68) as dZ_{it} , which can be interpreted as the growth rate of sales per value of capital. Then, the equation to be estimated is:

$$dx_{it} = \mu dZ_{it} + u_{it} \quad (69)$$

Different approaches can be followed to estimate Equation (69). Firstly, it can be estimated using panel data and obtaining an average markup of the economy (or the industries considered). This is the approach followed by Konings et al. (2005a and 2005b). Secondly, the availability of a long time period for each firm allows us to estimate individual markups. We are going to use this last one.

6.2 Data

Estimations are carried out with an unbalanced panel of Spanish manufacturing firms for the period 1990-2005. The variables were obtained from the Encuesta

Sobre Estrategias Empresariales (ESEE), a survey that was sponsored by the Ministry of Industry and carried out by the Fundación SEPI. The sampling scheme of this survey is conducted for each manufacturing NACE class (two-digit) level. Companies employing between 10 and 200 employees are chosen by a random sampling scheme and the rate of participation is around 4%. For firms employing more than 200 employees, the rate of participation is about 60%.

The set of variables required to estimate equation (69) are sales ($y_{it}p_{it}$), nominal labour cost ($l_{it}p_{it}^L$) -where l_{it} is the number of hours effectively worked and p_{it}^L refers to labour costs per employee and hour-, nominal value of intermediate consumption ($m_{it}p_{it}^M$) -where m_{it} is the intermediate consumption and p_{it}^M is a Paasche price index weighting the price variations of raw materials, energy and services purchased-, and the nominal value of capital services ($k_{it}p_{it}^K$).

Hours effectively worked are measured as the sum of the normal working time and overtime minus the non-worked hours, while intermediate consumption is defined as the sum of purchases and external services, minus the variation in the stock of purchases. For capital, we use the net stock of capital for equipment in real terms. It is calculated by using the perpetual inventory formula:

$$K_{it} = (1 - \delta_{it})K_{it-1} \frac{P_{it}^E}{P_{it-1}^E} + I_{it} \quad (70)$$

where P_{it-1}^E is the price index for equipment, δ_{it} is the depreciation rate, and I_{it} is the investment in equipment. The rental price of capital is calculated as the long-run debt interest rate paid by the firm minus the rate of change in capital goods for the equipment price index plus equipment good depreciation, multiplied by the investment good price index, $P_{it}^K = P_t^E(i_{it} - \pi_{it}^E + \delta_{it})$.

6.3 Results

Firstly, we began with the standard methodology using panel data. In Table A1, we present the estimation of equation (69) by pooled OLS. We have also run regressions with fixed effects but the null hypothesis that all fixed effects are equal to zero is not rejected. All of them are carried out by an unbalanced panel of manufacturing firms during the period 1990-2005.

In Column 1 we present the estimation for our sample observations with enough data to build the variables. As can be seen, the average markups is 0.169, although there is quite a lot of heterogeneity across industries.

To test some of the equations of the main text we will need information about the market share of the firm and other variables referring to the real data. Although the ESSE is not a census for the manufacturing firms – only firms with more than 10 employees are surveyed – , it is possible to use the sampling fractions to weight our sample observations. The sampling scheme was conducted for each manufacturing NACE class (two-digit) level and five size categories¹¹. The product of 20 sectors and 5 size classes defines the dimension of the weighting matrix, which has a total number of 90 sampling fractions. We have applied the weighting fractions of 2005 to update the observations of the firms.

The estimation for this sample is presented in Column 4. As can be seen, the results are quite similar to those presented with our sample.

The dependent variables of the equations of the main text are the Index Lerner. Thus, we need some estimation of the individual markups. As was previously stated, the availability of a sufficient time period for each firm allows us to estimate individual markups although the limitation of this approach is that we will only have one observation for each firm.

For this reason, and secondly, we estimate equation (69) by OLS for 947 firms with more than nine observations. That is:

$$dx_t = \mu_i dZ_t + u_t \quad t > 9 \dots 15$$

The average markup for these firms is 0.190. It is quite similar to the one obtained by the standard approach and presented in the last row of Table A1. Almost 10% of the sample presents negative markups. When these firms are not considered, the average markup increases to 0.224. Figure 1 presents the distribution of these markups for all firms. As can be seen, the dispersion is large and the distribution is slightly skewed, with a large proportion of firms on the right tail.

Table A2 below presents the average of the estimated individual markups for 20 manufacturing sectors. The average of estimated Lerner index ranges from 0.088 to 0.265, when all, positive and negative markups are considered.

Due to the fact that we have a small number of observations in some sectors, for the estimation of the equation of the main test, we aggregate some sectors: specifically, Food, Tobacco and Beverages, Timber and Paper, Office Machinery

¹¹See Segura et alia (1992) for the details.

and Electric Materials and Accessories, Vehicles and Motors and Other Transport Equipment, and Leather, Fur and Footwear and Miscellaneous.

Table A1. Pooled OLS estimations

	Sample observations (1990-2005)			Uprating the observations	
	μ	Firms	Obs	μ	Obs
Meat related products	0.108 (8.6)	77	603	0.106 (44.9)	17122
Food and tobacco	0.133 (15.6)	252	1971	0.130 (74.5)	48193
Beverages	0.189 (9.1)	55	395	0.138 (19.8)	4662
Textiles and clothing	0.170 (17.1)	256	1851	0.193 (81.5)	33377
Leather, fur and footwear	0.177 (8.5)	83	467	0.190 (44.2)	11008
Timber	0.144 (8.5)	77	516	0.158 (57.4)	21197
Paper	0.201 (13.5)	68	526	0.173 (48.0)	7820
Printing and Publishing	0.187 (15.8)	141	1029	0.194 (97.5)	31940
Chemicals	0.164 (16.8)	179	1379	0.146 (51.9)	20912
Plastic and rubber products	0.130 (11.6)	141	1031	0.129 (48.1)	19587
Non-mineral products	0.219 (19.7)	172	1340	0.230 (90.4)	29477
Basic metal products	0.166 (12.6)	77	653	0.169 (46.9)	9040
Fabricated metal products	0.192 (21.2)	261	1835	0.207 (118.8)	51330
Industrial and agricultural equipment	0.196 (19.0)	183	1407	0.221 (90.9)	27988
Office mach. data processing and similar	0.129 (5.0)	53	286	0.107 (15.3)	5040
Electric materials and accessories	0.168 (14.8)	169	1211	0.194 (60.7)	16752
Vehicles and motors	0.159 (14.7)	125	980	0.163 (42.1)	9327
Other transport equipment	0.164 (7.6)	50	371	0.147 (19.1)	3569
Furniture	0.162 (12.0)	134	949	0.166 (63.8)	26593
Miscellaneous	0.187 (9.6)	59	441	0.190 (49.3)	11786
All firms	0.169 (9.6)	2612	19241	0.177 (281.9)	406720

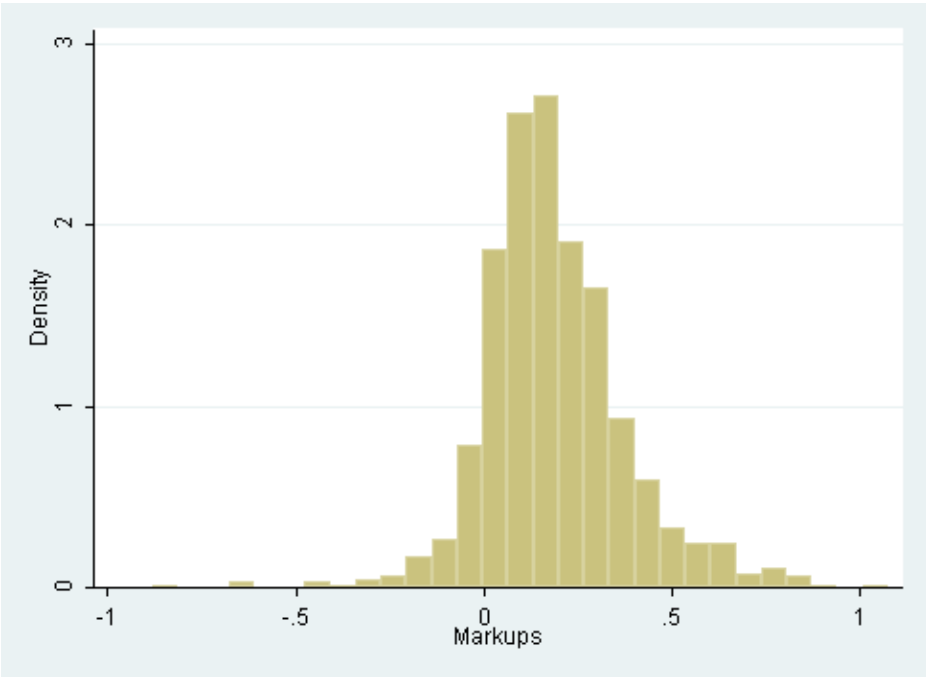


Figure 1. Markups distribution

Table A2. Average Individual Mark-ups by sectors

	Positive & negative			Positive markups		
	markups					
	Mean	S.D	Firms	Mean	S.D.	Firms
Meat related products	0.088	0.139	33	0.123	0.134	26
Food and tobacco	0.144	0.156	106	0.183	0.130	90
Beverages	0.203	0.202	19	0.258	0.165	16
Textiles and clothing	0.169	0.230	95	0.225	0.162	83
Leather, fur and footwear	0.227	0.178	20	0.255	0.165	18
Timber	0.201	0.165	23	0.212	0.161	22
Paper	0.209	0.194	27	0.226	0.191	25
Printing and Publishing	0.265	0.195	48	0.265	0.195	48
Chemicals	0.198	0.171	70	0.214	0.162	66
Plastic and rubber products	0.139	0.195	43	0.177	0.127	40
Non-mineral products	0.244	0.219	72	0.286	0.191	64
Basic metal products	0.177	0.137	37	0.184	0.132	36
Fabricated metal products	0.201	0.174	81	0.232	0.154	73
Ind. and agric. equipment	0.199	0.223	74	0.253	0.179	64
Office mach. and similar	0.257	0.257	10	0.328	0.237	8
Electric materials and accessories	0.219	0.187	57	0.253	0.165	51
Vehicles & motors	0.181	0.138	52	0.190	0.134	50
Other transport equipment	0.149	0.253	20	0.220	0.187	17
Furniture	0.194	0.184	38	0.230	0.157	34
Miscellaneous	0.233	0.185	22	0.245	0.182	21
All firms	0.191	0.192	947	0.224	0.164	852

6.4 Appendix II: Definitions of the variables

Capital stock of equipment goods: Net stock of capital for equipment goods in real terms. This is calculated by using the perpetual inventory formula: $K_t = (1 - d)K_{t-1}(P_t/P_{t-1}) + I_t$, where P is the price index for equipment, d is the depreciation rate, and I is the investment in equipment.

Effective hours of work: Normal hours plus overtime hours minus lost hours.

Intermediate consumption: Raw material purchases, energy and fuel costs and other external services.

Market share $\frac{p_i x_i}{\sum_{i=1}^n p_i x_i}$: The ratio of the nominal output of the firm over the total nominal of the sector output for each year. The denominator is calculated uprating the observations by the sampling fractions in 2005.

Output (x_i): Nominal output -Sales plus stock variation- is deflated using Paasche-type firm individual indices, constructed starting from the price changes in output reported by firms (*Price*).

Price (p_i): The surveyed firms give annual information about markets served (up to five), identifying their relative importance (in percentage) in total sales of the firm. This information allows us to calculate a price Paasche-type index for all markets and for each market, using the proportions with respect to total sales as weighting.

Table 1a. OLS cross-section estimation of Equation (20)

	Positive & negative markups			
	<i>a</i>	<i>b</i>	Firms	R^2
Meat related products	0.071*** (0.028)	4.7 (4.1)	33	0.01
Food, tobacco and beverages	0.160*** (0.016)	-9.7 (10.6)	125	0.0
Textiles and clothing	0.171*** (0.028)	-1.3 (10.2)	95	0.0
Timber and paper	0.168*** (0.028)	13.9** (5.2)	50	0.11
Printing and Publishing	0.276*** (0.031)	-4.1 (4.4)	48	0.0
Chemicals	0.231*** (0.024)	-24.0** (9.5)	70	0.07
Plastic and rubber products	0.142*** (0.035)	-1.5 (7.9)	43	0.0
Non-mineral products	0.253*** (0.030)	-4.6 (7.9)	72	0.0
Basic metal products	0.171*** (0.025)	1.1 (2.0)	37	0.0
Fabricated metal products	0.213*** (0.022)	-13.1 (11.2)	81	0.0
Ind. and agric. equipment	0.215*** (0.029)	-5.8 (4.6)	74	0.01
Office mach. and Electric mat.	0.234*** (0.028)	-5.2 (7.8)	67	0.0
Vehicles and transport equip.	0.172*** (0.022)	0.1 (4.5)	72	0.0
Furniture	0.174*** (0.033)	7.4 (5.7)	38	0.02
Leather, fur, footwear and Miscellaneous	0.213*** (0.032)	14.4 (13.5)	42	0.0
All firms	0.190*** (0.007)	0.04 (1.1)	947	0.0

Table 1b. OLS cross-section estimation of Equation (20)

	Positive markups			
	<i>a</i>	<i>b</i>	Firms	<i>R</i> ²
Meat related products	0.109*** (0.030)	3.8 (4.2)	26	0.00
Food, tobacco and beverages	0.205*** (0.016)	-15.1 (11.8)	106	0.01
Textiles & clothing	0.228*** (0.021)	-1.5 (7.4)	83	0.0
Timber & paper	0.178*** (0.027)	15.9*** (5.2)	47	0.16
Printing & Publishing	0.256*** (0.025)	-3.0 (3.5)	47	0.0
Chemicals	0.243*** (0.023)	-21.3** (9.5)	66	0.06
Plastic & rubber products	0.187*** (0.024)	-4.6 (5.2)	40	0.0
Non-mineral products	0.305*** (0.028)	-9.3 (6.9)	64	0.01
Basic metal products	0.178*** (0.025)	1.1 (2.0)	36	0.0
Fabricated metal products	0.247*** (0.020)	-16.6* (9.9)	73	0.02
Ind. and agric. equipment	0.271*** (0.025)	-6.3 (3.8)	64	0.03
Office mach. & Electric mat.	0.280*** (0.027)	-8.1 (7.0)	59	0.01
Vehicles & transport equip.	0.197*** (0.019)	0.0 (3.8)	67	0.0
Furniture	0.191*** (0.024)	19.3*** (4.7)	34	0.33
Leather, fur & Miscellaneous	0.237*** (0.032)	9.9 (13.1)	39	0.0
All firms	0.223*** (0.006)	0.07 (1.0)	851	0.0

Table 2. OLS cross-section estimation of Equation (26)

	Positive and negative markups				Firms	R^2
	a	b	d	e		
Meat related products	4.7* (2.5)	0.005 (0.197)	0.001 (0.001)	-0.0004* (0.0002)	33	0.06
Food, tobacco and beverages	-0.8 (0.6)	0.210** (0.109)	-0.0002 (0.0003)	0.000 (0.000)	125	0.05
Textiles and clothing	-0.8 (1.5)	0.245* (0.141)	-0.000 (0.001)	0.0001 (0.0001)	95	0.00
Timber and paper	-0.2 (0.4)	0.141 (0.175)	0.001** (0.001)	0.000 (0.000)	50	0.08
Printing and Publishing	0.6 (1.1)	0.118 (0.131)	-0.001 (0.001)	-0.000 (0.0001)	48	0.00
Chemicals	0.2 (0.7)	0.293** (0.125)	-0.001** (0.0003)	-0.000 (0.000)	70	0.14
Plastic and rubber products	-0.4 (0.5)	0.154 (0.265)	-0.0001 (0.001)	0.000 (0.000)	43	0.00
Non-mineral products	1.1 (0.6)	0.002 (0.164)	-0.0003 (0.001)	-0.0001 (0.000)	72	0.00
Basic metal products	0.6 (0.6)	0.250* (0.133)	-0.000 (0.000)	-0.000 (0.000)	37	0.05
Fabricated metal products	-0.1 (0.4)	0.376** (0.156)	-0.001 (0.001)	-0.000 (0.000)	81	0.06
Industrial and agricultural equipment	0.3 (0.9)	0.359* (0.206)	-0.0004 (0.0004)	-0.000 (0.0001)	74	0.02
Office mach. and Electric materials	-0.7 (0.8)	0.439** (0.187)	-0.000 (0.0003)	0.000 (0.0001)	67	0.06
Vehicles and transport equipment	-0.4 (0.4)	0.299** (0.169)	-0.000 (0.0001)	0.000 (0.000)	72	0.01
Furniture	0.2 (1.1)	0.563** (0.219)	0.002 (0.001)	-0.0002 (0.0003)	38	0.11
Leather, fur, footwear and Misc.	-1.8 (2.9)	0.165 (0.257)	0.003 (0.003)	0.0003 (0.0004)	42	0.00
All firms	-0.02 (0.01)	0.221*** (0.039)	-0.0001 (0.0001)	-0.0000* (0.0000)	947	0.04

Table 3. OLS cross-section estimation of Equation (32)

	Positive markups				
	a	b	Firms	ρ	R^2
Meat related products	27.2*** (8.1)	-6.7 (9.2)	26	0.963	0.00
Food, tobacco and beverages	14.1*** (4.2)	-2.6 (3.3)	106	0.929	0.00
Textiles and clothing	10.7*** (2.6)	-1.1 (2.0)	83	0.907	0.00
Timber and paper	9.4*** (1.7)	-1.0 (2.0)	47	0.893	0.00
Printing and Publishing	10.4*** (2.9)	-1.7 (1.9)	47	0.904	0.00
Chemicals	11.1*** (2.3)	1.8 (4.7)	66	0.910	0.00
Plastic and rubber products	12.7*** (3.2)	-1.5 (4.0)	40	0.921	0.00
Non-mineral products	10.4** (4.3)	2.1 (5.1)	64	0.904	0.01
Basic metal products	13.2*** (3.5)	-4.1 (13.4)	36	0.924	0.00
Fabricated metal products	9.7*** (1.5)	-2.8 (1.8)	73	0.896	0.02
Industrial and agricultural equipment	9.6*** (1.8)	-2.4 (2.1)	64	0.895	0.01
Office mach. and Electric materials	6.0*** (1.3)	3.1** (1.6)	59	0.833	0.04
Vehicles and transport equipment	19.7 (49.7)	259.6** (149.7)	67	0.949	0.03
Furniture	5.0** (2.5)	3.3 (2.2)	34	0.800	0.04
Leather, fur, footwear and Misc.	7.7*** (2.3)	0.7 (2.1)	39	0.870	0.00
All firms	14.0*** (4.4)	-0.4 (4.5)	851	0.929	0.00

Table 4a. OLS cross-section Estimation of ρ and Equation (43)

	Positive and negative markups			
	Estimated ρ	a	Firms	R^2
Meat related products	0.969	0.0009** (0.0004)	33	0.10
Food, tobacco and beverages	0.958	0.0006** (0.0003)	125	0.03
Textiles and clothing	0.972	0.002** (0.001)	95	0.04
Timber and paper	0.984	0.003*** (0.001)	50	0.35
Printing and Publishing	0.967	0.002* (0.001)	48	0.04
Chemicals	0.971	0.0006* (0.0003)	70	0.03
Plastic and rubber products	0.984	0.001* (0.0007)	43	0.04
Non-mineral products	0.994	0.002*** (0.0006)	72	0.09
Basic metal products	0.995	0.0004** (0.0002)	36	0.14
Fabricated metal products	0.959	0.002** (0.001)	81	0.05
Industrial and agricultural equipment	0.965	0.0006 (0.0004)	74	0.02
Office mach. and Electric materials	0.970	0.001*** (0.0004)	67	0.09
Vehicles and transport equipment	0.979	0.0002** (0.0001)	72	0.04
Furniture	0.955	0.004** (0.002)	38	0.13
Leather, fur, footwear and Misc.	0.957	0.010*** (0.003)	42	0.20
All firms	0.969	0.0005*** (0.0001)	947	0.04

Table 4b. OLS cross-section estimation of ρ and Equation (43)

	Positive markups			
	Estimated ρ	a	Firms	R^2
Meat related products	0.973	0.001** (0.0005)	23	0.16
Food, tobacco and beverages	0.958	0.001*** (0.0004)	101	0.10
Textiles and clothing	0.972	0.004*** (0.001)	81	0.16
Timber and paper	0.985	0.003*** (0.001)	47	0.38
Printing and Publishing	0.964	0.002** (0.001)	45	0.06
Chemicals	0.966	0.001*** (0.0005)	59	0.12
Plastic and rubber products	0.982	0.001* (0.0006)	39	0.09
Non-mineral products	0.997	0.002*** (0.0007)	64	0.10
Basic metal products	1.0	0.0004** (0.0002)	35	0.15
Fabricated metal products	0.959	0.003** (0.001)	69	0.07
Industrial and agricultural equipment	0.967	0.0009* (0.0005)	62	0.04
Office mach. and Electric materials	0.969	0.001** (0.0004)	57	0.10
Vehicles and transport equipment	0.976	0.0002** (0.0001)	66	0.06
Furniture	0.953	0.008*** (0.002)	33	0.40
Leather, fur, footwear and Misc.	0.953	0.010*** (0.003)	37	0.20
All firms	0.969	0.0006*** (0.0001)	812	0.06

Table 5. OLS cross-section estimation of Equation (45)

	Positive and negative markups			Positive markups		
	a	Firms	R^2	a	Firms	R^2
Meat related products	0.001** (0.0004)	33	0.16	0.001** (0.0005)	26	0.19
Food, tobacco and beverages	0.001*** (0.0002)	125	0.06	0.001*** (0.0003)	106	0.11
Textiles and clothing	0.002** (0.001)	95	0.05	0.004*** (0.001)	83	0.17
Timber and paper	0.003*** (0.001)	50	0.35	0.003*** (0.001)	47	0.37
Printing and Publishing	0.002* (0.001)	48	0.04	0.002** (0.001)	47	0.06
Chemicals	0.0006* (0.0002)	70	0.05	0.001** (0.0003)	66	0.07
Plastic and rubber products	0.001* (0.0006)	43	0.06	0.001* (0.0006)	40	0.10
Non-mineral products	0.002*** (0.0006)	72	0.10	0.002*** (0.0006)	64	0.10
Basic metal products	0.0004** (0.0002)	37	0.15	0.0004** (0.0002)	36	0.15
Fabricated metal products	0.002** (0.001)	81	0.07	0.003** (0.001)	73	0.09
Industrial and agricultural equipment	0.001* (0.0004)	74	0.03	0.001*** (0.001)	64	0.05
Office mach. and Electric materials	0.001*** (0.0004)	67	0.11	0.001** (0.0004)	59	0.12
Vehicles and transport equipment	0.0002** (0.0001)	72	0.05	0.0002** (0.0001)	67	0.06
Furniture	0.004*** (0.001)	38	0.15	0.007*** (0.0002)	34	0.36
Leather, fur, footwear and Misc.	0.010*** (0.003)	42	0.22	0.010*** (0.003)	39	0.21
All firms	0.0005*** (0.0001)	947	0.04	0.0006*** (0.0001)	852	0.06

Table 6a. OLS cross-section estimation of Equation (62)

	Positive markups				
	<i>a</i>	<i>b</i>	Firms	>0	R^2
Meat related products	0.004** (0.001)	-0.00001 (0.0)	26	3	0.25
Food, tobacco and beverages	0.004*** (0.001)	-0.00001*** (0.0)	106	10	0.20
Textiles and clothing	0.012*** (0.003)	-0.0001*** (0.0)	83	9	0.26
Timber and paper	0.005*** (0.001)	-0.00001** (0.0)	47	2	0.41
Printing and Publishing	0.006** (0.003)	-0.00003 (0.0)	47	3	0.09
Chemicals	0.003*** (0.001)	-0.000*** (0.0)	66	7	0.21
Plastic and rubber products	0.004** (0.001)	-0.00001** (0.0)	40	3	0.18
Non-mineral products	0.006*** (0.001)	-0.00002** (0.0)	64	1	0.25
Basic metal products	0.002*** (0.000)	-0.000*** (0.0)	36	1	0.36
Fabricated metal products	0.008*** (0.002)	-0.00005*** (0.00002)	73	3	0.18
Industrial and agricultural equipment	0.003*** (0.001)	-0.000** (0.0)	64	3	0.10
Office mach. and Electric materials	0.004*** (0.001)	-0.00001** (0.0)	59	7	0.19
Vehicles and transport equipment	0.001*** (0.0003)	-0.0000*** (0.0)	67	2	0.26
Furniture	0.028*** (0.005)	-0.0002*** (0.0001)	34	2	0.56
Leather, fur, footwear and Misc.	0.023*** (0.008)	-0.0003 (0.0002)	39	2	0.25
All firms	0.002*** (0.0001)	-0.000*** (0.0)	851	7	0.13

Table 6b. OLS cross-section estimation of Equation (62)

	Positive and negative markups				
	<i>a</i>	<i>b</i>	Firms	>0	R^2
Meat related products	0.002* (0.001)	-0.0000 (0.0)	33	9 2 (Li>0)	0.16
Food, tobacco and beverages	0.002*** (0.0005)	-0.000*** (0.0)	125	23 4 (Li>0)	0.14
Textiles and clothing	0.012*** (0.003)	-0.0001*** (0.0)	95	21 2 (Li>0)	0.15
Timber and paper	0.004*** (0.001)	-0.000 (0.0)	50	3	0.37
Printing and Publishing	0.006** (0.003)	-0.00003 (0.0)	48	4 3 (Li>0)	0.06
Chemicals	0.003*** (0.001)	-0.000*** (0.0)	70	11 7 (Li>0)	0.18
Plastic and rubber products	0.003** (0.002)	-0.00001 (0.0)	43	5 2 (Li>0)	0.08
Non-mineral products	0.005*** (0.001)	-0.00002*** (0.0)	72	9 1 (Li>0)	0.22
Basic metal products	0.002*** (0.0004)	-0.000*** (0.0)	37	2 1 (Li>0)	0.34
Fabricated metal products	0.007*** (0.002)	-0.00004*** (0.00002)	81	11 3 (Li>0)	0.14
Industrial and agricultural equipment	0.002** (0.001)	-0.000 (0.0)	74	13 3 (Li>0)	0.05
Office mach. and Electric materials	0.003*** (0.001)	-0.000** (0.0)	67	14 6 (Li>0)	0.15
Vehicles and transport equipment	0.001*** (0.0002)	-0.000*** (0.0)	72	7 2 (Li>0)	0.18
Furniture	0.010** (0.005)	-0.00008 (0.00006)	38	5 1 (Li>0)	0.17
Leather, fur, footwear and Misc.	0.022*** (0.008)	-0.0003* (0.0002)	42	5 2 (Li>0)	0.25
All firms	0.001*** (0.0001)	-0.000*** (0.0)	947	102 6 (Li>0)	0.09

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