



Munich Personal RePEc Archive

Political Uncertainty and Accounting Conservatism: Evidence from the U.S. Presidential Election Cycle

Dai, Lili and Ngo, Phong T. H.

Erasmus University Rotterdam, Australian National University

4 January 2013

Online at <https://mpra.ub.uni-muenchen.de/43606/>

MPRA Paper No. 43606, posted 07 Jan 2013 03:56 UTC

**Political Uncertainty and Accounting Conservatism:
Evidence from the U.S. Presidential Election Cycle**

Lili Dai

Erasmus University Rotterdam

Phong T. H. Ngo⁺

Australian National University

ABSTRACT

We document a positive association between political uncertainty and accounting conservatism. In the year prior to a U.S. presidential election, on average, accounting conservatism increases by nearly 20 percent. This election year effect is stronger when the election is closer, when the incumbent president is not seeking re-election, and when the incumbent party is Democrat. In the post-election year, conservatism is lower relative to the non-election period when the incumbent party wins, but remains higher under an opposition party victory. Moreover, the election year effect varies across industries and companies, and remains unchanged under different empirical specifications. For example, the impact of an election is greater for politically sensitive industries and for companies with less anti-takeover provisions, and is robust when we control for the business cycle. Collectively, we show the political process is important in determining accounting choices through the uncertainty channel.

JEL Classification: M41, D72, G34, G38

Keywords: accounting conservatism, political uncertainty, election cycle

⁺ Lili Dai is from Erasmus School of Economics, Erasmus University, PO Box 1738, Rotterdam, 3000 DR, the Netherlands, Tel, +31 (0)10 4081326, Email, ldai@ese.eur.nl; and Phong Ngo is from College of Business and Economics, Australian National University, Canberra, ACT 0200, Australia, Tel, +61 2 6125 1079, Email, phong.ngo@anu.edu.au.

**Political Uncertainty and Accounting Conservatism:
Evidence from U.S. Presidential Election**

ABSTRACT

We document a positive association between political uncertainty and accounting conservatism. In the year prior to a U.S. presidential election, on average, accounting conservatism increases by nearly 20 percent. This election year effect is stronger when the election is closer, when the incumbent president is not seeking re-election, and when the incumbent party is Democrat. In the post-election year, conservatism is lower relative to the non-election period when the incumbent party wins, but remains higher under an opposition party victory. Moreover, the election year effect varies across industries and companies, and remains unchanged under different empirical specifications. For example, the impact of an election is greater for politically sensitive industries and for companies with less anti-takeover provisions, and is robust when we control for the business cycle. Collectively, we show the political process is important in determining accounting choices through the uncertainty channel.

JEL Classification: M41, D72, G34, G38

Keywords: accounting conservatism, political uncertainty, election cycle

I. INTRODUCTION

We investigate the association between political uncertainty and accounting conservatism by exploiting the exogenous variation in political uncertainty induced by the U.S. presidential election cycle. The partisan view of politics argues that Republicans and Democrats differ with respect to economic, regulatory, and social policies, e.g., taxation rules, government expenditures, and welfare reforms (Alesina 1987). Therefore, elections bring about potential change in the status quo leading to an increase in uncertainty about future policy outcomes. For example, during the 2012 U.S. presidential election, Republican candidate, Mitt Romney, was quoted as saying that, if elected, he would repeal the two major reforms introduced under the Obama administration: the Dodd-Frank Act and the universal health care reforms. Referring to the health care reforms, he vowed: "What the Court did not do on its last day in session, I will do on my first day if elected President of the United States, and that is I will act to repeal Obamacare."¹

Accounting conservatism is defined as the asymmetric verifiability of accounting gains versus losses, i.e., the verifiability threshold for gains is greater than that for losses (e.g., Basu 1997; Watts 2003a, 2003b). To develop a *political uncertainty hypothesis*, we argue that in an election year, uncertainty about the election outcome and consequent changes in public policy can have an effect on conservatism for several reasons.² First, changes in policy might impact on the profitability of assets that are already in place – a direct impact on the expected cash flows of a company. For example, legislative and regulatory changes can alter compliance costs, while changes in government expenditures and taxation rules might increase the uncertainty of firms'

¹ See, e.g., news reports by *Washington Post - Wonkblog*, Khimm (2012), and *ABC News Blogs*, Friedman (2012).

² Another explanation for the effect of elections on accounting choices is based on *political cost hypothesis* (Watts and Zimmerman, 1978, 1986), i.e., manipulating accounting numbers to reduce costs related to adverse political scrutiny, especially during the election period (Ramanna and Roychowdhury 2010). We cannot rule out this alternative explanation, but indeed find evidence to support both political uncertainty and political cost hypotheses.

accounting performance. Moreover, the discount rate can also be affected by uncertainty about future government policies (Brogaard and Detzel 2012). As a consequence, an increase in political uncertainty implies an increase in the verifiability threshold for profits relative to losses, that is, an increase in conservatism.³ The increase in uncertainty about future cash flows for a firm's assets in place, *ceteris paribus*, also means that “shirking” is harder to detect leading to a rise in agency costs, and increased demand for conservatism.⁴

Further, policy changes can also influence the investment opportunity set of a company. Julio and Yook (2012) find that rising political uncertainty leads to a reduction in corporate investment before elections – firms defer investments until political uncertainty is resolved. Under the standard assumption that the investment, if the firm had undertaken it, does not increase the firm's investment opportunity set, a company should have more growth options because of deferred investment in an election year.⁵ Since growth options are positively related to agency costs (Smith and Watts, 1992), and conservatism is an efficient corporate governance response to agency costs (see Watts, 2003a), we expect an increase in conservatism in the year leading up to an election. Finally, political uncertainty might be positively associated with the cash holdings of a company. In addition to showing a reduction in investment, Julio and Yook (2012) find an increase of cash holdings in the pre-election period, which can raise agency costs, because the amount of “free” cash available increases for managers to pursue their personal interests (Jensen

³ The verifiability threshold for bad news might also increase. However, since investors are more concerned about the downside risk, the demand for conservatism will increase with uncertainty (Watts 2003a).

⁴ More formally, in standard models of outside financing with agency cost, the proportional difference between the probabilities of success if a manager ‘works’ and if the manager ‘shirks’ (known as the information ratio) is a key determinant of the size of agency costs. The larger the difference, the easier it is to detect shirking and the less likely the manager will shirk (Tirole 2001). All else equal, when political uncertainty lowers the probability of success if a manager works, then it will be more difficult to detect shirking (a fall in the information ratio) – leading to an increase in agency costs.

⁵ The classic papers analysing real options and investment (e.g., McDonald and Siegel 1986; Ingersoll and Ross 1992) assume that the investment does not increase the firm's investment opportunities.

1986). As argued previously, increasing agency costs in turn leads to an increase in accounting conservatism.

The empirical analysis is based on U.S. companies from *Compustat* for the period of 1971-2011, which covers 11 presidential elections from 1971 to 2012. We examine the association between political uncertainty and accounting conservatism using Kahn and Watts' (2009) conservatism measure, *C-Score*. A consistent picture emerges: accounting conservatism increases in the year prior to the election date. The point estimates imply that in the year leading up to an election, accounting conservatism increases by 17 to 19 percent (depending on specifications). The results are robust to different empirical specifications and estimation techniques. We also perform sensitivity tests, e.g., controlling for the business cycle, political connections, and audit quality, conducting the analysis in pre- and post-SOX periods, and using alternative pre-election windows. The results remain unchanged.

Further, we find that the effect of political uncertainty on conservatism is stronger when an election is more competitive. This result is intuitively appealing as a closer election is associated with greater uncertainty about the eventual electoral outcome. Specifically, a decrease in the margin for victory by 10 percent increases the overall impact of political uncertainty on conservatism by 25 to 59 percent. We also find that the impact of political uncertainty becomes greater when the incumbent president is not seeking a re-election and when the incumbent party is Democrat.⁶ In the post-election year, conservatism remains higher relative to the non-election period, but lower relative to the election year. Further tests indicate that this post-election effect is driven by opposition party victories, i.e., after the election, conservatism is actually lower compared to the non-election period when the incumbent party wins. This suggests that under an

⁶ When an incumbent president is not seeking re-election, the likelihood of party turnover tends to be higher leading to higher political uncertainty (Gelman and King 1990). A continuation of Democrat regime might be regarded as bad news by the markets (Santa-Clara and Valkanov 2003), and results in high conservatism.

opposition party victory, political uncertainty is higher than under an incumbent party victory, most likely since the policy stance of incumbent party is more predictable.⁷ Collectively, these results are consistent with the *political uncertainty hypothesis*.

Political uncertainty will impact on firms differently depending on factors that are associated with the political process and accounting conservatism, e.g., the firm's exposures to politics, government spending, contracting environment, litigation risk, taxation, and regulation.⁸ Accordingly, we investigate cross-sectionally how these industry and firm level factors can exacerbate or moderate the positive relation between political uncertainty and conservatism. Across industries, we find that the election year effect on conservatism is stronger for politically sensitive industries, heavily regulated industries, and industries that are sensitive to contract enforcement, while conservatism tends to be weaker for labor intensive industries and industries with higher international trade exposure. At the firm level, we show that the increase of conservatism in election year is larger for firms with better external governance (lower G-index) mechanisms but smaller for firms with better internal governance (higher ratio of independent directors), and higher litigation risk.

This paper contributes to three streams of literature. First, it contributes to the literature on accounting conservatism.⁹ Watts' (2003a) *regulatory explanation* argues that conservatism can be driven by the political process in which losses from overstated accounting numbers are more observable and usable for regulatory oversight. In this study, we explore the nature of conservatism in a particular political process, i.e., the U.S. presidential election, using a firm-year conservatism measure, *C-Score*, developed by Khan and Watts (2009). More importantly,

⁷ Please refer to Figure 1 for an illustration of conservatism cycles around the presidential election.

⁸ Watts (2003a) provides four explanations for conservatism, i.e., contracting, litigation, taxation, and regulation.

⁹ For a review, please see Watts (2003a, 2003b). Recent studies on conservatism include Ball and Shivakumar (2005), Ahmed and Duellman (2007), LaFond and Roychowdhury (2008), and LaFond and Watts (2008).

beyond the scrutiny role of the political process (*political cost hypothesis*), we document another channel associated with *political uncertainty*, through which the political process can also influence accounting conservatism.¹⁰ Second, it also contributes to the literature on the impact of election cycles on economic outcomes, which argues that politicians have incentives to induce favorable macroeconomic outcomes before elections (e.g., McRae 1977; Nordhaus 1975; Rogoff and Sibert 1988). We add to this line of literature, and show that the election cycle plays an important role not only at the macro-level, but also at the microeconomic level, i.e., accounting numbers manipulated by managers at company level.

Third, the paper contributes to an emerging literature on the role of politics in determining corporate performance and corporate policy. Several asset pricing studies examine the impact of political uncertainty on stock returns (e.g., Belo et al. 2012; Boutchkova et al. 2012; Brogaard and Detzel 2012). For example, Belo et al. (2012) show that highly government-exposed companies experience higher stock returns during Democratic presidencies. In a recent corporate finance study, Julio and Yook (2012) find a decline in corporate investment corresponding with the timing of national elections around the world.

From the perspective of accounting practices, Ramanna and Roychowdhury (2010) focus on a sample of 563 firms that are politically connected with U.S. congressional candidates, and find that in the election year of 2004, when a firm has more outsourcing activities, it tends to report lower discretionary accruals to avoid political oversight (an example of *political cost*), because outsourcing was a major campaign issue in 2004.¹¹ This paper differs from their work and complements the literature by showing how political forces in general (not only for the

¹⁰ We find strong evidence for the uncertainty explanation, e.g., the election effect is stronger for closer elections, but also evidence that may be consistent with both explanations, e.g., the effect is stronger for regulated industries.

¹¹ Other studies, examining the political role in accounting practices, include Ramanna (2008), Guay (2010), Chaney et al. (2011), Kido et al. (2012), and Lent (2012).

outsourcing issue) affect another dimension of financial accounting choices, accounting conservatism,¹² for a large sample of firms in a period related to 11 elections.¹³ More importantly, from a theoretical view, in addition to the political cost channel, we find that accounting conservatism can be influenced through another undocumented channel related to political uncertainty. Lastly, the dynamics of accounting conservatism around elections shed some light on several unanswered questions arising from Ramanna and Roychowdhury (2010): whether their election results related to discretionary accruals in 2004 can be generalized to other aspects of accounting choices, and to other election years; and whether there is a time-series or cross-sectional variation in this political impact depending on the nature of elections, industries and companies.

There is a growing interest in the impact of politics on firm performance and management decisions, since politics and public policy play central roles in determining the external boundaries in which a firm operates. In this paper, we document a new stylized fact – the tendency for firms to become more conservative with their accounting choices in the U.S. presidential election years. The paper improves our knowledge on how political forces can shape the nature of financial reporting. The rest of this paper is organized as follows. Section II discusses data sources and empirical design. Section III presents results of main and supplemental analyses. Section IV concludes.

II. DATA AND EMPIRICAL DESIGN

Our full sample includes 147,894 firm-year observations of all the U.S. public firms from *Compustat - Fundamentals Annual Database and CRSP Monthly Stock Database* for fiscal years

¹² Ball and Shivakumar (2006), and Goh and Li (2011) argue that lower discretionary accruals are not necessarily associated with higher conservatism.

¹³ Our sample covers both politically connected and non-connected firms.

1971-2011.¹⁴ We construct firm-year estimates of accounting conservatism, *C-Score*, developed by Khan and Watts (2009), which allows us to capture both cross-sectional and time-series variation of asymmetric earnings timeliness. Khan and Watts (2009) develop their firm-year measure of conservatism based on Basu's (1997) cross-sectional model:

$$X_i = \beta_1 + \beta_2 D_i + \beta_3 R_i + \beta_4 D_i R_i + \varepsilon_i \quad (1)$$

where X is earnings, R is returns, and D is denoted as one when $R < 0$ (bad news), and zero otherwise. The incremental timeliness for bad news over good news is represented by β_4 , while the timeliness measure for good news is β_3 . They further define G-Score (timeliness of good news) and C-Score (incremental timeliness of bad news) as functions of firm size, market-to-book ratio and leverage:

$$\text{G-Score} = \beta_3 = \mu_1 + \mu_2 \text{Size}_i + \mu_3 M/B_i + \mu_4 \text{Lev}_i \quad (2)$$

$$\text{C-Score} = \beta_4 = \lambda_1 + \lambda_2 \text{Size}_i + \lambda_3 M/B_i + \lambda_4 \text{Lev}_i \quad (3)$$

Substituting G-Score and C-Score into *equation (1)*, an annual cross-sectional regression model is derived to estimate the parameters (μ_1 - μ_4 and λ_1 - λ_4) in equations (2) and (3). Then, the estimated parameters (λ_1 - λ_4) can be used to calculate the C-Score in equation (3).

$$\begin{aligned} X_i = & \beta_1 + \beta_2 D_i + R_i(\mu_1 + \mu_2 \text{Size}_i + \mu_3 M/B_i + \mu_4 \text{Lev}_i) + D_i R_i(\lambda_1 + \lambda_2 \text{Size}_i + \lambda_3 M/B_i + \lambda_4 \text{Lev}_i) \\ & + (\delta_1 \text{Size}_i + \delta_2 M/B_i + \delta_3 \text{Lev}_i + \delta_4 D_i \text{Size}_i + \delta_5 D_i M/B_i + \delta_6 D_i \text{Lev}_i) + \varepsilon_i \end{aligned} \quad (4)$$

Election timing in the U.S. is determined by law. Since 1845, Election Day occurs on the first Tuesday after November 1. We use the U.S. presidential election cycle as an exogenous shock to political uncertainty to investigate its impact on accounting conservatism. Presidential elections follow a four year cycle on even numbered years. The sample period is associated with 11

¹⁴ We extract year-end earnings announcement dates from *Compustat - Fundamentals Quarterly Database*, which limits our sample period to start from 1971. Based on the differences between announcement dates and election dates, we define the pre-election and post-election dummy variables used in the analysis.

presidential elections from 1971 to 2012.¹⁵ Let the election date be day $t = 0$, we construct a *PRE-ELECTION* dummy which equals one if a firm's annual earnings announcement date falls in the window $[-360, 0)$, and zero otherwise.¹⁶ The baseline regression takes the form:

$$C\text{-SCORE} = \alpha + \gamma_1 \text{PRE-ELECTION} + \gamma_x \mathbf{X} + \text{Industry/Firm Fixed Effects} + \varepsilon \quad (5)$$

where α is the intercept and \mathbf{X} is a vector of control variables. The coefficient of *PRE-ELECTION*, γ_1 , captures the change in accounting conservatism in the period leading up to a presidential election and is expected to be positive and significant according to the *political uncertainty hypothesis*. The control variables include: 1) *SIZE*, natural log of the market value of equity; 2) *MARKET/BOOK*, market-to-book ratio; 3) *LEVERAGE*, long-term and short-term debt deflated by market value of equity; 4) *VOLATILITY*, standard deviation of daily firm-level returns; 5) *NOACC*, non-operating accruals scaled by lagged assets; 6) *CFOA*, cash flow from operations deflated by lagged assets; 7) *INVEST CYCLE*, a decreasing measure of the length of the investment cycle; 8) *BID-ASK*, bid-ask spread scaled by the midpoint of spread; and 9) *AGE*, firm age in years (Khan and Watts 2009).

Table 1 presents summary statistics of the variables for the full sample as well as the following sub-samples: 1) election year; 2) post-election year; and 3) non-election period. We find that conservatism as measured by C-Score is higher in election years (mean C-Score is 0.144) than in post-election years and non-election periods (mean C-Scores are 0.131, and 0.117, respectively). This suggests that compared to non-election period, on average, conservatism is 23 percent higher in the year prior to an election, and remains 12 percent higher in the year following the election in a univariate analysis.¹⁷ In the next section, we perform multivariate

¹⁵ Please refer to the *U.S. Election Statistics: A Resource Guide* from the *Library of Congress*.

¹⁶ We find results robust, when using different pre-election windows, e.g., 90 or 180 days.

¹⁷ An increase by 23 percent is equal to $(0.144-0.117)/0.117$, and 12 percent is equal to $(0.131-0.117)/0.117$.

regressions to examine whether the election-year effect holds after controlling for other factors that can influence accounting conservatism.

II. RESULTS

Elections and Conservatism

We present the results of our baseline model in Table 2. Consistent with our expectations, the coefficient estimate on PRE-ELECTION is positive and significant at the one percent level across all specifications. The result is robust to the inclusion of control variables, industry fixed effects as well as firm fixed effects. Conservatism increases in the year prior to elections, with coefficients ranging from 0.0214 (Model 3) to 0.0245 (Model 4). Given an unconditional mean C-Score equal to 0.127 in Table 1, these estimates translate to an increase in conservatism by 17-19 percent. In Model 6, we include control variables and firm fixed effects and find that *with-in* a firm, MARKET/BOOK is negatively correlated with conservatism while LEVERAGE, CFOA, INVEST CYCLE, VOLATILITY and AGE are positively correlated with conservatism.

To further test the *political uncertainty hypothesis*, we study whether the degree of electoral competition exacerbates the election year effect in conservatism. Closer elections entail more uncertainty about the eventual winner and therefore are associated with higher political uncertainty. We construct a variable called *MARGIN* to measure the degree of political competition as follows: Let c be the total number of electoral colleges, let w be the number of votes cast for the winning candidate, and let l be the number of votes cast for the losing candidate, then the *MARGIN* is denoted as:

$$MARGIN = (w - l) / c \tag{6}$$

The coefficient of interest is related to the interaction term between PRE-ELECTION and MARGIN ($PRE \times MARGIN$). We expect that strong political competition (small MARGIN) is associated with high political uncertainty and therefore greater conservatism. Table 3 presents the results from this analysis. As expected, the coefficient on PRE x MARGIN is negative and significant at the one percent level in all specifications. The coefficient estimates for the PRE x MARGIN range from -0.0594 (Model 6) to -0.145 (Model 4) implying that a 10 percent decrease in MARGIN results in an increase in the election effect by 25 to 59 percent.¹⁸

Incumbency Advantage and Party Affiliation

In this sub-section, we investigate whether: 1) the incumbency advantage; and 2) the party affiliation of the incumbent president exacerbate the election year impact on conservatism. Prior studies document the *incumbency advantage* phenomenon: *ceteris paribus*, incumbents have an advantage at the polls (e.g., Gelman and King 1990). Accordingly, we expect elections in which the incumbent president is running for re-election to be less competitive, and thereby lower political uncertainty. We capture exogenous variation in incumbency advantage using an indicator variable, *INCUMB*, equal to one if the incumbent president is seeking re-election,¹⁹ and expect the coefficient on the interaction term between *INCUMB* and PRE-ELECTION to be negative. Table 4 presents the results. We find a negative and significant coefficient on PRE x *INCUMB*, implying the conservatism in an election is reduced if the incumbent president is seeking re-election, which is consistent with the *political uncertainty hypothesis*.

Santa-Clara and Valkanov (2003) show that excess stock market returns are higher under Democratic than Republican presidencies after controlling for business-cycles and stock

¹⁸ An increase by 25 percent is equal to $(0.0594 \times 0.1)/0.0240$, and 59 percent is equal to $(0.145 \times 0.1)/0.0214$. The benchmarks of election effect, 0.0240, and 0.0214, are estimates of PRE-ELECTION in Table 2 for same models.

¹⁹ Legislation in the U.S. allows a president to serve a maximum of two consecutive terms, meaning an incumbent president can run for re-election only once. Therefore, to a large extent, *INCUMB* is exogenously determined by law - during our sample period, after their first presidential term, all the incumbent presidents seek re-election.

riskiness. Belo et al. (2012) further show that during Democratic presidencies, firms with high government exposure experience higher cash flows and stock returns. They also show that the reverse is true for Republican presidencies. To this end, we investigate whether incumbent party affiliation alters the pattern of conservatism around elections. We define an indicator variable, *DEM*, equal to one if the incumbent president is a Democrat and zero otherwise. The coefficient of interest is the interaction between *DEM* and *PRE-ELECTION* (*PRE x DEM*). We have no strong prediction for the coefficient on *PRE x DEM*, although we may expect the coefficient to be positive, because a victory by the Democratic party is regarded as bad news around the announcement date of election outcome (Santa-Clara and Valkanov 2003), which might make managers more conservative in the election year. In Table 4, after controlling for the effect of incumbency advantage (Models 7 – 9), we find that the coefficients on *DEM* are negative, and the coefficients on *PRE x DEM* are positive (both significant at one percent level). That is, prior to an election, managers are more conservative with financial reporting when a Democratic regime is very likely to be continued,²⁰ however, during a long-term Democratic mandate, companies tend to have lower accounting conservatism. These results are consistent with findings in Santa-Clara and Valkanov (2003) who show a negative market reaction to the announcement of a Democratic victory but higher stock returns under a Democratic regime.

Post-election Conservatism

Our primary focus up to this point has been to study whether firms become more conservative with their financial reporting leading up to an election. A natural question to ask is whether conservatism decreases when political uncertainty falls after the election. We construct a variable, *POST-ELECTION*, equal to one if firms' annual earnings announcement date falls in

²⁰ An incumbent party is highly likely to continue its administration. For example, for 29 elections from 1900 to 2012, the continuation rate of a political regime is 61 percent for Democrats and 62 percent for Republicans.

the window (0, +360] and zero otherwise. Panel A of Table 5 shows positive and significant estimates for POST-ELECTION consistently across specifications, implying that conservatism remains high in the year after elections relative to non-election periods. Focusing on the most restrictive model (Model 6) the coefficient estimates for PRE-ELECTION and POST-ELECTION are 0.0285 and 0.0137, respectively. Compared to the unconditional mean value for conservatism, these estimates imply that conservatism increases by about 22 percent in the year leading up to an election, and remains about 11 percent higher than all other years in the year following an election.²¹ These results suggest that while the end of an election may signal no more uncertainty about who forms office, the period immediately after an election may still be associated with higher than average political uncertainty.

The persistence of high conservatism after the election date might simply be driven by a change in political party as a result of the election. One might expect that the policy stance of the incumbent party is more predictable, relative to that of the opposition party, given that they have been in office for at least one term. Therefore, consistent with the *political uncertainty hypothesis*, an election that results in the opposition party forming office is associated with continued policy uncertainty relative to an election where the incumbent party retains power – leading to higher than average conservatism post-election. We investigate this possibility by creating an indicator variable, *OPPWIN*, equal to one if the opposition party wins the election, and interacting it with POST-ELECTION. We expect the coefficient of interest on *POST x OPPWIN* to be positive. In Panel B of Table 5, we find that the estimates on POST x OPPWIN are significantly positive, and after controlling for OPPWIN the coefficients on POST-ELECTION become negative and significant at one percent, which is consistent with the notion

²¹ Please note that these estimates are very close to those obtained from the univariate comparison of C-Score across election and non-election years in the previous section.

that the election year effect on conservatism is reversed if the incumbent party retains power, but persists at a high level when the election is won by the opposition party. In Model 6, the estimates for PRE-ELECTION and POST-ELECTION are 0.0288 and -0.00575 respectively, implying an increase in conservatism by approximately 23 percent in the year leading up to an election, and a reduction in conservatism of approximately five percent in the year following when the incumbent party retains power (relative to the unconditional mean C-Score of 0.127 in Table 1). Figure 1 illustrates the cycles in conservatism using the estimates from Model 6 of Table 5. The dashed line displays conservatism before and after a presidential election regardless of the electoral outcome (Model 6 of Panel A). The solid line represents the cycles in conservatism around elections that are won by the incumbent party (Model 6 of Panel B). A significant decline in conservatism is observed in the post-election year, however, the magnitude of the post-election decline is smaller than that of the initial increase – roughly one-fifth the size of pre-election increase.^{22, 23}

[Figure 1 is inserted here]

Industry Characteristics and Conservatism Cycles

Watts (2003a) argues that conservatism exists for four reasons: 1) to mitigate contracting problems (including debt and compensation contracts); 2) to reduce litigation risks; 3) to avoid regulatory costs; and 4) to minimise taxation obligations. Political uncertainty may affect accounting conservatism differently depending on firms' exposures to politics, government spending, contracting environment, litigation risk, taxation, and regulation. In this and the next sub-sections, we explore how the cross-sectional variation in political uncertainty across industries and firms affects the election year increase in conservatism. We first focus on industry

²² The relative magnitude of the post-election decline in conservatism is calculated as $(0.00575/0.0288)=0.20$.

²³ In Panel C, PRE-MARGIN is added into the specification. All the previous results remain qualitatively similar.

characterises including political sensitivity, regulatory oversight, contract enforcement, labor intensity, and international trade exposure. The results from industry cross-sectional analyses are presented in Table 6.

Political sensitivity and regulatory oversight

We perform tests on two industry factors, i.e., political sensitivity, and regulatory oversight, and expect the election effect to be stronger for more sensitive and regulated industries. Following Julio and Yook (2012), we categorise the following industries as sensitive sectors: tobacco products, pharmaceuticals, health care services, defence, petroleum and natural gas, telecommunications, transportation, and finance. We create a variable, *SENSITIVE*, equal to one if a firm belongs to a sensitive industry. From the regulatory oversight perspective, we construct another indicator variable, *REGULATE*, equal to one if a firm operates in the finance or utilities industries. The results presented in Models 1 to 4 confirm our conjecture. In particular, estimates on $PRE \times SENSITIVE$ and $PRE \times REGULATE$ are both positive and significant, suggesting two possibilities: in these two groups of industries, 1) there is a greater potential for policy change (*political uncertainty hypothesis*); or 2) costs related to political scrutiny are higher (*political cost hypothesis*).

Ramanna and Roychowdhury (2010) document a decline in discretionary accruals for companies connected with U.S. congressional candidates in the election year of 2004, especially for outsourcing-intensive firms, based on a *political cost hypothesis*. While we cannot rule out the possibility that the *political cost hypothesis* may also explain some of our election year effect, given our previous evidence in support of the *political uncertainty hypothesis*,²⁴ the findings here suggest that the two hypotheses may jointly explain the phenomenon of conservatism cycles.

²⁴ It is difficult to explain the results in the tests for electoral competition, incumbency advantage and post-election conservatism with the *political cost hypothesis*.

Contract enforcement, labor intensity, and international trade exposure

Following Boutchkova et al. (2012), we examine three additional industry characteristics, namely, contract enforcement, labor intensity, and international trade exposure and report the results in Models 5 to 10. Since the protection of property rights and contract enforcement depend on the quality of institutions, which are in turn affected by political forces, complex industries with high sensitivity to contract enforcement (also with a high demand for institutional quality) are expected to be more conservative with financial reporting in election years, when political uncertainty leads to uncertainty in future institutional quality. Boutchkova et al. (2012) use US Input-Output tables from 1998 to 2006 to calculate the Blanchard and Kremer (1997) measure of contract enforcement sensitivity for 50 U.S. industries at the two-digit SIC code level.²⁵ This measure is equal to one minus the Herfindahl index of industry input shares, and varies from zero (if the industry uses inputs from only one industry) to one. To study the impact of contracting environment, we construct a variable, *HIGH-CONTRACT*, equal to one if a firm belongs to an industry with an above median sensitivity to contract enforcement according to the Boutchkova et al. (2012) measure. In Models 5 and 6, we find that the coefficients on *PRE* \times *HIGH-CONTRACT* are positive and significant, implying that companies, which rely more on contract enforcement and institutional quality, tend to be more conservative, when there is an increase in political uncertainty and potential changes in institutional quality.

The labor intensity factor studied in Boutchkova et al. (2012) may be a rough indicator for *low-technology industries*, which are less beholden to the changes in government policies, and rely less on government funding for research and development purpose compared to high-technology industries (e.g., information technology and medical industries). Consequently, we might expect the election year effect to be smaller in the low-technology industries. For firms

²⁵ The estimates for sensitivity to contract enforcement are presented in column 7 of Table 2 in their paper.

with high exposure to international trade, revenues are less influenced by domestic politics, therefore a smaller election year effect is also predicted. The labor intensity measure in Boutchkova et al. (2012) is estimated as the value of labor inputs to the total value of production inputs using data from 1990 to 2005 for 32 sectors.²⁶ International trade exposure is the proportion of export flows for each industry to a particular trading partner, and then is summed across all trading partners, available for 29 industries.²⁷ We construct two dummy variables: *HIGH-LABOR* and *HIGH-TRADE* that are equal to one if a firm belongs to an industry with an above median value for labor intensity or international exposure according to Boutchkova et al. (2012). From Models 7 to 10, consistent with the predictions, we find smaller election-year effects in low-technology industries and in industries with higher international trade exposure (negative coefficients on *PRE x HIGH-LABOR* and *PRE x HIGH-TRADE*).

Firm Characteristics and Conservatism Cycles

In this sub-section, we examine the roles of corporate governance, tax avoidance and litigation risk in explaining the association between political uncertainty and accounting conservatism.

Corporate governance

Whether “good” governance has a moderating or exacerbating effect is unclear. As discussed, political uncertainty is positively associated with agency costs and thus demand for conservatism due to three reasons, i.e., greater difficulty detecting managers' shirking behavior (Tirole 2001), more growth options, and increased amount of the cash available to manager (Julio and Yook 2012). We may expect governance to have a negative (positive) influence on the

²⁶ The data are obtained from the input–output database using information from the *Bureau of Economic Analysis* and *Bureau of Labor Statistics*.

²⁷ The estimates for labor intensity and trade exposure are presented in columns 7 and 6 of Table 2 in their paper.

election year effect if it is a substitute (complement) to conservatism in solving agency problems. An exogenous shock to agency costs from the presidential election provides us an opportunity to explore these two possibilities.

We extract the corporate governance information from *RiskMetrics - Directors and Governance Databases*, and employ two governance proxies in this analysis: 1) the external governance index developed by Gompers et al. (2003) (*G-INDEX*) which is the anti-takeover provision index for the market for corporate control, and 2) the percentage of independent directors on the board, *INDEPENDENT*, as a measure of internal governance associated with the monitoring roles of investors and board of directors.²⁸ In Models 1 to 3 of Table 7, we find negative coefficients on *PRE x G-INDEX* and *PRE x INDEPENDENT*, which implies that external (internal) governance might be a complement (substitute) to accounting conservatism in mitigating the increased agency costs caused by rising political uncertainty. These results are also insightful for the literature on the relation between governance and conservatism for two reasons. First, we show that the relation between governance and conservatism may be different for internal versus external measures of governance. Second, previous studies such as Ahmed and Duellman (2007) and García Lara et al. (2009) who find a positive association between governance and conservatism take agency costs as given, while we provide a new piece of evidence on the governance-conservatism relation in a setting where there is an exogenous increase in agency costs due to the election cycle.

Tax and litigation

Finally, we consider the roles of tax avoidance and litigation risk. *TAX* is estimated as the effective tax rate under GAAP, an inverse measure of tax avoidance, equal to total tax expense

²⁸ A lower level of the G-index implies better external governance, while higher percentage of independent directors reflects stronger internal governance. The sample size is reduced by about 90 percent because *RiskMetrics* data are not available for earlier years and for small firms.

divided by pre-tax accounting income less special items. A higher value of TAX implies a lower incidence of tax avoidance (McGuire et al. 2012). We estimate the measure of probability of litigation, *LITIGATION*, an inverse logit of a linear combination of firm fundamentals (Shu 2000).²⁹ The results for TAX and LITIGATION are presented in Models 4 to 9. Focusing on Models 6 and 9 that control for firm fixed effects, we find that the coefficients are negative for the interaction terms, *PRE x TAX* and *PRE x LITIGATION*, but only significant for the latter. This indicates that the election year effect on conservatism is lessened for firms with higher litigation risk, which may represent the fact that managers of these firms are on average more conservative and accordingly do not respond as much to rising political uncertainty in the lead up to elections.

Robustness

There may be alternative explanations for our findings, and we take them into consideration by conducting several sensitivity tests in this section.

Political connections

Prior studies show that the political connectedness of a firm has implications for its behavior and performance due to preferential treatment, and that politicians can also potentially benefit from these connections. For example, using a sample of international firms, Chaney et al. (2011) find that the quality of earnings reported by politically connected firms is significantly poorer than that of similar non-connected companies. Ramanna and Roychowdhury (2010) document that politically connected companies report low discretionary accruals when their outsourcing

²⁹ Specifically, this combination equal to $\{-10.049 + 0.276 (\text{Size}) + 1.153 (\text{Inventory}) + 2.075 (\text{Receivables}) + 1.251 (\text{ROA}) - 0.088 (\text{Current ratio}) + 1.501 (\text{Lev}) + 0.301 (\text{Sales growth}) - 0.371 (\text{stock return}) - 2.309 (\text{stock volatility}) + 0.235 (\text{beta}) + 1.464 (\text{stock turnover}) + 1.060 (\text{Delist dummy}) + 0.928 (\text{Technology dummy}) + 0.463 (\text{Qualified opinion dummy})\}$. Please refer to Table 3 of Shu (2000) for more details.

activities are a major concern during the campaign.³⁰ Taking the Watts (2003a, 2003b) view that conservatism is a value enhancing mechanism to reduce agency costs, one may argue that the our conservatism results are largely driven by politically connected firms trying to help affiliated politicians win office. To address this issue, we include a dummy, *CONNECTED*, equal to one if a firm is politically connected based on the data in Faccio (2006) and re-run the analysis. In Models 1 and 2 of Table 8, we find that our results remain the same, and the interaction term, *PRE x CONNECTED*, has an insignificant coefficient.

The political business cycle

Conceptually, it is possible for the incumbent party to manipulate fiscal and monetary policy instruments to improve macroeconomic outcomes prior to an election in order to maximize the probability of re-election (e.g., McRae 1977; Nordhaus 1975; Rogoff and Sibert 1988).³¹ A rational firm may respond to this political business cycle by changing its accounting choices. If a policy change is favorable to the company, presumably managers may become less conservative, however when managers are aware of the transitory nature of favorable policy, they may even report earnings in a more conservative manner. To investigate this possibility, we introduce two macroeconomic indicators: economic growth, *GROWTH*, and inflation, *INFLATION*, proxied by growth in real gross domestic product (GDP), and growth in the consumer price index (CPI), respectively. In Models 3-5 of Table 8, results are robust to the inclusion of business cycle variables and their interactions with *PRE-ELECTION*.

³⁰ See other studies related to political connections, for example, include Johnson and Mitton (2003), Faccio (2006), and Faccio et al. (2006).

³¹ Empirical evidence in the U.S. to support this view is limited (Drazen 2000), while studies in emerging markets show some evidence that governments take actions to improve chances of re-election (Brown and Dinc 2005; Dinc 2005).

Auditor Quality

One may argue that the election effects might be driven by auditor quality. For example, firms may choose Big 4 auditors coincidentally in election years, and due to higher audit quality financial reports tend to be more conservative. To rule out this possibility, we add a dummy variable, *BIG 4*, into our analysis, which is equal to one if a firm is audited by a Big-4 CPA firm. The results presented in Models 6 to 8 show that our election effects remain robust to the inclusion of *BIG 4* and *PRE x BIG 4*.

Pre- versus post-SOX

After the introduction of the Sarbanes-Oxley Act (SOX), the impact of political process on conservatism can be stronger for two reasons. First, in the post-SOX period there has been increased public attention on accounting issues, firms may therefore react more to political uncertainty because of an increase in demand for conservatism from investors and politicians. Second, in the post-SOX period, the likelihood of a change in accounting standards due to a switch in political regime may be deemed as higher, therefore creating “accounting slack” in the form of conservatism might be desirable. For example, in a campaign event leading up to the 2012 election, when asked by a voter, Mr. Romney pledged to repeal the Sarbanes-Oxley Act.³²

We analyze the impact of SOX in two ways: 1) perform the analysis in two sample periods split by SOX; and 2) introduce an indicator variable, *SOX*, for the post-SOX period into the baseline model. The results presented in Table 9 show that the election year effects are robust in two sub-periods, and remain unchanged after the SOX dummy is included. In Models 7 to 9, we also find that, on average in post-SOX period, conservatism is higher (positive coefficient on *SOX*) and the election year effects are stronger (positive coefficient on *PRE x SOX*), which is consistent with our conjecture.

³² See, for example, a news report by *Wall Street Journal Blogs*, Murray (2012).

Alternative pre-election windows

There may be concern that our results are an artefact of our choice of pre-election window. Accordingly, we rerun the analysis using alternative windows of 90 days and 180 days. The results presented in Table 10 indicate that the findings of election effects are qualitatively the same.³³

IV. CONCLUSION

We exploit exogenous variation in political uncertainty induced by the U.S. presidential election cycle to study its impact on accounting conservatism. An interesting picture emerges: in the year leading up to an election, accounting conservatism increases by nearly 20 percent. Further tests reveal that this electoral phenomenon is indeed caused by political uncertainty and is stronger when the election is closer, and when the incumbent president is not seeking re-election. In the post-election year, conservatism is lower relative to the non-election period when the incumbent party wins, but remains higher under an opposition party victory. This new stylized fact adds to our knowledge regarding how political forces shape the nature of financial reporting, and contributes to a growing literature on the role of politics in determining corporate performance and corporate policy. One avenue for future research is to investigate the impact of elections, as an exogenous political shock to the flow of accounting information, on firms' information environment and the behavior of the users, preparers, auditors, and disseminators of financial reports.

³³ Note that we do not include firm fixed-effects in Table 10, as a pre-election dummy cannot capture all the firms in a short window (less than one year) prior to the election date.

REFERENCES

- Ahmed, A., and S. Duellman. 2007. Accounting conservatism and board of director characteristics: An empirical analysis. *Journal of Accounting and Economics* 43 (2–3): 411–437.
- Alesina, A. 1987. Macroeconomic policy in a two-party system as a repeated game. *Quarterly Journal of Economics* 102 (3): 651–678.
- Ball, R., and L. Shivakumar. 2005. Earnings quality in UK private firms: Comparative loss recognition timeliness. *Journal of Accounting and Economics* 39 (1): 83–128.
- , and ———. 2006. The role of accruals in asymmetrically timely gain and loss recognition. *Journal of Accounting Research* 44 (2): 207–242.
- Basu, S. 1997. The conservatism principle and the asymmetric timeliness of earnings. *Journal of Accounting and Economics* 24 (1): 3–37.
- Belo, F., V. D. Gala, and J. Li. 2012. Government spending, political cycles, and the cross section of stock returns. *Journal of Financial Economics* (forthcoming).
- Blanchard, O., and M. Kremer. 1997. Disorganization. *Quarterly Journal of Economics* 112 (4): 1091–1126.
- Boutchkova, M., H. Doshi, A. Durnev, and A. Molchanov. 2012. Precarious politics and return volatility. *Review of Financial Studies* 25 (4): 1111–1154.
- Brogaard, J., and A. Detzel. 2012. The asset pricing implications of government economic policy uncertainty. Working Paper, University of Washington.
- Brown, C. O., and I. S. Dinc. 2005. The politics of bank failure: Evidence from emerging markets. *Quarterly Journal of Economics* 120 (4): 1413–1444.
- Chaney, P. K., M. Faccio, and D. Parsley. 2011. The quality of accounting information in politically connected firms. *Journal of Accounting and Economics* 51 (1–2): 58–76.
- Dinc, I. S. 2005. Politicians and banks: Political influences on government-owned banks in emerging countries. *Journal of Financial Economics* 77 (2): 453–479.
- Drazen, A. 2000. The political business cycle after 25 years. *National Bureau of Economic Research Macroeconomics Annual* 15: 75–117.
- Faccio, M. 2006. Politically connected firms. *American Economic Review* 96 (1): 369–386.
- , R. Masulis, and J. McConnell. 2006. Political connections and corporate bailouts. *Journal of Finance* 61: 2597–2635.

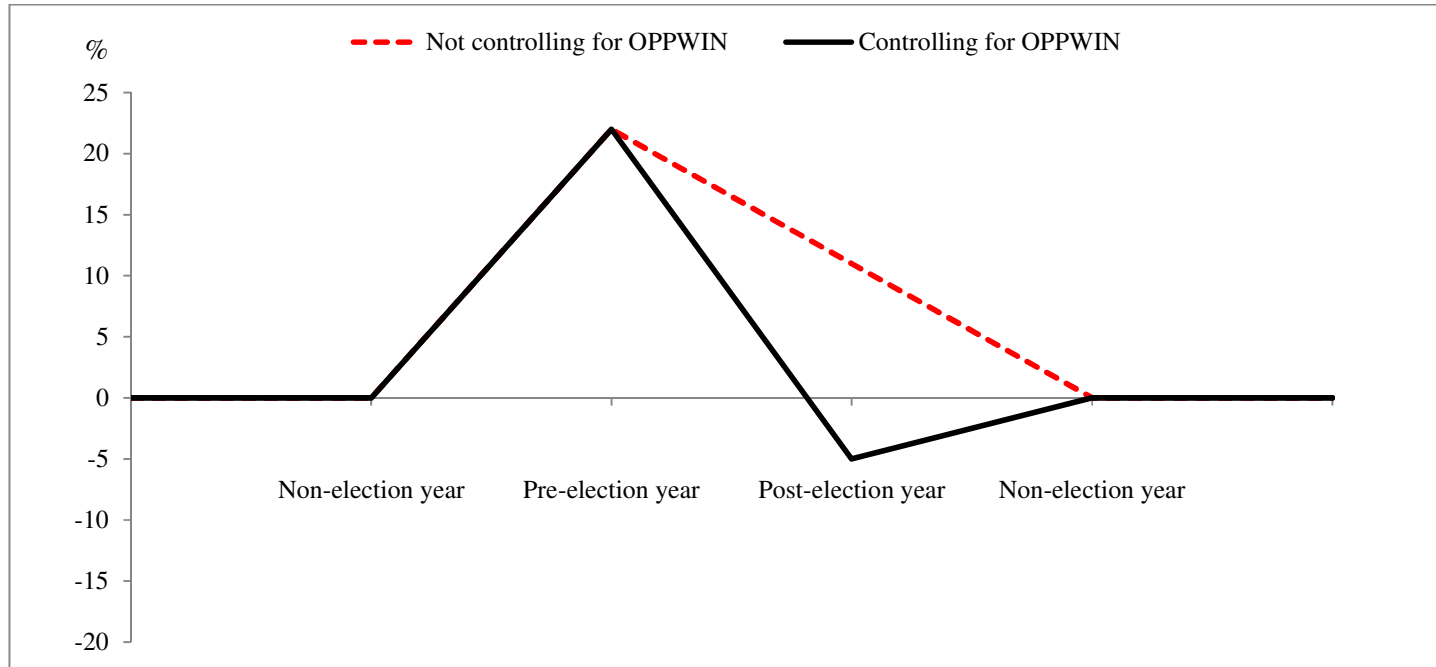
- Friedman, E. 2012. Romney calls for Obamacare repeal as ‘bad law’. *ABC News Blogs*, Jun 28, 2012.
- García Lara, J. M., B. García Osma, and F. Penalva. 2009. Accounting conservatism and corporate governance. *Review of Accounting Studies* 14 (1): 161–201.
- Gelman, A., and G. King. 1990. Estimating the incumbency advantage without bias. *American Journal of Political Science* 34 (4): 1142–1164.
- Goh, B. W., and D. Li. 2011. Internal controls and conditional conservatism. *The Accounting Review* 86 (3): 975–1005.
- Gompers, P., J. Ishii, and A. Metrick. 2003. Corporate governance and equity prices. *Quarterly Journal of Economics* 118 (1): 107–155.
- Guay, W. 2010. Discussion of elections and discretionary accruals: Evidence from 2004. *Journal of Accounting Research* 48 (2): 477–87.
- Ingersoll, J. E., and S. A. Ross. 1992. Waiting to invest: investment and uncertainty. *Journal of Business* 65 (1): 1–29.
- Jensen, M. C. 1986. Agency costs of free cash flow, corporate finance, and takeovers. *American Economic Review* 76 (2): 323–329.
- Johnson, S., and T. Mitton 2003. Cronyism and capital controls: evidence from Malaysia. *Journal of Financial Economics* 67: 351–382.
- Julio, B., and Y. Yook. 2012. Political uncertainty and corporate investment cycles. *Journal of Finance* 67 (1): 45–84.
- Khan, M., and R. Watts 2009. Estimation and empirical properties of a firm-year measure of accounting conservatism. *Journal of Accounting and Economics* 48: 132–150
- Khimm, S. 2012. Romney vows to repeal Dodd-Frank. The law’s biggest critics doubt that will happen. *The Washington Post, Wonkblog*, August 10, 2012.
- Kido, N., R. Petacchi, and J. Weber. 2012. The influence of elections on the accounting choices of governmental entities. *Journal of Accounting Research* 50 (2): 443–476.
- LaFond, R., and S. Roychowdhury. 2008. Managerial ownership and accounting conservatism. *Journal of Accounting Research* 46 (1): 101–135.
- , and R. L. Watts. 2008. The information role of conservatism. *The Accounting Review* 83 (2): 447–478.
- Lent, L. V. 2012. Discussion of the influence of elections on the accounting choices of

- governmental entities. *Journal of Accounting Research* 50 (2): 477–494.
- McDonald, R., and D. Siegel. 1986. The value of waiting to invest. *Quarterly Journal of Economics* 101 (4): 707–728.
- McGuire, S., T. Omer, and D. Wang. 2012. Tax avoidance: Does tax-specific industry expertise make a difference? *The Accounting Review* 87 (3): 975–1003.
- McRae, D. 1977. A political model of the business cycle. *Journal of Political Economy* 85 (2): 239–263.
- Murray, S. 2012. Romney, toughening stance, would repeal Sarbanes-Oxley. *The Wall Street Journal Blogs*, March 3, 2012.
- Nordhaus, W. D. 1975. The political business cycle. *Review of Economic Studies* 42 (2): 169–190.
- Ramanna, K. 2008. The implications of unverifiable fair-value accounting: Evidence from the political economy of goodwill accounting. *Journal of Accounting and Economics* 45 (2–3): 253–81.
- , and S. Roychowdhury. 2010. Elections and discretionary accruals: Evidence from 2004. *Journal of Accounting Research* 48 (2): 445–75.
- Rogoff, K., and A. Sibert. 1988. Elections and macroeconomic policy cycles. *Review of Economic Studies* 55 (1): 1–16.
- Santa-Clara, P., and R. Valkanov. 2003. The presidential puzzle: Political cycles and the stock market. *Journal of Finance* 58 (5): 1841–1872.
- Shu, S. 2000. Auditor resignations: Clientele effects and legal liability. *Journal of Accounting and Economics* 29 (2): 173–205.
- Smith, C., and R. L. Watts. 1992. The investment opportunity set and corporate financing, dividend and compensation policies. *Journal of Financial Economics* 32 (3): 263–292.
- Tirole, J. 2001. Corporate governance. *Econometrica* 69 (1): 1–35.
- Watts, R. L. 2003a. Conservatism in accounting part I: Explanations and implications. *Accounting Horizons* 17 (3): 207–221.
- . 2003b. Conservatism in accounting part II: Evidence and research opportunities. *Accounting Horizons* 17 (4): 287–301.
- , and J. Zimmerman. 1978. Towards a positive theory of the determination of accounting

standards. *The Accounting Review* 53 (1): 112–134.

———, and ———. 1986. *Positive Accounting Theory*. Upper Saddle River, NJ: Prentice Hall.

FIGURE 1
Conservatism around Presidential Elections



This figure displays estimates from the regression results reported in Model 6 of Table 5. OPPWIN is an indicator variable equal one if the election results in a opposition party's victory. The dashed line displays conservatism after presidential election regardless of the electoral outcome (Model 6 of Panel A). The solid line displays conservatism around presidential election when the incumbent party wins the campaign (Model 6 of Panel B). The vertical axis represents percentage change in conservatism relative to the unconditional mean.

TABLE 1
Summary Statistics

	Full Sample			Election Year			Post-Election Year			Non-Election Period		
	Obs	Mean	Std. Dev.	Obs	Mean	Std. Dev.	Obs	Mean	Std. Dev.	Obs	Mean	Std. Dev.
C-SCORE	147,894	0.127	0.154	38,259	0.144	0.206	35,327	0.131	0.144	74,308	0.117	0.124
SIZE	147,894	5.246	1.999	38,259	5.286	2.012	35,327	5.172	1.968	74,308	5.260	2.006
MARKET/BOOK	147,894	2.299	2.394	38,259	2.401	2.728	35,327	2.229	2.243	74,308	2.280	2.274
LEVERAGE	147,894	0.896	1.338	38,259	0.897	1.333	35,327	0.902	1.367	74,308	0.893	1.327
NOACC	147,894	-0.015	0.159	38,259	-0.014	0.122	35,327	-0.017	0.204	74,308	-0.015	0.151
CFOA	147,894	0.062	0.191	38,259	0.061	0.187	35,327	0.061	0.184	74,308	0.063	0.196
INVEST CYCLE	147,894	0.043	0.055	38,259	0.043	0.048	35,327	0.044	0.075	74,308	0.043	0.047
VOLATILITY	147,894	0.032	0.018	38,259	0.032	0.018	35,327	0.033	0.021	74,308	0.031	0.017
BID-ASK	147,894	0.015	0.016	38,259	0.015	0.016	35,327	0.016	0.018	74,308	0.015	0.015
AGE	147,894	13.282	10.160	38,259	13.375	10.205	35,327	13.061	9.976	74,308	13.339	10.223
INDEPENDENT	22,722	0.681	0.177	5,704	0.694	0.174	5,475	0.669	0.181	11,543	0.680	0.175
G-INDEX	21,952	9.054	2.753	5,092	8.998	2.787	5,093	9.084	2.739	11,767	9.065	2.743
TAX	116,411	0.325	0.165	30,314	0.324	0.167	27,969	0.327	0.164	58,128	0.325	0.165
LITIGATION	123,637	0.003	0.005	31,872	0.003	0.004	29,648	0.003	0.006	62,117	0.003	0.005
CONNECTED	147,894	0.001	0.035	38,259	0.001	0.035	35,327	0.001	0.036	74,308	0.001	0.035
BIG 4	131,403	0.850	0.357	33,881	0.848	0.359	30,585	0.858	0.349	66,937	0.848	0.359
SENSITIVE	147,894	0.314	0.464	38,259	0.320	0.466	35,327	0.309	0.462	74,308	0.313	0.464
REGULATE	147,894	0.226	0.418	38,259	0.229	0.420	35,327	0.222	0.416	74,308	0.226	0.418
GROWTH	124,111	0.028	0.020	31,417	0.026	0.016	29,959	0.034	0.021	62,735	0.026	0.021
INFLATION	124,111	0.038	0.026	31,417	0.039	0.025	29,959	0.042	0.026	62,735	0.036	0.026

This table presents summary statistics for the main variables used in the analysis. C-SCORE is Kahn and Watts (2009) measure of accounting conservatism. SIZE is defined as the natural log of the market value of equity. MARKET/BOOK is market-to-book ratio. LEVERAGE is defined as long-term and short-term debt deflated by market value of equity. VOLATILITY is standard deviation of daily firm-level returns. NOACC is non-operating accruals, scaled by lagged assets. CFOA is cash flow from operations, deflated by lagged assets. INVEST CYCLE is a decreasing measure of the length of the investment cycle. BID-ASK is bid-ask spread, scaled by the midpoint of the spread. AGE is firm age, in years. INDEPENDENT is proportion of independent directors. G-INDEX is Gompers et al. (2003) index of external governance. TAX is equal to total tax expense divided by pre-tax accounting income less special items. LITIGATION is Shu (2000) measure of the probability of litigation. CONNECTED is an indicator equal to one if the firm is politically connected. GROWTH is annual growth in real gross domestic product. INFLATION is annual growth in the consumer price index. BIG 4 is an indicator if a firm uses one of the Big-4 auditing firms. SENSITIVE is an indicator variable equal one if firm belongs to a politically sensitive industry. REGULATE is an indicator equal one if a firm is in the finance or utilities industry.

TABLE 2
Political Uncertainty and Conservatism

	(1)	(2)	(3)	(4)	(5)	(6)
PRE-ELECTION	0.0225*** (23.52)	0.0223*** (23.40)	0.0214*** (23.12)	0.0245*** (27.47)	0.0243*** (27.18)	0.0240*** (26.87)
SIZE				-0.0362*** (-97.40)	-0.0367*** (-98.24)	-0.0483*** (-76.66)
MARKET/BOOK				-0.00903*** (-41.00)	-0.00869*** (-40.34)	-0.00853*** (-30.49)
LEVERAGE				0.0415*** (64.58)	0.0381*** (53.95)	0.0419*** (39.26)
NOACC				-0.00946*** (-4.486)	-0.00970*** (-3.229)	0.00117 (0.605)
CFOA				-0.0184*** (-4.850)	-0.0129*** (-4.890)	0.0104*** (3.005)
INVEST CYCLE				-0.0776*** (-3.985)	-0.00710 (-1.129)	0.0269*** (3.046)
VOLATILITY				1.310*** (21.55)	1.406*** (20.57)	0.252*** (6.302)
BID-ASK				-0.787*** (-10.57)	-0.759*** (-10.45)	0.0164 (0.419)
AGE				0.000668*** (14.08)	0.00102*** (20.90)	0.00467*** (50.45)
Constant	0.122*** (135.1)	0.123*** (128.6)	0.122*** (508.1)	0.261*** (86.07)	0.188*** (86.64)	0.285*** (90.05)
Observations	147,894	147,894	147,894	147,894	147,894	147,894
R-squared	0.004	0.063	0.006	0.406	0.420	0.210
Fixed-Effects	None	Industry	Firm	None	Industry	Firm

This table presents our baseline results from the estimation of equation (5). The dependent variable is Kahn and Watts (2009) C-SCORE. Our independent variable of interest is PRE-ELECTION which is an indicator variable equal to one for all days in the year (360 days) prior to an election date. Our list of controls includes: (1) SIZE defined as natural log of the market value of equity; (2) MARKET/BOOK is market-to-book ratio; (3) LEVERAGE is sum of long-term and short-term debt deflated by market value of equity; (4) VOLATILITY is standard deviation of daily firm-level returns; (5) NOACC is non-operating accruals, scaled by lagged assets; (6) CFOA is cash flow from operations, deflated by lagged assets; (7) INVEST CYCLE is a decreasing measure of the length of the investment cycle; (8) BID-ASK is bid-ask spread, scaled by the midpoint of the spread; and (9) AGE is firm age, in years. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.

TABLE 3
Electoral Competition and Conservatism

	(1)	(2)	(3)	(4)	(5)	(6)
PRE-ELECTION	0.0621*** (31.60)	0.0587*** (30.29)	0.0485*** (26.55)	0.0782*** (44.84)	0.0742*** (42.59)	0.0460*** (27.13)
PRE x MARGIN	-0.107*** (-31.60)	-0.0989*** (-30.16)	-0.0734*** (-25.48)	-0.145*** (-52.14)	-0.135*** (-48.93)	-0.0594*** (-24.02)
SIZE				-0.0376*** (-109.1)	-0.0379*** (-108.7)	-0.0477*** (-76.47)
MARKET/BOOK				-0.00924*** (-44.25)	-0.00893*** (-43.30)	-0.00874*** (-31.67)
LEVERAGE				0.0413*** (65.39)	0.0383*** (54.99)	0.0419*** (39.28)
NOACC				-0.00875*** (-4.339)	-0.00906*** (-3.196)	0.00131 (0.691)
CFOA				-0.0178*** (-4.969)	-0.0128*** (-4.988)	0.00975*** (2.928)
INVEST CYCLE				-0.0681*** (-3.895)	-0.00705 (-1.158)	0.0251*** (2.966)
VOLATILITY				1.087*** (20.31)	1.186*** (19.67)	0.216*** (5.561)
BID-ASK				-0.699*** (-10.45)	-0.679*** (-10.35)	0.0177 (0.456)
AGE				0.000630*** (14.25)	0.000947*** (20.62)	0.00438*** (48.01)
Constant	0.122*** (135.1)	0.150*** (17,616)	0.122*** (510.9)	0.274*** (98.62)	0.326*** (124.6)	0.287*** (91.58)
Observations	147,894	147,894	147,894	147,894	147,894	147,894
R-squared	0.015	0.073	0.014	0.426	0.436	0.215
Fixed-Effects	None	Industry	Firm	None	Industry	Firm

This table presents results from the estimation of equation (5). The dependent variable is Kahn and Watts (2009) C–SCORE. Our independent variables of interest are: (1) PRE-ELECTION which is an indicator variable equal to one for all days in the year (360 days) prior to an election date; and (2) PRE x MARGIN which is the interaction term between PRE-ELECTION and the normalized margin for victory. Our list of controls includes: (1) SIZE is natural log of the market value of equity; (2) MARKET/BOOK is market-to-book ratio; (3) LEVERAGE sum of long-term and short-term debt deflated by market value of equity; (4) VOLATILITY is standard deviation of daily firm-level returns; (5) NOACC is non-operating accruals, scaled by lagged assets; (6) CFOA is cash flow from operations, deflated by lagged assets; (7) INVEST CYCLE is a decreasing measure of the length of the investment cycle; (8) BID-ASK is bid–ask spread, scaled by the midpoint of the spread; and (9) AGE is firm age, in years. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.

TABLE 4
Incumbency Advantage and Party Affiliation

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
PRE-ELECTION	0.0839*** (35.68)	0.0798*** (34.10)	0.0427*** (18.73)	0.102*** (40.59)	0.0985*** (39.39)	0.0658*** (27.04)	0.102*** (36.80)	0.0979*** (35.66)	0.0598*** (22.29)
PRE x MARGIN	-0.150*** (-47.69)	-0.140*** (-45.06)	-0.0591*** (-21.40)	-0.139*** (-53.28)	-0.129*** (-49.61)	-0.0553*** (-23.29)	-0.141*** (-48.17)	-0.131*** (-45.07)	-0.0516*** (-19.50)
DEM	-0.00576*** (-8.846)	-0.00775*** (-11.82)	-0.0196*** (-29.48)				-0.00811*** (-12.25)	-0.0101*** (-15.14)	-0.0219*** (-31.84)
PRE x DEM	-0.0123*** (-7.671)	-0.0117*** (-7.321)	0.00925*** (5.837)				0.00370*** (2.696)	0.00445*** (3.251)	0.0234*** (16.30)
PRE x INCUM				-0.0436*** (-22.85)	-0.0439*** (-23.05)	-0.0355*** (-18.69)	-0.0441*** (-24.27)	-0.0444*** (-24.47)	-0.0408*** (-22.12)
Constant	0.275*** (97.88)	0.331*** (124.4)	0.294*** (93.89)	0.274*** (98.97)	0.325*** (124.5)	0.285*** (91.06)	0.276*** (98.29)	0.333*** (125.2)	0.293*** (94.12)
Observations	147,894	147,894	147,894	147,894	147,894	147,894	147,894	147,894	147,894
R-squared	0.427	0.438	0.220	0.431	0.441	0.220	0.431	0.442	0.226
Fixed-Effects	None	Industry	Firm	None	Industry	Firm	None	Industry	Firm
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

This table presents our results from the estimation of equation (5). The dependent variable is Kahn and Watts (2009) C-SCORE. Our key independent variables of interest are: (1) PRE-ELECTION which is an indicator variable equal to one for all days in the year (360 days) prior to an election date; and (2) PRE x MARGIN which is the interaction term between PRE-ELECTION and the normalize margin for victory. DEM is an indicator variable equal to one if the incumbent president is a Democrat. INCUM is an indicator variable equal one if the incumbent president is seeking re-election. Control variables are from Table 2 but not reported. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.

TABLE 5
Post-Election Conservatism

PANEL A	(1)	(2)	(3)	(4)	(5)	(6)
PRE-ELECTION	0.0271*** (27.28)	0.0269*** (27.24)	0.0261*** (27.42)	0.0273*** (29.54)	0.0270*** (29.26)	0.0285*** (30.90)
POST-ELECTION	0.0141*** (20.20)	0.0141*** (20.35)	0.0143*** (21.92)	0.00845*** (14.61)	0.00846*** (14.56)	0.0137*** (24.50)
Constant	0.117*** (127.0)	0.118*** (119.7)	0.117*** (367.4)	0.258*** (84.94)	0.185*** (85.17)	0.282*** (89.45)
Observations	147,894	147,894	147,894	147,894	147,894	147,894
R-squared	0.005	0.065	0.009	0.407	0.420	0.212
Fixed-Effects	None	Industry	Firm	None	Industry	Firm
Controls	No	No	No	Yes	Yes	Yes

PANEL B	(1)	(2)	(3)	(4)	(5)	(6)
PRE-ELECTION	0.0271*** (27.27)	0.0269*** (27.25)	0.0261*** (27.42)	0.0274*** (29.72)	0.0272*** (29.46)	0.0288*** (31.17)
POST-ELECTION	-0.0152*** (-22.48)	-0.0147*** (-22.07)	-0.0148*** (-25.59)	-0.0137*** (-26.49)	-0.0126*** (-24.11)	-0.00575*** (-12.27)
POST x OPPWIN	0.0603*** (45.70)	0.0594*** (46.13)	0.0596*** (51.83)	0.0463*** (45.51)	0.0439*** (41.55)	0.0407*** (45.09)
Constant	0.117*** (127.0)	0.118*** (119.7)	0.117*** (371.3)	0.265*** (88.79)	0.189*** (87.12)	0.287*** (93.88)
Observations	147,894	147,894	147,894	147,894	147,894	147,894
R-squared	0.015	0.074	0.023	0.412	0.425	0.218
Fixed-Effects	None	Industry	Firm	None	Industry	Firm
Controls	No	No	No	Yes	Yes	Yes

PANEL C	(1)	(2)	(3)	(4)	(5)	(6)
PRE-ELECTION	0.0666*** (33.62)	0.0633*** (32.44)	0.0537*** (29.32)	0.0819*** (46.83)	0.0780*** (44.63)	0.0516*** (30.31)
PRE x MARGIN	-0.107*** (-31.60)	-0.0990*** (-30.18)	-0.0751*** (-26.01)	-0.147*** (-53.12)	-0.137*** (-50.02)	-0.0620*** (-25.04)
POST-ELECTION	-0.0152*** (-22.48)	-0.0147*** (-22.15)	-0.0154*** (-26.61)	-0.0146*** (-28.48)	-0.0135*** (-26.35)	-0.00674*** (-14.34)
POST x OPPWIN	0.0603*** (45.70)	0.0595*** (46.17)	0.0605*** (52.55)	0.0485*** (49.46)	0.0463*** (45.76)	0.0424*** (47.35)
Constant	0.117*** (127.0)	0.123*** (6.80)	0.117*** (372.9)	0.279*** (102.1)	0.332*** (129.7)	0.290*** (95.23)
Observations	147,894	147,894	147,894	147,894	147,894	147,894
R-squared	0.026	0.083	0.032	0.432	0.442	0.224
Fixed-Effects	None	Industry	Firm	None	Industry	Firm
Controls	No	No	No	Yes	Yes	Yes

This table presents our results from the estimation of equation (5). The dependent variable is Kahn and Watts (2009) C–SCORE. Our key independent variables of interest are: (1) PRE-ELECTION which is an indicator variable equal to one for all days in the year (360 days) prior to an election date; and (2) PRE x MARGIN which is the interaction term between PRE-ELECTION and the normalize margin for victory. POST-ELECTION is an indicator variable equal one for all days in the year (360 days) following an election. OPPWIN is an indicator variable equal one if the election results in the opposition party winning office. Control variables are from Table 2 but not reported. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.

TABLE 6
Industry-Level Characteristics and Conservatism Cycles

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
PRE-ELECTION	0.0551*** (27.03)	0.0744*** (41.71)	0.0522*** (25.99)	0.0705*** (39.76)	0.0575*** (24.47)	0.0749*** (36.14)	0.0692*** (26.45)	0.0810*** (34.60)	0.0491*** (16.36)	0.0636*** (24.36)
PRE x MARGIN	-0.102*** (-30.08)	-0.141*** (-50.85)	-0.101*** (-29.83)	-0.139*** (-49.51)	-0.108*** (-29.86)	-0.148*** (-49.99)	-0.112*** (-27.24)	-0.140*** (-42.54)	-0.0794*** (-17.92)	-0.111*** (-30.98)
SENSITIVE	0.0196*** (9.598)	0.0179*** (17.57)								
PRE x SENSITIVE	0.0170*** (7.396)	0.00757*** (3.717)								
REGULATE			0.0322*** (13.50)	0.0206*** (15.96)						
PRE x REGULATE			0.0359*** (14.39)	0.0255*** (11.04)						
HIGH-CONTRACT					0.0262*** (14.20)	0.00643*** (6.973)				
PRE x HIGH-CONTRACT					0.00867*** (4.273)	0.00803*** (4.257)				
HIGH-LABOR							-0.000379 (-0.175)	-0.0127*** (-11.88)		
PRE x HIGH-LABOR							-0.00899*** (-3.937)	-0.00766*** (-3.635)		
HIGH-TRADE									-0.000992 (-0.432)	0.00121 (1.175)
PRE x HIGH-TRADE									-0.00797*** (-3.239)	-0.00639*** (-2.844)
Constant	0.115*** (111.2)	0.269*** (95.72)	0.114*** (116.5)	0.266*** (91.12)	0.109*** (84.29)	0.275*** (99.49)	0.121*** (73.23)	0.285*** (105.1)	0.108*** (68.04)	0.256*** (89.90)
Observations	147,894	147,894	147,894	147,894	130,910	130,910	99,774	99,774	76,879	76,879
R-squared	0.021	0.429	0.029	0.431	0.025	0.429	0.018	0.420	0.010	0.410
Controls	No	Yes	No	Yes	No	Yes	No	Yes	No	Yes

This table presents our results from the estimation of equation (5). The dependent variable is Kahn and Watts (2009) C-SCORE. Key variables of interest are: (1) PRE-ELECTION, an indicator variable equal to one for all days in the year prior to an election date; and (2) PRE x MARGIN, an interaction term between PRE-ELECTION and normalize margin for victory; SENSITIVE equal one if firm belongs to a politically sensitive industry. REGULATE equal one if a firm is in the finance or utilities industry. HIGH-CONTRACT equal one if a firm is in an industry that has an above median value for the Boutchkova et al (2012) measure of sensitivity to contract enforcement. HIGH-LABOR is an indicator equal one if a firm is in an industry that has an above median value for the Boutchkova et al (2012) measure of labor intensity. HIGH-TRADE is an indicator equal one if a firm is in an industry that has an above median value for the Boutchkova et al (2012) measure of international trade exposure. Control variables are from Table 2 but not reported. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.

TABLE 7
Firm-Level Governance, Taxation, and Litigation and Conservatism Cycles

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
PRE-ELECTION	0.0594*** (4.407)	0.0592*** (4.379)	0.0447*** (3.157)	0.0679*** (31.31)	0.0637*** (29.55)	0.0323*** (15.13)	0.0758*** (37.02)	0.0749*** (36.55)	0.0464*** (23.65)
PRE x MARGIN	-0.0607*** (-4.665)	-0.0549*** (-4.237)	-0.0273 (-1.511)	-0.114*** (-36.54)	-0.104*** (-34.26)	-0.0391*** (-13.41)	-0.125*** (-41.37)	-0.122*** (-40.78)	-0.0507*** (-19.34)
INDEPENDENT	0.0460*** (11.48)	0.0478*** (11.73)	0.00947 (1.190)						
G-INDEX	0.000761*** (2.874)	0.000735*** (2.770)	-0.000115 (-0.128)						
PRE x INDEPENDENT	-0.118*** (-6.357)	-0.117*** (-6.319)	-0.107*** (-5.422)						
PRE x G-INDEX	-0.00441*** (-3.797)	-0.00445*** (-3.819)	-0.00427*** (-3.533)						
TAX				-0.0539*** (-23.07)	-0.0437*** (-17.79)	0.00130 (0.480)			
PRE x TAX				-0.0190*** (-3.547)	-0.0186*** (-3.496)	-0.000854 (-0.157)			
LITIGATION							-0.242** (-2.575)	-0.386*** (-3.835)	0.0110 (0.115)
PRE x LITIGATION							-3.397*** (-13.24)	-3.413*** (-12.87)	-2.857*** (-12.13)
Constant	0.400*** (75.65)	0.330*** (47.57)	0.252*** (17.18)	0.298*** (108.0)	0.265*** (98.68)	0.278*** (72.89)	0.253*** (70.99)	0.249 (0.00118)	0.286*** (84.92)
Observations	14,372	14,372	14,372	116,411	116,411	116,411	123,637	123,637	123,637
R-squared	0.533	0.535	0.238	0.420	0.432	0.166	0.416	0.420	0.217
Fixed-Effects	None	Industry	Firm	None	Industry	Firm	None	Industry	Firm
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

This table presents our results from the estimation of equation (5). The dependent variable is the Kahn and Watts (2009) C–SCORE. Our key independent variables of interest are: (1) PRE-ELECTION which is an indicator variable equal to one for all days in the year (360 days) prior to an election date; and (2) PRE x MARGIN which is the interaction term between PRE-ELECTION and the normalized margin for victory. INDEPENDENT is the proportion of independent directors. G-INDEX is the Gompers et al. (2003) index of external governance. TAX is equal to total tax expense divided by pre-tax accounting income less special items. LITIGATION is the Shu (2000) measure of the probability of litigation. Control variables are from Table 2 but not reported. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.

TABLE 8
Robustness: Political Connections, Business Cycle and Big 4 Auditor

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
PRE - ELECTION	0.0782*** (44.83)	0.0742*** (42.59)	0.0501*** (21.52)	0.0495*** (21.21)	0.0382*** (16.49)	0.0730*** (32.93)	0.0696*** (31.01)	0.0351*** (15.56)
PRE x MARGIN	-0.145*** (-52.12)	-0.135*** (-48.91)	-0.117*** (-35.52)	-0.112*** (-33.90)	-0.0560*** (-16.63)	-0.111*** (-36.53)	-0.107*** (-35.04)	-0.0413*** (-14.74)
CONNECTED	-0.0227 (-1.561)	-0.0198 (-1.274)						
PRE x CONNECTED	-0.00698 (-0.253)	-0.00801 (-0.288)						
GROWTH			-0.678*** (-42.87)	-0.612*** (-38.53)	0.218*** (16.49)			
INFLATION			-0.679*** (-41.92)	-0.580*** (-34.76)	0.842*** (52.41)			
PRE x GROWTH			0.00136 (0.0298)	-0.0437 (-0.961)	-0.363*** (-8.241)			
PRE x INFLATION			0.413*** (8.715)	0.403*** (8.482)	0.292*** (6.147)			
BIG 4						-0.0304*** (-23.22)	-0.0236*** (-18.82)	-0.00258 (-1.464)
PRE x BIG 4						-0.0197*** (-10.52)	-0.0173*** (-9.194)	-0.00388** (-2.003)
Constant	0.274*** (98.54)	0.326*** (124.5)	0.357*** (119.6)	0.371*** (108.2)	0.218*** (56.24)	0.294*** (94.03)	0.297 (0.00243)	0.282*** (78.72)
Observations	147,894	147,894	124,111	124,111	124,111	131,403	131,403	131,403
R-squared	0.426	0.436	0.447	0.455	0.209	0.452	0.460	0.221
Fixed-Effects	None	Industry	None	Industry	Firm	None	Industry	Firm
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

This table presents our results from the estimation of equation (5). The dependent variable is the Kahn and Watts (2009) C –SCORE. Our key independent variables of interest are: (1) PRE-ELECTION which is an indicator variable equal to one for all days in the year (360 days) prior to an election date; and (2) PRE x MARGIN which is the interaction term between PRE-ELECTION and the normalize margin for victory. CONNECTED is an indicator equal to one if the firm is politically connected. GROWTH is the annual growth in real gross domestic product. INFLATION is the annual growth in the consumer price index. BIG 4 is an indicator if a firm uses one of the Big-4 auditing firms. Control variables are from Table 2 but not reported. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.

TABLE 9
Robustness: Sarbanes-Oxley Act

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	Pre-SOX	Pre-SOX	Pre-SOX	Post-SOX	Post-SOX	Post-SOX	Full Sample	Full Sample	Full Sample
PRE-ELECTION	0.0518*** (34.66)	0.0493*** (32.18)	0.0325*** (19.17)	0.101*** (20.36)	0.101*** (20.31)	0.0577*** (10.64)	0.0564*** (36.64)	0.0536*** (34.15)	0.0378*** (22.85)
PRE x MARGIN	-0.0870*** (-45.08)	-0.0815*** (-40.83)	-0.0333*** (-14.62)	-0.278*** (-17.40)	-0.277*** (-17.32)	-0.105*** (-5.871)	-0.110*** (-49.35)	-0.103*** (-45.33)	-0.0520*** (-22.12)
SOX							0.0711*** (71.75)	0.0676*** (67.82)	0.00517*** (3.454)
PRE x SOX							0.0101*** (5.141)	0.0113*** (5.710)	0.0161*** (8.013)
Constant	0.215*** (130.6)	0.213*** (40.68)	0.245*** (87.53)	0.536*** (90.83)	0.538 (0.00626)	0.604*** (35.02)	0.277*** (124.7)	0.272*** (112.1)	0.286*** (90.00)
Observations	104,463	104,463	104,463	38,732	38,732	38,732	147,894	147,894	147,894
R-squared	0.436	0.443	0.220	0.647	0.648	0.264	0.462	0.469	0.216
Fixed-Effects	None	Industry	Firm	None	Industry	Firm	None	Industry	Firm
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

This table presents our results from the estimation of equation (5). The dependent variable is the Kahn and Watts (2009) C–SCORE. Our key independent variables of interest are: (1) PRE-ELECTION which is an indicator variable equal to one for all days in the year (360 days) prior to an election date; and (2) PRE x MARGIN which is the interaction term between PRE-ELECTION and the normalize margin for victory. SOX is an indicator equal to one for all years after the introduction of the Sarbanes–Oxley Act. Control variables are from Table 2 but not reported. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.

TABLE 10
Robustness: Alternative Windows

	(1)	(2)	(3)	(4)
90-PRE-ELECTION	0.0130*** (5.300)	0.0108*** (4.406)		
90-PRE x MARGIN	-0.0639*** (-14.90)	-0.0530*** (-12.21)		
180-PRE-ELECTION			0.0352*** (11.69)	0.0337*** (11.18)
180-PRE x MARGIN			-0.0980*** (-21.19)	-0.0880*** (-18.83)
Constant	0.267*** (87.35)	0.212*** (105.4)	0.269*** (88.37)	0.213*** (106.3)
Observations	147,894	147,894	147,894	147,894
Fixed-Effects	None	Industry	None	Industry
Controls	Yes	Yes	Yes	Yes

This table presents our results from the estimation of equation (5). The dependent variable is the Kahn and Watts (2009) C–SCORE. Our key independent variables of interest are: (1) 90-PRE-ELECTION which is an indicator variable equal to one all days up to 90 days prior to an election date; (2) 90-PRE x MARGIN which is the interaction term between 90-PRE-ELECTION and the normalize margin for victory; (3) 180-PRE-ELECTION which is an indicator variable equal to one all days up to 180 days prior to an election date; and (4) 180-PRE x MARGIN which is the interaction term between 180-PRE-ELECTION and the normalize margin for victory. Control variables are from Table 2 but not reported. Robust t-statistics are (clustered standard errors by firm) in parentheses. Significance levels of 10, 5, and 1 percent are represented by ***, **, and *.