



Munich Personal RePEc Archive

Resource curse: new evidence on the role of institutions

Sarmidi, Tamat and Siong Hook, Law and Jafari, Yaghoob

Universiti Kebangsaan Malaysia, Universiti Putra Malaysia,
Universiti Kebangsaan Malaysia

February 2012

Online at <https://mpra.ub.uni-muenchen.de/37206/>
MPRA Paper No. 37206, posted 09 Mar 2012 12:11 UTC

Resource curse: New evidence on the role of institutions

Tamat Sarmidi^{a,1} Siong Hook Law^b Yaghoob Jafari^a

^a *School of Economics, Universiti Kebangsaan Malaysia, Malaysia*

^b *Department of Economics, Universiti Putra Malaysia, Malaysia*

Abstract

This paper attempts to provide a probable answer to a longstanding resource curse puzzle; i.e., why resource-rich nations grow at a slower rate compared to less fortunate ones. Using an innovative threshold estimation technique, the empirical results reveal that there is a threshold effect in the natural resources – economic growth relationship. We find that the impact of natural resources is meaningful to economic growth only after a certain threshold point of institutional quality has been attained. The results also shed light on the fact that the nations that have low institutional quality depend heavily on natural resources while countries with high quality institutions are relatively less dependent on natural resources to generate growth.

JEL Classifications: O11, O13, Q32

Keywords: Economic development; Natural resource curse; Institutions.

¹ Corresponding Author: Tamat Sarmidi^a, School of Economic, Faculty of Economics and Management, Universiti Kebangsaan Malaysia, 43000 UKM Bangi, Selangore Darul Ehsan, Malaysia Tel. +60133949116, Fax:+60389215789, Email: tamat@pkriscc.ukm.my.

1. Introduction

We suppose natural resource-rich countries to enjoy better economic growth when compared to those countries that are less fortunate. Surprisingly, everyday experiences and empirical studies show the reverse (Frankel 2010). It seems that abundance of natural resources is detrimental to economic growth. This puzzling phenomenon is known as the natural resource curse (NRC) hypothesis, and the literature provides at least three theories explaining the NCR: Dutch disease models, rent seeking models, and institutional explanation (Sachs and Warner, 1995 and 2001). Empirical evidence from the last two decades consistently show the prevalence of the NRC (Leite and Weidmann (1999), Gylfason (2001), Gylfason and Zoega (2002), among others).

However, recent research – for instance, Brunnschweiler (2008), Brunnschweiler and Bulte (2009), and Boyce and Emery (2010) – finds new and contradicting empirical evidence to the existing NRC literature. Abundance of natural resources evidently has a positive relationship on economic growth as well as economic welfare. Isham et al. (2003) and Frankel (2010) argue that the probable reason for inconsistency in the empirical findings by previous researches could be the different type of resources, either point or diffuse, and different economic backgrounds in the area of level of human capital, level of debt overhang, and export diversification. Brunnschweiler (2008) postulates that the inconsistencies in the empirical finding originate from the inappropriateness of resource abundance measurement to proxy natural resources in the empirical estimation. Using two new variables from the World Bank database, namely the total natural capital and subsoil wealth, Brunnschweiler (2008) finds a positive and robust relationship between natural resource abundance and economic growth for more than 90 economies.

More interestingly, these two outwardly contradicting groups who investigate the NRC hypothesis unanimously agree on the importance of good institutions. It is found that economies with abundant natural resources and at the same time better institutional quality and governance such as strong democratic accountability, high law and order, lower corruption, or higher integration among government institutions are evident to have better economic growth and higher human welfare (Damania and Bulte, 2003 and Mehlum et al., 2006). This is because superior institutional quality could be very effective in nullifying the curse through avoidance of rent-seeking behavior (Auty 2001), reducing corruption (Ishan et al. 2005 and Robinson et al. 2006), lowering the risk of violent civil conflict (Collier and Hoeffler, 2005), and accelerating efficient resource allocation (Atkinson and Hamilton, 2003).

However, previous studies that delve into the institutional quality and resource curse hypothesis have imposed an important *a priori* restriction in their analysis; i.e., the impact of natural resource and institutional quality variables is set to be linear and monotonic to economic growth. The relationship might be non-linear, and only after a certain level of institution quality or any of its interaction terms will natural resources effectively contribute to economic growth. In other words, there may be a point only after a certain threshold of institutional quality at which the natural resources could have meaningful contribution to the economic growth. Therefore, this research affirms a possible answer to the NRC hypothesis puzzle. If natural resource abundance is indeed detrimental to economic growth, it might be true only at a low quality of institution. As institutional quality improves, the impact of natural resources on economic growth may be momentous. If so, policy makers should struggle to archive a high level of institutional quality. However, this raises another question: how high should institutional

quality be for natural resources to have a favorable effect on economic growth? At what level of institutional quality is the curse annulled?

The purpose of this paper is to examine the relationship between natural resources and economic growth by considering the threshold level of institutional quality using the Brunneweiller (2008) dataset. Specifically, the questions addressed in this paper are as follows.

(1) Why are the empirical findings on the issues of NRC hypothesis far from conclusive? This research offers analysis that favors the importance of high institutional quality and good governance that could provide an answer to the NRC puzzle. (2) Are different levels of institutional quality inherent in the economy contributors to the NRC effect? This study attempts to answer this second question by offering a threshold level at which the effect of natural resources on economic growth is positive or otherwise. In other words, is there any threshold level of institutional quality above which natural resources affect economic growth rate differently? This paper employs relatively new econometric methods for threshold estimation and inference, as proposed by Hansen (1996 and 2000).

The remainder of the paper is organized as follows. In Section 2, we briefly discuss the theory and recent evidence of the NCR hypothesis and institutions. Section 3 describes the dataset used in the empirical analysis and the layout of the econometric procedures. Section 4 discusses the estimation results, and finally, Section 5 offers concluding remarks.

2. Institutions and natural resource curse hypothesis

Economists generally agree that poor or good results from any growth policies are largely dependent on the level of institutional quality inherent in the economy and not natural resource

abundance (Barro 1991 and North 1994). Regardless of how good the policy is or how many resources they have, if the institutions – either public or private – are not accommodating, then the desired results from such good resource policy will shatter. Nelson and Sampat (2001) technically define institutions as ‘social technologies’ that are positively related to economic performance. They postulate that, when institutions are of low quality, due to frequent changes of rules, high levels of corruption, widespread nepotism, and weak law enforcement, the markets will not function well and may lead to high market volatility, and then the efficient allocation of resources may be severely affected. In contrast, high-quality market characteristics play an important role in promoting an efficient and low-risk investment opportunity that could be vital to providing a better environment for sustainable economic growth.

Rodrik et al. (2004) argue that a low-quality institution through which natural resources is channeled to an economic activity could aggravate information asymmetries and adversely affect resource allocation efficiency if used by the perverse politician. Then the decision made by the authority might be politically rational but economically inefficient. On the other hand, high-quality institutions could provide an efficient mean for channeling information about market conditions and participants by facilitating mutual cooperation between market actors that could eventually reduce transaction costs and increase efficiency. Therefore, institutions could act as a tool that reverses the negative association between natural resource wealth and poor outcomes. Good institutional arrangement is also crucial to the management of optimal and efficient resources².

² Leite and Weidmann (2002) explain that the negative associations among institutions, resource abundance, and economic growth are insufficient in establishing the direction of causality. Questions regarding the cause and the effect still remain unresolved. However, these researchers are in the position to argue that poor institutions are a result of resource abundance rather than the cause.

Mehlum, Moene, and Torvik (2006) explain that the high-quality institution is one that is producer friendly and not grabber friendly. In a similar argument, they divide institutions into two categories. The first group includes the institutions that are grabber friendly and the second group include those that are producer friendly. These researchers conjecture that the curse is only effective under a grabber-friendly institution and not for the latter. If the institution is producer friendly, then resource-rich countries are hypothesized to attract more producers to involve in production and then eventually increase growth; this is not the case for grabber-friendly institutions. Upon empirically testing the hypothesis for 87 countries, the result favors the idea that producer-friendly institutions could reduce significantly the effect of the resource curse. They find that countries like the U.S., Canada, Norway, and Australia are curse-free and enjoy high economic growth due to their producer-friendly institutions with exceptionally high quality (Larsen 2005). However, their study does not clearly differentiate between good qualities that are associated with the producer-friendly institutions in contrast to the grabber-friendly ones.

3. Empirical model

The empirical model is based on Brunnschweiler (2008), in which the empirical linkages between natural resources and growth use the following linear cross-country growth equation:

$$RGDPC_i = \alpha_0 + \alpha_1 R_i + \alpha_2 INS + \alpha_3 X_i + \varepsilon_i \quad (1)$$

where $RGDPC_i$ is the real GDP per capita in country i , R_i is the country's natural resource abundance, INS is institutional quality, X is a vector of controls (initial income per capita, latitude), and ε_i is a noise term. Since we use logs, the effect of natural resources on real GDP per capita is expressed as elasticity.

To test the hypothesis outlined in the previous section, we argue that the following Equation (2) is particularly well suited to capture the presence of contingency effects and to offer a rich way of modeling the influence of institutional quality on the impact of natural resources in economic growth. The model, based on threshold regression, takes the following form:

$$RGDPC_i = \begin{cases} \beta_0^1 + \beta_1^1 R_i + \beta_2^1 INS_i + \beta_3^1 X_i + e_i, & INS \leq \lambda \\ \beta_0^2 + \beta_1^2 R_i + \beta_2^2 INS_i + \beta_3^2 X_i + e_i, & INS > \lambda \end{cases} \quad (2)$$

where INS (i.e., level of institutional development) is the threshold variable used to split the sample into regimes or groups, and λ is the unknown threshold parameter. This type of modeling strategy allows the role of natural resources to differ depending on whether institutions are below or above some unknown level of λ . In this equation, institutions act as sample-splitting (or threshold) variables. The impact of natural resources on real GDP per capita will be β_1^1 and β_1^2 for countries with a low or high regime, respectively. It is obvious that, under the hypothesis $\beta^1 = \beta^2$, the model becomes linear and reduces to (1).

The first step of our estimation is to test the null hypothesis of linearity $H_0 : \beta^1 = \beta^2$ against the threshold model in Equation (2). We follow Hansen (1996, 2000), who suggests a heteroskedasticity-consistent Lagrange Multiplier (LM) bootstrap procedure to test the null hypothesis of a linear formulation against a threshold regression alternative. Since the threshold parameter λ is not identified under the null hypothesis of the no-threshold effect, the p values are computed by a fixed bootstrap method. Hansen (2000) shows that this procedure yields asymptotically correct p values. It is important to note that, if the hypothesis of $\beta^1 = \beta^2$ is rejected and a threshold level is identified, we should test again the threshold regression model

against a linear specification after dividing the original sample according to the threshold thus identified. This procedure is carried out until the null of $\beta^1 = \beta^2$ can no longer be rejected.

Even though natural resources may have a positive effect on growth, the results may have been driven by resource-rich countries with high institutional quality. In order to examine this possibility, Equation (2) is extended as follows to include an interaction term between institutions and natural resources:

$$RGDPC_i = \alpha_0 + \alpha_1 R_i + \alpha_2 INS + \alpha_3 (R \times INS)_i + \alpha_4 X_i + \varepsilon_i \quad (3)$$

If α_3 is negative and statistically significant, this implies that the negative growth effect increases as institutional quality improves. On the other hand, if α_3 is positive and significant, this indicates that the negative growth effect diminishes as institutional quality improves. Equation (3) is estimated using the threshold regression technique.

4. The data

This study employs cross-country estimations in order to estimate Equation (2). The number of countries is 90, and the sample period spans from 1984 to 2005.

Following Brunnschweiler (2008), three natural resource indicators are employed in the analysis, namely (i) primary exports over GDP (*sxp*); (ii) average total natural capital (*natcap*) – this measure includes subsoil assets, timber resources, non-timber forest resources, protected areas, cropland, and pastureland; and (iii) average subsoil assets (*subsoil*) – this measure includes energy resources and other mineral resources. The *sxp* dataset is obtained from Sachs and Warner (1995), whereas the *natcap* and *subsoil* dataset are gathered from World Bank.

The institutions dataset employed is from the International Country Risk Guide (ICRG), a monthly publication of Political Risk Services (PRS). In this study, four PRS indicators were used to measure the overall institutional environment: (i) corruption, which measures excessive patronage, nepotism, job reservation, ‘favour-for-favours,’ secret-party funding, and suspicious ties between politics and business. It is hypothesised that a high level of corruption distorts the economic and financial environment and reduces the efficiency of the government and businesses by enabling people to assume positions of power through patronage rather than ability. The index ranges between zero and six; the higher the corruption, the lower the index; (ii) rule of law, which reveals the degree to which citizens are willing to accept established institutions to make and implement laws and to adjudicate disputes; (iii) bureaucratic quality, which represents autonomy from political pressure, strength, and expertise to govern without drastic changes in policy or interruptions in government services, as well as the existence of an established mechanism for recruitment and training of bureaucrats; and (iv) government effectiveness, which measures the quality of public services, the quality of the civil service and the degree of its independence from political pressures, the quality of policy formulation and implementation, and the credibility of the government’s commitment to such policies. These four variables were scaled from 0 to 6, where higher values implied better institutional quality and vice versa.

The real GDP per capita is expressed in USD at constant 2000 prices, and latitude is the location of the country. All dataset are obtained from World Development Indicators.

5. Results

Equation (2) has been estimated using four different models depending on the institution indicator used (Model A: Rule of Law; Model B: Government Effect; Model C: Corruption; and Model D: Buraucratic Quality). Employing splitting sample threshold method from Hansen (1996 and 2000) to investigate the NRC hypothesis with three different measures of natural resources, namely the share of resources export to GDP (*sxp*) as in Sachs and Warner (1995), total natural capital (*natcap*) and subsoil wealth (*subsoil*) as used by Brunnschweiler (2008). The results of each model are presented in Tables 1a, 1b, 1c, 2a, 2b, and 2c. The letters *a*, *b*, and *c* after number 1 or 2 refer to an estimation using the share of natural resource export to GDP (*sxp*), total natural capital (*natcap*) and subsoil wealth (*subsoil*), respectively. The number 2 refers to the estimation of Equation (3) with an interaction of natural resources and institutional quality while number 1 is without an interaction term.

This study has revealed several interesting results. First, the result shows (as shown in Tables 1a, 1b, and 1c or 2a, 2b, and 2c) that the *p*-value of the hypothesis of the no threshold effect as computed by a bootstrap method with 1,000 replications and 15% trimming percentage are rejected at a very highly significant level with and without the interaction term irrespective of the models. The finding clearly indicates that the relationship between economic growth and natural resources are non-linear, and therefore the imposition of *a priori* monotonic restriction on the relationship also can be very misleading. The finding provides a better explanation for a dynamic rich relationship between natural resources and economic growth. Natural resources can effectively contribute to economic growth only after a certain level of institution quality or any of its interaction terms.

Second, the presence of threshold level also indicates that the sample can be split into two different groups depending on the level of institutional quality. Any country that has an institutional quality less than the threshold level is considered a low-quality institution, while one with quality greater than the threshold values is classified as a high-quality institution. The behavior of the relationships between natural resources and economic growth are different for low- and high-quality institutions. For instance, Table 1c depicts that the hypothesis of NRC is rejected at a lower level of institutional quality for Models A, B, C, and D. The coefficients of the subsoil variable for these models are 0.532, 0.505, 0.520, and 0.549, respectively, and significant at the 5 percent level at least. However, as institutions get better (above the threshold level) the contribution of natural resources is negligible. Another example is found in Table 2a where, at a lower level of government effect (< 0.47), the coefficient of β_1^1 is -34.1, while at a higher level (>0.47) of government integrity, the results dramatically change to only -9.47.

In addition, the regression's result from Equation 3 has provided new insight into the understanding of the resource curse. For instance, Table 2a Model B shows that the global as well as the threshold regression coefficient for natural resources is negative, thus confirming the NRC hypothesis. However, interestingly, the interaction term between natural resources and institutional quality from the regression is positive and significant. The negative coefficient of natural resources and then followed by a positive coefficient of interaction term is a sign of the NRC getting weaker as the government effect gets stronger. If the government effect reached 1.342 level, than it will cancel out the effect of the resource curse. Out of 90 countries, our sample shows that 64 countries have the sufficient institutional quality to insulate the economy from the resource curse.

6. Conclusion

In this paper, we re-examine the well-known empirical puzzle of the resource curse hypothesis using a threshold regression with reference to different institutional quality. In particular, we endogenously determine the threshold level of institutional quality and then use this threshold point to test the different effects of natural resources on economic growth at low-institutional quality in comparison to natural resources within high-institutional quality countries.

There are several major findings in this paper. First, *a priori* monotonic restriction on the study of NRC could lead to a premature conclusion. In this study, we consistently fail to reject the presence of the threshold effect in the estimation regardless of models. Furthermore, the study highlights the importance of good quality of institution to neutralize the effect of the natural resource curse. Resource policy will be effective only under a good institution. Abundance resource countries but with weak institutional quality will not be better off in economic growth if compared to the poor resource economies. Further, this study also shows that high quality institution nation is less dependent on the natural resources to generate economic growth.

In summary, a nation desiring to benefit fully from its natural resources should not neglect the important role of good institutions for a sustainable economic growth, and with good institutions, the NCR puzzle can also be challenged.

References

- Andersen, J.J., Aslaksen, S., 2008. Constitutions and the resource curse. *Journal of Development Economics*. 87(2), 227-246.
- Atkinson, G., Hamilton, K., 2003. Savings, Growth and the Resource Curse Hypothesis. *World Development*, 31(11), 1793-1807.
- Boyce, J.R., Herbert Emery J.C., 2011. Is a negative correlation between resource abundance and growth sufficient evidence that there is a "resource curse"? *Resources Policy*, 36(1), 1-13.
- Brunnschweiler, C.N., 2008. Cursing the blessings? Natural resource abundance, institutions, and economic growth. *World Development*. 36(3), 399-419.
- Brunnschweiler, C.N., Bulte, E.H., 2008. The resource curse revisited and revised: A tale of paradoxes and red herrings. *Journal of Environmental Economics and Management* 55(3), 248-264.
- Brunnschweiler, C.N., Bulte, E.H., 2009. Natural resources and violent conflict: Resource abundance, dependence, and the onset of civil wars. *Oxford Economic Papers*. 61(4), 651-674.
- Bulte, E., Damania, R., 2005. Resource intensity, institutions, and development. *World Development*. 33(7), 1029-1044.
- Damania, R., Bulte, E., 2003. Resource for sale: Corruption, democracy and the natural resource curse, Discussion paper no. 0320. Centre for International Studies, University of Adelaide.
- Frankel, J., 2010, The natural resource curse: A survey. Forthcoming in *Export Perils*, Brenda Shaffer, ed., (Philadelphia, PA: University of Pennsylvania Press). NBER WP No. 15836; HKS RWP10-005, Feb.

- Glaeser, E.L., Porta, R.L., 2004. Do institutions cause growth? *Journal of Economic Growth* 9(3), 271-303.
- Jonathan, I., Pritchett, L., Woolcock, M., Busby, G., 2003. *The Varieties of Resource Experience*. World Bank, Washington, D.C.
- James, A., Aadland, D., 2011. The curse of natural resources: An empirical investigation of U.S. counties. *Resource and Energy Economics*. 23(2), 440-453.
- Knack, S., Keefer, P., 1995, Institutions and Economic Performance: Cross-country tests using alternative institutional measures. *Economics and Politics*. 7, 207-227.
- Kolstad, I., Søreide, T., 2009. Corruption in natural resource management: Implications for policy makers. *Resources Policy*. 34(4), 214-226.
- Kolstad, I., Wiig, A., 2009. It's the rents, stupid! The political economy of the resource curse. *Energy Policy*. 37(12), 5317-5325.
- Larsen, E.R., 2005. Are rich countries immune to the resource curse? Evidence from Norway's management of its oil riches. *Resources Policy*. 30(2), 75-86.
- Mainguy, C., 2011. Natural resources and development: The gold sector in Mali. *Resources Policy*. 36(2), 123-131.
- Papyrakis, E., Gerlagh, R., 2007. Resource abundance and economic growth in the United States. *European Economic Review*. 51(4), 1011-1039.
- Sachs, J.D., Warner, A.M., 2001. The curse of natural resources: *European Economic Review*. 45(2001), 827-838.
- Stevens, P., Dietsche, E., 2008. Resource curse: An analysis of causes, experiences and possible ways forward. *Energy Policy*. 36(1), 56-65.

- Stijns, J.P., 2006. Natural resource abundance and human capital accumulation. *World Development*. 34(6), 1060-1083.
- Van der Ploeg, F., Poelhekke. S., 2010. The pungent smell of 'red herrings': Subsoil assets, rents, volatility and the resource curse. *Journal of Environmental Economics and Management*. 60(1), 44-55.
- Williams, A., 2011. Shining a light on the resource curse: An empirical analysis of the relationship between natural resources, transparency, and economic growth. *World Development*. 39(4), 490-505.
- Yuxiang, K., Chen. Z., 2011. Resource abundance and financial development: Evidence from China. *Resources Policy*. 36(1), 72-79.

TABLE 1a: Threshold estimates of equation $RGDP_i = \begin{cases} \beta_0^1 + \beta_1^1 R_i + \beta_2^1 INS_i + \beta_3^1 X_i + e_i, & INS \leq \lambda \\ \beta_0^2 + \beta_1^2 R_i + \beta_2^2 INS_i + \beta_3^2 X_i + e_i, & INS > \lambda \end{cases}$ using share of resources export to GDP (*sxp*)

	Model A			Model B			Model C			Model D		
	Institutions: Rule of Law			Institutions: Government Effect			Institutions: Corruption			Institutions: Bureaucratic Quality		
	Linear	<0.54	>0.54	Linear	<0.17	>0.17	Linear	<3.28	>3.28	Linear	<0.69	>0.69
Constant	11.389** (2.037)	8.955** (2.128)	26.8** (2.78)	10.42** (2.00)	6.118** (2.46)	18.9** (1.83)	5.221** (1.766)	4.617* (2.02)	-2.37 (7.234)	9.573** (0.767)	4.431** (1.138)	5.673** (1.439)
RGDP ₁₉₇₀	-0.764 (0.295)	-0.339 (0.302)	-2.85 (0.402)	-0.651** (0.292)	0.010 (0.344)	-1.77 (0.27)	0.048 (0.282)	0.086 (0.311)	1.309 (1.040)	-0.241 (0.099)	-0.266 (0.101)	0.143 (0.159)
Sxp	-5.637 (2.403)	-8.951 (2.083)	-2.73 (1.42)	5.178* (2.386)	-7.92 (2.279)	-0.811 (1.780)	-6.113 (3.07)	-9.223 (2.541)	3.008 (5.120)	-0.871 (0.094)	-0.555 (0.116)	-1.206 (0.248)
Rule of Law	1.442 (1.275)	1.735** (0.452)	1.42** (0.368)									
Government Effectiveness				1.297** (0.311)	1.204 (0.646)	1.220** (0.266)						
Corruption							0.110 (0.223)	0.257 (0.221)	-0.337 (0.656)			
Bureaucratic Quality										-0.810 (0.115)	7.498** (1.609)	-0.086 (0.129)
Latitude	-0.829 (1.275)	0.557 (2.006)	0.97 (0.996)	-0.347 (1.205)	1.265 (1.948)	0.260 (0.874)	1.586 (1.845)	3.875* (1.687)	-1.209 (3.814)	-0.445 (0.103)	0.223 (0.128)	-0.667 (0.097)
Boot (p-value)	0.000			0.000			0.483			0.000		
R-sq	0.367	0.396	0.638	0.343			0.196	0.348	0.080	0.572	0.868	0.748
Het(p-value)	0.089			0.035			0.015			0.011		
No. Obs	90	60	30	90	52		90	60	30	90	39	51

Notes: The standard errors are reported in parentheses (White corrected for heteroskedasticity). Results correspond to trimming percentage of 15%. *** and ** indicate significance at the 1% and 5% levels, respectively.

TABLE 1b: Threshold regression estimates of equation $RGDP_i = \begin{cases} \beta_0^1 + \beta_1^1 R_i + \beta_2^1 INS_i + \beta_3^1 X_i + e_i, & INS \leq \lambda \\ \beta_0^2 + \beta_1^2 R_i + \beta_2^2 INS_i + \beta_3^2 X_i + e_i, & INS > \lambda \end{cases}$ using total natural resource (*natcap*).

	Model A Institutions: Rule of Law			Model B Institutions: Government Effect			Model C Institutions: Corruption			Model D Institutions: Bureaucratic Quality		
	Linear	<0.45	>0.45	Linear	<0.48	>0.48	Linear	<1.6	>1.6	Linear	<1.95	>1.95
Constant	10.89** (2.45)	5.929** (1.920)	24.69** (3.37)	10.513** (2.567)	6.064** (2.23)	20.524** (3.465)	3.596 (1.899)	-5.772 (5.453)	3.833* (1.926)	6.761** (1.204)	7.352** (1.23)	-50.622 (9.31)
RGDP ₁₉₇₀	-1.314 (0.383)	-1.656 (0.471)	-2.327 (0.468)	-1.216 (0.395)	-1.585 (0.451)	-1.714 (0.469)	-0.342 (0.425)	4.080** (0.812)	-0.403 (0.379)	0.327** (0.089)	0.298** (0.091)	0.591** (0.222)
Natcap	0.383 (0.243)	1.267** (0.344)	-0.243 (0.130)	0.337 (0.254)	1.146** (0.335)	-0.159 (0.146)	0.424 (0.255)	-1.612 (0.654)	0.445 (0.237)	-0.646 (0.086)	-0.670 (0.090)	-0.559 (0.117)
Rule of Law	1.562** (0.372)	2.446** (0.458)	0.932** (0.342)									
Government Effectiveness				1.521** (0.423)	2.304** (0.626)	0.351 (0.397)						
Corruption							-0.208 (0.274)	4.078** (0.858)	-0.254 (0.311)			
Bureaucratic Quality										0.067 (0.278)	0.272 (0.345)	25.40** (4.69)
Latitude	0.370 (1.638)	2.004 (2.165)	2.37 (1.451)	0.522 (1.575)	2.91 (2.016)	1.287 (1.130)	5.058** (1.701)	-18.853 (4.652)	5.541** (1.873)	-0.115 (0.120)	-0.173 (0.125)	0.603** (0.238)
Boot (p-value)	0.000			0.000			0.070			0.606		
R-sq	0.315	0.398	0.614	0.301	0.361	0.615	0.170	0.819	0.212	0.453	0.457	0.880
Het(p-value)	0.395			0.381			0.280			0.001		
No. Obs	77	51	26	77	51	26	77	9	68	77	68	9

Notes: The standard errors are reported in parentheses (White corrected for heteroskedasticity). Results correspond to trimming percentage of 15%. *** and ** indicate significance at the 1% and 5% levels, respectively.

TABLE 1c: Threshold estimates of equation $RGDP_i = \begin{cases} \beta_0^1 + \beta_1^1 R_i + \beta_2^1 INS_i + \beta_3^1 X_i + e_i, & INS \leq \lambda \\ \beta_0^2 + \beta_1^2 R_i + \beta_2^2 INS_i + \beta_3^2 X_i + e_i, & INS > \lambda \end{cases}$ using subsoil wealth (*subsoil*)

	Model A Institutions: Rule of Law			Model B Institutions: Government Effect			Model C Institutions: Corruption			Model D Institutions: Bureaucratic Quality		
	Linear	<0.45	>0.45	Linear	<0.4	>0.4	Linear	<4.1	>4.1	Linear	<2.69	>2.69
Constant	16.4** (2.206)	16.6** (2.12)	23.6** (2.63)	13.509** (2.198)	17.586** (2.035)	19.07** (3.34)	11.055** (1.951)	13.342** (2.099)	18.756** (2.39)	11.76** (1.58)	11.36** (2.28)	10.26** (1.89)
RGDP ₁₉₇₀	-1.97 (0.324)	-2.185 (0.316)	-2.31 (0.37)	-1.981 (0.327)	-2.286 (0.261)	-1.64 (0.469)	-1.422 (0.374)	-2.053 (0.374)	-1.402 (0.343)	-1.58 (0.303)	-1.707 (0.401)	-0.596 (0.300)
Subsoil	0.363** (0.072)	0.532** (0.083)	0.023 (0.046)	0.341** (0.068)	0.505** (0.075)	-0.004 (0.051)	0.375** (0.090)	0.520** (0.094)	0.115 (0.146)	0.358** (0.077)	0.549** (0.072)	0.091 (0.065)
Rule of Law	1.58** (0.331)	1.417** (0.452)	0.461 (0.362)									
Government Effectiveness				1.703** (0.378)	2.020** (0.579)	0.106 (0.298)						
Corruption							0.262 (0.243)	0.561* (0.245)	-0.562 (0.275)			
Bureaucratic Quality										0.653* (0.284)	0.258 (0.379)	-0.247 (0.325)
Latitude	1.771 (1.337)	2.273 (1.763)	1.422 (0.961)	1.685 (1.263)	1.772 (1.832)	1.141 (1.022)	5.619** (1.545)	7.178** (1.703)	2.233* (1.020)	5.023** (1.409)	8.61** (2.13)	2.402* (1.080)
Boot (p-value)	0.000			0.000			0.001			0.004		
R-sq	0.497	0.578	0.792	0.524	0.634	0.663	0.366	0.481	0.645	0.423	0.519	0.188
Het(p-value)	0.887			0.817			0.676			0.557		
No. Obs	60	39	21	60	37	23	60	44	16	60	37	23

Notes: The standard errors are reported in parentheses (White corrected for heteroskedasticity). Results correspond to trimming percentage of 15%. *** and ** indicate significance at the 1% and 5% levels, respectively.

Table 2a: OLS and threshold regression estimates of equation $RGDPC_i = \alpha_0 + \alpha_1 R_i + \alpha_2 INS + \alpha_3 (R \times INS)_i + \alpha_4 X_i + \varepsilon_i$ using share of resources to export (sxp)

	Model A			Model B			Model C			Model D		
	Institutions: Rule of Law			Institutions: Government Effect			Institutions: Corruption			Institutions: Bureaucratic Quality		
	Linear	<0.09	>0.09	Linear	<0.47	>0.47	Linear	<3.29	>3.29	Linear	<2.08	>2.08
Constant	10.77** (2.00)	9.74** (2.83)	16.4** (2.06)	9.71** (2.02)	5.75 (3.84)	13.9** (2.27)	5.46** (1.819)	3.899* (1.936)	-4.300 (7.65)	0.699 (0.587)	0.757 (1.065)	-2.02 (1.29)
RGDP ₁₉₇₀	-0.630 (0.289)	-0.396 (0.379)	-1.434 (0.293)	-0.504 (0.296)	0.006 (0.57)	-1.001 (0.307)	0.139 (0.305)	-0.053 (0.320)	1.34 (0.99)	0.695** (0.054)	0.670** (0.175)	0.672 (0.805)
sxp	-6.454 (1.943)	-16.72 (3.438)	-0.972 (4.130)	-5.895 (1.725)	-34.1 (17.86)	-9.479 (2.355)	-13.25 (6.29)	6.809 (13.69)	16.806 (18.51)	-0.001 (0.037)	-2.334 (1.404)	-0.026 (0.106)
Rule of Law	0.964** (0.347)	3.506** (1.031)	1.277** (0.298)									
Government Effectiveness				0.786* (0.350)	2.40 (2.97)	0.613* (0.274)						
Corruption							-0.135 (0.27)	0.886 (0.638)	-0.009 (0.549)			
Bureaucratic Quality										-0.023 (0.067)	0.524 (0.560)	-0.154 (0.283)
Latitude	-1.331 (1.298)	2.650 (2.484)	-1.004 (1.130)	-0.815 (1.235)	3.46 (2.88)	-1.164 (1.097)	1.301 (1.914)	4.013** (1.588)	-0.834 (4.098)	-0.119 (0.090)	-0.126 (0.092)	-1.719 (2.418)
Interaction	4.511* (1.993)	-17.78 (2.484)	0.652 (3.010)	4.685** (1.825)	-41.92 (26.80)	7.062** (1.882)	2.284 (1.761)	-6.289 (5.149)	-3.14 (3.62)	-0.178 (0.080)	-0.091 (0.171)	-0.264 (0.030)
Boot (p-value)	0.018			0.017			0.395			0.198		
R-sq	0.403	0.329	0.462	0.382	0.317	0.403	0.208	0.368	0.102	0.865	0.808	0.961
Het(p-value)	0.018			0.093			0.052			0.193		
No. Obs	90	46	44	90	26	64	90	60	30	90	35	55

Notes: The standard errors are reported in parentheses (White corrected for heteroskedasticity). Results correspond to trimming percentage of 15%. *** and ** indicate significance at the 1% and 5% levels, respectively.

Table 2b: OLS and threshold regression estimates of equation $RGDPC_i = \alpha_0 + \alpha_1 R_i + \alpha_2 INS + \alpha_3 (R \times INS)_i + \alpha_4 X_i + \varepsilon_i$ using total natural resource (*natcap*).

	Model A			Model B			Model C			Model D		
	Institutions: Rule of Law			Institutions: Government Effect			Institutions: Corruption			Institutions: Bureaucratic Quality		
	Linear	<0.75	>0.75	Linear	<0.47	>0.47	Linear	<1.6	>1.6	Linear	<0.75	>0.75
Constant	9.39** (1.91)	62.3 (51.0)	7.85** (1.87)	9.179** (2.05)	52.15** (10.62)	7.038* (3.26)	-0.665 (4.25)	-9.92 (6.57)	-5.17 (4.39)	9.391** (1.913)	62.35 (51.01)	4.008 (42.32)
RGDP ₁₉₇₀	-1.75 (0.401)	-2.59 (0.51)	-1.601 (0.410)	-1.568 (0.383)	-1.071 (0.771)	-1.671 (0.331)	-0.351 (0.449)	4.564** (1.043)	-0.59 (0.441)	-1.75 (0.401)	-2.592 (0.513)	-2.96** (-0.735)
natcap	0.923** (0.282)	-3.92 (5.81)	0.959** (0.284)	0.784** (0.264)	-5.212 (1.374)	1.151** (0.393)	0.946 (0.688)	-1.411 (0.644)	1.609* (0.699)	0.923** (0.282)	-3.937 (5.819)	-2.896 (1.701)
Rule of Law	8.361** (1.55)	60.2 (48.1)	7.929** (1.668)									
Government Effectiveness				7.875** (1.541)	59.52** (12.96)	7.869** (1.891)						
Corruption							1.014 (1.362)	10.82 (9.06)	2.514 (1.48)			
Bureaucratic Quality										8.361** (1.550)	60.266 (48.101)	-12.67 (11.89)
Latitude	1.145 (1.508)	-21.7 (3.84)	2.599 (1.449)	1.318 (1.493)	0.091 (3.422)	2.977 (1.438)	5.611** (1.772)	-21.513 (6.33)	6.153** (1.857)	1.145 (1.508)	-21.74 (3.84)	-0.63 (5.508)
Interaction	-0.758 (0.162)	-6.32 (5.66)	-0.759 (0.169)	-0.711 (0.157)	-7.392 (1.574)	-0.797 (0.219)	-0.149 (0.145)	-8.801 (1.086)	-0.309 (0.150)	-0.758 (0.162)	-6.324 (5.662)	-1.167 (1.507)
Boot (p-value)	0.150			0.045			0.102			0.138		
R-sq	0.437	0.895	0.411	0.400	0.389	0.392	0.182	0.826	0.241	0.437	0.895	0.411
Het(p-value)	0.576			0.523			0.28			0.576		
No. Obs	77	10	67	77	21	56	77	9	68	77	10	67

Notes: The standard errors are reported in parentheses (White corrected for heteroskedasticity). Results correspond to trimming percentage of 15%. *** and ** indicate significance at the 1% and 5% levels, respectively.

Table 2c: OLS and threshold regression estimates of equation $RGDPC_i = \alpha_0 + \alpha_1 R_i + \alpha_2 INS + \alpha_3 (R \times INS)_i + \alpha_4 X_i + \varepsilon_i$ using subsoil wealth (*subsoil*)

	Model A			Model B			Model C			Model D		
	Institutions: Rule of Law			Institutions: Government Effect			Institutions: Corruption			Institutions: Bureaucratic Quality		
	Linear	<0.45	>0.45	Linear	<0.42	>0.42	Linear	<3.08	>3.08	Linear	<0.97	>0.97
Constant	16.905** (2.04)	16.67** (2.02)	26.49** (1.905)	17.17** (2.03)	17.5** (1.96)	24.4** (2.14)	9.659** (1.99)	14.59** (2.58)	-1.257 (3.198)	9.927** (1.706)	15.5** (4.84)	8.59** (2.17)
RGDP ₁₉₇₀	-2.012 (0.289)	-2.355 (0.356)	-2.323 (0.307)	-2.051 (0.282)	-2.35 (0.274)	-1.819 (0.344)	-1.424 (0.347)	-2.231 (0.435)	-0.402 (0.380)	-1.505 (0.294)	-2.664 (0.862)	-1.171 (0.322)
subsoil	0.357** (0.069)	0.705** (0.126)	-0.423 (0.135)	0.340** (0.064)	0.600** (0.09)	-0.671 (0.201)	0.630** (0.112)	0.517** (0.111)	2.017** (0.569)	0.622** (0.112)	1.953** (0.534)	0.681* (0.300)
Rule of Law	2.899** (0.438)	-0.092 (1.133)	-0.909 (0.456)									
Government Effectiveness				3.153** (0.504)	1.193 (0.885)	-1.947 (0.674)						
Corruption							0.746** (0.292)	0.596 (0.357)	1.272 (0.660)			
Bureaucratic Quality										1.232** (0.312)	5.219 (2.746)	0.948 (0.854)
Latitude	1.515 (1.17)	2.337 (1.669)	0.575 (0.971)	1.431 (1.218)	1.619 (1.720)	0.076 (0.948)	6.04** (1.61)	6.65** (2.214)	5.93** (1.38)	5.404** (1.396)	2.968 (3.443)	5.362** (1.374)
Interaction	-0.194 (0.050)	0.256 (0.164)	0.252** (0.074)	-0.208 (0.047)	0.148 (0.112)	0.392** (0.108)	-0.088 (0.033)	-0.012 (0.062)	-0.359 (0.104)	-0.119 (0.042)	-2.366 (0.850)	-0.148 (0.094)
Boot (p-value)	0.050			0.017			0.003			0.039		
R-sq	0.569	0.594	0.829	0.601	0.652	0.755	0.397	0.454	0.610	0.453	0.72	0.336
Het(p-value)	0.725			0.529			0.860			0.59		
No. Obs	60	39	21	60	38	22	60	36	24	60	10	50

Notes: The standard errors are reported in parentheses (White corrected for heteroskedasticity). Results correspond to trimming percentage of 15%. *** and ** indicate significance at the 1% and 5% levels, respectively.