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Ageing, government budgets, retirement, and growth

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Abstract

We analyze the short and long run effects of demographic ageing—increased longevity and reduced fertility—on per-capita growth. The OLG model captures direct effects, working through adjustments in the savings rate, labor supply, and capital deepening, and indirect effects, working through changes of taxes, government spending components and the retirement age in politico-economic equilibrium. Growth is driven by capital accumulation and productivity increases fueled by public investment. The closed-form solutions of the model predict taxation and the retirement age in OECD economies to increase in response to demographic ageing and per-capita growth to accelerate. If the retirement age were held constant, the growth rate in politico-economic equilibrium would essentially remain unchanged, due to a surge of social security transfers and crowding out of public investment.

JEL Codes: E62, H5, J26

Keywords: Ageing; Government Budgets; Retirement; Growth

1 Introduction

The prospect of “graying” populations in many developed economies raises concerns about the sustainability of economic growth. According to these concerns, rising old-age dependency ratios translate into growing tax burdens while generous pension and health care benefits crowd out public investment spending for infrastructure or education, with negative effects for capital accumulation and productivity growth. However, the demographic transition has been ongoing for a while—developed economies have experienced a marked decrease in fertility and increase in longevity for several decades—without producing clear evidence that this transition has caused a fall in per-capita growth. Rather to the opposite, Barro and Sala-i-Martin (1995) find in growth regressions that both a decrease in the fertility rate and an increase in longevity are associated with higher growth rates.¹

¹Falling fertility and increasing longevity can give rise to a temporary reduction in the (young-age) dependency ratio, generating a “demographic dividend” of higher growth. In developed economies, this growth dividend is predicted to be exhausted around the year 2010 (e.g. Bloom, Canning and Sevilla, 2003). Also, Acemoglu and Johnson (2007) estimate the effect of life expectancy at birth on economic growth. They find no evidence of a positive effect.

The evidence is similarly mixed as far as government budgets are concerned. While the GDP share of transfers to the elderly has increased, the share of public investment does not show a clear trend in most countries, see Figures 1 and 2.^{2 3} Moreover, most developed countries have started to increase the retirement age or tighten the conditions for early retirement, reducing the pressure on social security taxes.

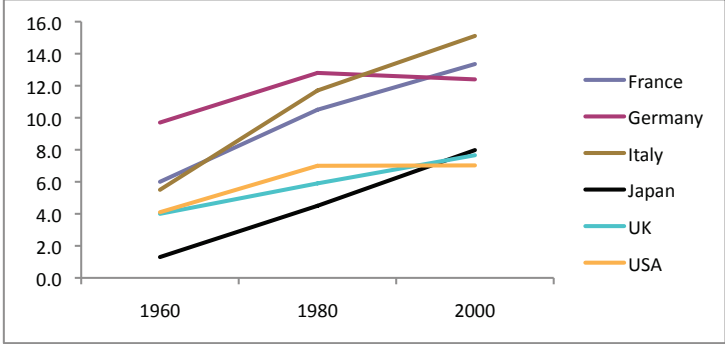


Figure 1: Public expenditure on pensions, percent of GDP. *Source:* Tanzi and Schuknecht (2000), OECD.

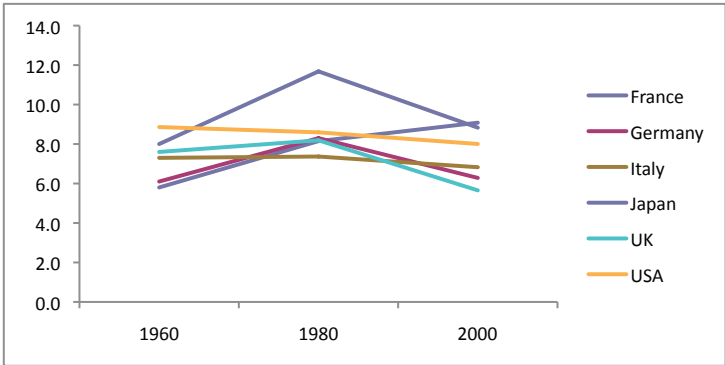


Figure 2: Public expenditure on education and infrastructure investment, percent of GDP. *Source:* Tanzi and Schuknecht (2000), OECD.

To interpret this data and gauge likely future developments, we develop a tractable model to analyze the effects of demographic ageing on government budgets and per-capita growth. Building on a standard overlapping generations setup with private and public capital formation sustaining endogenous growth, our framework features two demographic

²Data is taken from Tanzi and Schuknecht (2000) and OECD sources (*Society at a Glance*, 2009, old-age cash benefits, disability pensions and survivors’ pensions; *Economic Outlook*, 2008, government fixed capital formation; *Education At a Glance*, 2008, direct public expenditure plus subsidies to households and other private entities). The GDP share of infrastructure investment has fallen in some countries, in contrast to the GDP share of education spending. In this paper, we do not analyze the composition of public investment.

³In cross section data for the United States, the fraction of elderly residents in a district was negatively associated with education spending per child (Poterba, 1997).

driving forces—fertility and longevity—and a number of economic and political choices. In their role as economic agents, households in the model take prices, taxes, public investment, the retirement age and retirement benefits as given when choosing consumption, savings, and labor supply. In their role as voters, households choose among office motivated parties that offer policy platforms comprising labor income taxes, the expenditure shares for inter-generational transfers and public investment (reflecting spending components of central importance for developed economies), as well as the retirement age. The political process lacks commitment, and elections take place every period.

Policy choices in the model are of different concern to young and old voters: the exposure of households to labor income taxes changes over the life cycle; the old benefit from social security transfers to their group but are hurt by an increase in the retirement age; and only the young benefit from the returns to public investment. When evaluating the policy platforms on offer in the political arena, voters therefore disagree as to which platform should ideally be implemented. We model the resolution of the ensuing conflict under the assumption of probabilistic voting, reflecting a small degree of randomness in voters’ support for a party. In equilibrium, vote-seeking parties propose a policy platform maximizing average welfare of all voters, and changes in the economic or demographic environment give rise to a gradual adjustment of the policy instruments.

Policy choices do not only affect economic outcomes. Absent commitment, they also affect, indirectly, future policy decisions. In addition to the “economic” repercussions of their policy choices, voters internalize the “political” repercussions, reflected in the equilibrium relationship between future state variables and policy choices. We assume that only fundamental state variables enter this equilibrium relationship, excluding artificial state variables of the type sustaining trigger strategy equilibria. While we agree that the existence of intergenerational transfers or public investment may also owe to reputational arrangements, we focus on the Markov perfect equilibrium in order to identify the fundamental and robust forces that shape the size of these programs, and therefore growth.⁴

Under standard functional form assumptions, we characterize the politico-economic equilibrium in closed form.⁵ Changes in the demographic structure affect the equilibrium allocation both directly and indirectly, by inducing policy changes. The direct effect of changes in fertility and longevity works through modified private savings and labor-supply decisions which in equilibrium manifest themselves in faster capital accumulation. Indirectly, demographic ageing affects growth because it alters the relative political power of the old and the effect of later retirement on aggregate labor supply. At the same time, higher longevity increases the political support for public investment.

To quantify the equilibrium implications of demographic ageing, we analyze calibrated versions of the model representing a rich OECD economy, a rich European OECD economy, the United States, and Japan. For each of the countries and country blocks, the model predicts that the forecasted demographic changes give rise to a continued increase of the GDP share of social security transfers, a slightly higher GDP share of public in-

⁴For a discussion of Markov perfect equilibrium see Krusell, Quadrini and Ríos-Rull (1997).

⁵Our functional form assumptions imply a minimal amount of strategic interaction between policy makers in different periods.

vestment, a strong increase of the retirement age, and a rise in per-capita growth. In particular, annual per-capita growth is predicted to accelerate by approximately 35 basis points towards the end of the century.

Importantly, these findings hinge on the assumption that both fiscal policy and the retirement age are endogenous. With constant policy instruments, the growth rate would increase more strongly than in politico-economic equilibrium. With endogenous tax rates and budget shares but a fixed retirement age, the per-capita growth rate would essentially remain stuck at its current level in the medium run and increase only slightly in the long run, due to a surge of social security transfers and crowding out of public investment.

The central predictions of the model are robust to a variety of changes in the modeling assumptions. In particular, the results do not change if capital income taxes in addition to labor income taxes are introduced or if the balanced-budget assumption is relaxed. The results are also robust to replacing the endogenous growth specification by one of exogenous growth. In the model, the political process does not internalize the long-term benefits of public investment because these occur beyond the lifetimes of even the youngest voters. As a consequence, the exact specification of productivity growth has no effect on the evolution of the policy instruments and the government budget shares in politico-economic equilibrium. Moreover, as we show, it does not have a strong effect on the evolution of output per capita in the medium run either.

While broadly consistent with the evidence, the model predictions contradict the common view among policy makers that the political process will implement measures to raise productivity in order to “outgrow” the burden imposed by demographic change.⁶ According to the model, demographic ageing indeed induces the political process to raise public investment in order to foster productivity growth. However, the main positive growth effects arise directly while the net effect of endogenous policy on growth is negative.

Our work relates to the literature analyzing the effects of government policy on growth, see Barro (1990), Jones, Manuelli and Rossi (1993) or Glomm and Ravikumar (1997) for a review. Our contribution relative to these papers lies in modeling the determinants of policy and linking fertility and longevity to growth.⁷ Galor and Weil (1999), Cervellati and Sunde (2005) and Soares (2005) analyze the growth effects of demographic change due to its impact on private savings and education decisions, and Hazan (2009) introduces private retirement decisions in a model of human capital accumulation. Our model complements these papers by modeling the role of policy and its determinants in politico-economic equilibrium, and by focusing on the ongoing demographic transition in developed economies rather than historical developments.

Our work also relates to politico-economic models of redistribution and growth. Alesina and Rodrik (1994), Persson and Tabellini (1994) and Krusell et al. (1997) argue that inequality depresses growth because anticipated redistributive taxation reduces the incentive to accumulate, or because higher inequality pushes the median voter’s preferred level of public investment and taxes beyond the growth-maximizing level. Relative to these papers, we focus on inter- rather than intragenerational conflict, consider a larger set of

⁶See, for example, the discussions surrounding the European Union’s “Lisbon Agenda.”

⁷Azzimonti, Sarte and Soares (2009) analyze capital formation by a benevolent government without commitment in a representative agent economy.

policy instruments available to policy makers, and focus on the implications of fertility and longevity on growth. Our analysis therefore sheds light on the equilibrium size and composition of the government budget, and it emphasizes how demographic ageing affects both this composition and growth.^{8 9}

Like Bellettini and Berti Ceroni (1999) and Rangel (2003), our paper analyzes the choice of productive versus redistributive public spending in an overlapping-generations model. In these papers, voters support public investment even if they do not directly benefit from it because a trigger strategy links investment spending to the provision of public pensions in the future. Our model adopts a different perspective. Rather than emphasizing complementarities between investment spending and transfer payments, it focuses on the conflict over the size of these two spending components, and how the resolution of this conflict is shaped by fertility and longevity. The model also differs from these papers in that it features political and economic choices, embedded in the standard growth model. This allows us to model the macroeconomic consequences of population ageing in a rich setting without having to sacrifice analytical tractability. Gradstein and Kaganovich (2004) argue that public investment might rise in response to increased longevity. Our model incorporates the mechanism underlying Gradstein and Kaganovich’s (2004) argument. In addition, it features a role for fertility, the retirement age, a second government spending component that competes for funding, and—central to our analysis—growth effects of policy.

The remainder of the paper is structured as follows. Section 2 describes the model and characterizes the allocation conditional on policy. Section 3 solves for the politico-economic equilibrium and analyzes its properties. Section 4 contains the analysis of the short and long-run effects of demographic ageing on government budgets and macroeconomic outcomes, in particular the growth rate. Section 5 concludes.

2 Economic Environment

We consider an economy inhabited by two-period lived overlapping generations: young households and old households. Young households in period t supply labor, pay taxes, consume and save for retirement. They face idiosyncratic longevity risk: with probability $p_{t+1} \in (0, 1]$, they survive to become old households in period $t + 1$. Old households in turn consume the return on their savings, old-age benefits and the proceeds of their labor income in old age. The size of this labor income depends on the retirement age, ϱ_t , and the labor productivity of old relative to young workers, $\chi \geq 0$.¹⁰ If $\varrho_t = 0$ or $\chi = 0$, the effective per-capita labor supply of old households equals zero. Old households die at the end of the period.

⁸Our work shares with Krusell et al. (1997) the restriction to Markov perfect equilibrium. Methodologically, it is related to Gonzalez-Eiras and Niepelt (2008).

⁹Glomm and Ravikumar (1992) and Perotti (1993) analyze distributive conflict in models with human capital accumulation. They focus on the political choice of public versus private education and the effect of distortive redistribution in the presence of borrowing constraints, respectively.

¹⁰To be precise, ϱ_t equals the fraction of the period that an old household is required to work.

Each cohort consists of a continuum of homogeneous agents. The ratio of young to old households in period t equals ν_t/p_t , reflecting the gross rate of growth of the number of young households ν_t ($\nu_t > 0$), fertility for short, and longevity p_t .¹¹ Both these demographic parameters follow deterministic processes. On a balanced growth path, the survival probability is constant at value p and the gross population growth rate is given by ν . Savings of young households who die before reaching old age are distributed among their surviving peers, reflecting a perfect annuities market.

2.1 Technology

A continuum of competitive firms transforms capital and labor into output by means of a Cobb-Douglas technology. Capital depreciates after one period. The capital stock per young household, k_t , therefore corresponds to the per-capita savings of young households in the previous period, s_{t-1} , divided by the growth rate of the number of young workers, ν_t . We normalize the period time-endowment to unity and denote leisure consumption of the young and labor productivity by x_t and H_t , respectively. Labor supply by an old worker in period t equals $\varrho_t(1 - x_t)$ reflecting the assumption that young and non-retired old households work the same number of hours.¹² Labor supply per young household in period t is then given by $\psi_t(\varrho_t)(1 - x_t)$ with $\psi_t(\varrho_t) \equiv 1 + \varrho_t\chi p_t/\nu_t$. Output per young household in period t is given by

$$B_0 k_t^\alpha [H_t \psi_t(\varrho_t)(1 - x_t)]^{1-\alpha},$$

where $B_0 > 0$ and the capital share $\alpha \in (0, 1)$.

Production factors are paid their marginal products, due to perfect competition. The wage per unit of time, w_t , and the gross return on physical capital, R_t , therefore satisfy

$$\begin{aligned} w_t &= (1 - \alpha) B_0 H_t^{1-\alpha} k_t^\alpha [\psi_t(\varrho_t)(1 - x_t)]^{-\alpha}, \\ R_t &= \alpha B_0 H_t^{1-\alpha} k_t^{\alpha-1} [\psi_t(\varrho_t)(1 - x_t)]^{1-\alpha} = w_t \frac{\psi_t(\varrho_t)(1 - x_t)}{k_t} \alpha' \end{aligned}$$

with $\alpha' \equiv \alpha/(1 - \alpha)$. As a consequence of annuitization, the gross return on savings of a young household that survives to old age equals $\hat{R}_t \equiv R_t/p_t$.

Labor productivity H_t reflects productive public investment during previous periods.¹³ More specifically, productivity growth is a function of public investment per young household,

$$H_{t+1} = B_1 H_t^{1-\delta} I_t^\delta,$$

¹¹Net immigration also affects the rate of growth of the number of young households.

¹²This assumption is not restrictive as changes in the hours worked by old households may be undone by changes in the retirement age.

¹³Due to the Cobb-Douglas specification of the production function in the final good sector, H_t can equivalently be interpreted as total factor productivity. Since our objective is to analyze the link between changes in the size and composition of the population on the one hand and productivity growth on the other, we do not allow for scale effects as they are sometimes considered in endogenous growth models.

with $B_1 > 0$, $\delta \in (0, 1)$ and I_t denoting investment spending per young household. A specification of this type is standard in the literature.¹⁴

2.2 Government

The government taxes the labor income of young households in period t at rate $\tau_t + \sigma_t$. Revenues fund transfers to retired old households—the component corresponding to τ_t —as well as public investment—the component corresponding to σ_t .

Denoting the total transfer to an old household by b_t , we have¹⁵

$$\begin{aligned} b_t &= w_t(1 - x_t)\tau_t\nu_t/p_t, \\ I_t &= w_t(1 - x_t)\sigma_t. \end{aligned}$$

Public investment I_t , the transfer payment b_t and the retirement age must be non-negative (we exclude lump-sum taxes). The policy instruments therefore have to satisfy¹⁶

$$\tau_t, \sigma_t, \varrho_t \geq 0 \text{ for all } t. \quad (1)$$

We denote a combination of the policy instruments in period t by κ_t , $\kappa_t \equiv (\tau_t, \sigma_t, \varrho_t)$.

Note that we abstract from capital income taxes and impose a balanced budget restriction. Both assumptions are without loss of generality. As discussed in the working paper version (Gonzalez-Eiras and Niepelt, 2007), a capital income tax rate would equal zero in equilibrium since, from the perspective of political decision makers who set policy instruments *ex post*, capital income taxes and old-age transfers are close substitutes.¹⁷ Similarly, old-age transfers and the repayment of government debt are close substitutes as well, rendering a balanced-budget restriction non restrictive (see the discussion in Section 5).

2.3 Preferences

Young households value consumption during young age, c_1 , and in old age, c_2 , as well as leisure. Agents discount the future at factor $\beta \in (0, 1)$. Due to the risk of death, the effective discount factor of young households equals βp_{t+1} . For analytical tractability, we assume that the period utility functions of consumption and leisure are logarithmic. Maximizing expected utility, a worker in period t solves the following problem:

$$\begin{aligned} \max_{s_t, x_t} \quad & \ln(c_{1,t}) + m \ln(x_t) + \beta p_{t+1} [\ln(c_{2,t+1}) + m \ln(1 - \varrho_{t+1}(1 - x_{t+1}))] \\ \text{s.t.} \quad & c_{1,t} = w_t(1 - x_t)(1 - \tau_t - \sigma_t) - s_t, \\ & c_{2,t+1} = s_t \hat{R}_{t+1} + w_{t+1}(1 - x_{t+1})\chi\varrho_{t+1} + b_{t+1}, \end{aligned}$$

¹⁴For example, Boldrin and Montes (2005) use the above specification (which is a special case of Rebelo (1991)), with H_t interpreted as human capital and I_t interpreted as public education.

¹⁵The transfer received *while actually retired* is $w_t(1 - x_t)\tau_t\nu_t/(p_t(1 - \varrho_t))$. The total transfer received during old age is the product of this expression and $1 - \varrho_t$. Alternative normalizations do not affect the results.

¹⁶We suppress the upper bound of unity on the retirement age since it will not be binding.

¹⁷In the presence of within-cohort heterogeneity, this need no longer be the case.

where $m \geq 0$ characterizes the preference for leisure. Note that the treatment of leisure is symmetric over the life cycle. Leisure consumption during old age equals the difference between the time endowment, 1, and the time spent working which reflects the intensive and extensive (retirement) margins of labor supply.

2.4 Economic Equilibrium

The first-order conditions characterizing the households' savings and labor-supply decisions, respectively, are standard:

$$\begin{aligned}\frac{1}{c_{1,t}} &= \beta p_{t+1} \hat{R}_{t+1} \frac{1}{c_{2,t+1}}, \\ \frac{m}{x_t} &= w_t(1 - \tau_t - \sigma_t) \frac{1}{c_{1,t}}.\end{aligned}$$

Substituting, the Euler equation characterizing the optimal savings choice of an individual household is given by

$$\frac{s_t \hat{R}_{t+1} + w_{t+1}(1 - x_{t+1})\chi \varrho_{t+1} + b_{t+1}}{\beta p_{t+1} \hat{R}_{t+1}} = w_t(1 - x_t)(1 - \tau_t - \sigma_t) - s_t.$$

Using $s_{t-1}R_t/\nu_t = w_t\psi_t(\varrho_t)(1 - x_t)\alpha'$, simplifying the left-hand side of the equation and setting individual and average savings equal to each other, we find the aggregate savings function

$$s_t = z_{t+1}(\tau_{t+1}, \varrho_{t+1})w_t(1 - x_t)(1 - \tau_t - \sigma_t),$$

where $z_{t+1}(\tau_{t+1}, \varrho_{t+1})$ denotes the savings rate of young households in period t ,

$$z_{t+1}(\tau_{t+1}, \varrho_{t+1}) \equiv \frac{\alpha\beta}{\alpha(1 + \beta p_{t+1})/p_{t+1} + \frac{1-\alpha}{\psi_{t+1}(\varrho_{t+1})} \left(\frac{\tau_{t+1}}{p_{t+1}} + \frac{\chi\varrho_{t+1}}{\nu_{t+1}} \right)} \geq 0.$$

Note that the savings rate in period t depends on policy choices in period $t + 1$. (If $\tau_{t+1} > 0$ or $\chi\varrho_{t+1} > 0$, old households receive retirement benefits or labor income in addition to the return on their savings. This renders the savings rate endogenous to policy, even with logarithmic preferences.) If these policy instruments themselves depend on aggregate savings, then the above relation characterizes savings only implicitly. We return to this point when discussing the objective function maximized in the political process.

Combining the first-order condition for leisure with the expression for $c_{1,t}$ yields

$$x_t = \frac{m(1 - z_{t+1}(\tau_{t+1}, \varrho_{t+1}))}{1 + m(1 - z_{t+1}(\tau_{t+1}, \varrho_{t+1}))}. \quad (2)$$

Note that labor supply is independent of contemporaneous taxes as income and substitution effects cancel.

The endogenous state variables at time t are H_t and k_t . To simplify notation, we work with the state variables H_t and $q_t \equiv H_t^{1-\alpha}k_t^\alpha$ instead. Let $\mathcal{L}_t \equiv B_0(1 - \alpha)q_t(1 -$

$x_t)^{1-\alpha}\psi_t(\varrho_t)^{-\alpha} = w_t(1 - x_t)$ denote labor income of a young household. Combining $k_t = s_{t-1}/\nu_t$ and the aggregate savings function with the dynamic budget constraint and the expressions for factor prices, the equilibrium allocation can recursively be expressed in terms of the following functions of state variables and policy instruments:

$$\left. \begin{aligned} k_{t+1} &= \mathcal{L}_t (1 - \tau_t - \sigma_t) z_{t+1}(\tau_{t+1}, \varrho_{t+1})/\nu_{t+1} = s_t/\nu_{t+1}, \\ c_{1,t} &= \mathcal{L}_t (1 - \tau_t - \sigma_t) (1 - z_{t+1}(\tau_{t+1}, \varrho_{t+1})), \\ c_{2,t} &= \mathcal{L}_t \nu_t \left(\alpha' \psi_t(\varrho_t)/p_t + \frac{\tau_t}{p_t} + \frac{\chi \varrho_t}{\nu_t} \right), \\ x_t &= \frac{m(1-z_{t+1}(\tau_{t+1}, \varrho_{t+1}))}{1+m(1-z_{t+1}(\tau_{t+1}, \varrho_{t+1}))}, \\ H_{t+1} &= B_1 H_t^{1-\delta} (\mathcal{L}_t \sigma_t)^\delta, \\ q_{t+1} &= \left(B_1 H_t^{1-\delta} (\mathcal{L}_t \sigma_t)^\delta \right)^{1-\alpha} (\mathcal{L}_t (1 - \tau_t - \sigma_t) z_{t+1}(\tau_{t+1}, \varrho_{t+1})/\nu_{t+1})^\alpha. \end{aligned} \right\} \quad (3)$$

Conditional on initial values for the two endogenous state variables, (H_0, q_0) , as well as a sequence of policy instruments, $\{\kappa_t\}_{t=0}^\infty$, conditions (3) fully characterize the equilibrium allocation. Taking logarithms, we can express the laws of motion of the two endogenous state variables as

$$\begin{bmatrix} \ln(H_{t+1}) \\ \ln(q_{t+1}) \end{bmatrix} = \underbrace{\begin{bmatrix} 1 - \delta & \delta \\ (1 - \alpha)(1 - \delta) & \alpha + \delta(1 - \alpha) \end{bmatrix}}_{\equiv M} \begin{bmatrix} \ln(H_t) \\ \ln(q_t) \end{bmatrix} + \underbrace{\begin{bmatrix} \xi_t^H(\cdot) \\ \xi_t^q(\cdot) \end{bmatrix}}_{\equiv \xi_t} \quad (4)$$

where the definitions of $\xi_t^H(\sigma_t, \varrho_t, \tau_{t+1}, \varrho_{t+1})$ and $\xi_t^q(\tau_t, \sigma_t, \varrho_t, \tau_{t+1}, \varrho_{t+1})$ follow from (3).

In the special case with inelastic labor supply, $m = 0$, the equilibrium conditions (3) maintain their validity and $x_t = 0$.

2.5 Balanced Growth Path

Along a balanced growth path, all tax rates and demographic variables are constant, implying that labor supply is time-invariant as well. From (3), the growth rates of k_t , s_t , $c_{1,t}$, and $c_{2,t}$ then are equal to the growth rate of q_t . The laws of motion for the two endogenous state variables in (3) imply that, along the balanced growth path, the gross growth rate of H_t , γ_H , equals the gross growth rate of q_t . For any time-invariant choice of tax rates, the last two equations in (3) therefore pin down the ratio H_t/q_t on the corresponding balanced growth path. Given this ratio, the same two conditions pin down γ_H and thus, the balanced growth rates of q_t , k_t , s_t , $c_{1,t}$, and $c_{2,t}$:

$$\begin{aligned} \gamma_H &= \left((B_0 \psi(\varrho)^{-\alpha} (1 - \alpha) (1 - x)^{1-\alpha})^\delta B_1^{1-\alpha} \left((1 - \tau - \sigma) \frac{z(\tau, \varrho)}{\nu} \right)^{\alpha\delta} \sigma^{\delta(1-\alpha)} \right)^{\frac{1}{1-\alpha(1-\delta)}} \\ &\text{s.t. (2)}. \end{aligned} \quad (5)$$

Gross population growth ν has a direct negative effect on per-capita growth because it reduces the capital stock per young household for a given savings rate (the effect captured by ν in the denominator), and a positive effect because it reduces total labor supply and

increases wages given the stock of capital per young household (the effect captured by the term $\psi(\varrho)^{-\alpha}$). Longevity has a direct negative effect on growth by increasing total labor supply and reducing wages given the stock of capital per young household (the effect captured by the term $\psi(\varrho)^{-\alpha}$, again). In addition, changes in fertility and longevity affect the savings rate and thus also labor supply of young households.

Income taxes depress growth because they lower disposable income of young households (the effect captured by the term $1 - \tau - \sigma$), as do retirement benefits because they lower the savings rate ($z(\cdot)$ is decreasing in τ). At the same time, public investment fosters productivity growth (the effect captured by σ in the last term), in line with empirical evidence (e.g., Blankenau, Simpson and Tomljanovich, 2007). Later retirement has a negative effect on growth. It lowers the wages of workers (the effect captured by $\psi(\varrho)^{-\alpha}$) and also decreases the savings rate.

In the following, we sometimes write the growth rate as $\gamma_H((p, \nu), \kappa(p, \nu))$ to indicate that demographic change affects growth both directly and indirectly, by altering the choice of policy instruments κ_t . Growth theory commonly analyzes the direct effect of demographic change on growth, $\partial\gamma_H((p, \nu), \kappa)/\partial(p, \nu)$, or the direct effect of policy on growth, $\partial\gamma_H((p, \nu), \kappa)/\partial\kappa$. Our objective is to analyze the combined direct and indirect effects of demographic change on growth, $d\gamma_H((p, \nu), \kappa(p, \nu))/d(p, \nu)$. In Section 4, we quantitatively assess these effects for advanced OECD economies.

Physical capital along the long-run growth path satisfies $k_{t+1} = \mathcal{L}_t(1 - \tau - \sigma) z(\tau, \varrho)/\nu$. Since k_t grows at the gross rate γ_H , it follows that

$$\begin{aligned} \left(\frac{H_t}{k_t}\right)^{1-\alpha} &= \frac{\gamma_H \nu}{B_0 \psi(\varrho)^{-\alpha} (1-\alpha)(1-x)^{1-\alpha} (1-\tau-\sigma) z(\tau, \varrho)} \text{ s.t. (2),} \\ R &= \frac{\alpha \psi(\varrho) \gamma_H \nu}{(1-\alpha)(1-\tau-\sigma) z(\tau, \varrho)} \text{ s.t. (2).} \end{aligned}$$

2.6 Exogenous Growth Specification

The recent growth literature supports the notion that technology diffusion or trade linkages work towards an equalization of growth rates across regions and countries (see, e.g., Acemoglu, 2009, ch. 18). In the context of our model, technology diffusion across countries can be modeled by positing that a country's growth rate of H does not only depend on public investment in that country, but also on investment or productivity growth in other countries. For example, one may posit that

$$H_{t+1} = B_1 \bar{H}_t^{(1-\varepsilon)(1-\delta)} H_t^{\varepsilon(1-\delta)} I_t^\delta,$$

where $0 \leq \varepsilon < 1$ and \bar{H}_t denotes productivity in the rest of the world. (The model analyzed so far corresponds to the case $\varepsilon = 1$.) To the extent that "foreign" productivity growth is exogenous, "domestic" long-term productivity growth is exogenous as well. In particular, domestic growth dynamics may then be modeled by a closed-economy specification,

$$H_{t+1} = B_{1,t} H_t^{\varepsilon(1-\delta)} I_t^\delta,$$

where growth of H , k and q at rate γ_H is sustained by exogenous growth of $B_{1,t}$ at the rate $\gamma_{B_1} = \gamma_H^{(1-\delta)(1-\varepsilon)}$.

In this exogenous-growth specification, the equations for k_{t+1} , $c_{1,t}$ and $c_{2,t}$ in (3) remain valid while the law of motion for the state variables in (4) changes to

$$\begin{bmatrix} \ln(H_{t+1}) \\ \ln(q_{t+1}) \end{bmatrix} = \begin{bmatrix} \varepsilon(1-\delta) & \delta \\ \varepsilon(1-\alpha)(1-\delta) & \alpha + \delta(1-\alpha) \end{bmatrix} \begin{bmatrix} \ln(H_t) \\ \ln(q_t) \end{bmatrix} + \xi_t, \quad (6)$$

where ξ_t differs from the corresponding expression in the endogenous-growth specification insofar as $B_{1,t}$ increases over time. The equilibrium expressions for H_t/k_t and R given above still apply.

3 Politico-Economic Equilibrium

We assume that young and old households vote on candidates whose electoral platforms specify values for the policy instruments, κ_t . Voters do not only support a candidate for her policy platform, but also for other characteristics (“ideology”) that are orthogonal to the fundamental policy dimensions of interest. These characteristics are permanent and cannot be credibly altered in the course of electoral competition. Moreover, their valuation differs across voters (even if voters agree about the preferred policy platform) and is subject to random aggregate shocks, realized after candidates have chosen their platforms. This “probabilistic-voting” setup renders the probability of winning a voter’s support a continuous function of the competing policy platforms, implying that equilibrium policy platforms smoothly respond to changes in the demographic structure. This stands in contrast to the “median-voter” setup where, in a model with only a few generations, a small change in the demographic structure has large effects on policy outcomes if it alters the cohort the median voter is associated with.

In the Nash equilibrium of the probabilistic-voting game with two candidates choosing platforms to maximize their expected vote shares, both candidates propose the same policy platform.¹⁸ This platform maximizes a convex combination of the objective functions of all groups of voters, where the weights reflect the groups’ size and sensitivity of voting behavior to policy changes. Those groups that care the most about policy platforms rather than other candidate characteristics are the most likely to shift their support from one candidate to the other in response to small changes in the proposed platforms. In equilibrium, such groups of “swing voters” thus gain in political influence and tilt policy in their own favor. If all voters are equally responsive to changes in the policy platforms, electoral competition implements the utilitarian optimum with respect to voters.

Owing to political competition at the beginning of each period, policy makers cannot commit to future policy platforms. Voters therefore have to form expectations about the effect of current policy choices on future policy outcomes. Under the Markov assumption, future leisure and policy choices are functions of the fundamental state variables only, $x_{t+1} = \tilde{x}_{t+1}(H_{t+1}, q_{t+1})$ and $\kappa_{t+1} = \tilde{\kappa}_{t+1}(H_{t+1}, q_{t+1})$. (The state variables include

¹⁸See Lindbeck and Weibull (1987) and Persson and Tabellini (2000) for discussions of the probabilistic-voting setup.

demographic variables, thus the time indices of the policy functions.) If the policy functions are independent of (H, q) , $\kappa_{t+1} = \tilde{\kappa}_{t+1}$, then (2) implies that the leisure function is independent of (H, q) as well, $x_{t+1} = \tilde{x}_{t+1}$, and both the aggregate savings function and the economic equilibrium conditions (3) apply (recall the discussion of the aggregate savings function in subsection 2.4). In the following, we conjecture that the policy functions indeed are independent of (H, q) . We derive the equilibrium choice of policy instruments under this conjecture and show that this choice does not depend on (H, q) , thereby verifying the conjecture.

The political objective function, $W_t(\cdot)$, depends on the endogenous state variables (as well as the exogenous ones, thus the time index), the contemporaneous policy instruments, and the anticipated values of policy instruments and leisure in the following period. Letting ω denote the per-capita political influence of old relative to young households, we define

$$\begin{aligned} W_t(H_t, q_t, \kappa_t; \tilde{\kappa}_{t+1}, \tilde{x}_{t+1}) &\equiv \omega p_t \{ \ln(c_{2,t}) + m \ln(1 - \varrho_t(1 - x_t)) \} \\ &\quad + \nu_t \{ \ln(c_{1,t}) + m \ln(x_t) + \beta p_{t+1} [\ln(c_{2,t+1}) + m \ln(1 - \varrho_{t+1}(1 - x_{t+1}))] \} \\ \text{s.t. (3), } &\kappa_{t+1} = \tilde{\kappa}_{t+1}, \quad x_{t+1} = \tilde{x}_{t+1}. \end{aligned}$$

The program characterizing equilibrium policy choices in period t is given by

$$\max_{\kappa_t} W_t(H_t, q_t, \kappa_t; \tilde{\kappa}_{t+1}, \tilde{x}_{t+1}) \text{ s.t. (1), } H_t, q_t \text{ given.}$$

Political equilibrium requires that for any combination of state variables (H_t, q_t) , the κ_t solving this program is given by $\tilde{\kappa}_t$.

Using the equilibrium expressions for consumption from (3), the objective function can be expressed as

$$\begin{aligned} W_t(\cdot) &= \omega p_t \left\{ \ln \left[\psi_t(\varrho_t)^{-\alpha} \left(\alpha' \psi_t(\varrho_t) / p_t + \frac{\tau_t}{p_t} + \frac{\chi \varrho_t}{\nu_t} \right) \right] + m \ln(1 - \varrho_t(1 - x_t)) \right\} \\ &\quad + \nu_t \left\{ \ln[\psi_t(\varrho_t)^{-\alpha} (1 - \tau_t - \sigma_t)] \right. \\ &\quad \left. + \beta p_{t+1} \ln \left[(\psi_t(\varrho_t)^{-\alpha})^{\delta(1-\alpha)+\alpha} (1 - \tau_t - \sigma_t)^\alpha \sigma_t^{\delta(1-\alpha)} \right] \right\} \\ &\quad + \text{t.i.p. s.t. (2),} \end{aligned}$$

where t.i.p. denotes terms that are unaffected by contemporaneous policy choices (under the conjecture), due to the logarithmic preference assumption. In particular, t.i.p. includes H_t and q_t and, with an exogenous growth specification, the parameter ε determining the strength of the intertemporal spillover from H . Since the contemporaneous policy instruments do not interact with the state variables H_t or q_t , the equilibrium policy functions are independent of these state variables, confirming the initial conjecture.¹⁹ Similarly, since in the case with an exogenous growth specification the parameter ε does

¹⁹In related work, we analyze the sensitivity of a parallel result to changes in functional form assumptions (Gonzalez-Eiras and Niepelt, 2005). We find that, in the case of generalized CRRA preferences, state variables and policy instruments do interact. However, the quantitative implications for equilibrium policies are negligible.

not interact with the policy instruments, the equilibrium policy choices in the endogenous and exogenous growth specifications of the model coincide. This is a reflection of the fact that the political process does not internalize the long-term benefits of public investment because these occur beyond the lifetimes of even the youngest voters.

Letting $\Delta_{t+1} \equiv 1 + \beta p_{t+1}(\alpha + \delta(1 - \alpha))$ denote the semi-elasticity of young households' utility with respect to labor income, the first-order conditions with respect to τ_t , σ_t and ϱ_t , respectively, read

$$\begin{aligned} \frac{\omega p_t}{\nu_t} \frac{\frac{1}{p_t}}{\alpha' \psi_t(\varrho_t)/p_t + \frac{\tau_t}{p_t} + \frac{\chi \varrho_t}{\nu_t}} + \lambda^\tau &= \frac{1 + \alpha \beta p_{t+1}}{1 - \tau_t - \sigma_t}, \\ \frac{\delta(1 - \alpha) \beta p_{t+1}}{\sigma_t} + \lambda^\sigma &= \frac{1 + \alpha \beta p_{t+1}}{1 - \tau_t - \sigma_t}, \\ \frac{\omega p_t}{\nu_t} \left(\frac{\alpha' \psi_t'(\varrho_t)/p_t + \chi/\nu_t}{\alpha' \psi_t(\varrho_t)/p_t + \frac{\tau_t}{p_t} + \frac{\chi \varrho_t}{\nu_t}} - \frac{m(1 - x_t)}{(1 - \varrho_t(1 - x_t))} \right) + \lambda^e &= \frac{\alpha \psi_t'(\varrho_t)}{\psi_t(\varrho_t)} \left(\frac{\omega p_t}{\nu_t} + \Delta_{t+1} \right), \end{aligned}$$

where the λ s denote multipliers associated with the non-negativity constraints on the policy instruments.

If the tax rates τ_t and σ_t are interior (as is the case in the data and in the simulations we conduct later) and ϱ_t is in a corner then we can solve the former two first-order conditions for

$$\begin{aligned} \tau_t &= \frac{\frac{\omega p_t}{\nu_t} - \alpha' \Delta_{t+1}}{\frac{\omega p_t}{\nu_t} + \Delta_{t+1}}, \\ \sigma_t &= \beta \delta p_{t+1} \frac{1}{\frac{\omega p_t}{\nu_t} + \Delta_{t+1}}, \end{aligned}$$

implying that tax rates in period t depend on demographics in periods t and $t + 1$.

If the three instruments have an interior solution, solving the first-order conditions yields

$$\begin{aligned} \tau_t &= \frac{(1 - \alpha) \left(\frac{\omega p_t}{\nu_t} + \Delta_{t+1} \right) + \frac{\omega p_t}{\nu_t} m - \Delta_{t+1} \left(1 + \frac{\chi p_t}{\nu_t(1 - x_t)} \right)}{(1 - \alpha) \left(\frac{\omega p_t}{\nu_t} + \Delta_{t+1} \right) + \frac{\omega p_t}{\nu_t} m}, \\ \sigma_t &= \beta \delta p_{t+1} \frac{(1 - \alpha) \left(1 + \frac{\chi p_t}{\nu_t(1 - x_t)} \right)}{(1 - \alpha) \left(\frac{\omega p_t}{\nu_t} + \Delta_{t+1} \right) + \frac{\omega p_t}{\nu_t} m}, \\ \varrho_t &= \frac{(1 - \alpha) \left(\frac{\omega p_t}{\nu_t} + \Delta_{t+1} \right) / (1 - x_t) - \frac{\omega}{\chi} m}{(1 - \alpha) \left(\frac{\omega p_t}{\nu_t} + \Delta_{t+1} \right) + \frac{\omega p_t}{\nu_t} m}. \end{aligned}$$

Since x_t is a function of future policy choices (see (2)) it appears that, in this case, the policy instruments effectively depend on current and all future demographic shocks. This is not the case, however, because the combination of policy instruments entering into the expression for labor supply²⁰ in period t does *not* depend on $1 - x_{t+1}$ but only on

²⁰This combination is given by $\frac{1 - \alpha}{\psi_{t+1}(\varrho_{t+1})} \left(\frac{\tau_{t+1}}{p_{t+1}} + \frac{\chi \varrho_{t+1}}{\nu_{t+1}} \right)$, see equation (2).

parameters and demographic shocks in periods $t + 1$ and $t + 2$:

$$1 - x_t = 1 - \frac{m \frac{\omega}{\nu_{t+1}}}{(1 + m) \frac{\omega}{\nu_{t+1}} + \alpha \beta \left(\frac{\omega p_{t+1}}{\nu_{t+1}} + \Delta_{t+2} \right)}.$$

Accordingly, the equilibrium policy choices κ_t only depend on parameters and demographic shocks in periods t through $t + 2$.

Demographic change affects the policy instruments through several channels. (Unless otherwise noted, the comparative statics results for τ_t and σ_t hold even if ϱ_t is in a corner or fixed.) First, by altering the relative political power of the old in the current period, $\frac{\omega p_t}{\nu_t}$. Higher relative political power of the old (reflecting a higher number of old relative to young households due to lower fertility or higher longevity in the previous period) raises τ_t and lowers σ_t and ϱ_t .²¹ Intuitively, more powerful elderly voters secure higher intergenerational transfers and earlier retirement. Because taxes are distorting, funding for other government outlays is reduced.

Second, demographic change affects the policy instruments by altering the effect of later retirement on aggregate labor supply, $\frac{\chi p_t}{\nu_t}$. A stronger such effect (reflecting a higher number of old relative to young households, again) reduces τ_t and increases σ_t and ϱ_t . Intuitively, a higher number of old relative to young households strengthens the extent to which an increase in the retirement age translates into available resources. The additional resources generated due to later retirement reduce the need for intergenerational transfers and limit tax distortions. This allows to increase funding for public investment. The combined effect of the first two channels (reflecting the number of old relative to young households) is to raise τ_t and ϱ_t and lower σ_t in response to demographic ageing. That is, the retirement age rises during the demographic transition although the relative political power of the old increases when society ages.

Third, demographic change affects the policy instruments by altering the longevity of young households, p_{t+1} , and thus, the weight the political process attaches to the future (see also Gradstein and Kaganovich, 2004). With a higher such weight, public investment becomes easier to sustain politically. As a consequence, an increase in longevity reduces τ_t but increases σ_t and ϱ_t .

Finally, with an interior retirement age, demographic change affects the policy instruments by altering future policy choices and thus, contemporaneous labor supply, $1 - x_t$. In particular, increases in p_{t+1} , ν_{t+1} and p_{t+2} all raise contemporaneous labor supply which in turn increases τ_t and reduces σ_t and ϱ_t .

The comparative statics results working through the first channel correspond with conventional wisdom. Often overlooked are the consequences of the other channels, as well as the fact that their interaction generates non-monotone dynamics. For example, a permanent shock to longevity may give rise to a fall in the tax rate τ in the period preceding the shock (where longevity fosters the incentive to invest) followed by a recovery thereafter (where it strengthens the political power of the old). In response to the same shock, the tax rate σ may display the opposite dynamics while the retirement age

²¹We interpret the ratio ω/χ as $\frac{\omega p_t}{\nu_t} / \frac{\chi p_t}{\nu_t}$.

rises both in the short and the long run. Before the background of these non-monotone dynamics, data about the short-run evolution of government budget shares cannot easily be extrapolated to predict the direction of long-run change in these shares.²²

The equilibrium policy functions κ_t are *unique* in the limit of the finite horizon economy. To see this, consider the final period T and note that labor supply is inelastic in this final period (from (2)). The political objective function in period T therefore depends on the consumption of young and old households,

$$c_{1,T} = \mathcal{L}_T(1 - \tau_T) \quad \text{and} \quad c_{2,T} = \mathcal{L}_T \nu_T \left(\frac{(\alpha' + 1)\psi_T(\varrho_T) + \tau_T - 1}{p_T} \right),$$

respectively. Note that $\sigma_T = 0$ since there is no benefit of public investment in the final period. Differentiating the political objective function in period T with respect to τ_T and ϱ_T yields two equations in the policy instruments that are independent of (H_T, q_T) . This implies that κ_T is not a function of (H_T, q_T) . Moving to period $T - 1$, the forgoing analysis shows that the policy functions κ_{T-1} are independent of (H_{T-1}, q_{T-1}) as well, and given by the equilibrium expressions reported earlier. The result then follows by induction.

In Appendix A, we derive as a criterion for *production efficiency* along a balanced growth path the requirement that

$$1 \geq \alpha' \frac{I}{s} > \delta.$$

If either of the two inequalities is violated then a reallocation of investment spending between I and s may weakly increase output in all future periods, and strictly in some (see Cass, 1972). In particular, if the left inequality is violated, the economy accumulates too much H and if the right inequality is violated, the economy accumulates too much k . As shown in Appendix A, the economy necessarily over accumulates k relative to H in politico-economic equilibrium. As a consequence, the allocation in politico-economic equilibrium necessarily differs from the allocation supported by any Ramsey policy.²³

4 Quantitative Implications of Demographic Ageing

Based on the analytical results derived earlier, we compute quantitative predictions for a synthetic “rich OECD economy,” representing the population weighted average of Australia, Canada, Denmark, France, Germany, Italy, Japan, New Zealand, Sweden, the

²²The comparative statics with respect to the structural parameters of the model are intuitive. For example, an increase in the preference for leisure, m , raises the marginal cost of working for the old and induces the political process to reduce the retirement age and shift government spending from investment to social security benefits. The effect working through changes in labor supply reinforces this adjustment. If the retirement age is fixed, changes in m do not affect the two tax rates. An increase in the elasticity of productivity growth to public investment, δ , leads to a reduction of τ_t and an increase in σ_t and ϱ_t .

²³With exogenous growth ($\varepsilon < 1$), the production inefficiency result remains valid, see Appendix A. The political process does not internalize the dynamic externality from current to future H , in contrast to a Ramsey planner. As a consequence, the strength of this dynamic externality is irrelevant for the equilibrium κ_t .

United Kingdom and the United States; a synthetic “rich European OECD economy,” representing the population weighted average of the European countries in the above list; the United States; and Japan. We take one period in the model to correspond to 30 years in the data. Accordingly, we compute three sequences of model predictions with a period length of 30 years each. In the first sequence, the periods correspond to the years 1970, 2000, 2030, . . . ; in the second sequence, to the years 1980, 2010, 2040, . . . ; and in the third sequence, to the years 1990, 2020, 2050, When reporting time series predictions, we merge the three sequences in a single time series.

We use the 30-year population growth rate as a measure of ν_t , the number of young households in period t relative to the number in the preceding period. For p_t , the number of old households in period t relative to the number of young households in the preceding period, we use estimates for life expectancy at age 65 divided by 30 years.²⁴ Figures 3 and 4 plot the demographic series underlying the model predictions.

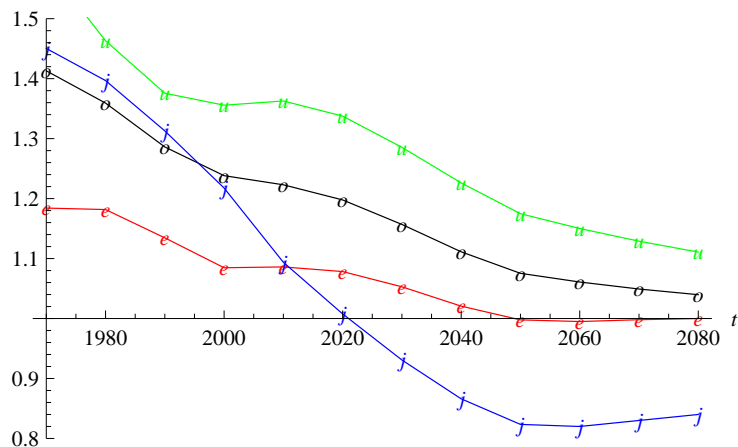


Figure 3: ν_t for the rich OECD economy (black, “o”), the rich European OECD economy (red, “e”), the United States (green, “u”), and Japan (blue, “j”).

We set α to 0.3, a standard value in the literature, normalize B_0 to unity, and let $\chi = 1.135$, based on estimates of labor productivity over the life cycle in Heathcote, Storesletten and Violante (2008).²⁵ To calibrate $\beta, \delta, \omega, B_1$ and m , we impose model restrictions. First, we fix the GDP-shares of transfers and public investment in the year 2000, $(1 - \alpha)\tau_{2000}$ and $(1 - \alpha)\sigma_{2000}$ respectively, at the values 0.0796 and 0.0727, the corresponding averages in the rich OECD economy.²⁶ Second, we fix the balanced-growth-

²⁴Demographic data from the year 1950 onward is taken from Population Division of the Department of Economic and Social Affairs of the United Nations Secretariat (2005, 2007): *World Population Prospects*, New York. Population data for the year 1940 is taken from <http://www.populstat.info>. Data about life expectancy at age 65 is available up to the year 2000. Data for 2010 to 2050 is extrapolated from data about life expectancy at birth. Demographic data for the year 2060 and later is extrapolated under the assumption that p and ν converge to unity in the long run.

²⁵We fit a polynomial to Heathcote et al.’s (2008) estimates for the United States and compute the average values before and after age 65.

²⁶We proxy the former by the GDP-share of old-age and survivors pensions and the latter by the

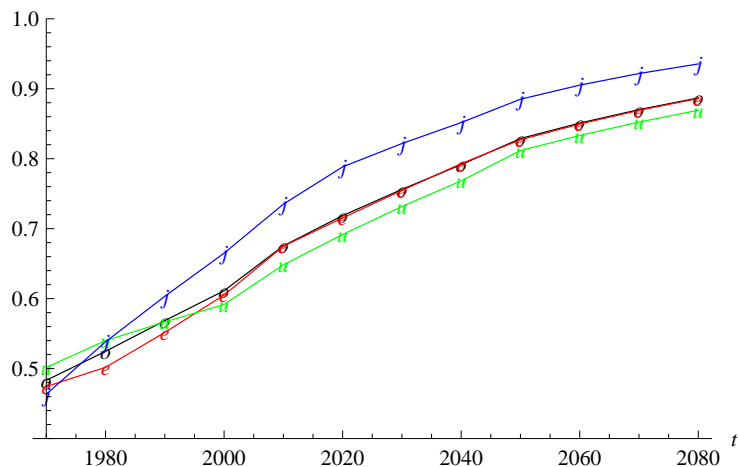


Figure 4: p_t for the rich OECD economy (black, “o”), the rich European OECD economy (red, “e”), the United States (green, “u”), and Japan (blue, “j”).

path growth rate and interest rate when evaluated at the year-2000 demographics at the observed values in the rich OECD economy.²⁷ Finally, we fix the labor supply of a young household in the year 2000 at $1/3$. These restrictions imply $\beta = 0.7226$ (0.9892 on an annual basis), $\delta = 0.4039$, $\omega = 1.8256$, $B_1 = 10.7738$ and $m = 2.7011$.

Figures 5–9 display the predicted policy responses to demographic change in the rich OECD economy, the rich European OECD economy, the United States, and Japan. All simulations are based on the calibration described above and differ only with respect to the demographic series fed into the model. As a consequence, the actual budget shares in the year 2000 are exactly matched in the case of the rich OECD economy.

According to the model predictions displayed in Figure 5, τ_t more than doubles in the rich OECD economy between the years 1970 and 2000, flattening out thereafter and increasing further by approximately three percentage points up to the year 2080. Tax rate σ_t increases much slower, rising by slightly more than two percentage points between the years 1970 and 2080. Retirement age ϱ_t rises by more than 6 years between 2000 and 2080, to be compared with an increase in life expectancy at age 65 of more than 8 years. Labor supply of young households (not displayed) rises by about one percent. In contrast with the savings rate of young households, the national savings rate (not displayed) is predicted to fall from more than 7 percent in 1970 to roughly 6 percent in 2010 and roughly 4 percent in 2080, due to the increased fraction of the elderly who are dissaving.

If the retirement age were not allowed to rise beyond its level in the year 2000, the tax

GDP-share of government fixed capital formation and (other) government expenditures for education (all levels of government). Data is taken from OECD sources. (Due to data limitations, the components underlying the historical shares reported in the Introduction differ slightly from the ones we choose for the calibration. In particular, the historical pension-share series contains disability benefits and the historical investment-share series may be slightly biased due to double counting.)

²⁷We calibrate γ_H based on the average annual multifactor productivity growth rate of the rich OECD economy between the years 1985 and 2005, 1.0113 (OECD sources), and R based on Gonzalez-Eiras and Niepelt’s (2008) estimate of the annual gross interest rate in the United States, 1.0483.

rate τ_t would increase steeply and σ_t would decline, see Figure 6. Intuitively, with a capped ϱ_t the growing number of non-working old relative to young households would require increasingly high social-security contributions per young household and the induced rise in τ_t would render taxation more costly, triggering a fall in the tax rate σ_t and crowding out of government investment. Interestingly, this scenario closely corresponds with fears voiced in the policy debate. While the model encompasses the mechanisms underlying such fears, it predicts a different resolution of intergenerational conflict because of adjustments along the retirement margin.

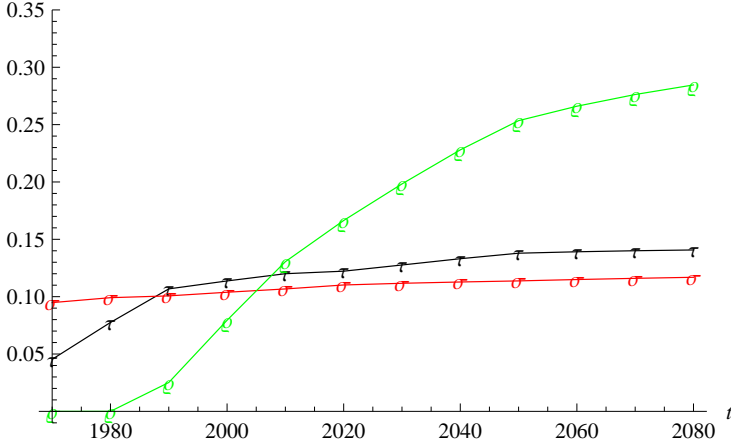


Figure 5: Predicted policies for the rich OECD economy: τ_t (black), σ_t (red), ϱ_t (green).

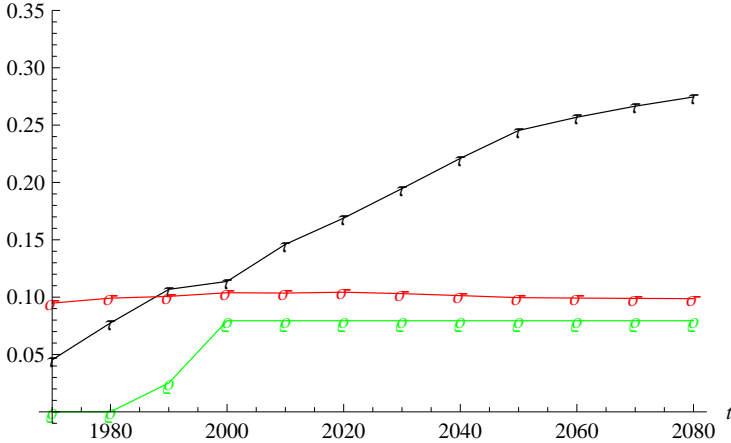


Figure 6: Predicted policies for the rich OECD economy if ϱ_t is capped at its year-2000 value: τ_t (black), σ_t (red), ϱ_t (green).

Returning to the scenario where all three policy instruments are free to adjust, the predicted policy responses in the rich European OECD economy, the United States and Japan are similar as far as the public investment share is concerned, see Figures 7–9. The main differences between the three economies concern the budget share for social security transfers on the one hand and the retirement age on the other: the social-security budget

share in the United States and Japan starts out from a lower level than in the rich European OECD economy but catches up during the early years of the simulation; and the retirement age increases earlier in the rich European OECD economy and Japan than in the United States.

In broad terms, these predictions about the relative performance of the three countries and country blocks are consistent with the evidence. In particular, the model predicts the GDP-share of social-security transfers in the year 2000, $(1 - \alpha)\tau_{2000}$, to be highest in the rich European OECD economy (nearly 9 percent), followed by Japan and the United States (more than 7 percent). In the data, the corresponding shares equal roughly 10, 8, and 6 percent respectively. Similarly, the model correctly predicts that the GDP-share of public investment, $(1 - \alpha)\sigma_{2000}$, is higher in Japan than in the United States and the rich European OECD economy: The model predicts GDP-shares of nearly 8 percent for Japan and more than 7 percent for the United States and the rich European OECD economy, in line with the data. The model also performs well in predicting a sharp increase of retirement age in Japan around the year 1990 and a smoother and later response in the United States.²⁸ It performs less satisfactory in predicting a robust increase of retirement age in the rich European OECD economy by the year 1990.²⁹

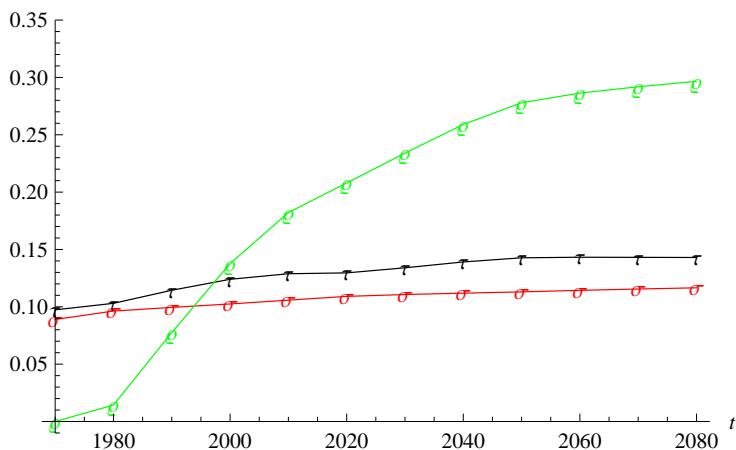


Figure 7: Predicted policies for the rich European OECD economy: τ_t (black), σ_t (red), ρ_t (green).

²⁸In 1994, Japan enacted a rapid increase of retirement age and a reduction of effective tax rates for workers close to retirement (Yashiro and Oshio, 1999). The model captures this rapid increase, but not the precise timing. In the United States, the retirement age started to increase around the year 2000, at a slower rate.

²⁹Empirically, it has only been recently that many European countries have moved towards delaying the statutory retirement age and reducing the incentives for early retirement (e.g., Galasso, 2006, pp. 23–25), or to discussing proposals of such policy changes. This suggests that the model does not capture certain institutional frictions that delay adjustment along the retirement margin, or other motives for changes in the retirement age. For example, starting in the late 1960s and 1970s, early retirement provisions were introduced in many OECD countries in response to high levels of unemployment among middle-aged workers (see, e.g. Conde-Ruiz and Galasso, 2004). Our framework is silent about these developments.

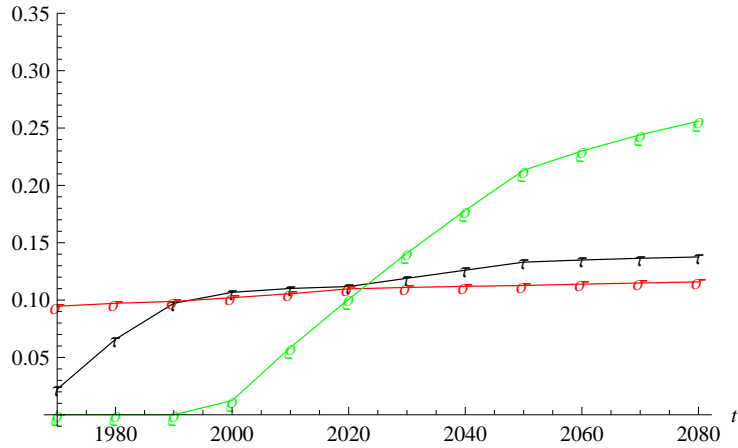


Figure 8: Predicted policies for the United States: τ_t (black), σ_t (red), ρ_t (green).

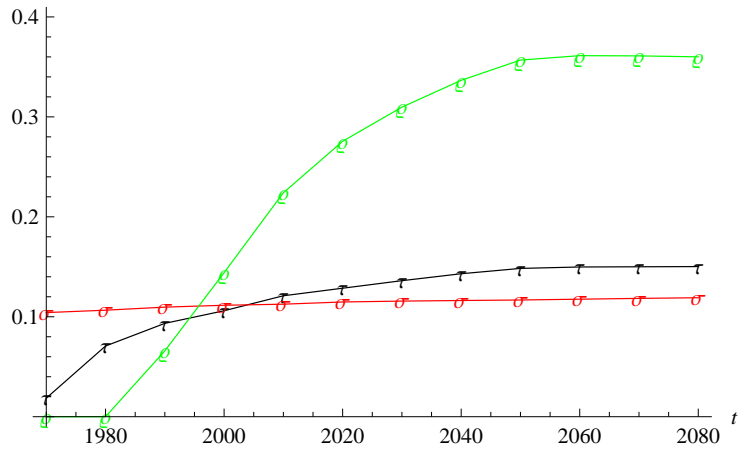


Figure 9: Predicted policies for Japan: τ_t (black), σ_t (red), ρ_t (green).

Figure 10 displays the predicted annual per-capita output growth rates for the four countries and country blocks under consideration. These growth rates are reported as deviations from the balanced-growth rates subject to the year-2000 demographics. Annual growth accelerates by roughly four basis points per decade. By the year 2080, the growth rates have increased by 30 to 35 basis points. The growth accelerations in the rich OECD economy, the rich European OECD economy and the United States are very similar and slightly exceed the one in Japan. If the retirement age were capped at its level in the year 2000 (such that τ_t would rise steeply and σ_t decline), per-capita growth would essentially remain stuck at its current level.

To understand the sources of these predicted growth effects, we compare the per-capita balanced growth rate along the initial balanced growth path subject to year-2000 demographics with the one along a new balanced growth path subject to $p^* = \nu^* = 1$. Recall from the discussion in Subsection 2.5 that the total growth effect of demographic

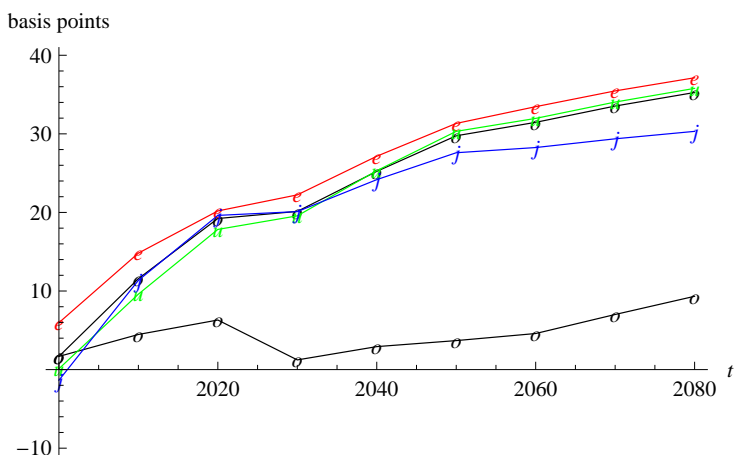


Figure 10: Predicted annual per-capita output growth rates for the rich OECD economy (black, “o”), the rich European OECD economy (red, “e”), the United States (green, “u”), and Japan (blue, “j”) as deviations from the respective initial balanced-growth rates. The series at the bottom displays the growth deviation for the rich OECD economy if ϱ_t is capped at its year-2000 value.

change can be decomposed into a direct and an indirect, policy induced effect:

$$\frac{d\gamma_H((p, \nu), \kappa(p, \nu))}{d(p, \nu)} = \frac{\partial\gamma_H((p, \nu), \kappa(p, \nu))}{\partial(p, \nu)} + \frac{\partial\gamma_H((p, \nu), \kappa(p, \nu))}{\partial\kappa} \frac{\partial\kappa(p, \nu)}{\partial(p, \nu)}.$$

The first term on the right-hand side includes the direct effect of demographic change on economic growth as described in Section 2.5. The second term includes the indirect effect working through induced policy adjustments as discussed in Section 3.

Figure 11 illustrates the relative importance of these two effects. The leftmost bar (denoted by “initial”) indicates the annual per-capita growth rate along the balanced growth path subject to year-2000 demographics. The other bars indicate the predicted annual per-capita growth rates along the new balanced growth path subject to p^* and ν^* under different assumptions about the adjustment of policy instruments. In particular, the bar denoted by “new (1)” indicates the new growth rate if only the direct effect is accounted for, $\gamma_H((p^*, \nu^*), \kappa(p_{2000}, \nu_{2000}))$, and the bar denoted by “new (2)” indicates the new growth rate if direct and indirect effects are accounted for, $\gamma_H((p^*, \nu^*), \kappa(p^*, \nu^*))$. The rightmost bar denoted by “new (3)” indicates the new growth rate if direct and indirect effects are accounted for but the retirement age is held fixed at its year-2000 value.

Figure 11 shows that the direct effect is positive, summing to 57 basis points of annual growth. In contrast, the indirect effect working through adjustments in policy is negative and amounts to roughly 10 basis points, due to higher transfers and later retirement and in spite of higher public investment. With a capped retirement age, the indirect growth effect would be much more negative (37 rather than 10 basis points), leaving a net growth increase of only 20 basis points. We emphasize this last point: while in isolation, the increase of the retirement age works towards reducing growth, fixing the retirement age would not improve growth prospects; to the contrary, it would go hand in hand with a

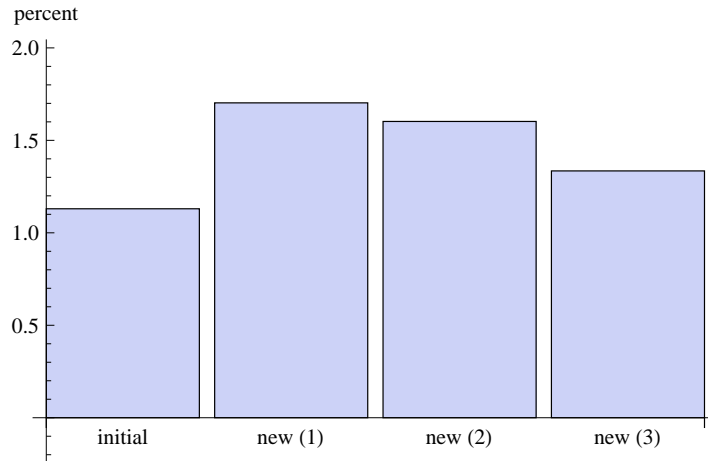


Figure 11: Predicted annual per-capita balanced-growth rate for the rich OECD economy, around the year 2000 (leftmost bar) and along the new balanced growth path (middle and rightmost bars).

very steep increase in taxation and thus, an even stronger fall in growth.

In summary, the picture that emerges is only partly consistent with the view promoted by policy makers according to which the political process will implement measures to raise productivity in order to “outgrow” the burden imposed by demographic change. According to the model, demographic ageing indeed induces the political process to raise public investment in order to foster productivity growth. However, the main positive effects on growth arise directly, through capital deepening, a higher savings rate, and slightly increased labor supply and the net effect of endogenous policy on growth is negative, due to higher social security transfers and an increased retirement age. Nevertheless, in politico-economic equilibrium, flexibility along the retirement margin plays a positive role for growth. For with a capped retirement age, social security transfers would increase much more steeply and the induced negative growth effects would be even stronger.

The central predictions of the model are robust to changes in the calibration. Modified values for the targeted balanced-growth-path interest rate are mainly reflected in adjusted values for β and δ , leaving the simulation results largely unchanged. Reducing χ renders the predicted increase of retirement age less pronounced and more delayed. The effects on the decomposition of the long-run change in γ_H are small but the growth cost of capping ϱ is reduced. The most important parameter for calibration purposes is the capital share α . Increasing α from 0.3 to 0.35 leaves the net long-run effect on γ_H largely unchanged but amplifies the positive and negative contributions to this long-run change through the different channels discussed earlier. The predicted increase in retirement age is smaller if the capital share is high. If α is reduced to 0.25, the retirement age steeply increases already before the year 2000. Starting from the high base value in the year 2000, the further increase of the retirement age does not depress growth as strongly as in the baseline simulation and the net effect of endogenous policy on long-run growth becomes positive. With a capped retirement age, the net effect of endogenous policy

remains strongly negative.

The model predictions are also robust to replacing the endogenous growth specification by one of exogenous growth. As discussed earlier, the equilibrium policy choices remain *identical* in such a variant of the model since the first-order conditions with respect to κ_t are unchanged.³⁰ To evaluate the robustness of the implied growth results, we simulate the model with the exogenous growth specification introduced in Subsection 2.6. This model specification can be calibrated based on the same moment restrictions used previously, except for the one relating to the endogenous balanced growth rate. This latter restriction (which does not apply any longer if γ_H is determined by the exogenous γ_{B_1}) can now be dropped and the previously calibrated parameter in the production function for productivity growth imposed exogenously in the base year 2000. As a result, the numerical values for the model parameters m , β , δ and ω are given by the values calibrated previously and the parameter B_1 in the endogenous growth specification is replaced by the sequence $\{B_{1,t}\}$ whose values grow at the exogenous rate γ_{B_1} . Based on this modified calibration, we can compute the balanced exogenous growth values for the state variables in the base year, H_{2000} and q_{2000} , and use the modified law of motion (6) to analyze the effect of the parameter ε on the growth implications of the demographic transition.

As illustrated in Figure 12, which corresponds to $\varepsilon = 0.5$, the medium-term growth implications are very similar to those in the endogenous growth specification. With endogenous growth, annual per-capita growth in the rich OECD economy is predicted to accelerate by 25–30 basis points in the year 2050 and 30–35 basis points in the year 2080 (see Figure 10). With exogenous growth and $\varepsilon = 0.5$, in contrast, growth is predicted to accelerate by 25 basis points in the year 2050 before the economy begins to revert to its long-run growth rate of $\gamma_H = (\gamma_{B_1})^{\frac{1}{(1-\delta)(1-\varepsilon)}}$. Lower values for ε imply a smaller maximal growth acceleration around the year 2050 and faster reversion thereafter. If ε rises towards the limiting value of unity, the growth dynamics increasingly mimic those of the endogenous growth specification. Finally, if the retirement age is restricted not to rise beyond its value in the year 2000, per-capita growth remains stuck at its level in the year 2000, in parallel to the outcome with endogenous growth (see Figure 12).

In conclusion, the specification of productivity growth in the model does not have a major bearing on the medium-term transition dynamics of the economy and no effect at all on the short-, medium- and long-run evolution of the policy instruments and the government budget shares in politico-economic equilibrium.³¹

5 Concluding Remarks

We have presented a rich, yet tractable framework to analyze the impact of demographic ageing on economic growth. Building on a standard overlapping generations model, our framework combines various channels discussed in the literature and referred to in the political debate. On the one hand, it captures the implications of rising longevity and falling fertility in general equilibrium, including adjustments in the savings rate, labor

³⁰The latter result hinges on the assumption that households live for only two periods.

³¹With exogenous growth, the long-run effect of the demographic transition on output is positive.

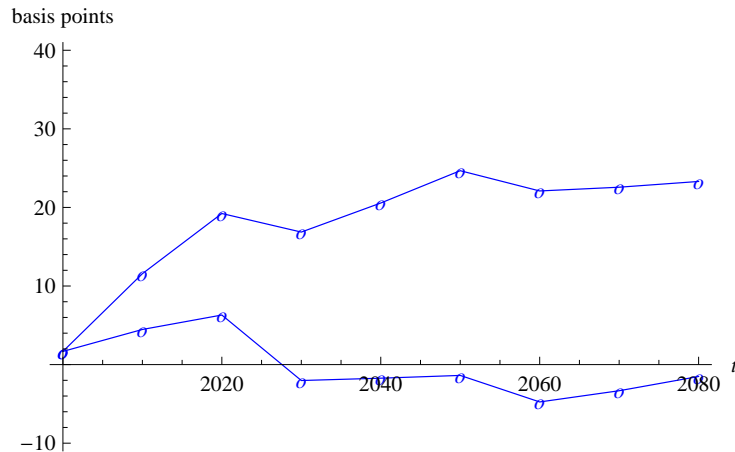


Figure 12: Predicted annual per-capita output growth rate for the rich OECD economy as deviation from the initial balanced-growth rate if growth is exogenous and $\varepsilon = 0.5$. The series at the bottom displays the growth deviation if ϱ_t is capped at its year-2000 value.

supply, factor prices and capital deepening. On the other hand, it captures responses by the political system, in particular adjustments of the size of the government budget and its composition between investment and transfer spending as well as changes in the retirement age.

Calibrated versions of the model predict that annual per-capita growth in rich OECD economies will increase by roughly 30–35 basis points during the twenty-first century, with the positive direct growth effects of demographic ageing partly being reversed by the consequences of endogenous policy responses. The model predictions support the view that rising longevity paired with falling fertility increases the GDP-share of social security transfers, with negative implications for growth. However, they do not support the common view that rising social security transfers crowd out productive public investment (as a share of GDP). Crowding out only results in an extreme scenario where the political process adjusts tax rates and the composition of government spending, but not the retirement age. In the more plausible scenario where the political process adjusts instruments along all three margins, both social security transfers and public investment as a share of GDP rise, and the increase of the former is much more moderate than with a fixed retirement age. These results are robust to changes in the specification of the source of economic growth.

Throughout the paper, we have assumed that the government runs a balanced budget, excluding government deficits and debt. This assumption is not very restrictive. In our model, unlike in Bassetto and Sargent (2006) who assume commitment, public underinvestment cannot be overcome by letting voters finance investment expenditures out of government debt. For lack of commitment implies that the economic equivalence of social-security and debt policies largely extends to the political sphere.³² We have also

³²Gonzalez-Eiras and Niepelt (2010) analyze the economic and politico-economic equivalence of fiscal policies.

assumed that longevity and fertility are exogenous. While this assumption is useful for the purpose of studying the long-run effects of demographic ageing on growth, there are clearly potential feedback effects from government budgets to demographics, for example via investments in public health (see Hall and Jones, 2007). We leave an analysis of such feedback effects for future research.

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A Production Efficiency

For generality, productivity growth is specified as $H_{t+1} = B_1 H_t^{\varepsilon(1-\delta)} I_t^\delta$ with $0 < \varepsilon \leq 1$. For $\varepsilon < 1$, the model does not display endogenous growth.

Consider a path with constant ν and p and let y denote output per worker. Conditional on H_t and aggregate labor supply, we have

$$\ln(y_{t+i+1}) \simeq \alpha \ln(k_{t+i+1}) + \delta(1-\alpha) \sum_{j=0}^i (\varepsilon(1-\delta))^j \ln(I_{t+i-j}), \quad i \geq 0.$$

Starting from the investment policy $\{k_{t+i+1}, I_{t+i}\}_{i=0}^\infty$, consider a sequence of small reallocations of investment spending. This sequence involves, in each period i , a small change in public investment of Δ_i and a corresponding change in physical investment of $-\Delta_i$ (per worker in period i). If this policy change weakly increases output in all subsequent periods, then it amounts to a Pareto improvement and the initial allocation is production inefficient. Formally, the conditions for production inefficiency are given by

$$\begin{aligned} d \ln(y_{t+i+1}) &= -\alpha \frac{\Delta_{t+i}}{k_{t+i+1}\nu} + \delta(1-\alpha) \sum_{j=0}^i (\varepsilon(1-\delta))^j \frac{\Delta_{t+i-j}}{I_{t+i-j}} = \\ &= -\alpha \frac{I_{t+i}}{k_{t+i+1}\nu} \epsilon_{t+i} + \delta(1-\alpha) \sum_{j=0}^i (\varepsilon(1-\delta))^j \epsilon_{t+i-j} \geq 0 \quad \text{for all } i \geq 0, \end{aligned}$$

where we define $\epsilon_{t+i} \equiv \Delta_{t+i}/I_{t+i}$, and where at least one inequality must hold strictly. Since the initial allocation corresponds to a balanced growth path, the recurrent term

$$a \equiv -\alpha \frac{I_{t+i}}{k_{t+i+1}\nu} + \delta(1-\alpha)$$

is time-invariant. The conditions for production inefficiency can therefore be summarized as

$$\begin{aligned} a\epsilon_t &\geq 0, \\ a\epsilon_{t+i} + \delta(1-\alpha) \sum_{j=1}^i (\varepsilon(1-\delta))^j \epsilon_{t+i-j} &\geq 0 \quad \text{for all } i \geq 1, \end{aligned}$$

where at least one inequality must hold strictly.

Intuitively, the term a (multiplied by the amount of physical investment) represents the effect of an infinitesimal reallocation from k to H investment on output in the subsequent period. To increase output in period $t+1$, ϵ_t must have the same sign as a . To increase output in periods later than period $t+1$, the combined effect of the lagged changes in k - and H -investment must be positive.

When $a > 0$, capital is over accumulated in the initial allocation. As is apparent from the above conditions, one can generate a Pareto improvement in this case by reallocating

resources from k to H (corresponding to $\epsilon_{t+i} > 0$). Over accumulation of capital is also present if $a = 0$ and $\epsilon(1 - \delta) > 0$, corresponding to the allocation in an economy without government intervention, but with markets for the provision of the “public investment.”³³ In such a setting, savings is allocated across H and k investment in such a way that output in the *subsequent* period cannot be increased. However, if $\epsilon(1 - \delta) > 0$, the level of productivity contributes to future productivity growth, and a slight reallocation from k to H investment therefore increases output in all *later* periods, as is apparent from the above conditions. The allocation satisfying $a = 0$ is not Pareto optimal in this case because it does not properly account for the dynamic productivity externality.

When a is negative and large in absolute value, the allocation again is production inefficient. In this case, a reallocation of resources from H to k accumulation (corresponding to $\epsilon_{t+i} < 0$) generates a Pareto improvement. For example, if $a = -1$, a sequence of $\epsilon_{t+i} = \epsilon < 0$ for all $i \geq 0$ increases production in all future periods because the positive effect from additional physical investment, $a\epsilon = -\epsilon > 0$, dominates the cumulative negative effect from reduced productivity growth, $\delta(1 - \alpha) \sum_{j=1}^i (\epsilon(1 - \delta))^j \epsilon < \epsilon$. To characterize the largest $a < 0$ allowing for a persistent increase in output, we consider a sequence $\{\epsilon_{t+i}^*\}_{i=0}^{\infty}$ with $\epsilon_t^* < 0$ where $\{\epsilon_{t+i}^*\}_{i=1}^{\infty}$ is recursively defined by the requirement that $d \ln(y_{t+i}) = 0$ for all $i \geq 2$. If such a sequence is bounded then production is inefficient. For $i \geq 1$, the terms of such a sequence satisfy $a\epsilon_{t+i}^* + \delta(1 - \alpha) \sum_{j=1}^i (\epsilon(1 - \delta))^j \epsilon_{t+i-j}^* = 0$. This implies

$$\epsilon_{t+1}^* = \frac{\delta(1 - \alpha)}{-a} \epsilon(1 - \delta) \epsilon_t^*$$

and

$$\begin{aligned} \epsilon_{t+i}^* &= \frac{\delta(1 - \alpha)}{-a} \sum_{j=1}^i (\epsilon(1 - \delta))^j \epsilon_{t+i-j}^* \\ &= \frac{\delta(1 - \alpha)}{-a} \epsilon(1 - \delta) \epsilon_{t+i-1}^* + \frac{\delta(1 - \alpha)}{-a} \sum_{j=2}^i (\epsilon(1 - \delta))^j \epsilon_{t+i-j}^* \\ &= \frac{\delta(1 - \alpha)}{-a} \epsilon(1 - \delta) \epsilon_{t+i-1}^* + \frac{\delta(1 - \alpha)}{-a} \epsilon(1 - \delta) \sum_{j=1}^{i-1} (\epsilon(1 - \delta))^j \epsilon_{t+i-j-1}^* \\ &= \frac{\delta(1 - \alpha)}{-a} \epsilon(1 - \delta) \epsilon_{t+i-1}^* + \epsilon(1 - \delta) \epsilon_{t+i-1}^* \\ &= \epsilon(1 - \delta) \left(1 - \frac{\delta(1 - \alpha)}{a} \right) \epsilon_{t+i-1}^*, \quad i > 1. \end{aligned}$$

Boundedness of the sequence and thus, production inefficiency requires $-1 < \epsilon(1 - \delta) \left(1 - \frac{\delta(1 - \alpha)}{a} \right) < 1$ which simplifies (due to $a < 0$) to the condition $a < -\frac{(1 - \alpha)\epsilon\delta(1 - \delta)}{1 - \epsilon(1 - \delta)}$.

In conclusion, if $\epsilon(1 - \delta) > 0$ (such that $a = 0$ is not efficient), the criterion for production efficiency is given by $-\frac{(1 - \alpha)\epsilon\delta(1 - \delta)}{1 - \epsilon(1 - \delta)} \leq a < 0$ for $a \equiv \delta(1 - \alpha) - \alpha \frac{I}{kV}$. If the left

³³Boldrin and Montes (2005) and Docquier, Paddison and Pestieau (2007) characterize such an economy; they interpret public investment as public education.

inequality is violated, then the economy accumulates too much H ; if the right inequality is violated, then the economy accumulates too much k .

In politico-economic equilibrium, $\frac{I_{t+i}}{s_{t+i}} = \frac{\sigma}{z(1-\tau-\sigma)}$ and the production efficiency criterion subject to $\varepsilon = 1$ therefore reduces to

$$1 \geq \frac{\alpha' \sigma}{z(1-\tau-\sigma)} > \delta.$$

From the first-order condition for σ tax rates in politico-economic equilibrium satisfy

$$\frac{\alpha' \sigma}{z(1-\tau-\sigma)} = \frac{\delta \alpha \beta p}{(1+\alpha \beta p)z} = \delta \frac{\alpha(1+\beta p) + (1-\alpha)^{\frac{\tau+\psi-1}{\psi}}}{1+\alpha \beta p} < \delta$$

where the last inequality follows from $\frac{\tau+\psi-1}{\psi} < 1$. We conclude that, in politico-economic equilibrium, the economy necessarily over accumulates k relative to H .