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Assessing Municipal Bond Default Probabilities

Matthew Holian and Marc Joffe

San Jose State University, Public Sector Credit Solutions

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Matthew J. Holian, Ph.D.

Associate Professor of Economics

San Jose State University

1 Washington Square Hall

San Jose, CA 95192-0114

Tel: (408) 457-4367

E-mail: matthew.holian@sjsu.edu

Marc D. Joffe

Principal Consultant

Public Sector Credit Solutions

1955 N. California Blvd.

Walnut Creek, CA 94596

Tel: (415) 578-0558

Email: marc@publicsectorcredit.org

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Abstract

In response to a request from the California Debt and Investment Advisory Commission, we propose a model to estimate default probabilities for bonds issued by cities. The model can be used with financial data available in Comprehensive Annual Financial Reports that cities are required to publish. The study includes modeled default probability estimates for 261 California cities with population over 25,000. Our model relies on case study evidence, logistic regression analysis of major city financial statistics from the Great Depression – the last time a large number of cities defaulted – as well as logistic regression analysis of more recent city financial statistics. Independent variables in our model include (1) the ratio of interest and pension expenses to total revenue, (2) the annual change in total revenue, (3) the ratio of general fund surplus (or deficit) to general fund revenues and (4) the ratio of general fund balance to general fund expenditures.

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Executive Summary

California local agencies have faced substantial fiscal stress in the aftermath of the financial crisis. Several cities have filed for bankruptcy, defaulted on bond payments or declared fiscal emergencies. However, the vast majority of California local bond issuers continue to perform on their obligations.

The fiscal troubles faced by individual governments typically receive substantial publicity. Such news reports reinforce dire predictions from high profile analysts that a municipal market crisis is imminent. As a result, bondholders may be dissuaded from investing in the obligations of all municipal bond issuers – even those that are relatively healthy. This phenomenon threatens to exacerbate municipal bond market illiquidity, which, according to Ang & Green (2011), already costs issuers an extra 1.1% in annual interest.

With the collapse of the municipal bond insurance business and questions concerning the credibility of bond ratings, new methods of credit risk assessment are required. In response to a request by the California Debt and Investment Advisory Commission, we have created an empirically-based methodology for assigning credit scores to municipalities, using quantitative techniques that are resistant to bias. These scores take the form of default probabilities and are based on a modeling procedure often applied to corporate borrowers. Using this methodology, we have assigned default probability scores to over 260 California cities with population over 25,000 that have filed 2011 or 2012 Comprehensive Annual Financial Reports.

In bond market terms, a default is usually defined as the failure on the part of an issuer to pay principal and/or interest in full and on a timely basis. It is this definition of

default that we use in this study. This means we do not consider the concept of a technical default which often relates to the failure of an issuer to carry out other obligations under the bond agreement, such as the prompt filing of continuing disclosures. Further, we do not consider failure to pay contractors, employees, retirees or beneficiaries promised sums as defaults for the current purpose – the concept narrowly applies to bondholders.

Since our model applies to cities themselves, it does not consider the specific attributes of their individual bond issues. Thus, general obligation bonds issued by a city should be expected to have less risk than our estimates suggest, while certificates of participation and other securities not explicitly backed by a diverse stream of tax revenues may be more risky.

Our model of municipal default risk is based on four fiscal indicators. These are: (1) the ratio of interest and pension expenses to total governmental fund revenue, (2) the annual change in total governmental fund revenue, (3) the ratio of the city's general fund surplus or deficit to its general fund revenue and (4) the ratio of the city's end of year general fund balance to its general fund expenditures. In the study, we provide statistical and case study evidence to support the choice of these variables and appropriate coefficients.

We hope that the proposed model will help municipal bond investors and other stakeholders in city government solvency better comprehend the risks faced by municipal issuers. We also hope that other researchers and practitioners will become interested in this topic and our analytical approach, so that they will build upon and improve our findings.

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Matthew Holian and Marc Joffe

April 30, 2011

Introduction

The issue of municipal solvency has frequently made the headlines in recent years. Meredith Whitney's 2010 appearance on *60 Minutes* was but one of a number of dire predictions for municipal bondholders. In 2012, the bankruptcies of Stockton and San Bernardino attracted significant media attention, as has the state takeover of Detroit, Michigan earlier this year.

Unfortunately for bondholders and the many other stakeholders in city solvency, the debate about municipal credit has been often generated more heat than light. Whitney's analysis fed into a narrative about skyrocketing public employee pension costs triggering a tsunami of municipal bankruptcies.

These politically charged predictions have yet to be borne out by the facts on the ground. In the *60 Minutes* interview, Whitney predicted 50-100 sizable defaults (CBS News, 2010). She later stated that this would be a "something to worry about" within 12 months of her appearance which occurred in December 2010. When it became apparent that this dire forecast was failing to materialize, Michael Lewis (2011) wrote an influential piece in *Vanity Fair* quoting Whitney as saying "who cares about the stinking muni-bond market?" and attempting to rehabilitate her by turning the reader's attention to fiscal problems in California cities, public employee pensions and the risk of "cultural" as opposed to financial bankruptcy.

For those who do care about the "stinking" municipal bond market, the discussion left much to be desired. Investors are still wondering how much risk they actually shoulder when purchasing municipal bonds issued by California cities and how much extra interest they should expect to receive in compensation for taking on this risk. The question of the

appropriate interest rate resonates far beyond the municipal bond market, since it directly affects municipal debt service costs, which in turn impact tax rates, service levels and cities' ability to add infrastructure by borrowing.

As we discuss in this study, defaults by cities have been quite rare since the Great Depression. Doty (2012) estimates that annual default rates on general obligation bonds have been consistently below 0.1% in recent decades. Indeed, a researcher is compelled to unearth 80-year old data just to obtain a statistically meaningful sample of general obligation bond defaults on the part of US cities. Even when this dark period in the history of municipal finance is investigated, we find the defaults were often the result of idiosyncratic factors that don't portend ill for modern investors. Finally, pension underfunding is not a new phenomenon: as Munell (2012) documents, it was also a serious concern in the 1970s – a period that witnessed some highly publicized city financial emergencies, but no spate of municipal bond defaults.

All that said, defaults have occurred and will continue to occur, perhaps with somewhat greater frequency than they have in recent decades. Clearly, some cities are more at risk than others, and so stakeholders would benefit from objective, widely available measures of municipal credit risk.

While credit rating agencies have the potential to better inform the public's understanding of municipal credit risk, they face several barriers in doing so. First, since they rely primarily on bond issuers for their revenue, they have limited incentive to evaluate cities they are not paid to rate. Second, much of their investor-oriented research is sold as premium content and thus cannot be freely distributed. Third, rating agencies have lost credibility in the aftermath of the 2008 financial crisis. And, finally, the three major

rating agencies were sued by the Connecticut attorney general – also in 2008 – for assigning overly harsh ratings to municipal bond issuers relative to corporate and structured finance issuers. Two of the three agencies recalibrated their municipal ratings in response to the suit, which was settled in 2011 with no admission of responsibility but the extension of credits to the state of Connecticut for future ratings services.

We believe that the informational vacuum created by the rating agency problem can be filled by academic research. This study represents our initial contribution to this academic project, and we hope that it will motivate others to add their insights. Our approach involves the use of statistical and case study analysis to create a municipal bond default probability model targeted at California cities with population greater than 25,000.

The discussion proceeds as follows. First, we provide a literature review which discusses previous efforts to model municipal credit quality. We find that most of the literature uses ratings or bond yields as a proxy for credit risk, and offer objections to these approaches.

Next, we review the Depression-era municipal bond default experience and propose a simple logit model based on a set of data collected from this period. This analysis identifies two significant variables intuitively related to default risk: the ratio of interest to revenue and the change in annual revenue.

After this, we provide a comprehensive review of California city bond defaults and bankruptcies with case studies of the most recent payment difficulties. The case study evidence suggests that exhaustion of the general fund – an element that is not available in the Great Depression data set – has been a major driver in recent bankruptcy filings and attendant defaults.

We then survey post-Depression defaults in US cities outside California. As part of this discussion, we see how New York and Cleveland – both of which defaulted in the 1970s – ranked against peer cities with respect to variables of interest. We also provide information that supplies much needed context to popular media reporting about municipal bond distress. Specifically, we find that bankruptcy does not necessarily involve default (and *vice versa*) and that most bankruptcies have occurred in small towns, many of which did not have significant volumes of outstanding municipal bonds (if any).

Finally, we propose a second model based on more recent data that enables us to test the relative significance of general fund surplus/deficit and general fund balance variables, which we hypothesize to be default drivers in the modern context. We then hybridize this model with the Depression-era model proposed earlier to create a final specification.

Given the relative dearth of municipal bond defaults, we found a necessity to depart from orthodox modeling techniques in creating our specification. To further aid future research, we also provide an appendix that analyzes the Depression-era data using academic best practices.

A second appendix investigates the possibility of using data from the California State Controller's Office Cities Annual Report as a basis for municipal bond default probability estimation and the final appendix provides our default probability scores for 261 California cities with population greater than 25,000.

The study is supplemented by a web site that contains supporting data and analytics. This web site will become publicly available on or about May 8, 2013. The

address for this site is <http://www.publicsectorcredit.org/ca>. An Excel workbook providing an implementation of our model is available from the authors upon request.

Literature Review

Previous Depression-Era Municipal Default Research

Dr. George Hempel's contribution to our understanding of Depression-era municipal defaults is widely regarded in the municipal bond industry. Aside from his most commonly cited study, *The Post War Quality of State and Local Debt* (1971), some of Hempel's other work is relevant. Particularly noteworthy was his contribution to a 1973 study published by the now-defunct US Advisory Commission on Intergovernmental Relations (ACIR). This work contains a wealth of statistics as well as detailed case studies of eight high profile defaults from the Depression era.

In addition to default counts and descriptive material, Hempel also presented an econometric default model in his 1971 NBER study. Unfortunately, the model was based on data from only 24 municipal issuers in the State of Michigan, 17 of which defaulted. This sample has three shortcomings: small overall size, geographic distribution not representative of the nation as a whole and an in-sample default rate inconsistent with population default rates. Contemporaneous estimates published in *The Bond Buyer* (1938) indicate that there were about 30,000 municipal issuers in the 1930s. The approximate default count of 4800 issuers in that decade implies a population default rate of 16%. This contrasts to a rate of 71% in Hempel's sample.

Hempel collected 11 independent variables for the sample issuers. These were:

- Population
- Dollar Amount of Notes Outstanding
- Dollar Amount of Debt Outstanding
- Per Capita Debt
- Total Assessed Property Values
- Dollar Amount of Taxes Levied

- Tax Levy Per \$1,000 Assessed Value
- Debt / Assessed Property Values
- Percentage of Current Taxes Delinquent
- Tax Levy Per Capita
- Assessed Property Values Per Capita

This set of variables captures many of the factors theorized to cause municipal bond defaults including size of the issuer, debt burden as well as the willingness and ability of local government and the citizenry to generate required tax revenue. No variables capture other costs that municipal leaders might choose to pay instead of debt service – such as municipal employee salaries or pensions. Also, some of Hempel’s variables are derived from others, introducing a risk of multicollinearity. For example, Per Capita Debt is the quotient of Dollar Amount of Debt Outstanding and Population.

After collecting the data, Hempel subjected it to factor analysis, multiple discriminant analysis and multiple regression analysis. He reported a multiple regression equation that contained 8 of the 11 variables, which were significant at $p < .1$. While the overall regression had an r^2 of 64%, a number of the variables had signs inconsistent with theory, perhaps due to multicollinearity. Hempel addressed multicollinearity by further reducing the set of independent variables to the following four (shown here with their coefficients and standard errors):

Hempel’s (1971) Municipal Default Model

Variable	Coefficient	Standard Error
Tax Levy per \$1,000 Assessed Value	-0.00310	0.00247
Tax Levy per Capita	-0.00115	0.00108
Debt / Assessed Property Values	+0.3521	0.17000
Percentage of Current Taxes Delinquent	+0.07209	0.07277

Hempel does not report any goodness of fit measures for the overall equation, but notes that it had a higher r^2 than other alternatives he evaluated, and that all variables have

the expected sign. On the other hand, two of the four variables are not significant at $p < .05$, while the two best predictors are theoretically related.

In the interest of using Depression-era data to predict future defaults, it is fortunate that certain variables fell out of Hempel’s specification. Given the substantial change in prices and wealth since the 1930s, it would be difficult to use the Dollar Value of Notes Outstanding, the Dollar Value of Debt Outstanding or Per Capita Debt to model current issuers. Tax Levy per Capita, which remained in Hempel’s specification, has a similar challenge. Variables that take the form of ratios, such as Debt/Assessed Property Values or Tax Levy per \$1000 Assessed Value are more appropriate for analysis and forecasting independent of time period.

Hempel (1973) later expanded the sample to 45 Michigan cities – 28 of which defaulted – and 23 independent variables. Many of the added variables were 1922 values most likely obtained from that year’s Census of State and Local Governments. He identified a regression equation with nine exogenous variables significant at $p < .05$.

Hempel's (1973) Municipal Default Model

Variable	Coefficient	Standard Error
Log of 1932 Population	-0.07678	0.0321
Assessed Property Value Per Capita in 1932	+0.0001585	0.0000523
Growth of Population from 1922 to 1932	-0.02146	0.0113
Growth of Debt Relative to Population Growth	-0.007912	0.00213
Debt/Assessed Property Values in 1932	+0.4885	0.258
Tax Levy Per \$1000 Assessed Value in 1932	+0.00919	0.00242
Tax Levy Per Capita in 1932	-0.007197	0.00322
Percentage of Current Taxes Delinquent in 1932	+0.2095	0.0962
Notes Outstanding Per Capita in 1932	+0.009159	0.00246

Hempel noted the presence of multicollinearity but did not present an alternative equation that addressed it. Two of the nine variables presented above – Growth of Debt

Relative to Population Growth and Tax Levy Per Capita in 1932 – have coefficient signs that are inconsistent with intuition. Hempel reported that the nine variable regression had an adjusted r^2 of 51%, while alternatives that remedied multicollinearity had adjusted r^2 of between 39% and 45%.

In his discussion of Hempel's findings, Forbes (1973) questions the use of Depression-era data for modeling purposes, while admitting that the paucity of more recent defaults forces this choice. In particular he noted that local governments received more state aid – at the time of his writing – than they did in the 1930s. This institutional change could reduce the relevance of the historic default data.

Predicting Credit Ratings as a Proxy for Estimating Default Risk

Rubinfeld (1973) proposed a multiple regression model for predicting credit ratings. Since credit ratings are intended to convey information about the likelihood of default, exogenous variables that explain credit ratings could also be used as predictors of default. Using a sample of 128 New England municipal bond issuers, he found that the following independent variables were predictive of the credit rating at the 10% significance level:

- Percentage of Taxes Uncollected in the Previous Year
- Ratio of Direct Net Debt to Assessed Valuation
- Median Family Income
- Full Valuation of the Property Tax Base
- Overlapping Debt

The first two of these exogenous variables are consistent with those in Hempel's 1973 study. Overlapping Debt refers to the indebtedness of other issuers who rely on the same tax base. For example, if property owners pay taxes to both their city and county and if both governmental entities carry debt, the county's debt would be considered overlapping debt

viz.-a-viz. the city and *vice versa*. This variable, along with Median Family Income and Full Valuation of the Property Tax Base, would have to be restated as a ratio to be useful in a default prediction model.

Carelton & Lerner (1969) attempted to use statistical techniques to match Moody's bond ratings using a random sampling of issuers extracted from Moody's 1967 Municipal and Government Bond manual. They tested six variables – all of which they found to be significant. These were:

- Whether the issuer was a School District
- Ratio of Debt to Assessed Valuation
- Ratio of Debt to Population
- Log of Population
- Log of Debt
- Average Collection Rate

Using a large sample of 976 cities, Farnham & Cluff (1984) tested 23 variables to determine whether they were predictive of Moody's bond ratings. They found 12 of the variables to be significant $\alpha = .05$. The method used was an "N-chotomous" probit analysis. The authors chose this method because the four possible ratings in the dependent variable were thought to be of unequal lengths, i.e. many more cities fell into the A rating category than into the Aaa category. Their analysis included several variables not considered by other authors – including housing stock attributes, form of government and geographical location. Four of the housing stock attributes proved to be significant. Farnham & Cluff's variables are listed in the following table.

Farnham & Cluff's (1984) Independent Variables

Variable	Significant at 5% Level?
Gross Debt / 1000 Population	*
Total General Revenue	*
Percent Change in Total Revenue	*
Assessed Valuation	*
Population	
Percent Change in Population	
Percent Nonwhite	*
Percent Eighteen Years and Under	
Population Density	*
Income Per Capita	
Ratio of Non-Workers to Workers	*
Number of Manufacturing Establishments	
Percent One-Unit Housing Structures	*
Percent Housing Units Occupied	
Percent Housing Units Owner Occupied	*
Percent Housing Units Built Before 1940 (as of 1970)	*
Median Value of Owner Occupied Housing Units	*
Median Years of Education	*
Local Documents Available	
Council-Manager Form of Government	
City Located in Northeast Region	
City Located in Northcentral Region	
City Located in South	

The papers reviewed above are part of a large literature that attempts to estimate municipal bond ratings. Loviscek & Crowley (1990) compare the studies described here with eleven others that had the same objective.

Since Loviscek & Crowley published their review, at least two additional papers modeling municipal bond ratings have appeared. Moon & Stotsky (1993) analyzed data for 892 US cities with population over 25,000, of which 727 were rated. They first modeled the decision by city officials to seek a rating and then factors determining the ratings actually

assigned. This methodology highlights the fact that by choosing to be rated, cities self-select into the samples used in previous studies. This suggests that studies which use ratings as a proxy for default probability suffer from selection bias.

Moon & Stotsky (1993) found that cities choosing to remain unrated were likely to receive a low rating. They tested twenty variables potentially affecting rating levels, and found 15 to be significant. The variables they evaluated were as follows:

Moon & Stotsky's (1993) Independent Variables

Variable	Significant at 5% Level?
Median Housing Value	*
Proportion of Housing Units that were Built Before 1940	*
Proportion of Housing Units that were Built After 1970	
Proportion of Housing Units that are Owner-Occupied	*
Per Capita income	*
Percentage Change in Population from 1970 to 1980	*
Proportion of the Population that is Non-White	*
Population Density	*
Total Debt	
Per Capita Debt	*
Ratio of Debt to Income	*
Ratio of Surplus Revenues to General Revenues	
Ratio of Intergovernmental Revenues to General Revenues	
Council-Manager form of government	*
Commission Form of Government	
City Located in Midwest	*
City Located in South	*
City Located in West	*
Population Between 100,000 and 500,000	*
Population Greater Than 500,000	*

Most recently, Palumbo & Zaporowski (2012) analyzed ratings for 965 county and city governments rated by Moody's in 2002. This population encompassed all such units that issued rated full faith and credit debt and that could be matched against Census, Bureau of

Economic Analysis (BEA) and Bureau of Labor Statistics (BLS) data sets. Of the 15 variables they examined, 13 proved to be significant at the 5% level as shown below.

Palumbo & Zaporowski's (2012) Variables

Variable	Significant at 5% Level?
Per Capita Income	*
Percentage Change in Population 1990-2000	*
Unemployment Rate	*
Percentage Change in Earnings Per Worker 1986-2001	*
Economic Diversity Index from BEA	*
State Aid Per Capita	*
State General Obligation Bond Rating	*
Debt to Market Value (Ratio of Full Faith and Credit Debt to Population Weighted Median Value of Housing)	*
Non-Guaranteed Debt Per Capita	
Per Capita Interest Payments for Nonutility Debt	
Per Capita General Revenues	*
State Imposed Taxation Limit	*
State Imposed Expenditure Limit	*

Objections to Rating Based Analysis

Researchers who model ratings rather than defaults, implicitly assume that the former predict the latter.¹ However, if ratings do not change in response to underlying credit conditions experienced by municipal bond issuers, they may not be an effective proxy for default risk. Under SEC rules, rating agencies are required to publish transition matrices showing the distribution of rating changes over a given period. A review of the transition matrices published by Moody's Corporation (2012), Standard & Poors Corporation (2012a)

¹ In fairness to the authors of these studies, it is worth pointing out that most do not make the claim that ratings proxy default probability. When modeling credit ratings, researchers may have goals other than estimating default probability. For example, they may be interested in modeling rating agency behavior.

and Fitch, Inc. (2012) suggests that about 90% of municipal bond ratings remain unchanged within a given year.

For example, an S&P transition matrix (for non-housing municipal issuers) shows that 89.11% of AA rated issuers remained AA the following year, while 0.18% were upgraded to AAA, 1.62% were upgraded to AA+ and a total of 9.09% were downgraded to various rating categories ranging from AA- down to BB+.

The S&P matrix represents all rating change activity that occurred between 1986 and 2011. During most of this period, a substantial proportion of municipal bond ratings reflected insurance “enhancements”. So-called monoline insurers like Ambac, FGIC and MBIA – which were rated AAA – sold bond insurance policies to municipalities guaranteeing that any missed bond payments would be covered by the insurer. Consequently, the ratings assigned to these insured issuers were AAA – reflecting the estimated credit quality of the insurer. During the 2007-2008 financial crisis, all monoline bond insurers went out of business or suffered ratings downgrades (Palumbo & Zaporowski, 2012).

While the insurance was in place, ratings might have appeared to remain stable despite changes in municipal credit conditions, simply due to the stability of the insurer’s credit rating. However, Fitch’s NRSRO ratings transition exhibit states that the ratings analyzed are “unenanced” which means they reflect the underlying credit quality of the issuer excluding any insurance benefit. We expect that that is also the case for the S&P and Moody’s tables.

Insurance coverage aside, municipal ratings stability could be explained by some combination of three factors. First, underlying credit conditions for most issuers do not materially change from year to year. Second, ratings grades are too coarse to capture many

credit quality changes. And, third, rating agencies do not perform sufficient surveillance activities to detect and respond to many changes in issuer credit quality. To the extent that the second and third causes are explanatory, they pose challenges to the use of ratings as a proxy for default probability.

Little evidence is available to determine the relative weight of each of these three factors. One item that may be relevant is the criticism rating agencies received for their inadequate monitoring of Residential Mortgage Backed Securities (RMBS) and Collateralized Debt Obligations (CDO) prior to the financial crisis of 2007 and 2008. The United States Senate Permanent Subcommittee on Investigations (2011) found that:

Resource shortages impacted the ability of the credit rating agencies to conduct surveillance on outstanding rated RMBS and CDO securities to evaluate their credit risk. The credit rating agencies were contractually obligated to monitor the accuracy of the ratings they issued over the life of the rated transactions. CRA surveillance analysts were supposed to evaluate each rating on an ongoing basis to determine whether the rating should be affirmed, upgraded, or downgraded. To support this analysis, both companies collected substantial annual surveillance fees from the issuers of the financial instruments they rated, and set up surveillance groups to review the ratings. In the case of RMBS and CDO securities, the Subcommittee investigation found evidence that these surveillance groups may have lacked the resources to properly monitor the thousands of rated products. At Moody's, for example, a 2007 email message disclosed that about 26 surveillance analysts were responsible for tracking over 13,000 rated CDO securities. (p. 314).

Since these findings relate to structured securities rather than municipal bonds, it is possible that they are not relevant. On the other hand, it is reasonable to think that if rating companies under-invested in surveillance for their most profitable asset class – structured finance - (Cornaggia, Cornaggia & Hund, 2011), they probably made similar under-

investments in the surveillance of other asset classes. It is the co-author's contention, - based on his experiences at a major rating agency - that surveillance procedures for structured assets were actually superior to those undertaken for municipal bonds.

Estimating Default Probability from Market Prices

A number of researchers have attempted to derive default probabilities from bond yields or Credit Default Swap (CDS) spreads (Longstaff, Mithal & Neis, 2004). In theory, bond yields should be a function of their credit risk. More specifically, yields should compensate investors for the expected loss arising from a potential default. In the literature, expected loss is defined as the product of default probability and loss given default (LGD). LGD is simply the complement of a bond's rate of recovery, and is also called "loss severity".

Theoretical bond yields contain a number of components aside from expected loss. Bohn, Arora and Agarwal (2004) propose an equation for corporate bond yields that includes the risk free rate of interest, the level of investor aversion to risk, the bond's maturity date, issuer size (as a proxy for liquidity) and the correlation of the bond's default risk with that of other instruments. Yields may also be affected by call provisions that give issuers the option to redeem their bonds prior to maturity.

With respect to municipal bonds, a further complexity arises as a result of their tax status. Since interest on most municipal bonds is exempt from federal, state and local income taxation, their yields are not comparable to those on taxable securities. Some adjustment to the municipal bond yield must be made in order to make it "taxable equivalent". One approach is to convert the tax free yield to a taxable yield based on the highest prevailing marginal tax rate, on the assumption that municipal investors are predominantly high income individuals. However, given the complexities of the tax code, the heterogeneity of

individual investors and the participation of institutional investors (with different tax considerations), the use of the top marginal rate is a relatively strong assumption. Chalmers (1998) finds that interest rate differentials between long term US Treasuries and federally insured municipals (which are assumed to have no default risk) are not consistent with the tax benefits available to individuals in the top tax bracket.

The literature includes a number of efforts to decompose municipal bond yields into default risk and other components. Wu (1991) found that the risk aversion factor was not significant, but his functional form excluded recovery rates. Wu, Wang & Zhang (2006) offered a more comprehensive model that included a static recovery rate assumption. The authors attributed a substantial portion of municipal bond yields to liquidity factors.

In corporate credit markets, analysts often derive default probabilities from CDS spreads rather than bond yields. Credit Default Swaps are insurance contracts against default. If the issuer defaults, the CDS seller (or insurer) pays the protection buyer the face value of the bond and takes the bond in exchange. Deriving default probabilities from CDS spreads is easier than using bond yields because CDS have fewer complexities, such as call provisions. The applicability of CDS implied default probabilities to the municipal market is greatly limited, however, by the fact that CDS trades against a relatively small number of municipal issuers, and trading volume is low even for those issuers for which CDS are available.

A final concern regarding market implied default probabilities pertains to how efficiently markets price credit risk. Decomposing yields into default probabilities and other components implicitly assumes that bond prices are efficient, i.e. that they accurately reflect all available information. This assumption is consistent with the strong form of the Efficient

Markets Hypothesis (EMH) markets advanced by Fama (1970). More recently EMH generally, and the strong form of the hypothesis in particular, have come under attack (Summers, 1986; Crotty, 2011). Most tests of EMH have involved equities rather than bonds. In a 2003 survey of EMH literature, Malkiel (2003) identified only one study addressing bond market efficiency, and that paper found inefficiency in the pricing of corporate bonds (Keim & Stambaugh, 1986). Since large capitalization stocks experience much higher trading volumes than municipal bonds, it is not clear that EMH applies at all to the latter asset class. Indeed, there is a substantial literature documenting the lack of liquidity and transparency in the municipal bond market – suggesting the existence of substantial inefficiencies (Ang & Greene, 2011).

In summary, the task of deriving default probabilities from municipal bond yields is impeded by both the complexities of decomposing yields into their components and the likelihood that observed yields do not efficiently incorporate credit risk insight.

Default Probability Modeling Using Logit and Probit Techniques

More recent efforts to model bond default probabilities have used logit and probit techniques. An obvious advantage of logit and probit over Ordinary Least Squares (OLS) for default probability modeling is that the dependent variable is restricted to a range of 0 to 1. In addition, the use of a binary endogenous variable, like default/non-default, violates a number of assumptions of the OLS model. (Menard, 2002).

Because corporate bankruptcy has been much more common than municipal default, the academic literature contains many more efforts to model the former. Ohlson (1980) was first to apply a logit model to corporate bankruptcy modeling.

Shumway (2001) built upon previous logit models by using panel rather than cross sectional data. This approach addresses the fact that most bankrupt firms were solvent for many years before going into distress, and that it is thus useful to analyze a time series of data for each firm.

The literature also contains applications of probit models to corporate bankruptcy starting with Zmijewski (1984). Moody's RiskCalc is a commercial private firm default probability model that uses probit. The RiskCalc methodology document written by Falkenstein, Boral & Carty (2000) suggests that the choice of probit over logit was not a significant one, as the two models usually produce similar results. On the other hand, Altman & Sabato (2007) assert that logit models have outperformed probit models in the corporate bankruptcy field.

Probit and logit models are functionally similar, with the key difference being the fact that probit is based on a cumulative normal probability density function, whereas logit uses a logarithmic distribution. This latter distribution has more observations in its left and right tails and fewer observations at its center. Ameniya (1980), in his extensive survey of binary choice and other discrete choice models concludes that "it does not matter much whether one uses a probit model or a logit model, except in cases where data are heavily concentrated in the tails due to the characteristics of the problem being studied (p. 1487)."

Although the published literature does not appear to include general obligation municipal bond default probability models that employ logit and probit techniques, Bialaszewski (1985) applied a logit model to a set of municipal revenue bonds – issues which are supported by user fees and other operating revenues collected by the issuing agency rather than with tax revenues. Bialaszewski collected financial, economic and demographic

data for 36 defaulted revenue bonds and for 36 comparable bonds that did not default. She then created models using data at issuance, two years prior to default, one year prior to default and at the time of default. Different variables were significant in each model. She reported that her one year prior to default model accurately classified 87% of the observations into defaulting and non-defaulting categories, where these categories were defined in terms of a “cut point” in the calculated probabilities. Her cut point of 65.8% was set to produce the highest degree of accurate classification. It may be more appropriate to use a fixed cut point of 50%, since probability estimates over that level could be reasonably characterized as default predictions, while probabilities under this level could be seen as predictions of non-default. The significant variables in Bialaszewski’s regression were: (1) Total Population, (2) Percentage of Population that is Non-White, (3) Debt Service as a Percentage of Total Revenue, (4) Welfare Payments as a Percentage of Total Revenue and (5) Short Term Debt as a Percentage of Cash and Security Holdings.

Since the observations involved revenue bonds, the theoretical case for some of the variables in this specification is not immediately apparent. For example, welfare payments are financed by a municipality’s general fund, and should thus not be expected to compete with revenue bondholders for priority. On the other hand, non-white population and welfare dependency levels may be indicators of poverty. Impoverished residents may be less able to pay fees required to service debt incurred by the facilities that defaulted.

Finally, the use of race-based criteria for evaluating municipal bonds has been subject to criticism. Yinger (2010) finds that general obligation municipal bond ratings penalize communities with relatively high non-white populations despite the lack of evidence that

these communities are more likely to default. He characterizes this result as a form of redlining and argues for municipal bond rating regulation to curtail this practice.

Great Depression Review and Analysis²

Since 1940, interest or principal payment defaults on US municipal bonds have been rare. This is especially true of general obligations bonds – those backed by the full faith and credit of a state, county, city or other governmental unit with taxing authority. By contrast, there were about 4800 reported municipal bond defaults during the 1930s (ACIR 1973; Fons, Randazzo & Joffe, 2011).

With the assistance of colleagues and a data entry vendor, Joffe (2012) collected information on approximately 5000 defaults from the period 1920 to 1939. The primary sources were contemporary *Moody's Municipal and Government Bond Manuals* (now published and owned by Mergent Corporation), and back issues of the *Daily Bond Buyer* and weekly *Bond Buyer*. Joffe (2012) also found and catalogued defaults from state-level bond listings and other documents housed in state archives, Reconstruction Finance Corporation records, local newspaper accounts and other sources.

In their book, *This Time is Different*, Carmen Reinhart and Kenneth Rogoff (2009) marshaled older data in their analysis of banking and sovereign debt crises. Due to the paucity of recent defaults, a similar approach may be applicable to US municipal bonds. In contrast to some areas of fixed income - such as mortgage backed securities - institutional change in the municipal sphere over the last century has been incremental rather than revolutionary. Political and budgetary processes at the state and local level have evolved relatively slowly in the context of a stable national political framework. Older municipal

² This section contains previously published research that originally appeared in Fons, Randazzo and Joffe (2011) and Joffe (2012). However the statistical analysis presented below has been updated for this report.

defaults are thus more relevant to modern experience than older defaults in other asset classes.

The goal of this chapter is to mine the Depression-era municipal bond default record to learn whatever insights it can offer for present day credit research. This is done by providing a brief description of the 1930s municipal credit crisis and by developing a quantitative default probability model.

The Depression Era Municipal Default Wave

According to US Treasury statistics reported by *The Bond Buyer*, the dollar volume of municipal bonds outstanding more than quadrupled between 1913 and 1931 – a period during which the CPI rose 54%. The boom in municipal issuance during this period is largely attributable to the inception of the federal income tax and the popularization of automobile travel. Municipal bond interest was exempt from income taxes since the levy's 1913 inception, creating demand for these securities among high income investors. On the supply side, automobiles created a need for paved roads – which states, counties and cities often financed with bonds. Communities also used bonds to finance drainage, irrigation and levee projects to support agricultural developments and to fund school construction.

Those concerned about today's municipal credit quality correctly point to the rapid growth in municipal bonds outstanding in recent years. But the growth in municipal bonds outstanding between 1913 and 1931 far exceeded the rate of increase over the eighteen years up to 2010 – and both of these booms were outpaced by growth following World War II, during the years 1946 to 1964. While the pre-Depression municipal bond boom ended with a spike in defaults, the post-War expansion was not followed by a similar circumstance.

Bank Closings, Bank Holidays and Municipal Bond Defaults

It is also worth considering that the peak in estimated municipal default rates coincided with a nationwide outbreak of bank failures and bank holidays. In a 1933 survey of 1,241 state, city and county financial officials, Martin Faust (1934, 1936) found that slightly more than half of their governmental units had funds in closed banks. The municipalities surveyed had a total of over \$98 million tied up in these failed institutions. Faust estimates that the aggregate balance in failed banks for all state and local governments would have been \$450 million – more than 2% of the principal outstanding on municipal bonds at the time. Contemporary accounts attributed many of the defaults to the closure of banks in which funds intended for bondholders had been deposited.

A major source of distress for municipalities in North Carolina, Louisiana, Arkansas, Tennessee and other southern states was the November 1930 collapse of Caldwell & Company and its affiliates. Founder Rogers Caldwell, dubbed the “J.P. Morgan of the South” had built a large business marketing municipal bonds issued by southern states. Bond proceeds were typically held at Caldwell’s Bank of Tennessee until they were required by the issuer. According to John McFerrin’s (1939) history of Caldwell and Company, most issuers required that their deposits be supported by high quality collateral – typically other municipal bonds. Caldwell often pledged such bonds as collateral initially, and then substituted illiquid, high-risk real estate bonds without notifying the issuer. In addition to following deceptive practices, Caldwell looted bank assets to finance an extravagant lifestyle.

On November 7, 1930, a Tennessee state audit declared Caldwell & Company insolvent. News of this declaration triggered runs on Caldwell and numerous affiliated banks throughout the South. In Tennessee alone, \$9 million in county and municipal deposits were

lost. Caldwell's failure triggered a run on affiliates, including Central Bank and Trust Company in Asheville, North Carolina, which was followed by runs on other area banks.

Property Tax Delinquencies

While the vast majority of the enumerated defaults occurred in special districts, school districts and small towns, the Depression-era did witness several spectacular defaults by large issuers including Cleveland and Detroit. New York City, the nation's largest municipality back then, also experienced a brief default in December 1933. Chicago, then the nation's second largest city, narrowly avoided default by refinancing its bonds at lower interest rates. Cook County – which encompasses the city – failed to make scheduled interest and principal payments, as did a number of independent taxing districts within the city's limits.

As statistics collected at the time by Dun & Bradstreet (Bird, 1936) suggest, major city defaults during the Great Depression were preceded by substantial spikes in tax delinquency rates. For example, the tax delinquency rate in Detroit rose from 10.8% in fiscal year 1930 to 17.2% in 1931, 25.0% in 1932 and 34.8% in 1933 – the year in which it defaulted. In New York and Chicago, delinquency rates peaked at 26.5% and 42.4% respectively.

Although many of the property tax delinquencies were undoubtedly the result of economic distress, the early 1930s was also a period of organized tax revolts. This long-forgotten tax resistance movement is described in David Beito's 1989 book *Taxpayers in Revolt*. Beito argues that the resistance was in large measure a reaction to substantial increases in property taxes during the preceding decade. This increased burden was often accompanied by stable or falling property values, since the 1920s was a time of weak real estate prices.

Beito traces the history of the property tax resistance movement in Chicago where anti-tax activism was most potent. The Chicago resistance was led by the Association of Real Estate Taxpayers (ARET), an organization originally formed by relatively affluent investors, but which later attracted broad support among the City’s skilled blue collar workers worried about maintaining their foothold in the middle class. At its peak, ARET leaders hosted a thrice-weekly radio program and the organization had 30,000 members.

As the following table indicates, large cities were especially vulnerable to property tax delinquencies due to their heavy reliance on real estate taxes. According to 1931 Census data on which this table is based, the average city received about two thirds of its revenue from this one source.

Share of Total Revenues from Property Taxes, Cities Over 300,000, 1931

City	Property Tax Revenue Share
Baltimore, MD	66.8%
Boston, MA	69.3%
Buffalo, NY	69.6%
Chicago, IL	67.2%
Cincinnati, OH	58.9%
Cleveland, OH	67.4%
Detroit, MI	61.2%
Houston, TX	74.6%
Indianapolis, IN	86.8%
Jersey City, NJ	70.9%
Kansas City, MO	64.2%
Los Angeles, CA	52.5%
Louisville, KY	69.3%
Milwaukee, WI	67.1%
Minneapolis, MN	72.3%
New Orleans, LA	61.8%

City	Property Tax Revenue Share
New York, NY	70.8%
Newark, NJ	69.1%
Philadelphia, PA	71.8%
Pittsburgh, PA	81.2%
Portland, OR	65.7%
Rochester, NY	66.5%
San Francisco, CA	59.5%
Seattle, WA	47.5%
St. Louis, MO	62.5%
Washington, DC	56.1%

While over-reliance on one revenue source can be attributed to the relative lack of municipal finance sophistication at the time, part of the problem was beyond the control of city governments. According to Census statistics reported by C. E. Rightor (1938) in *Municipal Finance*, roughly 4-1/2% of major city revenue was derived from alcohol taxation in 1916. This revenue source disappeared with Prohibition, and did not return until the 18th Amendment was repealed in 1933. Additional policing costs associated with Prohibition-related organized crime must have further contributed to the cities' fiscal distress.

Public Employee Pensions

Contemporary concerns about municipal bond defaults are often linked to public pensions, but underfunding is not unique to our era. During the Great Depression, many retired government workers were eligible for pensions. Buck (1936) notes that before the establishment of pensions, older municipal employees would continue to report for work even though they could no longer perform their jobs (at least not to the satisfaction of contemporary management). Supervisors, guided by a humanitarian impulse rather than a

concern for the bottom line, were reluctant to fire these older employees. Administrators thus reached the conclusion that it would be less expensive to pension off the older workers at a percentage of their former salary.

Many cities had not yet created pension funds and those that did often failed to make actuarially appropriate contributions. A 1937 National Municipal League Consulting Service survey of Atlanta's finances reported serious underfunding in the city's three pension funds: .

It is obvious from these figures that the firemen's fund with a cash balance of \$491.38 is no fund at all. Nor are the reserves of either the general or police funds even a faint approximation of what they should be to guarantee the payment from the fund of its probable obligations. ... Firemen this year who paid money into their pension fund saw it go out again immediately to pay other firemen's pensions. Their sacrifice in no way built up for them any protection. They have in fact nothing to rely on but the naked promise of the city as their security for old age. We would recommend therefore that in all the pension funds the employee's contribution be treated as a trust fund and invested for him in securities or in the purchase of an annuity.

That said, the NML consultants were not advocates of full funding:

We believe on the other hand that it is not necessary for a public body deriving its income from taxes to accumulate a fund as if it were a private insurance company. Unless there are some predictable sharp upturns in the curve of natural retirement there is no reason why the City should not pay pensions out of income. The integrity and solvency of the city should be a sufficient guarantee to the employee that the city will fulfill its pension contract. In fact if the city went bankrupt any fund it might have accumulated would probably disappear in the crash.

Atlanta public employee pensions at the time were generous – at least by the standards of today's private sector. Employees could retire on 50% of their salary after 25 years of service, regardless of age. Survivor benefits were also provided. Atlanta avoided default during the Depression and evidence reviewed thus far does not attribute any case of

municipal default during the 1920-1939 timeframe to employee pensions. Although pensions were available to Depression-era public employees, legal protections for these benefits have increased in recent decades. It may be appropriate to conclude that pension benefits were junior to debt service in a government's priority of payments during the 1930s, while today these two types of obligation appear to be almost *pari passu*, that is, on equal footing.

Data Selection

Today, the municipal bond market covers a broad range of issuers. This diversity was also present – albeit to a lesser extent – in the years prior to World War II. The municipal bond default list compiled in Joffe (2012) includes 5079 issuers who failed to make timely and complete principal or interest payments (or who obliged investors to accept refunding bonds in lieu of cash at maturity) at some time between 1920 and 1939. Most of the defaulting issuers were school districts, small towns and special tax districts – created to build roads and other infrastructure.

Financial data for special assessment districts and for school districts is more limited than for other issuer categories. Moody's bond manuals provide some data, but it is incomplete and not in a consistent format. The best data is available for states and large cities because they reported their financial statistics to annual Censuses at the time. Comprehensive financial data for smaller cities and counties was collected by decennial Censuses in 1922 and 1932.

Since annual Census data is available for a substantial number of larger cities, and since these cities experienced a significant number of defaults, the statistical analysis is most readily applicable to this subset of issuers.

For fiscal years 1930 and 1931, the Census Bureau reported financial statistics for 311 US cities with population over 30,000 (as of April 1, 1930). After 1931, the collection effort was scaled back, perhaps due to budgetary pressures at the federal level. In fiscal years 1932 and 1933, the Bureau reported similar statistics for 94 cities with population over 100,000 (also as of April 1, 1930). In fiscal 1934, Honolulu was added to the annual data set. Thus annual time series of fiscal data are available from the Census for 94 cities during the Depression period while more limited data is available for an additional 217 cities. In all, a total of 1000 city/year observations are available for the period FY 1930-1935. Data reported for each entity include revenues by category, expenditures by category, as well as various classifications of assets and debt.

Of the 311 cities in the sample, 46 had defaults on general obligation bonds between 1930 and 1936, implying a cumulative default rate of 15% for this sample. The overall municipal default rate during this period was about 16%. Among the non-defaulting cities, some had “forced refundings” in which investors were obliged to exchange maturing bonds for new ones with later maturities. Many others had defaults on special assessment bonds which were not general obligations of the cities. In the following analysis, none of these instances are classified as a default – but adjusting the default classifications in light of these circumstances is a reasonable task for future research.

Some defaults were attributed at the time to bank closures or bank holidays. Since FDIC insurance is now available, it would be reasonable to exclude defaults that really were the result of banking issues. However, reclassifying such defaults should only be done after an intensive reading of contemporary newspapers to confirm that they were fully attributable to banking problems. In certain cases, city officials may have used bank closures

or holidays as a pretext to obscure fiscal problems that rendered the city unable or unwilling to pay even if funds had not been temporarily frozen. Thus, these classification adjustments are also left to future research.

Once a city defaults, its data may become idiosyncratic as it suspends interest payments and possibly writes down principal. For example, Miami's interest costs fell from \$2.2 million in 1929 to \$0.3 million in 1933. News sources indicate that the city first defaulted in 1930. Since the purpose of the analysis is to predict default, post-default observations are dropped from the data set, resulting in the loss of 43 observations. Of the remaining 957 city/year pairs in the sample, 125 are associated with defaulting cities.

Although several hundred series are available in the Census data, most of them relate to small components of revenue and expenditure. This still leaves a number of aggregate revenue, expenditure, debt and asset series that may yield useful explanatory variables. Below, variables are evaluated in ratio form to maximize their modern relevance despite the substantial increase in population, price levels and per capita economic output that have occurred over the last 80 years.

Conceptual Model and Variable Selection

The ratio most commonly used in discussions of sovereign credit is the debt-to-GDP ratio, and that concept is sometimes applied to lower levels of government. The fiscal Census data do not include GDP or any other indicia of economic activity. More recently, regional income account statistics have been reported for states, metropolitan areas and counties in recent years, but most series do not extend back to the Depression era.

Although reliable measures of total economic output are not available at the municipal level, other demographic and macroeconomic variables can be employed. Previous

studies have used population, assessed valuation and per capita income as independent variables in default probability or rating prediction models.

Given fiscal data is available, the use of demographic and macroeconomic variables may not be necessary. The choice of whether to pay or default upon debt service obligations is made by the political leaders of a governing unit. The most immediately accessible data available to these officials include the size of the interest or principal payment that needs to be made, what financial resources are available to the government to make the payment and what other spending priorities are competing with the debt service obligation.

Consequently, the model constructed here is derived solely from fiscal measures. We present several alternative models in Appendix 1 that do include socioeconomic variables. The models presented in the Appendix 1 are also subjected to alternate specifications and estimation strategies as checks on the robustness of the results presented here.

Rating agencies use a number of purely fiscal metrics that can be estimated directly from the municipal Census data set. One commonly used metric is the ratio of interest costs to revenue. The rationale for including this ratio is that a default becomes likely when interest costs become so onerous that they threaten to crowd out other spending priorities. When the interest burden is low, it is not rational for a political leader to default, because he or she then loses access to capital markets and is thus compelled to reduce spending or raise taxes. As interest expenses rise, this disincentive is increasingly likely to be outweighed by the near term political costs of cutting spending on popular programs.

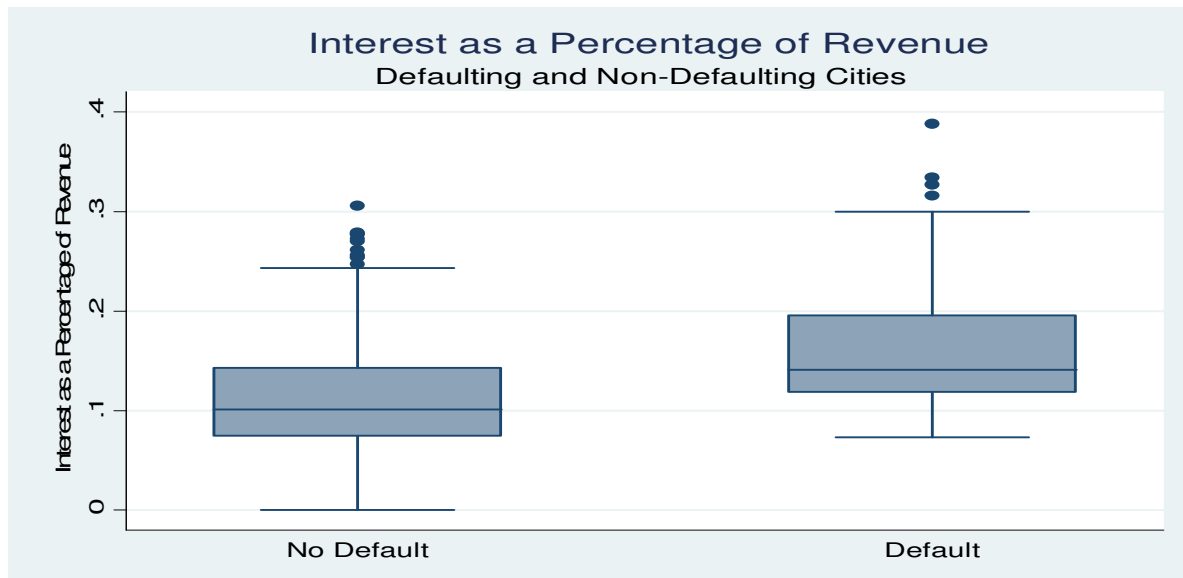
This theoretical underpinning does have a couple of limitations that should be noted. First, defaults often occur when a principal payment – rather than an interest payment – becomes due. In the Depression era, cities were more vulnerable to principal repayment

defaults because the concept of serialized maturities had yet to become popular. Large bond issues were typically scheduled to mature all at once. Many obligors accumulated revenues in “sinking funds” to meet these large debt repayments, while others expected to pay off the maturing bonds by issuing new ones. When sinking fund assets declined in value and the new issue market dried up, many governments were unable to redeem or roll over maturing issues. In the aftermath of the Depression experience, public finance specialists began to advocate serialized maturities, under which a large bond issue is broken down into a number of smaller tranches whose principal becomes due at varying dates – often one year apart.

Second, revenue may not be an ideal denominator, since political leaders may have the option of running surpluses or deficits. While many state and local governments are and have been subject to balanced budget requirements, these are typically prospective rather than retrospective and are often subject to evasion. On the other hand, using expenditures rather than revenues as a denominator is also an imperfect measure. Local governments cannot sustain large annual deficits indefinitely, so revenues appear to be a better indicator of their long term fiscal capacity.

A review of the 957 city/year observations shows that defaulting cities tended to have higher interest to revenue ratios than those that did not default. This is reflected in the box and whisker plot below. Further, a one sided t test of the defaulting and non-defaulting sample means fails to reject the null hypothesis that defaulting cities have higher interest to revenue ratios than non-defaulting cities at $p < .0001$. The sample mean for defaulting cities is 16.1% versus 11.0% for cities that did not default.

Figure: Box and whisker plot highlighting differences between defaulting and non-defaulting cities



Several of the Non-Defaulting observations with high Interest to Revenue Ratios were associated with two cities in Virginia: Norfolk and Portsmouth. Ackerson and Chamberlain (1935) report that Virginia implemented unique municipal default legislation in 1932. The Virginia law – which remains on the books³ – authorizes the Governor to investigate and withhold state aid to defaulting municipalities and to directly pay the withheld money to bondholders. Since gubernatorial action under this law is triggered by a petition from holders of affected bonds, the law does not appear to directly stop defaults, but does create an incentive on the part of local officials to avoid default.

³ Section 15.2-2659 of the Virginia Statutes states: Whenever it appears to the Governor from an affidavit filed with him by or on behalf of the owner or owners of any general obligation bonds of any locality, or by any paying agent for the bonds that the locality has defaulted in the payment of the principal of or premium, if any, or interest on any of its outstanding general obligation bonds, the Governor shall immediately make a summary investigation into the facts set forth in the affidavit. If it is established to the satisfaction of the Governor that the locality is in default in the payment of its bonds or the interest on them, the Governor shall immediately make an order directing the Comptroller to withhold all further payment to the locality of all funds, or of any part of them, appropriated and payable by the Commonwealth to the locality for any and all purposes, until the default is cured. The Governor shall, while the default continues, direct in writing the payment of all sums withheld by the Comptroller, or as much of them as is necessary, to the owners of the bonds in default, or the paying agent for the bonds, so as to cure, or cure insofar as possible, the default as to the bonds or interest on them.

Aside from the absolute burden of debt services, changes in available resources may be expected to enter into the default decision. For example, if revenues are declining, officials may face the choice of reducing public services below baseline levels or defaulting. Thus year-on-year revenue changes should be predictive of default. This analysis is supported by the Depression-era data. A one sided t test for defaulting and non-defaulting governments rejects the null hypothesis that annual revenue changes for the former group are not less than the latter at $p < 0.001$. The mean revenue change for defaulting observations is -2.3% versus +0.1% for the non-defaulting cases. Unfortunately, the use of this variable entails the loss of some observations. Revenue change is not directly observable from the Census data of any one year; for any given annual Census; it must be calculated by comparing revenues from the current Census to the prior one. The first Census used in the data set is that of 1930, so 1929 revenues are required to make data from that year usable. While 1929 census data is available, it only included 250 of the 311 cities in the 1930 census, resulting in the loss of 61 observations. Of the remaining 896 observations, 117 are associated with a defaulting city.

A city's liquid assets may be expected to act as a cushion against default. The Census data contains several categories of cash. Given the great variance of city size in the sample and the need to produce a time-independent model, any cash balance must be scaled. Since cash may be used to pay interest or principal, it can be reasonably scaled by converting it to a proportion of outstanding debt.

The t-test shows that Cash in Public Trust Funds as a Percentage of Gross Debt fails to reject the null hypothesis that this factor does not correctly differentiate defaulting and non-defaulting cities at $p=.01$. Since data is available for all cities in 1930, the full 957 observations can be included in the test. The mean cash balance for the 832 non-default

observations is 0.92% of debt, while the mean for the 125 default observations is only 0.16%. The t statistic for the Public Fund Cash to Debt Ratio was somewhat higher than two other ratios tested - All Cash Assets as a Percentage of Gross Debt and All Assets as a Percentage of Gross Debt - although both of these ratios are significant at the 5% level.

Theoretically, a city could sell fixed assets to remedy a shortfall, but it may not be feasible to do so quickly enough to avert a default. Other cash balances, such as those in pension funds, could in theory be borrowed to make debt service payments, but there may be legal barriers to accomplishing that. While controlling for these factors may add explanatory power to the model, doing so does not appear to be feasible in this current analysis.

Finally, it is reasonable to argue that a cash balance three or four years prior to a default is less relevant than the amount of cash available closer to the time of the actual default. The data set contains 21 observations in which the data's as of year is coincident with the city's default date, i.e. the fiscal data is from the year in which the city actually defaulted. (The other 25 defaults occurred in smaller cities after 1931, when the census stopped collecting their financial data.) The average Cash in Public Trust Funds as a Percentage of Gross Debt for these 21 observations is only 0.06% versus 0.83% for the remaining 936 observations.

If data for the smaller cities post-1931 could be collected, it may be possible to implement a balanced-panel logit regression like that described by Shumway (2001). This approach would allow the cash variable to play a stronger role in the model. In recognition of this potentiality, the cash variable (identified as the cash to debt ratio) is included in two of the models described in the next section, but it falls out in the final specification.

Population appears in a number of the models found in the literature and is also considered by rating agencies. In theory, population should be inversely related to default on the grounds that smaller cities are likely to be less diversified (and thus more risky) and that they are also more likely to have operational problems that would trigger a default.

This theory does not appear to be supported by summary statistics from the 311-city data set. The 46 cities that defaulted had an average population of approximately 321,000 while the 265 non-defaulting cities had an average population of roughly 127,000. Population averages provide a distorted view because they overweight larger cities. At the time (much like now), New York was the nation's largest city by a considerable margin. The fact that it defaulted adds substantially to the average population of defaulting cities.

One way to address this bias is to rank the cities from 1 to 311 and then compute the average rank of defaulting and non-defaulting cities. This analysis yields an average rank of 128 for defaulting cities and 161 for non-defaulting cities - also counter to theory. Population is excluded from the fiscal-ratio model presented in the next section, however as mentioned above, Appendix 1 contains extensions of the model that include population and other socioeconomic and demographic variables.

Statistical Methodology

The current study applies a binary response technique to a larger number of geographically representative issuers than Hempel (1971, 1973) could access. Since municipal bankruptcy is a rare event, it is preferable to use a model that differentiates between observations in the tail – suggesting a choice of logit over probit.

As detailed above, analysis of the data set identified three fiscal variables that have both empirical and theoretical support. These are: (1) interest to revenue ratio, (2) revenue

percentage change and (3) cash (in public trust funds) to (gross) debt ratio. Summary statistics for these variables are presented in the table below.

Table: Summary Statistics, Depression Era Sample

Variable	Obs	Mean	Std. Dev.	Min	Max
DEFAULT	957	0.131	0.337	0	1
Interest_as_Pct_of_Revenues	957	0.117	0.054	0	0.388
Receipts_Percent_Change	896	0.005	0.124	-0.333	1.282
Public_Cash_Pct_of_Gross_Debt	957	0.008	0.079	0	1.455

As mentioned earlier, there are fewer observations for the variable Receipts_Percent_Change as this variable requires receipts (revenue) data from two years, and the 1929 Census contained data for fewer cities than did the 1930 Census.

The equation below states the model described above mathematically:

$$Prob(DEFAULT_{it}) = F(X_{it}\beta)$$

where $DEFAULT_{it}$ is the binary dependent variable indicating whether or not city i is in default in year t , $F(\cdot)$ is the cumulative standard logistic distribution function, X_{it} is a matrix that contains the three fiscal-ratio variables summarized in the table above. Finally, β is a vector of coefficients to be estimated by maximum likelihood. The results of estimating this equation are given in the first column of the table below.

Table: Logit Analysis, Depression Era Sample

VARIABLES	DEFAULT	DEFAULT	DEFAULT	DEFAULT
Interest_as_Pct_of_Revenues	14.48*** (1.765)	15.03*** (1.618)	14.94*** (1.815)	25.47*** (3.857)
Receipts_Percent_Change	-2.212** (0.878)	-2.167** (0.861)	-2.104** (0.973)	-1.182 (1.242)
Public_Cash_Pct_of_Gross_Debt	-31.44 (35.060)			
Constant	-3.746*** (0.301)	-3.884*** (0.253)	-3.862*** (0.283)	-4.767*** (1.068)
Estimation strategy	Logit	Logit	Firthlogit	Firthlogit

Standard errors in parentheses	Robust	Robust	Homoskedastic	Homoskedastic
State and year fixed effects?	No	No	No	Yes
Observations	896	896	896	896
Chi2	95.6	95.14	73.97	148.2

In the first column of the table above, one can see that the estimated coefficients on Interest_as_Pct_of_Revenues, Receipts_Percent_Change, and Public_Cash_Pct_of_Gross_Debt have the theoretically predicted sign; namely, high interest expense relative to revenues increases probability of default, while a positive revenue change or high cash reserves reduces the probability of default. However, the coefficient on Public_Cash_Pct_of_Gross_Debt has a high standard error and the hypothesis that the true coefficient is zero cannot be rejected at conventional confidence thresholds.

In the interest of developing a parsimonious model, the next column estimates a restricted version of the equation; in particular in this new specification the variable Public_Cash_Pct_of_Gross_Debt is dropped. As one can see from comparing the coefficient estimates across the first two columns, dropping the cash variable does not substantially change the magnitude of the other coefficients.

It is possible that a failure to include certain unobserved factors, for example state rules that affect a city's likelihood of defaulting, are biasing the estimates. It is also possible that the fiscal ratio variables had differential impacts on cities in different states (that is, there may be heterogeneous effects as well as fixed effects.) While a more detailed analysis of these possibilities is left for future research, given the panel nature of the data, it is straightforward to implement a fixed-effect strategy as a first pass at controlling for unobserved state and time effects (even if this does not address the possibility of heterogeneous effects.) A complication that arises in a fixed-effect logit context, due to the

so-called separation problem (where a lack of variation in the dependent variable for cities in some states prevents the possibility of estimation) requires that a new estimation strategy must be employed if fixed effects are to be included.

The specification reported in column three includes the same variables that were included in the specification in column two. However, the estimation strategy differs (the coefficients in column three were estimated using Stata's *firthlogit* package; note, this package does not allow for estimating robust standard errors.) Comparing the coefficient estimates across columns two and three, one can see that the different estimation strategies produce very similar estimates. That is, the coefficient estimates are virtually identical, whether the estimation strategy is logit or *firthlogit*.

Finally, the column four presents estimates from a model with the same estimation strategy as in column three, but which includes state and time fixed effects along with two fiscal ratio variables. From the results in column four, one may conclude that the results in columns two and three may understate the magnitude of the effect of the variable *Interest_as_Pct_of_Revenues* while the magnitude of *Receipts_Percent_Change* may be overstated.

The coefficient estimates from column two can be used to estimate the default probabilities for cities in-sample and out-of-sample as shown below. The default probability formula implied by the model is:

$$Prob(DEFAULT|X) = \frac{1}{1 + e^{-(-3.884 - 2.167 * \Delta R + 15.03 * IR)}}$$

Where *IR* = *Interest_as_Pct_of_Revenues*, and ΔR = *Receipts_Percent_Change*. Although one may prefer to use estimates from the fixed-effect model from column four, we presented default probability predictions using the results from column two because they are simpler

and aid in exposition. If one wanted to use column four's results, the constant term should then reflect the fixed-effect for California, which is -2.073 (so the constant would be $-4.767 - 2.073 = -6.84$), and the coefficients on ΔR and IR would be -1.182 and 25.47 respectively.

Limitations

The use of varying numbers of observations per city (i.e. an unbalanced panel) gives greater weight to the larger cities in estimating the model's coefficients. In addition, the exclusion of socioeconomic variables which have been used in much of the previous municipal credit literature is unorthodox. As already mentioned, we address these issues by providing a set of alternative models using the Depression-era data set in the Appendix.

Further data supplementation may facilitate comparisons between city financials in the fiscal year of default to those one, two or three years away from defaulting. It may also permit model specifications that could incorporate a measure of the city's cash reserves.

The explanatory variables in the foregoing analysis were based on composite financials for all governmental units within the city. A better approach would involve using only data from the municipal government itself – which is identified as the “city corporation” in the Census reports.

Despite these limitations, the empirical analysis presented here supports the common sense view that revenue declines and large debt service costs lead to defaults. The model also allows us to quantify these vulnerabilities in a way that may permit reasonable estimation of default probabilities for current issuers. Although the data is many decades old, it addresses a timeless issue affecting municipal leaders in a liberal democracy: how to trade off service demands from the electorate with the need to maintain credibility among bond investors.

Municipal Bond Defaults in California: History and Case Studies

Over California's 160 year history, city bond defaults have been relatively rare. Default activity has varied over time, peaking during the Depression. This chapter surveys the entire history of payment defaults by California cities, with sections on each of the five cities filing municipal bankruptcy petitions since 2001. It concludes with a brief discussion of defaults attendant to the dissolution of Redevelopment Agencies and a review of city fiscal emergencies.

Pre-1930

For the period from statehood through 1930, Hillhouse (1935) lists eight defaults ascribed to California cities, but a review of original sources indicates that not all of these are legitimate defaults in the contemporary sense.

The highest profile municipal bond default in California's early years was that of Placerville in 1866. Three years earlier, the city issued \$100,000 in bonds to finance a rail connection to Folsom, where service from Sacramento terminated at the time. The funds were invested in the newly created "Placerville and Sacramento Valley Railway Company". The company was unable to complete the railroad, terminating it 12 miles short of the city. Meanwhile, Placerville suffered declining population due to the end of the Gold Rush. In 1866, the city stopped making interest and principal payments. In 1873, city leaders decided to avoid personal liability by disbanding the municipal government. In 1900, the city government was re-established and an accommodation was made with the holders of \$34,500 still outstanding.

Holders of \$7000 in Placerville fire department bonds issued earlier in 1863 did not fare as well. In the case of *Wichman v. City of Placerville* (1905), 81 Pac. 537, the California Supreme Court invalidated their bonds. Placerville was incorporated by the State legislature in 1859 and then reincorporated on April 6, 1863 when it was authorized to issue the railroad bonds. The previous city corporation issued the fire department bonds on April 3. The court held that the bonds were obligations of a liquidated city corporation and thus not the responsibility of the new government. California courts invalidated a number of other city bonds listed among Hillhouse’s defaults. These cases are listed below:

City	Issue	Case	Description
Sacramento	General obligations issued in 1854.	<i>Bates v Gregory</i> (1891) 26 Pac 891	City was reincorporated in 1863. Bonds matured in 1874 but were not presented for payment until 1887. Court ruled that under the statute of limitations, the city’s liability for the bonds lapsed in 1878.
San Diego	Railroad Aid bonds issued in 1874	<i>McCoy v. Briant</i> (1878), 53 Cal 247 and <i>Lehman v. City of San Diego</i> (1897), 83 Fed 669	In <i>McCoy v. Briant</i> , bonds were invalid because Board of Trustees failed to pass a resolution authorizing them. In <i>Lehman v. City of San Diego</i> , bonds were issued after statutory authorization was repealed but antedated so that they would appear to be valid.
San Francisco	Dupont Street and Montgomery Avenue street construction bonds issued in 1872 and 1877 respectively	<i>Shapter v. San Francisco</i> (1901), 110 Fed. 615 and <i>Liebman vs. City and County of San Francisco</i> (1885), 24 Fed. 705	The Dupont Street bonds were held to be a special assessment and thus an obligation of neighborhood property owners rather than a general obligation of the city. The Montgomery Avenue bonds were determined to be obligations of a public works corporation, rather than city general obligations. The public works corporation’s taxing authority had previously been invalidated in <i>Mulligan v. Smith</i> (1881), 59 Cal 206.
Santa Cruz	Water works bonds issued in 1885	<i>Santa Cruz Water Company v. Kron</i> (1887), 15 Pac 772	Bond issued after enabling state legislation was passed but prior to its effective date.

Sources: Dean (1912), Court Filings

Hillhouse also lists a San Francisco default on Improvement Obligations, but his reference - Sakolski (1932) - suggests that the payment failures dating to the 1850s were not related to bonds. Apparently, San Francisco, like many gold rush era boomtowns issued scrip to compensate vendors before taxes were collected. San Francisco was unable to redeem the scrip with cash in a timely manner and its value depreciated. Sakolski describes the exploits of one speculator, Peter Smith, who bought a substantial amount of the scrip at a discount and then profited by obtaining judgments against the city – compelling it to redeem the IOUs by selling its real estate holdings. State and local governments across the country have resorted to issuing scrip at various times. The state of California did so in 2009 during a budget impasse. Although the use of scrip may be characterized as a default in the popular press, it does not meet the definition of a payment default used by contemporary bond market participants or in this study.

The last default listed by Hillhouse involved the City of Stockton in 1870. In that year, voters overwhelmingly approved \$300,000 to fund the proposed Stockton & Visalia Railroad (Tinkham, 1923). The Council expected the new railroad to lay 15 miles of new track terminating at the San Joaquin River and connecting it to the rest of the Central Valley. The railroad laid one mile of new track connecting Stockton to a pre-existing railway line, and then demanded payment from the City (Burrill, 2011). Stockton's City Council refused to levy taxes required to service the bonds, which appear to have been held by the railroad. After extended litigation, the city and the railroad reached an out of court settlement (Tinkham, 1923).

Depression Era

California local government issuers experienced a substantial number of defaults during the Great Depression. Most of these defaults affected irrigation and reclamation districts as well as special assessment districts (older equivalents of today's Mello Roos issuers). Prior to the Depression, many special assessment districts had been created to pave roads and build other local infrastructure under the state's 1915 Improvement Act.

Available documents contain reports of 11 California towns and cities that defaulted on general obligation bond payments. These defaults are listed in the following table.

City	Population	Date	Comments	Source
Alturas	2,400	As of 11/1936	Defaults on both General Obligation and 1915 Act Bonds. Cause not given. Voters approved issuance of refunding bond issue in December 1936.	Moody's Government Securities Manual 1937, p.228
Arcadia	5,216	1/1/1932	Defaults on water, street and library general obligation bonds. Three corporate taxpayers who accounted for 30% of the town's property tax revenue became delinquent.	Bond Buyer, 4/7/1934, p. 825.
Benicia	2,913	10/1/1932	Cured by May 1933. Cause not given.	Bond Buyer, 5/27/1933, p. 1115.
Blythe	1,020	1930	Population fell after a 1922 flood. Special assessment defaults began in 1927. General obligations defaulted in 1930. City clerk told <i>The Bond Buyer</i> that "the taxes were so prohibitive that no taxes could be collected. At one time our tax rate was over \$36 per hundred."	Bond Buyer, 7/2/1935, p. 1763
Brawley	11,300	As of 10/1/1935	26% of fiscal 1935 tax levy uncollected within the fiscal year.	Moody's Government Securities Manual 1936, p.173
Calexico	6,299	As of 9/11/1933	Property tax revenue decline due to lower assessments and higher delinquencies.	Moody's Government Securities Manual 1934, p.202
Calipatria	1,554	As of 11/18/33	Only 65% of tax levy was collected in 1931.	Moody's Government

				Securities Manual 1934, p.202
Culver City	7,500	As of 6/30/34	Fall in property tax collections	Moody's Government Securities Manual 1936, p.175
Imperial	1,943	As of 11/17/1933	50% of fiscal 1933 tax levy uncollected within the fiscal year.	Moody's Government Securities Manual 1934, p.204
Lynwood	10,000	As of 7/25/39	Tax delinquencies rate exceeded 30% in the mid-1930s but had fallen substantially by the time the default was reported.	Moody's Government Securities Manual 1940, p.150
Oroville	3,698	1/1/1933	Cured by May 1933. Cause not given.	Bond Buyer, 5/25/33, p. 1093.
Vacaville	1,556	11/1/1932	Default not due to lack of funds, but rather a clerical oversight. Missed coupon paid in full less than one month late.	Bond Buyer, 11/30/32, p. 2892.

None of the defaulting cities had population of more than 12,000. According to the 1932 Census of State and Local Governments, 40 of the state's cities had populations in excess of 12,000 at the time – suggesting that solvency problems during the Depression were confined to smaller cities. However, it should be noted that several larger cities experienced defaults in special assessment districts. Los Angeles, already the state's largest city, contained thirty defaulting districts. San Diego and Oakland also had assessment bond defaults.

Property tax delinquencies appear to be the primary cause of Depression era defaults. As noted in the table, a number of cities experienced delinquency rates in excess of 30%. It is reasonable to assume that the bonds were issued with the expectation that the bulk of property taxes would be paid on a timely basis. When this did not occur, towns and cities with substantial debt burdens became unable to meet them. This is consistent with

the nationwide trend described in the previous chapter. It is why we include Annual Revenue Change as a factor in our models.

Post 1940

After the Depression, benign credit conditions returned. Only three defaults by a city (as opposed to a special assessment district) appear to have occurred between 1940 and 1993. In 1965, Redondo Beach defaulted on an unrated \$9 million harbor bond issue. According to a contemporary Moody's Government Securities Manual (1968), the bonds were secured by revenues from the boat harbor and taxes collected in the harbor area in excess of the amount realized in fiscal 1958. The bonds were thus tax supported but not a general obligation or general fund obligation of the city. According to the *Wall Street Journal* (1965), "The city blamed the default on a number of factors including construction delays, unusually severe winter storms in 1962 and the failure of expected offshore oil revenue to materialize." The *Journal* report went on to state that the 1959 feasibility study for the harbor project had been "overly optimistic".

In 1982, Parlier defaulted on general obligation bond payments, as well as debt service obligations to a bank and the Farmers Home Administration. The city's population was reported as 2902 or 5093 by different sources perhaps because it was in the process of annexing an unincorporated area known as West Parlier. According to the US Advisory Commission on Intergovernmental Relations (1984), the city had \$110,000 in general obligation bonds outstanding and defaulted on a \$6000 debt service payment. As of December 15, 1982 the city had total debts of \$819,089 and cash of only about \$2000. A *Los Angeles Times* article (Taylor, 1982a) attributed the city's financial distress to a cost overrun on a recently completed community center, and the initial failure of an industrial

park project to attract tenants. Interest in the park suffered after a carcinogen was found in the town's well water. The *Times* coverage also cited fiscal mismanagement which included the lack of a "meaningful audit" of the city's books between 1975 and 1982. The ACIR report notes that the city ran persistent deficits ahead of the financial crisis, culminating in a fiscal 1982 shortfall that amounted to 36% of revenues.

Once the extent of the crisis was determined, the city took drastic action to avoid a municipal bankruptcy and cure its various defaults. Measures included terminating 16 of the 22 city employees. The terminations included all of Parlier's police officers as Fresno County took responsibility for local law enforcement (Taylor, 1982b). By June 30, 1983 the city had resolved all of its defaults.

In 1993, the small city of Arvin defaulted on \$2.945 million of Certificates of Participation used to finance temporary housing for farmworkers. In early 1994, it defaulted on \$7.89 million in Certificates of Participation used to finance a golf course (Altman, 1994). Although the city considered a chapter IX filing, it appears to have reached an accommodation with certificate owners. According to Mysak (2010), bondholders received about 28 cents on the dollar for the defaulted golf course COPs.

In 1998, the City of Healdsburg issued \$7 million in bonds on behalf of Nuestro Hospital Group to purchase the local hospital (California Healthline, 1998). The bonds did not constitute a claim on tax revenues and apparently went into default when the hospital generated insufficient revenue to service them. We were unable to locate further information about this default, which is listed in Mergent Corporation's municipal bond database.

Since 2001, five California cities have made municipal bankruptcy filings under Chapter IX of the federal bankruptcy: Desert Hot Springs (2001), Vallejo (2008), Stockton (2012), Mammoth Lakes (2012) and San Bernardino (2012). Four of these situations also involved municipal bond defaults. We describe each of these cases in separate sections below.

Desert Hot Springs

Although Desert Hot Springs December 19, 2001 bankruptcy filing (case number 6:01-bk-30756-DN Central District of California) followed an adverse court decision, its financial problems were not wholly attributable to the lawsuit. In fact, the city's 1999 default on unrated revenue anticipation notes occurred well before the bankruptcy filing.

According to financial statements attached to offering documents available on the MSRB EMMA system, the city experienced large and growing all-fund deficits in fiscal 1995, 1996 and 1997. The city also experienced a large general fund deficit in fiscal 1997 – amounting to 36% of revenues.

On August 27, 1997, the City issued \$1,275,000 in Revenue Anticipation Notes bearing a 4.75% interest rate and due the following year. These notes, rated F-2 by Fitch, were apparently redeemed with proceeds from a second set of Revenue Anticipation Notes issued on September 30, 1998. This second issue, totaling \$1,415,000 was unrated and carried an 8.25% interest rate – a clear signal of the city's financial distress.

Selected fiscal statistics for Desert Hot Springs gathered from contemporaneous financial statements are provided in the accompanying tables. All governmental fund revenues fell 4% in fiscal 1997 but rose 7% and 8% in 1998 and 1999 respectively. These robust increases were primarily attributable to intergovernmental revenue; real estate

assessment and property tax revenues were relatively weak during this period, despite the strong national housing market. A later filing on EMMA (City of Desert Hot Springs, 2004) shows that assessed valuations rose 3% in fiscal 1998 and fell 1% in fiscal 1999. Overall, valuations remained stagnant throughout the mid and late 1990s, dropping slightly from \$490.6 million in fiscal 1994 to \$487.0 million in 1999.

Interest expense accounted for 17% of all fund revenue in 1996 and 15% in 1997 and 1998. Also, in 1998, Pension Contributions became a significant budgetary factor, following the city's entry into the CalPERS. In 1999, the city's interest expense fell due to the default.

While aggregate governmental fund balances remained positive ahead of the city's default, Desert Hot Springs' general fund balance fell below zero in fiscal 1996 and became increasingly negative ahead of the 1999 default. General fund exhaustion appears to be the key driver of the Desert Hot Springs default – a theme that repeats in the Vallejo, Stockton and San Bernardino cases described below.

City of Desert Hot Springs - All Governmental Funds					
Statement of Revenues, Expenditures and Changes in Fund Balances					
		Total Governmental Funds			
		1996	1997	1998	1999
Revenues					
	Taxes	3,910,047	3,778,247	3,921,255	3,461,885
	Permits and fees	186,355	158,992	250,835	549,967
	Intergovernmental	956,230	1,008,842	1,491,961	2,089,252
	Licenses	145,208	120,710	115,212	113,607
	Interest	582,585	323,448	343,186	256,268
	Miscellaneous	268,671	398,663	56,404	202,337
	Total Revenues	6,049,096	5,788,902	6,178,853	6,673,316
Expenditures					
	General government	1,743,931	1,872,688	1,775,738	2,052,264
	Public safety	2,078,876	2,987,568	2,666,403	2,393,606
	Public works	467,096	525,026	911,993	730,159
	Parks and recreation	466,688	576,999	74,700	41,053
	Capital outlay	851,832	425,726	771,725	60,478
	Principal	918,297	2,523,996	416,657	22,253
	Interest	1,009,056	860,125	895,990	365,347
	Economic development	192,688	100,204	193,161	307,922
	Prior year expenditures	603,889	207,344	165,894	652,175
	Total Expenditures	8,332,353	10,079,676	7,872,261	6,625,257
	Excess (Deficiency) of Revenues over Expend.	(2,283,257)	(4,290,774)	(1,693,408)	48,059
	Other Financing Sources (Uses) + Adjustments	427,784	424,579	118,580	(302,887)
	Net Change in Fund Balances	(1,855,473)	(3,866,195)	(1,574,828)	(254,828)
	Beginning Fund Balances	11,246,073	9,390,600	5,524,405	3,949,577
	Ending Fund Balances	9,390,600	5,524,405	3,949,577	3,694,749
	Pension Contributions		40,957	203,951	200,000
	Annual Revenue Change		-4.30%	6.74%	8.00%
	(Interest + Pension) / Revenue	16.68%	15.57%	17.80%	8.47%
<i>Source: Audited Financial Reports, FY 1996-1999.</i>					
<i>Pension contributions for 1999 not available; estimated value entered.</i>					

City of Desert Hot Springs - General Fund				
Statement of Revenues, Expenditures and Changes in Fund Balances				
General Fund				
	1996	1997	1998	1999
Revenues				
Taxes	2,235,026	2,200,461	2,337,528	2,160,812
Permits and fees	168,770	141,281	169,281	512,408
Intergovernmental	582,359	606,870	598,647	740,025
Licenses	145,208	120,710	115,212	113,607
Interest	35,048	3,297	30,161	5,503
Miscellaneous	233,125	335,812	53,369	171,578
Total Revenues	3,399,536	3,408,431	3,304,198	3,703,933
Expenditures				
General government	1,375,135	1,417,084	1,448,594	1,617,127
Public safety	2,075,894	2,955,166	2,263,685	1,690,046
Public works	93,952	81,959	233,801	177,311
Parks and recreation	53,277	121,020	74,700	21,867
Capital outlay	-	-	134,974	-
Principal	-	46,095	42,152	-
Interest	-	3,994	122,047	38,000
Economic development	-	-	-	-
Prior year expenditures	10,297	-	-	1,126
Total Expenditures	3,608,555	4,625,318	4,319,953	3,545,477
Excess (Deficiency) of Revenues over Expends.	(209,019)	(1,216,887)	(1,015,755)	158,456
Other Financing Sources (Uses) + Adjustments	(663,407)	417,079	1,026	105,441
Net Change in Fund Balances	(872,426)	(799,808)	(1,014,729)	263,897
Beginning Fund Balances	214,503	(657,923)	(1,457,731)	(2,472,460)
Ending Fund Balances	(657,923)	(1,457,731)	(2,472,460)	(2,208,563)
Pension Contributions		40,957	203,951	200,000
Annual Revenue Change		0.26%	-3.06%	12.10%
(Interest + Pension) / Revenue	0.00%	1.32%	9.87%	6.43%
<i>Source: Audited Financial Statements, FY 1996-1999.</i>				
<i>Pension contributions for 1999 not available; estimated value entered.</i>				

In 2000, voters ratified additional property taxes and utility taxes that were projected to raise \$1.8 million in additional revenue (Deborah, 2000). The city may have been able to cure its default and avoid bankruptcy had it not suffered a legal setback in 2001.

As reported by Gold (2001) in the *Los Angeles Times*, Desert Hot Springs chose to enter Chapter 9 after losing a decisive battle in an 11 year fight with Silver Sage Developers. The litigation began in 1990 after the City Council threw out the company's plan to build a mobile home park. The developer sued, claiming that the Council's decision violated the Fair Housing Act by discriminating against low income families. The initial jury award to Silver Sage of \$3 million was later reduced to \$1 by a second jury. But, in July 2001, the 9th Circuit Court of Appeals reinstated the \$3 million award and added another \$3 million for interest and legal fees. On December 18, a federal judge declined to block Silver Sage from seizing city assets to satisfy the judgment. Desert Hot Springs filed for Chapter IX bankruptcy to forestall the asset seizure.

Gold's 2001 article also notes that developers had been hesitant to start projects in Desert Hot Springs because of the legal uncertainty. This effect may explain the relatively stagnant real estate assessments mentioned earlier.

Vallejo

On May 23, 2008, Vallejo became the largest city to file a Chapter IX bankruptcy petition (case number 2:08-bk-26813 Eastern District of California) since Congress first allowed municipal bankruptcies in 1934. Vallejo's bankruptcy involved a default on Certificates of Participation (COPs). These certificates, unlike General Obligations or General Fund Obligations, are not senior claims on a city's tax revenue. Instead, they

represent the investor's share in lease revenues the city agrees to pay on certain facilities. As noted in the COPs offering materials, "the City could choose to fund other services before making Lease Payments" and that holders have limited recourse in the event of a default or bankruptcy (Wulf, Hansen and Company, 2003).

According to documents filed with the court and posted on the city's web site, bankruptcy was necessitated by the fact that the city's general fund had been exhausted and was expected to continue running large deficits in FY 2009. Media coverage in early 2008 also attributed the situation to high police and fire employee costs as well as unwillingness on the part of public safety unions to make concessions (see, for example, Jones, 2008 and Rohrs, 2008).

Data retrieved from CAFRs and other reports during this period confirm the exhaustion of the general fund, but show large positive balances in other funds. A February 28, 2008 staff memo showed \$137 million in cash balances across all funds (Mayer, 2008). Governmental Fund balances reported as of June 30, 2008 in the CAFR were \$105 million.⁴

It would appear that the city could have avoided or postponed bankruptcy by lending money from other governmental funds to the general fund. The city may not have chosen this option because of a strict reading of Governmental Accounting Standards. Mayer (2008) states that these standards "as applied by the City and examined by our external auditors, permit short-term interfund borrowing ... only to the extent that there is a demonstrated ability to repay these loans." Further, staff appears to have underestimated the costs of the Chapter IX process. In a May 6, 2008 staff memo, bankruptcy costs were

⁴ As discussed by GASB (2006), the concept of fund balance varies across fund types, so aggregating balances across all funds may be misleading.

estimated at \$750,000 to \$2 million (City of Vallejo, 2008). A 2011 *Wall Street Journal* article put the final cost at \$9 million (White, 2011).

During each year between 1999 and 2003, the city issued COPs totaling more than \$54 million. After the city filed its bankruptcy petition, Vallejo capped interest payments below the contractual rate. As the bankruptcy progressed, the City completely suspended interest and principal payments twice (City of Vallejo, 2011). At the end of the bankruptcy, the COPs were replaced with new lease agreements. Creditors are still entitled to receive all principal originally lent, but later than expected and with less interest than required by the original COPs. The city's remaining debt, composed mostly of revenue bonds issued by Enterprise (i.e., business type or non-governmental) funds, was not adjusted.

The accompanying tables show Total Governmental and General Fund values for FY 2005-2009. The bankruptcy filing occurred at the end of FY 2008. Annual government-wide revenue fell 5% in 2006, rose 4% in 2007 and fell almost 10% in 2008. While this last observation is consistent with the idea that falling revenue is predictive of default, it would not have been available until well after the bankruptcy filing since the CAFR in which it appeared was published several months after the fiscal year end. Some cities, including Vallejo, publish interim financials, so it may have been possible to estimate this value during the fiscal year.

Interest and retirement costs as a proportion of total revenue were not especially high. In the year of the default, the ratio peaked at 11.36%. Overall, the parameters derived from the Depression data do not effectively predict the Vallejo default. On the other hand, large general fund deficits and low balances seem to have been the major contributors. As

suggested above, city management did not take all possible measures to stave off insolvency and appear to have been surprised by the high cost of the bankruptcy filing.

City of Vallejo - All Governmental Funds					
Statement of Revenues, Expenditures and Changes in Fund Balances					
Total Governmental Funds					
	2005	2006	2007	2008	2009
Revenues					
Taxes	55,161,490	57,550,479	59,919,619	58,729,898	55,815,958
Licenses, permits and fees	11,852,968	6,000,434	4,677,963	3,327,060	2,440,614
Fines and forfeitures	1,450,743	3,887,337	1,683,911	1,827,945	1,560,809
Intergovernmental	55,423,739	49,780,261	48,824,041	38,546,090	36,743,052
Use of money and property	6,419,396	6,702,584	7,789,616	6,891,090	4,520,458
Charges for services	14,692,748	11,326,489	17,925,042	20,234,105	20,026,448
Other	2,717,589	4,448,399	4,543,088	1,833,312	841,040
Total Revenues	147,718,673	139,695,983	145,363,280	131,389,500	121,948,379
Expenditures					
Current:					
Legislative and advisory	311,994	323,174	270,743	292,370	215,485
Executive	1,413,534	1,374,916	1,425,841	736,846	621,003
Finance	1,286,935	1,864,644	1,926,168	1,159,374	1,696,164
Human resources	973,717	1,282,442	1,319,304	723,493	553,254
Law	628,677	764,614	871,733	892,284	863,429
Development services	3,078,353	3,650,863	2,786,231	3,248,627	2,633,028
Community development	52,217,190	39,856,041	46,712,048	33,707,475	39,056,197
Fire services	22,742,661	24,723,674	28,072,502	28,063,568	23,107,207
Police services	32,898,573	36,630,148	40,252,109	41,185,818	34,354,261
Public works	12,204,692	13,793,065	13,486,711	14,077,984	12,713,924
Nondepartmental	6,120,498	6,826,980	7,284,198	11,045,020	11,426,993
Capital outlay	5,867,421	5,218,215	18,761,691	14,039,215	7,499,257
Debt Service:					
Principal	5,056,682	1,657,337	3,966,314	1,497,254	1,464,697
Interest and fiscal agent fees	2,463,555	2,705,865	2,667,960	2,589,723	2,336,172
Total Expenditures	147,264,482	140,671,978	169,803,553	153,259,051	138,541,071
Excess (Deficiency) of Revenues over Expenditures	454,191	(975,995)	(24,440,273)	(21,869,551)	(16,592,692)
Total Other Financing Sources (Uses)	509,605	(1,613,209)	724,900	3,666,290	(513,905)
Net Change in Fund Balances	963,796	(2,589,204)	(23,715,373)	(18,203,261)	(17,106,597)
Beginning Fund Balances	148,931,220	149,895,016	147,305,813	123,590,440	105,387,179
Ending Fund Balances	149,895,016	147,305,812	123,590,440	105,387,179	88,280,582
Pension Contributions	9,599,955	11,293,291	11,734,043	12,332,457	10,372,222
Annual Revenue Change		-5.43%	4.06%	-9.61%	-7.19%
(Interest + Pension) / Revenue	8.17%	10.02%	9.91%	11.36%	10.42%

Source: Comprehensive Annual Financial Reports, FY 2005-2009.

City of Vallejo - General Fund					
Statement of Revenues, Expenditures and Changes in Fund Balances					
General Fund					
	2005	2006	2007	2008	2009
Revenues					
Taxes	51,579,991	53,083,876	55,617,416	53,821,263	51,071,916
Licenses, permits and fees	7,436,293	3,440,959	2,749,888	2,431,928	1,846,301
Fines and forfeitures	1,372,316	1,410,820	1,483,923	1,428,818	1,430,689
Intergovernmental	11,191,393	14,321,658	11,553,159	12,718,335	11,908,871
Use of money and property	558,743	546,224	262,559	473,351	599,651
Charges for services	6,782,717	2,193,226	5,867,269	7,682,796	8,061,261
Other	213,559	2,826,607	3,017,544	117,944	131,494
Total Revenues	79,135,012	77,823,370	80,551,758	78,674,435	75,050,183
Expenditures					
Current:					
Legislative and advisory	311,994	323,174	270,743	247,668	215,485
Executive	1,170,568	1,310,691	1,312,440	736,846	621,003
Finance	1,274,935	1,864,644	1,926,168	1,159,374	1,696,164
Human resources	973,717	1,282,442	1,319,304	723,493	553,254
Law	628,677	764,614	871,733	892,284	863,429
Development services	2,199,270	2,338,949	2,726,448	2,812,419	2,457,265
Community development	-	416,290	559,173	795,304	744,688
Fire services	20,715,988	22,533,874	25,238,098	25,286,403	20,424,746
Police services	32,013,022	35,264,688	38,050,873	38,204,475	31,487,056
Public works	4,615,794	5,055,339	4,052,169	4,210,768	3,987,009
Nondepartmental	5,902,455	5,041,656	6,228,746	9,868,036	9,894,757
Capital outlay	114,776	-	-	-	-
Debt Service:					
Principal	127,330	581,866	543,912	286,700	219,899
Interest and fiscal agent fees	5,962	125,755	78,727	99,577	50,157
Total Expenditures	70,054,488	76,903,982	83,178,534	85,323,347	73,214,912
Excess (Deficiency) of Revenues over Expend.	9,080,524	919,388	(2,626,776)	(6,648,912)	1,835,271
Total Other Financing Sources (Uses)	649,886	(3,919,060)	(1,208,499)	3,140,555	(1,022,443)
Net Change in Fund Balances	9,730,410	(2,999,672)	(3,835,275)	(3,508,357)	812,828
Beginning Fund Balances	4,125,934	13,856,344	10,856,672	7,021,397	3,513,040
Ending Fund Balances	13,856,344	10,856,672	7,021,397	3,513,040	4,325,868
Pension Contributions	9,599,955	11,293,291	11,734,043	12,332,457	10,372,222
Annual Revenue Change		-1.66%	3.51%	-2.33%	-4.61%
(Interest + Pension) / Revenue	12.14%	14.67%	14.66%	15.80%	13.89%

Source: Comprehensive Annual Financial Reports, FY 2005-2009.

Mammoth Lakes

The Town of Mammoth Lakes filed a Chapter IX bankruptcy petition with the Eastern District of California Bankruptcy Court (Case Number 2:12-bk-32463) on July 3, 2012. The town, which had 8234 residents in 2010, did not default on any of its bonded indebtedness and the bankruptcy case was dismissed on November 16, 2012.

Mammoth Lakes filed for bankruptcy after it lost a law suit to Mammoth Lakes Land Acquisition LLC (MLLA), exhausted its appeals and failed to convince the plaintiff to reduce the amount of the judgment. MLLA sued the town because it reneged on a 1997 development agreement in which MLLA improved the municipal airport in exchange for the right to build a hotel/condominium project at the site. The town did not grant approval for the hotel/condominium project because of safety concerns expressed by the Federal Aviation Administration, but officials were aware of these concerns when they signed the 1997 agreement (Goodwin Proctor, 2011). The original judgment of \$30 million awarded in 2008 increased to over \$42 million by early 2012 due to attorneys' fees and interest.

According to the Mammoth Lakes financial statements, the town had no general obligation bonds or revenue bonds outstanding when it filed, but did have about \$2.5 million in Certificates of Participation issued in 2000 and 2004. The 2000 COPs were rated while the 2004 issue was unrated and sold privately to Citizen's Bank. After the bankruptcy filing, S&P downgraded the 2000 COPs issue from BB to C (Standard & Poors, 2012a).

In the city's initial plan of adjustment, it stated its intention to continue servicing the 2000 COPs on time and in full. On the other hand, it planned to extend the term of the 2004 Citizen's Bank COPs by three years (Town of Mammoth Lakes, 2012a).

Ultimately, no default occurred because the Town reached a settlement with MLLA. Under the settlement, MLLA agreed to accept \$29.5 million plus interest paid over 23 years. The town accommodated the annual cost of the settlement through budget cuts and by increasing its revenue forecast (Town of Mammoth Lakes, 2012b). Once the settlement was reached, the bankruptcy case was dismissed and S&P upgraded the 2000 COPs to BB+ (Standard & Poors, 2012b).

Stockton

The City of Stockton filed a Chapter IX petition on June 28, 2012 (case number 2:12-bk-32118) after it was unable to secure concessions from creditors during the AB 506 mediation process.⁵ Most of Stockton's municipal bonds are insured by Ambac, National Public Finance Guarantee and Assured Guarantee.

During the AB506 discussions the City stopped making debt service payments on 2004 Lease Revenue Bonds. These bonds are secured by parking garage revenues and are not a general obligation of the City. The bond insurer, National Public Finance Guarantee, initially received payments from a reserve fund administered by the bond trustee. Once that fund was exhausted the trustee took possession of the three parking facilities covered by the lease agreement and diverted a portion of the proceeds to debt service (Wells Fargo, 2013).

In February 2013, the City reached an agreement with Ambac to scale back payments on 2003 Certificates of Participation, but the other two insurers continued to

⁵ Assembly Bill 506 requires cities contemplating a bankruptcy filing to engage in a neutral evaluation process with creditors. During the neutral evaluation process, which can last up to 90 days, the city is shielded from legal action on the part of creditors.

press their objections to the bankruptcy filing in court. They were joined by Franklin Templeton, which holds uninsured City obligations. On April 1, 2013, Judge Christopher Klein upheld the City's bankruptcy filing clearing the way for it to reduce debt service payments without facing legal action.

Stockton's reasons for filing a bankruptcy were given in a recent news release:

On July 1, 2012, the City had less than \$2 million in cash, all of which would have been entirely depleted within the first few days of the fiscal year. It was not even enough to make the City's July payroll. Thus, the City would not have been able to pay its employees, let alone its creditors, during any month of the 2012-2013 fiscal year. ... The City could not balance its budget outside of chapter 9, absent massive giveups by creditors who refused to make such concessions. The proposed 2012-2013 budget that City staff submitted to the City Council in May 2012 confirmed what the City already knew: It could not close its \$26 million "gap" and balance its budget without chapter 9 protection. Its anticipated revenues remained depressed. And while it had already made across-the-board reductions in employee costs, the costs of healthcare, pensions, and debt service in particular were trending upward (City of Stockton, 2013).

As of this writing, the City's 2012 CAFR has not been filed and no data for total governmental funds is available for the most recent fiscal year. The accompanying tables show data for fiscal 2008-2011. General fund estimates for 2012 and 2013 were derived from the city's most recent budget report submitted to the City Council in March 2013. No more recent financial statistics appear to be available on Stockton's web site.

The 2011 data reflect \$15 million in adverse prior period adjustments to the city's general fund balance – part of \$109.7 million in such adjustments across all funds. The adjustments are described in footnote 14 of the CAFR starting at page 129 (although the adjustments described in the note do not appear to account for the full \$109.7 million in adjustments mentioned on page 4).

The General Fund adjustments included \$12.3 million in allowances for doubtful accounts, \$1.2 million in accrual adjustments, and \$0.5 million for double counted parking citation revenue. Most of the adjustment to other Governmental Funds was attributable to a change in accounting method for city loan programs. Although the note does not explicitly say this, the adjustments appear to address the possibility that a large proportion of these loans will not be fully repaid.

The fact that audited financials had to be adjusted post-filing compromises their usefulness for analyzing a city's credit risk. If the data are not reliable, they may not provide meaningful insight. In this connection, it is worth noting that the vast majority of the adjustments are related to accrual accounting issues such as allowances for doubtful loans. These concepts are less concrete than cash values like total revenue, interest expense and pension cost. Thus they are more vulnerable to restatement.

As in the case of Vallejo, Stockton's filing is closely associated with general fund exhaustion. Although the city reported a \$12 million general fund balance at the end of fiscal 2011 – just prior to its default – this amounted to less than 7% of annual expenditures.

City of Stockton - All Governmental Funds				
Statement of Revenues, Expenditures and Changes in Fund Balances				
Total Governmental Funds				
	2008	2009	2010	2011
Revenues				
Taxes				
Property	63,998,000	58,640,000	45,549,000	41,051,105
In lieu of sales tax	10,164,000	9,823,000	7,087,000	8,118,132
Utility user	30,861,000	30,854,000	30,717,000	30,993,997
Sales (levied by City)	9,409,000	7,921,000	7,652,000	7,875,429
Franchise fees	11,537,000	11,608,000	11,354,000	11,502,735
Business license	10,772,000	9,699,000	9,717,000	9,855,031
Hotel/motel room	2,287,000	1,962,000	1,749,000	1,798,740
Document transfer	686,000	702,000	559,000	583,418
Other	246,000	234,000	203,000	154,983
Licenses and permits	5,273,000	4,335,000	4,257,000	3,584,311
Federal grants and subsidies	13,617,000	12,976,000	26,034,000	33,243,873
Other shared revenue (sales and use tax levied by state)	36,098,000	31,245,000	28,856,000	30,060,798
Other governmental	59,976,000	53,498,000	47,779,000	47,929,416
Charges for services	55,244,000	31,462,000	26,174,000	21,261,669
Fines and forfeitures	3,321,000	4,499,000	5,090,000	3,538,020
Use of money and property	12,922,000	13,234,000	11,962,000	14,966,292
Investment income:				
Interest income	13,100,000	11,375,000	5,352,000	1,338,707
Refunds and reimbursements	4,253,000	4,113,000	5,186,000	9,789,326
Miscellaneous	8,515,000	13,429,000	8,449,000	7,594,326
Total revenues	352,279,000	311,609,000	283,726,000	285,240,308
Expenditures				
Current:				
General government	22,285,000	24,272,000	21,818,000	30,900,316
Public safety	168,372,000	163,339,000	152,714,000	152,526,746
Public works	18,464,000	16,113,000	14,029,000	13,528,150
Library	13,432,000	12,485,000	11,041,000	10,252,107
Parks and recreation	27,185,000	22,376,000	17,948,000	19,669,013
Capital outlay	135,071,000	105,384,000	84,194,000	66,974,739
Debt service:				
Principal retirement	1,017,000	3,973,000	11,739,000	22,661,216
Cost of issuance	777,000	99,000	846,000	0
Interest and fiscal charges	10,771,000	11,938,000	12,523,000	12,705,728
Total expenditures	397,374,000	359,979,000	326,852,000	329,218,015
Excess (Deficiency) of Revenues over Expenditures	(45,095,000)	(48,370,000)	(43,126,000)	(43,977,707)
Total Other Financing Sources (Uses)	53,608,000	7,122,000	39,911,000	4,869,469
Special Items	0	(8,736,000)	(288,000)	3,269,612
Net Change in Fund Balances	8,513,000	(49,984,000)	(3,503,000)	(35,838,626)
Prior Period Adjustments	0	0	0	(109,666,067)
Beginning Fund Balances	303,721,000	312,234,000	262,250,000	258,748,200
Ending Fund Balances	312,234,000	262,250,000	258,747,000	113,243,507
Pension Contributions	17,715,000	20,512,027	21,110,516	21,030,435
Annual Revenue Change (Interest + Pension) / Revenue	8.09%	-11.54%	-8.95%	0.53%
		10.41%	11.85%	11.83%
<i>Source: CAFRs</i>				

City of Stockton - General Fund						
Statement of Revenues, Expenditures and Changes in Fund Balances						
General Fund						
	2008	2009	2010	2011	2012	2013
Revenues						
Taxes						
Property	37,077,000	33,030,000	29,170,000	28,318,427	26,375,894	25,988,000
In lieu of sales tax	10,164,000	9,823,000	7,087,000	8,118,132	8,392,001	9,937,924
Utility user	30,861,000	30,854,000	30,717,000	30,993,997	31,504,354	31,943,600
Sales (levied by City)						
Franchise fees	11,537,000	11,608,000	11,354,000	11,502,735	12,464,835	11,611,700
Business license	10,134,000	9,197,000	9,289,000	9,249,774	8,915,457	8,900,000
Hotel/motel room	2,287,000	1,962,000	1,749,000	1,798,740	1,932,630	1,940,000
Document transfer	686,000	702,000	559,000	583,418	603,313	495,000
Other	2,000	1,000	2,000			
Licenses and permits	377,000	641,000	392,000	339,636	395,949	370,109
Federal grants and subsidies		467,000	55,000	44,417		
Other shared revenue (sales and use tax levied by state)	31,900,000	27,522,000	25,623,000	26,550,862	29,504,817	29,696,242
Other governmental	24,872,000	25,299,000	27,160,000	26,370,481	27,624,762	31,872,634
Charges for services	10,213,000	11,894,000	13,043,000	10,763,721	1,907,657	1,890,668
Fines and forfeitures	3,302,000	4,492,000	5,045,000	3,452,493	1,729,835	1,115,605
Use of money and property	2,462,000	3,669,000	7,082,000	7,417,175	6,651,317	
Investment income:						
Interest income	1,316,000	1,126,000	888,000	(387,403)	260,885	(6,328)
Net increase (decrease) in value of investments	302,000	593,000	178,000			
Refunds and reimbursements	3,709,000	3,583,000	2,300,000	9,092,383	872,486	298,596
Miscellaneous	6,086,000	10,763,000	6,091,000	1,075,037	(291,469)	(60,500)
Total Revenues	187,287,000	187,226,000	177,784,000	175,284,025	158,844,723	155,993,250
Expenditures						
Current:						
General government	15,089,000	13,871,000	11,469,000	12,665,758	16,812,203	16,512,454
Public safety	143,955,000	141,427,000	133,901,000	134,539,420	123,753,893	115,287,031
Public works	13,936,000	11,965,000	3,541,000	3,515,999	7,438,423	6,805,947
Library			10,695,000	9,937,259	3,977,759	3,907,000
Parks and recreation	8,904,000	6,724,000	15,814,000	17,323,254	10,374,653	8,742,603
Capital outlay	86,000	46,000	60,000	158,851	500,000	575,000
Debt service:					3,013,468	978,560
Principal retirement						
Cost of issuance	30,000	99,000	177,000			
Interest and fiscal charges						
Contingency					850,000	2,000,000
Total expenditures	182,000,000	174,132,000	175,657,000	178,140,541	166,720,399	154,808,595
Excess (Deficiency) of Revenues over Expenditures	5,287,000	13,094,000	2,127,000	(2,856,516)	(7,875,676)	1,184,655
Total Other Financing Sources (Uses)	(11,198,000)	(7,097,000)	(2,392,000)	7,053,572	1,578,515	836,528
Special Items		(6,340,000)	(4,793,000)			
Net Change in Fund Balances	(5,911,000)	(343,000)	(5,058,000)	4,197,056	(6,297,161)	2,021,183
Prior Period Adjustments			5,124,000	(15,088,027)		
Beginning Fund Balances	28,992,000	23,081,000	28,205,000	8,059,178	12,256,234	4,342,349
Ending Fund Balances	23,081,000	22,738,000	23,147,000	12,256,234	5,959,073	6,363,532

Sources: CAFRS, 2012-13 Adopted Budget (http://www.stocktongov.com/files/2012-13_Final_Budget.pdf); 2012-13 Budget Update 2Q (<http://www.stocktongov.com/clerk/granicusagendas/citycouncil/20130319.pdf>)

San Bernardino

The City of San Bernardino filed a Chapter IX petition on August 1, 2012 (case number 6:12-bk-28006-MJ) after defaulting on a general fund debt service payment due July 20, 2012.

Citing the exhaustion of the city's general fund and an estimated fiscal year 2013 general fund deficit of \$45.8 million, San Bernardino staff recommended that the city declare bankruptcy and adopt an emergency budget that deferred debt service payments, retiree health contributions and other items. Staff argued that these steps were necessary to meet the City's payroll on August 15. The affected obligations included Taxable Pension Obligation Bonds Series 2005-A and Refunding Certificates of Participation issued in 1999. All city obligations were insured so the defaults did not directly affect bondholders. The affected insurers, Ambac and National Public Finance Guarantee filed objections to the City's bankruptcy filing.

Ultimately, the City defaulted on a July 20, 2012 pension obligation bond debt service payment but appears to be continuing to perform on its COPs⁶. The City has defined the Pension Bond as part of its overall pension expense, which it has chosen to defer.

As of this writing, the bankruptcy case is still being litigated, so the ultimate outcome is unknown. The court docket and media reports (summarized by Shafroth, 2012-2013) suggest that the San Bernardino case is particularly contentious for a number of reasons (Reid, Podkul & McNeill, 2012):

⁶ An August 15, 2012 court filing on behalf of National said that the city had informed the insurer it would not be making an \$861,000 payment due that date. However, the EMMA system shows no payment default for the affected bond.

- The City skipped the 60-day AB 506 creditor mediation process before filing, arguing that it has a right to do so because the City Council declared a fiscal emergency.
- The City suspended payments to CalPERS at the time of its filing.
- A total of 51 city employees (about 4% of the work force) retired in the three months prior to the filing, receiving \$2 million for unused sick and vacation time

In October 2012, the SEC announced an informal investigation into San Bernardino's finances. In January 2013, both the interim city manager and finance manager resigned. Finally, in March 2013, the State Controller's office accused the City of improperly transferring \$529 million in former Redevelopment Agency assets to the San Bernardino Economic Development Corporation. Given these troubles, the City is very likely to face very large legal costs associated with the bankruptcy process. This outcome may deter other cities from filing Chapter IX petitions.

The accompanying tables show San Bernardino's total government and general fund revenues and expenditures leading up to the bankruptcy filing. As of this writing, the 2012 CAFR has not appeared; incomplete data was obtained from budget documents. Since budget documents do not use the modified accrual basis of accounting employed in CAFRs the 2012 numbers are likely to vary somewhat from the final audited amounts that will appear in the CAFR.

For the general fund, we provide both the original 2013 budget estimates presented to the Council in July 2012 when it authorized the Chapter IX filing and estimates presented in February 2013 as part of the 2013-14 budget presentation. The projected deficit shrank from \$45.8 million to \$6.3 million. While most of this change is the result of the city's pendency plan implementation, the original deficit calculation appears to have been pessimistic.

City of San Bernardino - All Governmental Funds								
Statement of Revenues, Expenditures and Changes in Fund Balances								
		All Governmental Funds						
		2008	2009	2010	2011	2012 unaudited	2013 revised	
Revenues								
	Taxes	135,605,049	138,027,508	120,443,480	123,896,615			
	Licenses and permits	11,116,513	10,048,833	8,796,052	8,516,516			
	Impact fees	5,268,475	1,065,305	2,036,352	618,030			
	Fines and forfeitures	3,880,674	4,721,725	5,850,072	2,338,684			
	Investment Income	8,413,713	5,708,816	4,191,386	2,485,053			
	Intergovernmental	36,683,544	37,970,647	42,318,633	48,775,893			
	Charges for services	13,465,003	11,020,644	10,932,430	12,886,715			
	Other	6,803,988	6,213,253	8,603,639	8,925,459			
	Total revenues	221,236,959	214,776,731	203,172,044	208,442,965	197,344,437	206,698,054	
Expenditures								
	Current:							
	General Government	24,983,025	23,468,564	23,815,033	17,234,673			
	Public Safety	100,534,357	105,613,213	96,130,768	101,657,184			
	Streets	27,129,532	25,150,386	20,267,012	24,433,688			
	Culture and Recreation	9,480,585	6,842,658	5,228,540	6,579,287			
	Community Development	9,363,769	11,777,679	12,817,428	15,901,097			
	Community Service	6,914,615	4,329,133	9,150,336	10,266,256			
	Economic Development	18,165,689	22,276,482	34,168,768	29,735,854			
	Debt service:							
	Principal retirement	10,759,184	11,223,004	12,285,742	12,627,234			
	Interest and fiscal charges	13,146,478	14,101,348	13,420,944	13,745,859			
	Total expenditures	220,477,234	224,782,467	227,284,571	232,181,132	213,429,031	209,268,929	
	Excess (deficiency) of revenues over expends	759,725	(10,005,736)	(24,112,527)	(23,738,167)	(16,084,594)	(2,570,875)	
	Total other financing sources (uses)	7,325,889	5,873,219	10,191,103	29,139,935			
	Net change in fund balances	8,085,614	(4,132,517)	(13,921,424)	5,401,768			
	Beginning Fund Balances	208,819,975	216,905,589	212,773,072	198,851,648			
	Ending Fund Balances	216,905,589	212,773,072	198,851,648	204,253,416			
	Pension Contributions	13,696,000	15,923,153	15,763,362	15,817,310			
	Annual Revenue Change (Interest + Pension) / Revenue	12.13%	-2.92%	13.98%	-5.40%	2.59%	-5.32%	4.74%
						NA	NA	

Sources: CAFRs and FY 2013-14 Budget Message (<http://www.sbcity.org/civica/filebank/blobdload.asp?BlobID=14807>)

City of San Bernardino - General Fund								
Statement of Revenues, Expenditures and Changes in Fund Balances								
General Fund								
	2008	2009	2010	2011	2012 unaudited	2013 budget	2013 revised	
Revenues								
Taxes	100,443,781	94,030,428	83,518,733	85,428,247	87,209,311	89,775,443	89,326,711	
Licenses and permits	10,122,997	9,385,470	8,387,017	8,091,822	9,045,223	9,441,900	9,221,900	
Fines and forfeitures	1,499,214	2,250,060	3,379,135	2,283,426	1,904,360	2,104,300	2,204,300	
Investment income	1,441,416	736,536	789,438	609,721	794,158	733,000	733,000	
Intergovernmental	9,181,679	8,916,249	7,213,053	7,718,864	2,614,369	7,297,722	1,734,259	
Charges for services	6,388,869	6,419,995	6,509,637	7,423,815	6,008,881	6,898,400	5,499,000	
Other	4,181,440	4,122,007	6,051,308	4,341,597	6,317,022	4,173,400	7,461,600	
Total revenues	133,259,396	125,860,745	115,848,321	115,897,492	113,893,323	120,424,165	116,180,770	
Expenditures								
Current:								
General government	24,307,456	22,936,346	23,540,159	16,910,683	17,486,830	38,659,593	18,400,960	
Public safety	95,611,918	100,610,784	89,121,424	92,732,629	96,657,228	106,754,372	86,475,347	
Streets	9,666,812	8,280,754	7,356,336	8,318,267	8,127,566	9,971,142	8,259,249	
Culture and recreation	6,899,521	5,770,269	4,301,541	5,067,528	5,551,123	5,425,725	4,656,966	
Community development	2,482,040	2,039,117	-	-	-	-	-	
Community service	-	-	1,426,189	1,244,529	-	-	-	
Economic development	-	-	-	-	-	-	-	
Debt service:								
Principal	1,780,591	1,824,372	2,290,508	1,623,576	5,551,123	5,425,725	4,656,966	
Interest and fiscal charges	2,219,639	2,590,600	3,054,448	2,516,407	-	-	-	
Total expenditures	142,967,977	144,052,242	131,090,605	128,413,619	133,373,870	166,236,557	122,449,488	
Excess (deficiency) of revenues over expends	(9,708,581)	(18,191,497)	(15,242,284)	(12,516,127)	(19,480,547)	(45,812,392)	(6,268,718)	
Total other financing sources (uses)	7,264,977	4,746,772	12,944,258	10,924,230	8,708,983	4,829,642	10,371,754	
Net change in fund balances	(2,443,604)	(13,444,725)	(2,298,026)	(1,591,897)	(10,771,564)	(40,982,750)	4,103,036	
Beginning Fund Balances	18,596,648	16,153,044	2,708,319	410,293	(1,181,604)	(11,953,168)	(11,953,168)	
Ending Fund Balances	16,153,044	2,708,319	410,293	(1,181,604)	(11,953,168)	(52,935,918)	(7,850,132)	

Sources: CAFRs, Attachment A to City of San Bernardino Budgetary Analysis and Recommendations for Budget Stabilization (<http://www.sbcity.org/civica/inc/displayblobpdf2.asp?BlobID=13856>) and FY 2013-14 Budget Message (<http://www.sbcity.org/civica/filebank/blobload.asp?BlobID=14807>)

The cases presented above do not account for all recent California city defaults. The following two sections address defaults which were not accompanied by municipal bankruptcy filings.

Special District Bond Defaults

Over the past 20 years, there have also been a number of special assessment district bond defaults. While special assessment districts are often administered by a city, our focus is on city-wide tax revenue supported obligations so a detailed study of these situations is beyond the current scope. California cities experiencing recent payment defaults by special assessment districts include Borrego, Lathrop, Palmdale and Ione. Descriptions of these situations may be found on the MSRB EMMA system.

Redevelopment Agency Defaults

In June 2011, the legislature passed and the governor signed ABX1 26, a law that mandated dissolution of local redevelopment agencies (RDAs). The Supreme Court upheld the law and allowed the dissolutions to take effect on February 1, 2012 (California State Association of County Auditors, 2012). Redevelopment agency assets and liabilities mostly reverted to the cities and counties that created them.

Many California cities took on significant amounts of bonded debt as a result of the dissolution act, but also began to receive incremental property tax revenues necessary to service them. While most of these transitions did not impact RDA bondholders, the cities of Hercules and Monrovia did experience temporary defaults.

On February 1, 2012, Hercules defaulted on \$2.4 million of interest payments on RDA Tax Allocation bonds. The default did not directly impact municipal bondholders because payment was made by Ambac, the agency's municipal bond insurer. Ambac filed suit against the city claiming it had failed to remit RDA related property tax collections to the bond trustee as required. Instead the proceeds were placed in a Pooled Cash Account (Hercules Redevelopment Agency, 2012). In March 2012, Ambac and the City settled the litigation with the City pledging two parcels of land to the insurer (Kearney, 2012). The City further agreed to place these two properties on the market, apparently to offset the \$4.05 million property tax remittance the city had failed to make earlier (City of Hercules, 2012).

The fact that city assets had to be sold to clear the RDA default situation shows that the Hercules episode is indicative of a fiscal insolvency. Although the City's population is slightly below 25,000, its fiscal indicators may be relevant to the larger cities in this study. Unfortunately, the City had yet to publish 2011 or 2012 CAFRs at the time of this writing.

The city's 2010 CAFR showed total governmental funds revenue of \$37,740,183 and interest expenses of \$10,268,495. Pension expenses were \$1,596,456. Interest and pension expenses thus accounted for 31.44% of revenue. 2010 total revenues were 10.04% below prior fiscal year revenues of \$41,667,224. The city's general fund balance was relatively healthy but it was running a substantial deficit.

The failure to file CAFRs on a timely basis is part of a larger financial management issue in Hercules. In May and November 2012, the State Controller's Office issued three audits highly critical of the city's fiscal controls. One report "found the City of Hercules' administrative and internal accounting control deficiencies to be serious and pervasive." (State Controller's Office, 2012a). These insufficient controls may explain why RDA tax revenues could be directed away from debt service, thereby subjecting the city to costly litigation.

On June 1, 2012, Monrovia failed to redeem \$11,750,000 in maturing RDA bonds. On February 22, 2013, the City paid off the overdue principal with 12% interest (which included a 3% default penalty rate). Funds to pay off the defaulted bonds came from the proceeds of a refunding issue floated by the city (EMMA, 2013). According to news accounts, the bonds were not refunded upon maturity because state law did not permit it (Smith, 2012). On June 27, 2012, Governor Brown signed AB 1484 which specifically allowed successor agencies to issue RDA refunding bonds.

City Fiscal Emergencies

Press reports indicate that 13 California cities have declared fiscal emergencies of which eight have more than 25,000 residents. The cities are listed below.

City Declaring Emergency	2010 Population
Arvin	19,304
Atwater	28,168
Culver City	38,883
El Monte	113,475
Fairfield	105,321
Grover Beach	13,156
La Mirada	48,527
Lancaster	156,633
Monrovia	36,590
Riverbank	22,678
San Fernando	23,645
Stanton	38,186
Tehachapi	14,414
<i>Sources: Buchanan (2012), Garcia (2013), Taxis and Dreier (2012), White (2012)</i>	

Some press accounts suggest that these declarations constitute an official notification to the municipal bond market (see, for example, Buchanan, 2012). But the MSRB EMMA system does not have a section for fiscal emergency declarations, and it does not appear that the cities in question have made event disclosures to this effect.

Instead, these declarations should be understood in the context of the state constitution. Normally, tax increases may only be approved by voters during a general election. However, Article 13 Section 2 permits one exception. If the city council (or equivalent legislative body in another local agency) unanimously declares a fiscal emergency, a special election may be held to consider a tax increase.⁷

Fiscal emergency declarations may also be employed to alter collective bargaining agreements and other contracts. Although this prerogative is not specifically granted in the

⁷ Section 13B also allows a local agency to exceed its appropriation limit by declaring a fiscal emergency.

state constitution, Holtzman, Dickey & Cikes (2011) identify cases in which it has been invoked by local agencies and supported by courts.

Consequently, fiscal emergency declarations may be seen as a way of balancing budgets or avoiding spending reductions when a regular election is not imminent. A fiscal emergency declaration is thus not necessarily a precursor to default or bankruptcy.

Concluding Comments

General fund exhaustion – a factor not considered in the Depression-era survey seems to be a significant driver of recent city bankruptcies and their attendant bond defaults. Other factors accounting for recent default activity such as adverse court judgments and the dissolution of redevelopment agencies should be less relevant for the purpose of modeling major city defaults. Cities below the 25,000 population threshold, like Mammoth Lakes, are more vulnerable to lawsuit-driven defaults or bankruptcies because their revenue base is less able to absorb multi-million dollar awards. The two RDA-related defaults appear to be, at least in part, transitional issues.

While we have enumerated a significant number of payment difficulties in this section, it is important to put these in context. California currently has 480 cities, and it has had at least 200 cities through most of its history as a state. With the exception of 1933, it appears that the municipal default rate has not exceeded 1% in any given year. In the vast majority of years, the rate has been zero.

City Bond Defaults and Bankruptcies outside California

Since California's city bond default experience is relatively limited, we supplement the previous chapter with information regarding cities in other states. First, we consider two high profile defaults from the 1970s and then we provide an overview of municipal credit quality issues in other states over the last 25 years.

The 1970's

Between the Depression and Vallejo's 2008 bankruptcy, we are aware of only two defaults by major US cities – both of which occurred in the 1970s. These two cases are described in some depth in a 1984 report by the US Advisory Commission on Intergovernmental Relations entitled *Bankruptcies, Defaults and Other Local Government Financial Emergencies*. After summarizing these narratives we review these cities' financial statistics as reported by the US Census.

In November 1975, New York City temporarily suspended debt service payments on short term obligations after the state legislature passed a Moratorium Act shielding the city from bondholder lawsuits. The ACIR narrative attributes the city's fiscal crisis to persistent operating deficits starting in fiscal 1971. The city became increasingly dependent on short term borrowing to fund its operations. After declaring the moratorium, New York State took a number of actions to improve the city's finances including the imposition of a financial emergency control board, provision of short term loans and establishing the Municipal Assistance Corporation to issue new bonds on behalf of the city. As a result of these actions, the city's budget was balanced and short term debt was replaced by longer term obligations.

In December 1978, Cleveland defaulted on a \$15 million bond anticipation note as part of a larger fiscal crisis that enveloped the Ohio city. The ACIR report attributed Cleveland's fiscal distress to persistent operating deficits as well as poor accounting and fiscal management practices. A June 30, 1978 audit revealed that the city had used capital funds and other restricted funds to pay general fund obligations and that the city's financial records were in disarray. Cleveland's bond ratings were then lowered and suspended, preventing the city from rolling over its bond anticipation notes, thereby triggering the December default. The city's default was cured after the state auditor declared a fiscal emergency (which allowed for the provision of state loans) and the city raised its income tax, producing surpluses in fiscal 1980 and 1981.

The Census Bureau has been reporting city financial data since 1904. Report format, scope and measurement definitions have changed over time, so data are not necessarily comparable across periods. Earlier, we used Census data to obtain independent variables for the Depression era default modeling. While two defaults are not sufficient for to create a model, it may be useful to see how the defaulting cities ranked against peers according to selected metrics.

During the 1970s, the Census published detailed financial statistics for cities with population greater than about 300,000 – a class that includes New York and Cleveland. In fiscal 1975-76, New York had the 11th highest interest-to-revenue ratio out of 48 cities in that year's survey. In 1978-79, Cleveland ranked 8th out of 46. While the two defaulters were not the most indebted relative to revenue, they did place in the upper quartile of peer cities on this measure of debt burden.

The Annual Revenue Change rankings were less indicative – New York was 30th and Cleveland 21st in their respective years of default. In this case, a lower ranking (i.e., a lower revenue change) should reflect greater risk, but in both cases, the impacted cities were near the middle of the distribution.

The Census also reported Employee Retirement Expenditure, but the numbers do not appear to be accurate. For example, Cleveland’s Employee Retirement Expenditure was listed as 0 for 1978-79. Several other cities in the survey also reported zero values. According to ACIR (1984), Cleveland did not have its own employee retirement system during the 1970s, apparently relying on the Ohio Public Employee Retirement System (OPERS) to manage its pension obligations. It appears that employer contributions to state systems were excluded from the Census figures. In 1975-1976, New York’s Employee Retirement Expenditure ranked 15th of 42 cities with non-zero values.

While General Fund balances don’t appear in the Census data, General Revenues and General Expenditures are reported thereby allowing a calculation of General Fund surpluses or deficits. While the Census shows New York running a surplus during the year of its default, the city experienced substantial deficits in three of the four prior fiscal years – consistent with the ACIR report. Cleveland had a substantial deficit in its default year and in each of the four preceding years. In three of those years, the City’s deficits exceeded 10% of revenue, placing it near the bottom of peer cities in the measure of general fund balance relative to revenue.

More Recent Experience in Other States

Contemporary data related to municipal bond defaults from other states could potentially be used for default probability modeling. However, there have been few relevant defaults over the last 25 years. Since media reports often conflate state takeovers and bankruptcies with defaults, we include these kinds of municipal fiscal crises in this section. In the next chapter, we use some of these observations to supplement California data to perform additional logistic regression analysis.

Although the list includes 32 municipal bankruptcies, many involved small towns that did not issue municipal bonds. In other cases, a larger city filed a petition but the case was dismissed without a default or rescheduling of debt. The list also includes a number of defaults that occurred without a bankruptcy filing as well as several state takeovers, most of which occurred in Michigan. Since takeover situations may have resulted in defaults absent intervention from a higher level of government, the financial statistics of cities requiring takeovers might also be considered for modeling purposes.

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Allen Park, MI (2012) Population: 28,210 State appointed emergency financial manager	Yes	<p>In November 2009, the city issued \$31 million in long term General Obligation bonds to finance the creation of a movie studio which failed. Debt service on these bonds has contributed to persistent general fund deficits.</p> <p><i>Sources:</i> <i>Report of the Allen Park Financial Review Team (2012).</i> http://www.michigan.gov/documents/treasury/AllenPark-ReviewTeamReport- 8-8-12_417419_7.pdf Burton. Paul (2013 Mar 20). Michigan Treasurer: Orr's the Right Man. <i>The Bond Buyer</i>. http://search.proquest.com/newsstand/docview/1317819784#</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Alorton, IL (2005) Population: 2,549 Bankruptcy	No	<p>The closure of Alcoa Aluminum in the 1960s and the loss of the Cahokia Downs Race Track in 1978 prompted the economic decline of the Village. Lawsuits from various individuals compounded fiscal issues, leading to the Village's bankruptcy.</p> <p><i>Source:</i> Levin, Richard, Jonathan Solomon & Campell Ayapong. (2011). Some Causes of Municipal Distress and Bankruptcy. http://html.documation.com/cds/NCBJ2011/assets/PDFs/VI.D.pdf</p>
Benton Harbor, MI (2013) Population: 10,047 State appointed emergency financial manager	No	<p>Formerly an economically vibrant manufacturing center of 20,000, Benton Harbor's per capita income is now roughly \$10,000 and 60% of its population is on public assistance.</p> <p><i>Sources:</i> <i>Report of the Benton Harbor Financial Review Team</i> (2010). http://www.michigan.gov/documents/treasury/BentonHarbor-ReviewTeamReport-1-29-10_417426_7.pdf</p> <p>Mahler, Jonathan (2011 Dec. 15). Now that the factories are closed, it's tee time in Benton Harbor, Mich. <i>New York Times</i>. http://www.nytimes.com/2011/12/18/magazine/benton-harbor.html?pagewanted=all</p>
Bridgeport, CT (1991) Population: 141,719 Bankruptcy	Yes	<p>Dismissed. Court found that the city was not insolvent. No interruption in payments to creditors. The state backed "\$53 million in bonds to balance Bridgeport's books".</p> <p><i>Sources:</i> Mills, M. (2011). Bridgeport – Distressed but not insolvent. <i>Bankruptcy Blog</i>. http://business-finance-restructuring.weil.com/chapter-9/bridgeport-%E2%80%93-distressed-but-not-insolvent</p> <p>Duby, Christopher. (1995 Oct 9). Bridgeport finally sheds Financial Review Board. <i>Fairfield County Business Journal</i>. http://search.proquest.com/docview/216380696</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Brighton, AL (2011) Population: 2,947 Default	Yes	<p>The city was unable to make a \$22,783 interest payment or comply with a mandatory redemption. The default was attributed to the failure of many residents to pay bills due to the city. Some residents had recently been laid off by Jefferson County – also in bankruptcy.</p> <p><i>Sources:</i> Sigo, Shelly (12 Aug 2011). Brighton Ala. Defaults on General Obligation Warrants from 2003. The Bond Buyer. http://search.proquest.com/docview/884326244</p>
Brooklyn, IL (2003) Population: 626 Bankruptcy	No	<p>The small town was suffering from declining population, internal corruption, and the closure of several strip clubs. The strip clubs had provided much of the tax base for the village but were shut down. Initial estimates showed the town had \$100,000 in assets and \$500,000 in debts -- but a 2006 news report places total debt at \$1,600,000.</p> <p><i>Sources:</i> Shaw, Michael (2003 Oct 15). Brooklyn Goes Broke, Files for Bankruptcy. <i>St. Louis Post-Dispatch</i>. http://search.proquest.com/docview/402336061</p> <p>Hollinshed, Deneice (2006 Dec. 29). Allegations of corruption cast pall over Brooklyn. <i>St. Louis Post-Dispatch</i>. http://search.proquest.com/docview/403054028</p>
Camden, NJ (1999) Population: 79,904 Bankruptcy	Yes	<p>An act of brinksmanship during a debate over a potential state takeover. Petition withdrawn shortly after filing. No default.</p> <p><i>Source:</i> Couloumbis, Angela and Dwight Ott (1999 July 25). Camden’s Bankruptcy Drama Ends But Self-sufficiency is still far off. <i>The Philadelphia Inquirer</i>. http://articles.philly.com/1999-07-25/news/25521825_1_mayor-milton-milan-camden-residents-aid-agreement</p>
Camp Wood, TX (2005) Population: 822 Bankruptcy	Yes	<p>Camp Wood Convalescent Center did not generate sufficient revenue to service certificates of obligation. “The municipality refinanced its debt with bonds and other obligations, but was unable to make payments on its debt due to continued underperformance of the Convalescent Center.”</p> <p><i>Source:</i> Levin, Solomon & Ayapong (2011).</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Central Falls, RI (2011) Population: 19,376 Bankruptcy	Yes	<p>The city was placed into receivership in 2010 under a Financial Stability Act passed by the state legislature. The receiver filed a Chapter IX bankruptcy petition in 2011. Central Falls had about \$21 million of outstanding general-obligation bonds at the time of its filing and faced a \$4.8 million budget gap for fiscal 2012. The city continued to service its bonds in bankruptcy, but raised health insurance deductibles and copayments for city employees and retirees. By altering collective bargaining agreements, the city was able to emerge from bankruptcy within a year. Avoidance of default was credited to a 2011 state law giving bondholders the right to place liens on Rhode Island municipal tax revenues.</p> <p><i>Sources:</i> Bidgood, Jess (2012 Sep. 6). Plan to End Bankruptcy in Rhode Island City Gains Approval. <i>New York Times</i>. http://www.nytimes.com/2012/09/07/us/central-falls-ri-to-emerge-from-bankruptcy.html City of Central Falls (2012 June 30). Continuing Disclosure Report Rhode Island General Obligation Debt. http://emma.msrb.org/ER644731-ER500139-ER902895.pdf Nolan, Kelly (2011 Aug 1). Rhode Island City Files for Bankruptcy. <i>Wall Street Journal</i>. http://search.proquest.com/docview/880200621</p>
Copperhill, TN (1988) Population: 450 Bankruptcy	No	<p>Factory closings and a declining population left the town with no way to pay even the interest on a \$400,000 construction loan for a sewage plant.</p> <p><i>Source:</i> Uzelac, Ellen (1991 June 17). A year after bankruptcy, Tenn. town was flooded. <i>The Baltimore Sun</i>. http://search.proquest.com/docview/407130678</p>
Detroit, MI (2013) Population: 706,585 State appointed emergency financial manager	Yes	<p>A long term population decline, political corruption and inflexible union contracts are cited as general causes for the city's secular fiscal decline. The financial review team identified insufficient cash, eight consecutive general fund deficits, long term liabilities including pension and OPEB obligations and bureaucratic inflexibility as causes for the state takeover.</p> <p><i>Sources:</i> Afford, Harry C. (2013 Mar 15). Long in decline, Detroit can't outrun its past. <i>The Philadelphia Tribune</i>.</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
		http://search.proquest.com/docview/1321681728 <i>Report of the Detroit Financial Review Team</i> (2013). http://www.michigan.gov/documents/treasury/Review-Team-Report-2-19-13_415662_7.pdf
East St. Louis, IL (1990) Population: 40,944 State Supervision	No	<p>City was declared to be financially distressed by the state, received a state loan and was placed under supervision by a Financial Advisory Authority. Subsequent bond issuance, starting in 1994, has taken place under state supervision - most recently by the Illinois Finance Authority.</p> <p><i>Source:</i> Harrison, Eric (1990 Aug. 9). East St. Louis: Illinois Bails Out Troubled City Close to Bankruptcy. Los Angeles Times. http://articles.latimes.com/1990-08-09/news/mn-375_1_east-st-louis.</p>
Ecorse, MI (2009) Population: 9,512 State appointed emergency financial manager	Yes	<p>City mayor and controller both arrested for corruption related to public works contracts. The financial review team noted four consecutive general fund deficits and a negative general fund balance in its report.</p> <p><i>Source:</i> <i>Report of the Ecourse Financial Review Team</i> (2009). http://www.michigan.gov/documents/treasury/Ecorse-ReviewTeamReport-8-19-09_417433_7.pdf Egan, Paul (2009 Sep 26). Bribery scandal rattles Ecorse: Mayor, controller arraigned on federal corruption charges. <i>Detroit News</i>. http://search.proquest.com/docview/404426108</p>
Flint, MI (2011) Population: 101,558 State appointed emergency financial manager	Yes	<p>Review team declared a fiscal emergency because the city was running persistent and increasing general fund deficits, had insufficient cash to meet short term obligations and lacked a credible plan for addressing its financial problems. The city was also under state emergency financial control from 2002 to 2006.</p> <p><i>Sources:</i> <i>Michigan Radio</i> (2011). 7 things to know about Michigan's financial emergency law. http://www.michiganradio.org/post/7-things-know-about-michigans-emergency-manager-law <i>Report of the Flint Financial Review Team</i> (2011 Nov 7). http://www.michigan.gov/documents/treasury/Flint-ReviewTeamReport-11-7-11_417437_7.pdf</p>
Gould, AR (2008) Population: 1,305 Bankruptcy	No	<p>The town owed more than \$900,000 to the IRS, the Arkansas Department of Finance and Administration, Arkansas Natural Resources, a mosquito-control</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
		<p>company, the Lincoln Country jail, and the US Agriculture Department's Rural Development agency in St. Louis. With assets totaling only \$300,000, Gould filed a Chapter IX petition.</p> <p><i>Source:</i> Hale-Shelton, Debra. Bankruptcy filed, tiny town hopes to rise again. <i>Northwest Arkansas Times</i>. http://freerepublic.com/focus/chat/2007387/posts</p>
<p>Harrisburg, PA (2011) Population: 49,528 Bankruptcy and Default</p>	<p>Yes</p>	<p>A failed incinerator project generated roughly \$300 million in city guaranteed debt, while the city relied on sewerage charges to offset a persistent general fund deficit. The city filed a Chapter IX petition in October 2011 but the filing was dismissed because it violated a state moratorium on certain municipal bankruptcies. The city has defaulted on three general obligation bond debt service payments since March 15, 2012.</p> <p><i>Sources:</i> Barnes, Tom (2011 Nov 24). Harrisburg's Petition for Bankruptcy Protection Rejected. <i>Pittsburgh Post-Gazette</i>. EMMA (2013). Continuing Disclosure for Harrisburg Refunding Notes. http://emma.msrb.org/IssueView/IssueDetails.aspx?id=F68DABD8F80120CF1482B5AD2CB1D695 Unkovic, Steve (2013). <i>Municipal Financial Distress: Causes and Solutions</i>. Bond Buyer Distressed Municipalities Conference. http://www.bondbuyer.com/media/pdfs/BBdistressed13-presentations-Unkovic-Municipal-Physical-Distress.pdf</p>
<p>Hillsdale, MO (2001) Population: 1,477 Bankruptcy</p>	<p>No</p>	<p>The city had over \$250,000 in debt and under \$100,000 in assets. Upon being ordered to pay \$88,000 to an officer after he slipped on ice the city filed chapter 9.</p> <p><i>Source:</i> O'Neil, Tim (2001 Dec 11). Hillsdale files for bankruptcy after order for injury award; village has been scraping by, lawyer says. <i>St. Louis Post-Dispatch</i>. http://search.proquest.com/docview/402025863</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Inkster, MI (2012) Population: 25,111 State appointed emergency financial manager	No	<p>Financial review team cited negative cashflow, unrealistic budgets and high debt levels. City laid off 20% of its police force shortly before the state takeover.</p> <p><i>Sources:</i> Hullett, Sarah (2011 Dec 14). <i>Michigan Town Grapples with Shrinking Public Sector</i>. National Public Radio. http://www.npr.org/2011/12/14/143705814/michigan-town-grapples-with-shrinking-public-sector Report of the Inkster Financial Review Team (2012). http://www.michigan.gov/documents/treasury/Inkster-ReviewTeamReport-3-1-12_417444_7.pdf</p>
Kendleton, TX (2001) Population: 466 Bankruptcy	No	<p>Texas officials seized the town's bank account and withdrew what cash was left -- about \$18,600. "The seizure of the city's money was based on a 1997 court ruling that Kendleton owed the state \$660,000 as its portion of traffic fines collected between 1990 and 1996."</p> <p><i>Source:</i> Hanson, Eric (2001). Kendleton files for bankruptcy. <i>Houston Chronicle</i>. http://search.proquest.com/docview/395856622</p>
Kinloch, MO (1994) Population: 2,699 Bankruptcy	No	<p>City population decreased from 2699 to 449 between 1990 and 2000 due to the airport buying up homes in the town as part of its expansion. The bankruptcy petition was a response to a dispatching firm's move to garnish the city's income from sales tax.</p> <p><i>Source:</i> Bryant, Tim. Bankruptcy will help Kinloch, Mayor Says. <i>St. Louis Post-Dispatch</i>. http://search.proquest.com/docview/303919247</p>
Lipscomb, AL (1991) Population: 2,800 Bankruptcy	Yes	<p>Declining tax base. Defaulted in 1985. "Defaulted on \$832,000 GO bond and \$353,000 GO refunding warrant from Farmer's Home Administration." The FmHA filed suit in 1987 and 1988 with awards totaling \$120,000. "In 1991, FmHA filed suit again asking for city to turn over keys to City Hall, 3 police cars and its 1976 and 1954 fire trucks."</p> <p><i>Source:</i> Deal, Keren (2010). <i>Municipal Bankruptcy in Alabama</i>. http://www.gfoaa.org/docs/CGAT/CGAT%20Muni%20BR%20Presentation.pdf</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Macks Creek, MO (2000) Population: 267 Bankruptcy	No	<p>Macks Creek was financing operations using traffic fines (they accounted for 75-85% of revenue). This was deemed excessive and the state enacted the so-called Macks Creek Law in 1995, capping the maximum revenue permissible from fines at 45% (excess going to county schools). A 1997 state audit found "major financial problems". After this was revealed, almost every town official resigned.</p> <p><i>Source:</i> Frankel, Todd C. (2009 May 17). Speed trap law is full of loopholes in Macks Creek, the town that inspired the measure, has passed into oblivion. <i>St. Louis Post-Dispatch</i>. http://search.proquest.com/docview/403205030</p>
Marion, MS (2007) Population: 1,305 Bankruptcy	No	<p>Filed bankruptcy petition to avoid paying a \$400,000 judgment won by neighboring Meridian, MS for waste water treatment. Dismissed.</p> <p><i>Source:</i> Brown, Ida (2009 July 27). Editorial board. <i>Meridian Star</i>. http://meridianstar.com/local/x681086861/Editorial-Board</p>
Marshall Creek, TX (2006) Population: 430 Bankruptcy	No	<p>Loss of contract to patrol Marshall Creek Park and the loss of federal police grants resulted in bankruptcy consolidation with neighboring Roanoke, Texas.</p> <p><i>Source:</i> McGowen, Lorraine (2011). Presentation to Sovereign & Municipal Debt Roundtable. http://bankrupt.com/DI2011/Docs/doc/0840McGowen.pdf</p>
McCurtain Municipal Authority, OK (2007) Population: 466 Bankruptcy	Yes	<p>The authority, which provides water and sewer services to the town, lost a contractor lawsuit over a disputed bill. Case dismissed after the Authority reached an agreement with the contractor.</p> <p><i>Sources:</i> Levin, Solomon & Agyapong (2011), McGowen (2011). U.S. Bankruptcy Court Eastern District of Oklahoma (Okmulgee) (2007). Court Documents related to Bankruptcy Petition #: 07-80363. Available on PACER at https://ecf.okeb.uscourts.gov/.</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Menasha, WI (2009) Population: 17,442 Default	Yes	<p>The city defaulted on debt from the spiraling construction costs associated with a steam plant that would not be profitable. Three years later, agreements were made allowing the city to repay \$17.5 million in debt over the course of twenty years.</p> <p><i>Source:</i> King, Michael (2012 Mar 6). Menasha steam plant debacle, uncertainty wind down. <i>Appleton Post-Crescent</i>. http://www.postcrescent.com/article/20120316/APC030208/120305182/Menasha-steam-plant-debacle-uncertainty-wind-down</p>
Millport, AL (2004) Population: 1,000 Bankruptcy	Yes	<p>Defaulted on \$1.3 million of general fund obligations, as well as a \$2 million loan from the US Department of Agriculture to improve the town's water and sewer systems. Default attributed to declining population and employment opportunities as well as financial mismanagement under previous administration.</p> <p><i>Sources:</i> Deal (2010).</p> <p>US Fed News Service. Information issued by U.S. Attorney's Office for the Northern District of Alabama on April 15; US Settles Action to Appoint Receiver for Millport, Alabama's Sewer, Water System. http://search.proquest.com/docview/472147666</p>
Moffett, OK (2006) Population: 400 Bankruptcy	No	<p>Town was recognized as a speed trap and no longer allowed to finance its operations using traffic fines as its primary revenue source. "Court records show that the town owe[d] nearly \$200,000 in secured and unsecured claims from nearly 50 businesses. Moffett generate[d] only about \$20,000 in annual sales-tax revenues, a 2004 filing from the state Auditor's Office shows."</p> <p><i>Sources:</i> Walton, Rod. Moffett files for bankruptcy. <i>Tulsa World</i>. http://search.proquest.com/docview/400285035</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Muldrow, OL (2005) Population: 3,104 Bankruptcy	No	<p>“The immediate cause of the chapter IX filings was the likelihood that Muldrow faced significant fines and penalties from ODEQ [Oklahoma Department of Environmental Quality] for past and ongoing violations of the Clean Water Act and its Oklahoma counterpart.”</p> <p><i>Source:</i> Levin, Solomon & Ayapong (2011).</p>
North Bonneville, WA (1991) Population: 350 Bankruptcy	No	<p>In the 1970s the town was condemned and relocated to allow for construction of a dam. The Army Corp of Engineers moved the town and 20 years later claimed that it still owed \$365,000 in maintenance and operations costs for the municipal facilities. Due to a declining tax base since the move and city assets totaling only \$258,000, the town filed for bankruptcy. The issue was settled with the signing of the 1993 Defense Appropriations Act. The "measure calls for the corps to cancel the city's debt, convey title to the town for the relocation lands and facilities and clean up a hazardous waste site on Hamilton Island, a peninsula on the town's south side. In return, North Bonneville agreed to accept the facilities "as is." The city also loses its right to sue the corps for failure to perform according to the terms of the relocation agreement."</p> <p><i>Source:</i> Senior, Jeanie. North Bonneville's fight with the army ends. <i>The Oregonian</i>. http://search.proquest.com/docview/416587177</p>
North Courtland, AL (1992) Population: 1,000 Bankruptcy	No	<p>Court awarded \$100,000 to a former employee of the city on a discrimination claim. Plaintiff began garnishing the city's tax revenues. City could not pay the judgment which amounted to over one third of the town's annual \$290,000 revenue.</p> <p><i>Source:</i> Deal (2010)</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Ozan, AK (1995) Population: 69 Bankruptcy	No	<p>City in decline due to a water system that yielded undrinkable water. The town financed the construction of a new water system, but two unexpected stop orders delayed the project. The contractor sued for \$55,000 in lost income and the town did not have the resources to pay or fight the suit. "The water system is being paid for with a \$291,700 loan from the Rural Economic Development Corp., some \$645,782 in grants from the Arkansas Industrial Development Commission and \$7,300 from town coffers."</p> <p><i>Source:</i> Copeland, Larry (1995 July 23). Ozan's solution became problem. <i>Tulsa World</i>. http://search.proquest.com/docview/399523993</p>
Pontiac, MI (2009) Population: 59,515 State appointed emergency financial manager	Yes	<p>GM plant closings resulted in fewer jobs and a declining population. The financial review team cited persistent large general fund deficits and a deteriorating cash position.</p> <p><i>Source:</i> Holeywall (2012 May). Emergency Financial Managers: Michigan's Unwelcome Savior. <i>Governing</i>. http://www.governing.com/topics/mgmt/gov-emergency-financial-managers-michigan-municipalities-unwelcome-savior.html <i>Report of the Pontiac Financial Review Team</i> (2010). http://www.michigan.gov/documents/treasury/Pontiac-ReportToGovernor-6-23-08_417450_7.pdf</p>
Prichard, AL (1999/2009) Population: 28,633 Bankruptcy	No	<p>Dwindling population, persistent deficits and substantial pension obligations forced Prichard to file for bankruptcy in 1999. After emerging from bankruptcy in 2007, the city filed again in October of 2009, in an effort to further reduce pension payments. While the city's petition was dismissed in 2010, it drastically reduced pension benefits. Although the city does not have any municipal bonds listed on EMMA, the bankruptcy court docket indicates that it had a lease arrangement with Region's Bank at the time of its 1999 filing.</p> <p><i>Sources:</i> Chang, Semoon (2012). A tale of the Prichard (AL) pension program, <i>Pensions</i>, 17(2), 112-120. http://media.al.com/live/other/Prichard%20Pension%20Article%20Semoon%20Chang.PDF</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
		<p>Deal (2010).</p> <p>Heck, Hannah (2011). Solving Insolvent Public Pensions: The Limitations of the Current Bankruptcy Option. <i>Emory Bankruptcy Developments Journal</i> 28(1), 89-133. http://search.proquest.com/docview/923754470</p> <p>U.S. Bankruptcy Court Southern District of Alabama. Court Documents related to Bankruptcy Petition #: 99-13465. Available at http://ia600400.us.archive.org/26/items/gov.uscourts.alsb.49664/gov.uscourts.alsb.49664.docket.html.</p> <p>Watson, Douglas, Donna Handley and Wendy Hassett. Financial Distress and Municipal Bankruptcy: The Case of Prichard, Alabama. <i>Journal of Budget, Accounting and Financial Management</i>, 17(2), 129-150.</p>
Reeds Spring, MO (2002) Population: 465 Bankruptcy	No	<p>A 1998 lawsuit won by a woman who slipped on a city sidewalk left the city owing \$100,000 to her and \$25,000 in legal fees. The town was also running a deficit: its 2002 revenue of \$205,000 was \$30,000 less than expenditures.</p> <p>Source: Bengali, Shashank (2002 Nov. 17). Suit pushes town into bankruptcy. <i>Charleston Sunday Gazette-Mail</i>. http://search.proquest.com/docview/332227135</p>
Rio Bravo, TX (2002) Population: 5,553 Bankruptcy	No	<p>Town was unable to pay on a \$180,000 loan and \$800,000 owed to former police chief. 50% of residents not paying taxes or fees.</p> <p>Taylor, Erinn (2003 Sept. 18). Rio Bravo mayor wants taxes paid. <i>Laredo Morning Times</i>. http://madmax.lmtonline.com/textarchives/091803/s5.htm</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Scranton, PA (2012) Population: 76,089 Default	Yes	<p>On June 1, 2012 the City failed to make a required lease payment to the Scranton Parking Authority causing authority bonds to go into default. In addition, the city temporarily reduced employee salaries to the statutory minimum wage in order to conserve cash. Later in the year, the city's cash crisis was alleviated by state aid, a loan from a union-owned bank and proceeds from additional bond issues.</p> <p><i>Sources:</i> Shafroth, Frank (2012 July 13) <i>The Week that Was</i>. Singleton, D. (2012 Dec. 30) Scranton's Financial Crisis Tops 2012 News, <i>The Times-Tribune</i>, http://thetimes-tribune.com/news/scranton-s-financial-crisis-tops-2012-news-1.1422801. Wells Fargo Bank (2012 Aug. 17), <i>Notice of Defaults, Event of Default and Appointment of Receiver</i>. http://emma.msrb.org/EP678369-EP528610-EP929851.pdf.</p>
Tyrone, OK (2000) Population: 880 Bankruptcy	No	<p>"The Texas County town of about 880 has filed for Chapter IX bankruptcy, which will allow it to stave off claims and attorney's fees that now equal the town's annual budget of about \$150,000." The suits were filed by two police officers over the town's noncompliance with federal wage laws.</p> <p><i>Source:</i> <i>Oklahoma City Journal Record</i> (2000 Oct 6). Lawsuits force Oklahoma Panhandle town into bankruptcy. http://search.proquest.com/docview/259442432</p>
Vadnais Heights, MN (2012) Population: 12,302 Default	Yes	<p>The city issued \$27 million in bonds on behalf of an entity that was to build and operate a sports arena. The city was to then lease the facility "for a rental payment equal to its annual operating budget, which includes debt-service costs." The facility brought in significantly less revenue than expected and the city terminated its lease for 2013, triggering a default event.</p> <p><i>Source:</i> Shields, Yvette (2012 Sept 11). Minnesota City Cancels Sports Lease Backing \$27M of Bonds. http://www.bondbuyer.com/issues/121_176/moodys-downgrades-Vadnais-Heights-to-junk-status-1043840-1.html</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Warrens, WI (2010) Population: 366 Default	Yes	<p>The village of Warrens defaulted on general obligations as well a \$3.6 million sewer bond held by the state of Wisconsin. It is in forbearance until April 2013. The bonds were issued in part to finance infrastructure associated with a new hotel and water park complex. The development went into foreclosure before it could be completed, significantly impacting tax revenues.</p> <p><i>Sources:</i> Village of Warrens (2012). Financial Statements as of December 31, 2011. http://emma.msrb.org/ER586883-ER456192-ER858928.pdf Warrens Finance Committee Report (2012 June 20). http://www.co.monroe.wi.us/wp-content/uploads/2012/01/20120625100401063.pdf</p>
Washington Park, IL (2004) Population: 5,451 Bankruptcy	No	<p>The 2004 filing after losing an employee harassment lawsuit but was dismissed after the town's finances temporarily improved. The city filed again in 2009 claiming assets of less than \$50,000 and debts of over \$1 million. The second filing occurred after two village workers were convicted of embezzling a total of over \$300,000 from the town. Second filing was dismissed on the grounds that Chapter IX filings were not authorized under Illinois state law.</p> <p><i>Sources:</i> McGowen (2010). Suhr, Jim (2009 Aug. 4). Illinois village seeks bankruptcy protection. http://dailyreporter.com/2009/08/04/illinois-village-seeks-bankruptcy-protection/</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Westfall Township, PA (2009) Population: 2,500 Bankruptcy	No	<p>"Supervisors in rural Westfall Township., with annual revenues of about \$1 million, sought Chapter IX protection . . . to force negotiations on a \$20 million federal judgment granted to a developer. The compromise under the bankruptcy plan allows the township to make \$75,000 quarterly payments over 20 years, funded through a dedicated property tax hike that raises taxes by about \$200 a year for the average homeowner."</p> <p><i>Source:</i> McConnell, Steve (2010 Mar. 22). Westfall Township's first-in-the-state bankruptcy may not be Pennsylvania's last. <i>Scranton Times Tribune</i>. http://search.proquest.com/docview/458423160</p>
Westlake, TX (1997) Population: 250 Bankruptcy	No	<p>"Westlake's woes began this spring after [Ross] Perot -- whose family owns the 2,500-acre Circle T ranch that makes up more than half of Westlake -- was unable to reach agreement on development plans with city officials, notably former Mayor Scott Bradley. The flap ultimately resulted in two Westlake aldermen and Perot sympathizers being voted out of office. But before they left office, the aldermen removed Mr. Bradley from office, approved Mr. Perot's request that his property be disannexed and then approved a separate disannexation request by the owners of the Solana [office] complex." That office complex was responsible for 99% of the town's revenue. A state pool and local bank both temporarily froze the town's accounts while the disannexations were adjudicated. However, because the town had had \$1,895,321 in cash and only \$122,199 in outstanding obligations, the bankruptcy court dismissed the case.</p> <p><i>Sources:</i> <i>Wall Street Journal</i> (1997 June 10). Town near Dallas files for bankruptcy protection. http://search.proquest.com/docview/398562382</p> <p><i>In Re: Town of Westlake, Texas</i>. US Bankruptcy Court ND Texas (1997 July 25). http://www.leagle.com/xmlResult.aspx?page=2&xmldoc=19971071211BR860_1943.xml&docbase=CSLWAR2-1986-2006&SizeDisp=7</p>

City, State (Year) Population Type of Event	Bonds	Notes Source(s)
Westminster, TX (2000) Population: 390 Bankruptcy	No	<p>"Twice, the city of Westminster sought to declare for bankruptcy; the first effort, in 2001, failed after creditors rejected a payout plan. In early 2004, the state agencies agreed to relinquish their claims, provided that Westminster disincorporate."</p> <p><i>Source:</i> Collin County Station (2013). <i>Westminster, Texas History and Information.</i> http://collincountystation.com/westminsterh.html</p>
Winstonville, MS (1997) Population: 277 Bankruptcy	No	<p>"Court records show Winstonville has had financial problems for at least two decades, and filed for bankruptcy in 1997." "The town also owed \$323,759 to the USDA for a community facilities loan. After negotiations, the USDA agreed to let Winstonville pay \$100 to clear its obligation for the loan".</p> <p><i>Mississippi Business Journal</i> (2011 Dec. 2). Delta town finally gets gas after failing to pay bill. http://msbusiness.com/blog/2011/12/02/delta-town-finally-gets-gas-after-failing-to-pay-bill/</p>

Recent Data Analysis and a Hybrid Model

The case study evidence provided in earlier suggests that general fund exhaustion played a pivotal role in recent California defaults and bankruptcies. While interest over revenue and annual revenue change remain intuitively attractive, they seem to have had a lesser role in these more contemporary situations. Outside California, pension costs appear to have played a major role in the Prichard and Central Falls bankruptcy filings.

Addressing Public Employee Pensions and OPEBs

As discussed in the Great Depression Review chapter, pensions were not seen as a significant contributor to the spate of municipal bond defaults that occurred in the 1930s. Since that time, California courts have repeatedly ruled that public employees and their beneficiaries have a right to receive pension benefits according to the rules that prevailed at the time of their employment. For example, in *Kern v. City of Long Beach* (29 Cal.2d 848, 1947), the California Supreme Court ruled that a fireman could not be deprived of his pension benefits by a change to the city charter. In *Betts vs. Board of Administration* (21 Cal.3d 859, 1978), the Court ruled that a public agency could not apply a new benefit formula lowering benefits to a vested employee unless it also provided a comparable, offsetting advantage.

Pension benefits earned by existing employees thus enjoy legal protections similar to debt service payments. Although we lack strong statistical evidence of pensions causing municipal bond defaults, it seems appropriate to add pension expenses in the numerator of the interest to revenue ratio. This new ratio of “uncontrollable costs” to revenues provides

a parsimonious way to incorporate a variable that has been the subject of extensive political discussion.

On the other hand, we do not apply this approach to OPEB costs. As the Government Accountability Office noted in a 2007 report, relative to pensions, “state and local law provides much less protection for retiree health benefits. Retiree health benefits are generally treated as an operating expense for that year’s costs on a pay-as-you-go basis and managed together with active employee benefits.” Since OPEBs appear to be junior to debt service and pension obligations, it seems inappropriate to include them in a ratio intended to predict municipal bond defaults.

Since much of today’s municipal solvency discussion focuses on public employee pension costs, it is worth taking a moment to explain how they enter into our model. We use actual employer contributions obtained from the city’s CAFR. For California cities, this amount is generally the same as Actuarially Required Costs (ARC) because most cities belong to CalPERS which sets employer contribution rates based on ARC.

Since our goal is to estimate one year default probabilities, a city’s Unfunded Actuarially Accrued Liabilities (UAAL) are less relevant. The UAAL represents the present value of future payments needed to bring a pension system to full funding. It is not necessary for cities to remedy pension underfunding in one year. Indeed, in the case of a single employer system, it is not necessary for a city to remedy underfunding at all. Pensions, like OPEBs, may be funded on as a pay-as-you-go basis, although doing so may place an unsustainable burden on the city’s revenue base.

The model implementation we present later may be used with current balances and annual cash flows or with *projected* future balances and cash flows. The latter option

provides forward looking default probability estimates which should be more useful to investors. For CalPERS member cities, it is possible to obtain future employers contribution rates from actuarial reports published on the system's web site at <http://www.calpers.ca.gov/index.jsp?bc=/about/forms-pubs/calpers-reports/actuarial-reports/home.xml>. These reports also include a sensitivity analysis showing the effect of different portfolio return rates on future employer contribution rates.

In April 2013, CalPERS (2013) changed its method for amortizing UAAL with the goal of achieving full funding within 30 years. For cities and other local agency CalPERS members, this new approach results in higher annual pension costs relative to what they would have been under the existing actuarial method. The change will be phased in starting in fiscal year 2016. At that time, the employer contribution rates for the median public safety employee plan will increase from 29.3% to 30.8%. In FY 2020, the expected increase will be from 33.9% to 39.5%. This means that, all other things being equal, an average city will face an increase in pension costs of 5% in FY 2016 and 17% in FY 2020. In fiscal 2011, pension costs accounted for about 7% of total governmental revenues, so the budgetary impact of this change should average roughly 1% of revenues by FY 2020.

Statistical Analysis of Contemporary Data

We collected financial statistics for 261 California cities with population greater than 25,000 for fiscal year 2011. Two of these cities, Stockton and San Bernardino, defaulted the following year. Unfortunately a sample containing 2 defaults out of 261 cities does not contain enough variation to be conducive to modeling.

We addressed this problem by making a number of additions and alterations to the sample. Each of these changes adds default observations at the cost of compromising the

consistency of the data set. To preserve variation, we also added non-defaulting cities to the sample that are comparable to the defaulting cities that were added. These alterations are reviewed in the following table.

Alteration	Defaults Added	Drawbacks
Reclassified cities declaring fiscal emergencies as defaults	Atwater, Culver City, El Monte, Fairfield, La Mirada, Lancaster, Monrovia, Stanton (8 cities in all)	As discussed in the California survey, a fiscal emergency is not necessarily a precursor to default. While it is reasonable to assume that these cities are closer to default than most others, the degree of relative proximity is unclear.
Added data for Vallejo from the time of its 2008 default	Vallejo	Inconsistent time period; database contains no other 2008 California data.
Added 19 cities in Pennsylvania including 2 that defaulted	Harrisburg and Scranton	Each state has unique laws governing municipal finance, so these observations may not be comparable.
Added 16 cities in Alabama including 1 that filed a bankruptcy petition	Prichard	Same as above; also Prichard did not have municipal bonds outstanding at the time of its 2010 bankruptcy filing.
Added 14 cities in Rhode Island including 1 that filed a bankruptcy petition	Central Falls	Same as above; also Central Falls did not default during its bankruptcy and has less than 25,000 people. Some Rhode Island cities have large general fund surpluses because they collect tax revenues on behalf of school districts whose operations are consolidated on their CAFRs.
Added 35 cities in Michigan including 5 that had state takeovers	Allen Park, Detroit, Flint, Inkster, Pontiac	None of the Michigan cities actually defaulted, and may not have defaulted in the absence of state intervention. The timing of a state takeover is likely to differ from the timing of a bankruptcy filing or default if state intervention had not occurred.

In addition to reclassifying "default" for California cities, we also added cities from four other states. For Alabama, Michigan, Pennsylvania and Rhode Island we added 2011 CAFR data for all cities with population above 25,000. These states are similar to California

in that they too have cities that have experienced serious financial stress. Data for non-defaulters were obtained from 2011 CAFRs. Data for “default” observations came from the fiscal year prior to the default or emergency. Since all independent variables are stated in ratio form, changes in CPI across years do not need to be considered.

We tested four fiscal ratios with this new sample of 346 observations, which contain a total of 20 observations classified as defaults (which include 13 fiscal emergency declarations not associated with actual defaults or bankruptcies as well as two bankruptcies not associated with defaults). The ratios were:

- Interest plus Pension Expenses over Revenue
- Annual Revenue Change
- General Fund Surplus over General Fund Revenue (a measure of the relative size of a city’s general fund surplus or deficit)
- General Fund Balance over General Fund Expenditure (a measure of the relative size of a city’s general fund balance)

The first two variables are identical to those we included in the final Great Depression-era model. The third and the fourth variables listed above are meant to be analogs to this cash variable; accounting differences across the two time periods make it difficult to select a single current-period variable to correspond to the Depression-era variable, thus we selected two that are similar.

It is worth noting that the final variable, General Fund Balance over General Fund Expenditure, is highly correlated with the San Bernardino and Stockton defaults. Of the 261 California observations from fiscal 2011, these cities rank 259th and 251st respectively. In other words, the two defaulters both rank in the bottom 4% of this variable’s distribution. It would be tempting to use this single variable as a standalone indicator of impending default, but the very limited record of municipal defaults in California suggests caution.

The following tables present summary statistics for the focus variables for the entire 346 city sample as well as the subsample consisting of just the 261 California observations from 2011.

Category	n	(Interest + Pension) / Revenue	Annual Revenue Change	GF Surplus / GF Revenue	GF Balance / GF Expenditure
Entire Multi-State Sample	346	11.96%	2.52%	1.65%	54.12%
"Defaults"	20	11.60%	-1.06%	-10.56%	24.95%
Non-Defaults	326	11.98%	2.74%	2.40%	55.91%

Category	n	(Interest + Pension) / Revenue	Annual Revenue Change	GF Surplus / GF Revenue	GF Balance / GF Expenditure
California 2011 Sample	261	12.44%	2.69%	-0.44%	63.67%
Stockton & San Bernardino	2	13.01%	1.56%	-6.21%	2.98%
All except Stockton and San Bernardino	259	12.39%	2.71%	-0.33%	65.42%
Cities Declaring Emergency	8	15.24%	0.14%	-13.13%	77.68%
Cities Not Defaulting or Declaring Emergency	251	12.30%	2.80%	0.08%	65.03%

In most cases, relative means are consistent with expectations. One exception is (Interest + Pension) / Revenue for the overall sample. The defaults and quasi-defaults have a somewhat lower ratio than the cities that performed without incident. In the California sub-sample, Stockton, San Bernardino and the emergency declarants have higher uncontrollable cost burdens than the remaining cities.

The table below shows summary statistics for the sample constructed in the manner described above.

Table: Summary Statistics, Current Period Sample

Variable	Obs	Mean	Std. Dev.	Min	Max
DEFAULT	346	0.06	0.23	0	1
Annual Revenue Change	346	0.03	0.10	-0.36	0.64
GF Surplus / GF Expenditure	346	0.02	0.16	-0.70	0.68
GF Balance / GF Revenue	346	0.54	0.54	-0.60	3.72

Next, we present estimates of logit models. Other than the addition of the two general fund variables, this is the same model presented in the section on the Great-Depression. The separation problem did not prevent including state fixed-effects and so all four specifications presented below make use of logit estimation.

Table: Logit Analysis, Current-Period Sample

VARIABLES	DEFAULT	DEFAULT	DEFAULT	DEFAULT
Annual Revenue Change	-4.539*			-3.9
	(2.439)			(2.678)
GF Surplus / GF Expenditure		-5.217***		-4.468***
		(1.390)		(1.565)
GF Balance / GF Revenue			-1.908	-1.032
			(1.806)	(1.371)
Constant	-1.877***	-1.925***	-1.448***	-1.746***
	(0.515)	(0.551)	(0.561)	(0.658)
Estimation strategy	logit	Logit	logit	logit
State fixed effects?	Yes	Yes	Yes	Yes
Observations	346	346	346	346
chi2	10.08	21.29	5.123	21.76

Robust standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1

As one can see from the table above, all of the coefficient estimates have the theoretically expected sign. As for statistical significance, only GF Surplus / GF Expenditure is consistently significant at the 1% level. However, while the standard errors on GF Surplus / GF Expenditure are large, it is important to remember that the magnitude of the effect of each variable is a product of the coefficient's estimate and the variation in the

variables. Mathematically, a one standard deviation increase in GF Surplus / GF Expenditure (which is 0.16) turns out to increase the number inside the logit function by -4.468×0.16 or 0.72. In contrast, although the coefficient on GF Balance / GF Revenue is only -1.032 , a one standard deviation in this variable (equal to 0.54) increases the number inside the logit function by -1.032×0.54 , or -0.56 . Thus although GF Surplus / GF Expenditure has a lower standard error and is thus statistically significant, the effects of both GF Surplus / GF Expenditure and GF Balance / GF Revenue are similar in magnitude when evaluated at their standard deviations.

In summary, the general fund surplus/deficit variable is the strongest predictor of a fiscal emergency in the multi-state data set. However for the two California defaults in 2012, general fund balance was the strongest indicator, and even in the multi-state sample the magnitude of the effect is nearly as large as the surplus variable. Given the limited evidence available it appears that some combination of accumulated deficits from prior years (expressed in the general fund balance) and current deficits trigger the type of fiscal crisis that leads to default.

Creating a Hybrid Model

We now have two statistical models: (1) a Depression era model that uses an internally consistent data set with a relatively high number of defaults, but which is 80 years old, and (2) a contemporary model that relies on convenience sampling to get a smaller proportion of default observations – most of which are not actual defaults.

Neither of these models are entirely satisfactory, but both highlight the role of independent variables that are supported by strong intuition and/or recent case study evidence. Consequently, we believe that the best approach under the circumstances is to

create a hybrid of the two models – one that uses all four of the variables we have highlighted: (1) interest plus pension expense over revenue, (2) annual revenue change, (3) general fund surplus/(deficit) over general fund revenue and (4) general fund balance over general fund expenditure.

Creating a hybrid model is challenging. One possibility would be to combine the Depression-era and recent data sets and then run regression analysis on the aggregate data set. Unfortunately, this is complicated by the lack of comparability between the older and newer data sets. One source of incompatibility is the accounting basis used – Census data in the 1930s was cash based while modern CAFRs use the modified accrual standard. However, it is not clear that this accounting standard issue would introduce any systematic bias in the ratios we are studying and could arguably be assumed away. More problematic is the fact that 1930s era Census data does not separate general fund revenues, expenditures and balances from those in other funds.

While fund accounting was developed and considered to be a best practice well before the Depression (Bureau of Municipal Research, 1913), it is not clear that it was widely used in the early 1930s. To the extent that fund level data from that period are available at all, they would have to be collected from financial reports produced by individual cities which may no longer be available and would lack the standardization imposed by modern Government Accounting Standards.

Finally, even if comparable Depression-era data were available, there would be a question of how to weight them against modern data. For that matter, even within our contemporary data set, there is the question of how to weight the California 2011 data against the out-of-state data and the 2008 Vallejo default observation which we added.

After discussing several possible ways to select coefficients for our default probability model, we decided that the best option is to assign coefficients subjectively based on our reading of the case studies and statistical evidence. Ultimately, any choice of coefficients is subjective but we believe our choice of coefficients is strongly supported by the evidence we have presented.

The general fund balance variable most effectively predicts the San Bernardino and Stockton models. It is also strongly associated with the 2008 default in Vallejo and the 1999 default in Desert Hot Springs and was cited (conceptually if not by name) as a justification by a couple of the defaulting cities. Therefore this variable is dominant in our model.⁸ The other three variables are assigned coefficients intended to give them roughly equal weights in the default probability calculation. The interest plus pension to revenue variable is assigned a higher coefficient than the other two variables (general fund surplus and annual revenue change) because its range and standard deviation are lower.

We also exclude constants from the subjectively determined model. Although constants were highly significant in the empirical models shown above, there is no obvious way to combine them in a meaningful way. Instead of using constants, we apply a calibration procedure to the model generated default probabilities so that they fall within a reasonable range.

This last step requires some elaboration. In 2012, two of the 265 cities in the target population (California cities with over 25,000 residents) defaulted on city obligations

⁸ We also note that fund balance factors were assigned higher weights in a recent Fiscal Stress Monitoring System published by the New York State Comptroller's Office (2013).

yielding a default rate of just over 0.75%. The goal of the calibration is to ensure that the mean city default probability matches this rate.

Admittedly, this is a controversial assumption. On the one hand, the 2012 default rate among California cities with population greater than 25,000 is the highest in 140 years. On the other hand, many observers believe that the recent defaults represent the beginning of a much larger wave.

We are skeptical of this latter view. City pension reforms and the recent stock market rally are likely to reduce pressure on city retirement costs in the intermediate term. Further, the recent rebound in property prices that started in 2012, lower unemployment and continued low interest rates are all positive influences on city solvency. Finally, as of this writing, no city in this class has initially defaulted or filed for bankruptcy in 2013.

Thus, we believe that 0.75% is a pessimistic mean assumption, but we see the benefits of conservatism in this regard and lack a rigorous approach to estimating a future default rate. Thus, this 2012 mean appears to be the best option for model calibration.

We calibrate the modeled default probabilities by raising them to a power such that the mean probability of the 261 city 2011 sample equals the 2012 default rate of 0.75%. This exponent was 5.76. Thus, if the initial calculation yielded a default probability of 50%, the calibrated probability would be $0.5^{(5.76)} = 0.0185 = 1.85\%$. This process is more attractive than simply dividing all the probabilities by a fixed amount because it still allows for a full range of probabilities between zero and one, since one raised to any power remains one.

The hybrid model presented here places both Stockton and San Bernardino in the top decile of default probability estimates published in Appendix 3 – an unsurprising

outcome given that their default data is in-sample and the variable most associated with their defaults has been over weighted in our model. All but one of the reported probabilities are below 3% due to the calibration process.

Limitations and Future Research

We have marshaled case study and statistical evidence to produce a quasi-empirical logistic model of city default risk based on fiscal accounting ratios. Since ours' appears to be the first recent attempt to produce such a model, we are confident that other researchers can improve upon our results. In this section, we briefly consider some opportunities for enhancement.

First and foremost, replacing our subjective coefficient estimates with ones deriving from a more rigorous empirical method would be attractive. As we have seen, this is a challenge complicated by the relative paucity of defaults and heterogeneity of fiscal data across states and time periods. We also caution researchers against succumbing to the temptation of using ratings as a proxy for default probabilities in order to obtain a larger pool of current observations. Earlier we raised questions about the responsiveness of ratings to changing municipal financial conditions.

Second, other researchers may find opportunities to alter our list of independent variables, through either substitution or addition. Two variables that we would have liked to have analyzed further were cash and unrestricted general fund balances. While cash has a very strong intuitive basis, we are concerned that cash levels may be quite volatile on a daily basis. Thus, the cash on a city's balances sheet reported at the end of the fiscal year may be unrepresentative of the amount of cash it will have on hand when it has to make interest or pension payments a few weeks later. We would of course caution against

creating a long list of predictors given the possibility that imperfect multicollinearity may lead to imprecise estimates. In preparing this study, our team has gained substantial experience with collecting and handling CAFR data, and we welcome inquiries from other researchers on how to cost-effectively gather this data.

A third opportunity for improvement is to use forward looking inputs with our models. In Appendix 3, we report default probabilities derived from CAFR data as of June 30, 2011 and June 30, 2012. Since our default probability estimates have been designed to cover one year, the numbers in the Appendix are essentially historical values. It is possible to collect data from various sources that would support more current and even forward looking estimates of the independent variables we have highlighted. These include city budgets, monthly or quarterly cashflow reports, pension system actuarial reports (which include future employer contribution rates) and socioeconomic variables. These forecast values could be used with our model specifications to produce forward looking default probability estimates.

A common objection to the use of budgets in gathering forecast independent variables is that the quality of budget estimates can suffer from political manipulation or the inexperience of financial analysts who prepare them. These concerns are likely to apply differently across cities: some cities have highly professionalized budgeting processes resistant to political interference, while others do not. Also, tools are available to researchers to assess the validity of budgets. First, CAFRs contain comparisons of budgeted and actual results, so it is possible to gauge the effectiveness of the budget process in prior years. Second, the interim financial results produced by many larger cities can enable a closer to real time assessment of budget accuracy.

Conclusion

While statistical models of corporate credit risk have become quite common over the past fifty years, municipal credit risk modeling has remained relatively undeveloped. Our hope is that this situation begins to change with the release of this study.

Many observers emphasize that municipal default and bankruptcy is a political decision. But political decisions are not immune to modeling. Substantial research has explained and demonstrated how political actions can be predicted based on the conditions faced by political actors.

Standing at the intersection of financial modeling and political analysis, we suggest that a model based on fiscal indicators can improve our ability to predict municipal credit crises. Gaining the ability to predict such events is the first step toward minimizing them. Once a municipal scoring mechanism gains acceptance, it can be used to proactively identify the cities most at risk, thereby creating the opportunity for remedial action.

In this connection, it is worth dwelling on the “to do” list created by our project. Our default probability estimates rely on Comprehensive Annual Financial Report data typically filed 4-8 months after the end of the fiscal year. While these reports are more standardized and thus easier to exploit, there is no reason that the model cannot be used with forecast variables. All cities publish budgets and many provide interim financials that can be used to project current and out year fiscal results. Ratios based on these projections can be loaded into our model to obtain more forward looking default probability estimates.

Our analysis focuses only on the issuer level, abstracting from the variations in risk associated with different bond issues. While we know that general obligations are less risky

than other types of issues, it would be useful to quantify these risk differentials for any given city.

Thus, this admittedly long study just scratches the surface of what is possible in the area of quantitative municipal default probability modeling. We believe that further work will benefit not only municipal bondholders, but also the political leaders, taxpayers, public employees and beneficiaries who are all victims in a city credit crisis.

Appendix 1: Extending the Depression-Era Default Model

The Depression-era model presented in the main text did not include socioeconomic variables and contained a varying number of observations per city. In this Appendix, we present additional models derived from the Depression data set that explore the effect of altering these characteristics of the model. We estimate the following cross-sectional model by ordinary least squares (OLS):

$$DEFAULT_i = X_i\gamma + \alpha_j + \varepsilon_i$$

where $DEFAULT_i$ is an indicator variable equal to one if city i defaulted in 1931, 1932 or 1933,⁹ X_i is a matrix of independent variables for city i and γ is a vector of coefficients to be estimated, α_j is a state-specific intercept (i.e. all cities in the same state have a common intercept) and ε_i is an error term with the usual properties. Although using OLS to estimate this linear probability model has well-known problems it also has the advantage that it allows for each state i to have its own intercept α_i (i.e. state fixed effects). In the main text we used firthlogit to enable fixed-effect estimation, although this came at the expense of not being able to estimate robust standard errors. Therefore we estimate linear probability models here so that we can include both fixed effects and estimate robust standard errors. In addition OLS coefficients are easy to estimate and to interpret. We have performed various robustness checks on the functional form and estimation strategy of this model and these results are available upon request. In short, the choice of OLS over logit does not alter our general conclusions.

⁹ 3 cities defaulted in 1930, 6 in 1934, 1 in 1935 and 1 in 1936. We dropped these cities from our sample.

We have identified eight independent variables that all have theoretical justification for including in the matrix X_i . Four of these variables are socioeconomic, and four are fiscal ratios. We discuss all of these in detail below. A major goal of this research is to develop a municipal default model that is *parsimonious*, i.e. a model that explains default well, using the fewest possible variables. To this end, our strategy is to estimate various restricted versions of equation (1), in particular to estimate eight versions of equation (1) which include each of the independent variables separately, and then to retain only those variables that are statistically significant.¹⁰

Our first task is to develop two “semi-final” models, where one semi-final model includes socioeconomic variables only, and the other includes fiscal ratio variables only. The motivation for this approach is that, as we described in the literature review, previous studies have used both sets of variables. We are interested in running a “horse race” between these two sets of variables to determine whether socioeconomic or fiscal ratio variables have greater explanatory power. After estimating these two semi-final models, we pull together all significant variables into a final model, which we will then use as the basis for predicting current-era and future defaults.

We begin by estimating the socioeconomic-only model. Table A1 below describes each socioeconomic variable, as well as the source from which we obtained it.

¹⁰ We understand this method of variable selection is not ideal for several reasons, including that it runs the risk of excluding variables that are not significant in the restricted versions due to omitted variable bias. In future work we will think more carefully about better approaches to variable selection. However the advantage of retaining only significant variables is that it will tend to result in a model with the greatest predictive power.

Table A1: Variable Descriptions, Socioeconomic Models

Variable	Description	Source
DEFAULT	<i>An indicator of whether city defaulted in 1931, 1932 or 1933</i>	Authors calculations
lnPOP	<i>The natural logarithm of city population in 1930</i>	IPUMS
HOMEVALUE	<i>Average value of owner-occupied housing in city in 1930</i>	Authors calculations using data from IPUMS
HOMEOWNER	<i>Fraction of households in city living in owner-occupied housing in 1930</i>	Authors calculations using data from IPUMS
SEI	<i>Average level of Duncan's socioeconomic index among city residents, 1930</i>	Authors calculations using data from IPUMS

As mentioned above, the variable DEFAULT is the dependent variable and it is an indicator equal to one if the city defaulted; extensive review of media reports was used to determine whether or not each of the largest 311 cities in the United States (as per the 1930 Census) defaulted over the subsequent years. The remaining four variables are the socioeconomic variables we selected based on theoretical appeal, their use in previous research, and data availability. The source for each of these is the 1930s integrated public use microdata series (IPUMS) available from the University of Minnesota (<https://usa.ipums.org>; see Ruggles et. al., 2010). This data set includes 1930 population and the individual responses to Census interviewers from the 1930 Census. We use these individual-level data to construct aggregate city-level measures for HOMEVALUE, HOMEOWNER and SEI.

The first socioeconomic variable is lnPOP. This is the natural logarithm of city population based on the 1930 Census. Rather than including this variable in levels, we take the log transformation primarily so extreme values (e.g. New York City) do not overly-influence the resulting estimates. This is an alternative approach to the method of ranking cities by population that we employed in the main text. Literature we reviewed included suggestions that large cities were less likely to default than smaller ones. Larger cities would generally be expected to have more diverse economies rendering them less

vulnerable to the collapse of any given industry. Also, larger cities may be more able to attract state or federal bailouts, due to their greater importance and political power. Next, HOMEVALUE is the average value of owner-occupied housing in the city. We expect that cities with high values will be less likely to default as property tax revenue will be greater in these cities allowing governments to more easily service debt.

The variable HOMEOWNER is the fraction of residents that own their home (either outright or have a mortgage.) It is possible that voters in cities with a high homeownership rate will pressure politicians to not default, as they are worried that bad publicity resulting from a default will lower the value of what for most of them is their single largest asset. On the other hand, if default allows cities to “wipe the slate clean”, it could be in the financial interest of homeowners for the city to default. Thus determining the effect of homeownership on default probability is an empirical question. Finally, SEI stands for socioeconomic index. This particular index is the Duncan SEI and is contained in the IPUMS data. The 1930 Census did not ask questions about income, and so SEI is the closest available proxy for income at the city level that is available. Like HOMEVALUE, we expect that richer cities will be less likely to default, as governments will be able to more easily raise revenue from richer citizens.

Table A2 below presents summary statistics for these variables. As can be seen, our sample size is 299, which excludes the eleven cities that defaulted in 1930, 1935 or 1936, and also Honolulu for which IPUMS data was not available. Twelve percent of cities in this sample are classified as defaulting. About 46% of citizens were homeowners, and this value ranged from 22% to 74%. These and other facts about our sample can be seen in the table below.

Table A2: Summary Statistics, Socioeconomic Models

Variable	Obs	Mean	Std. Dev.	Min	Max
DEFAULT	299	0.12	0.32	0	1
lnPOP	299	11.25	0.86	10.32	15.77
HOMEVALUE	299	7.37	3.55	2.11	35.80
HOMEOWNER	299	0.46	0.10	0.22	0.74
SEI	299	33.84	4.05	21.94	51.13

We next present the results of estimating the socioeconomic-only versions of Equation (1) in Table A3. This table lists the coefficient estimates with the (heteroskedastic-robust) standard errors in parentheses.

Table A3: Socioeconomic Linear Probability Models

VARIABLES	DEFAULT	DEFAULT	DEFAULT	DEFAULT	DEFAULT	DEFAULT
lnPOP		0.0549** (0.026)				0.0587** (0.026)
HOMEVALUE			-0.00213 (0.004)			
HOMEOWNER				-0.500* (0.280)		
SEI					-0.00962** (0.005)	-0.0109** (0.005)
Constant	0.5 (0.382)	-0.148 (0.514)	0.512 (0.384)	0.684* (0.401)	0.809* (0.420)	0.155 (0.537)
Observations	299	299	299	299	299	299
R-squared	0.27	0.29	0.27	0.28	0.28	0.30
adjusted R-squared	0.15	0.17	0.14	0.16	0.16	0.18
State fixed effects?	Yes	Yes	Yes	Yes	Yes	Yes
Robust standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1						

The first column of Table A3 presents a model with no independent variables but does include state-specific intercept terms (i.e. fixed effects.) Here we see that the adjusted R-squared is 0.15. This is an indication of how well the model explains defaults and this value

serves as a benchmark for the incremental explanatory power of the independent variables. In the second column, we see that the coefficient on $\ln\text{POP}$ is positive (suggesting greater default risk with increased population; not the expected sign) which is significant at the 5% level. The adjusted R^2 is 0.17, a slight improvement over the fixed-effect only specification in column one.

In column three we see that the coefficient on HOMEVALUE is -0.00213. However, while the sign is consistent with our expectations, the estimate is not statistically significant. In column four, the coefficient on HOMEOWN is -0.5. The sign is consistent with a situation where homeowners view defaulting as eroding the value of their single largest investment, and the estimate is marginally significant. However, as our cutoff value for including variables in the semi-final model is the 5% level, we consider the estimated coefficient on HOMEOWN to be statistically insignificant (i.e. we cannot reject the hypothesis that the true population coefficient is zero with 95% confidence.)

In column five the coefficient on SEI is -0.00962. This estimate is significant at the 5% level. The sign is consistent with our expectations. The adjusted R^2 is 0.16, an even slighter improvement over the fixed-effect only specification in column one than the model that included only $\ln\text{POP}$.

After reviewing the estimates on each of the coefficients from columns two through five, we determined that two socioeconomic variables meet the threshold for inclusion in the socioeconomic-only semi-final model, and we present this in column six. Both estimates retain their statistical significance, and the magnitudes of the estimates do not change much. The adjusted R^2 increases to 0.18, the highest of any of the models presented in Table A3. However, this is still a rather modest increase over the fixed-effect model from

column 1. In sum, we find that two socioeconomic variables are statistically significant. However, as measured by adjusted R^2 , they do not explain defaults very well.¹¹

Do fiscal ratio variables predict default better than socioeconomic variables? To answer this question we turn next to the estimating the fiscal ratio-only models. Table A4 below describes the four fiscal ratio variables we selected for this analysis.

Table A4: Variable Descriptions, Fiscal Ratio Models

Variable	Description	Source
INT_BY_REV	$\frac{\text{total interest costs in 1931}}{\text{total receipts in 1931}}$	CENSUS
REV_CHANGE	$\frac{\text{total receipts in 1931} - \text{total receipts in 1930}}{\text{total receipts in 1930}}$	CENSUS
SURPLUS	$\frac{\text{total receipts in 1931} - \text{total government costs in 1931}}{\text{total receipts in 1931}}$	CENSUS
ASSETS_BY_REV	$\frac{\text{aggregate assets in 1931}}{\text{total receipts in 1931}}$	CENSUS

The first two fiscal ratio variables INT_BY_REV and REV_CHANGE are the same variables analyzed and described in the Depression-era discussion of the main text. The variables SURPLUS and ASSETS_BY_REV are different than the cash variable described in the main text, but we include them here given their theoretical appeal. The SURPLUS variable reflects the difference between government-wide revenues and expenditures divided by government-wide receipts; ASSETS_BY_REV is similar to the cash variable from

¹¹ A few words about the interpretation of the constant term are in order. As cities in each state all have a common intercept (or constant) term, the reported constant has meaning for only one state. To avoid perfect multicollinearity, we excluded one state fixed-effect, namely Alabama. So, the reported Constant term is correctly interpreted as the intercept for cities in Alabama. Although reporting the intercept for cities in Alabama (and not Ohio, California or another state) is arbitrary, we report it to remind readers that our model includes state fixed effects.

the main text in that it provides a measure of the financial assets possessed by the city, scaled by its revenue. Table A5 below presents summary statistics for the variables described above.

Table A5: Summary Statistics, Fiscal Ratio Models

Variable	Obs	Mean	Std. Dev.	Min	Max
INT_BY_REV	299	0.110797	0.055343	0	0.334087
REV_CHANGE	299	-0.00735	0.098319	-0.2547	0.5215
SURPLUS	299	-0.07465	0.176743	-1.0212	0.247061
ASSETS_BY_REV	299	3.632558	0.999222	0.139465	9.781581

Table A5 documents a large amount of variation in the fiscal health of city governments during the start of the Great Depression. The average city in our sample had interest expenses equal to 11% of revenue, however some cities had interest to revenue ratios three times this value. Moving to REV_CHANGE, we see that the average city saw its receipts fall by about one percent from 1930 to 1931, however some cities saw as much as a 25% fall in revenues, while still others saw revenues rise by as much as 50%.

Table A6 below presents the results of estimating restricted versions of equation (1). We follow the same approach below as we took with the socioeconomic-only semi-final model, namely we include each of the four fiscal ratio variables one at a time, then evaluate the statistical significance of the variables, and then combine all variables with statistically significant estimates into the semi-final fiscal ratio model.

Table A6: Fiscal Ratio Linear Probability Models

VARIABLES	DEFAULT	DEFAULT	DEFAULT	DEFAULT
INT_BY_REV	2.840***			

	(0.600)			
REV_CHANGE		-0.174		
		(0.198)		
SURPLUS			-0.0226	
			(0.130)	
ASSETS_BY_REV				0.0156
				(0.025)
Constant	-0.0429	0.497	0.494	0.433
	(0.322)	(0.392)	(0.382)	(0.391)
Observations	299	299	299	299
R-squared	0.35	0.27	0.27	0.27
adjusted R-squared	0.24	0.15	0.14	0.15
State fixed effects?	Yes	Yes	Yes	Yes
Robust standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1				

In the first column we see that the coefficient on INT_BY_REV is 2.84. The positive value of this is consistent with our expectations and the estimate is highly significant. If a city's interest expenses rose and caused INT_BY_REV to increase by 0.055 (one standard deviation from Table A5), then the estimate in column one suggests its default probability will increase by 0.055×2.84 , or 15.6%. The adjusted R² is 0.24 and this is much higher than we saw for any of the socioeconomic models in Table A3.

None of the other fiscal ratio variables are statistically significant. Therefore, the semi-final fiscal ratio model is the one presented in column one of Table A6. As already noted, as measured by adjusted R², this model explains defaults much better than the semi-final socio-economic model.

Before concluding our discussion of fiscal ratios and presenting the final model, we consider an extension of the fiscal ratio only model. This extension recognizes that default may be a function of multiple aspects of city government financial statements. In particular, default may not happen when just interest payments are high, or just when

revenues fall, for example, but instead may result when both of these conditions occur at the same time. In other words, we are interested in determining if city defaults in the Great Depression were the result of a “perfect storm”, evidenced by poor indicators of multiple fiscal measures.

To explore this possibility, we first construct dummy variables for each city, indicating whether they were in the upper or lower quintile of the distribution for each of the four fiscal ratio variables. We have given these dummy variables intuitive names and they appear in the descriptions in Table A7. Then, we interact each of these dummy variables that indicate poor fiscal state. Table A7 describes the six variables that result from interacting the poor quintile indicator variables.

Table A7: Variable Descriptions, Perfect Storm Models

Variable	Description
INTERACT1	HIGH_INT_BY_REV * NEG_REV_CHANGE
INTERACT2	HIGH_INT_BY_REV * NEG_SURPLUS
INTERACT3	HIGH_INT_BY_REV * LOW_ASSETS_BY_REV
INTERACT4	NEG_REV_CHANGE * NEG_SURPLUS
INTERACT5	NEG_REV_CHANGE * LOW_ASSETS_BY_REV
INTERACT6	NEG_SURPLUS * LOW_ASSETS_BY_REV

An explanation of one of these interaction variables should serve to clarify our approach. Take for example the variable INTERACT1, which is equal to one if a city was one of the 60 worst cities in our sample with regard to interest over revenues, *and was also* one of the worst 60 cities in our sample with respect to revenue change. That is, if the value of INTERACT1 is equal to 1 for a city, that means this city was in the top quintile for interest over revenues, and also was in the bottom quintile for revenue change. If these two conditions occur simultaneously for a city, then it could be that this is a perfect storm, and a city in this situation will be very likely to default. To provide a picture of how many cities

were in these potential perfect storm situations, we present summary statistics for the six interaction variables below in Table A8.

Table A8: Summary Statistics, Perfect Storm Models

Variable	Obs	Mean	Std. Dev.	Min	Max
INTERACT1	299	0.064	0.244	0	1
INTERACT2	299	0.047	0.212	0	1
INTERACT3	299	0.014	0.115	0	1
INTERACT4	299	0.040	0.197	0	1
INTERACT5	299	0.027	0.162	0	1
INTERACT6	299	0.027	0.162	0	1

Table A8 reveals that between 1.4 and 6.4 percent of cities found themselves in at least one of the six possible perfect storm situations we have identified. To examine which of these situations really was a perfect storm, we include each of these indicator variables in a model that is similar to those presented above. These results are presented in Table A9.

Table A9 reveals that only one of the six possible perfect storm situations we have identified is a statistically significant predictor of default during the Great Depression. As a result we restrict our attention to this variable, INTERACT2. As shown in Table A7, the variable INTERACT2 is equal to one if the city is in a situation where it had one of the highest interest over revenue values in our samples, and had one of the lowest surpluses. Although each of the interaction variables represents reasonable perfect storm situations, the situation behind INTERACT2 is one where a city faces high interest expenses relative to its revenues, and it had a negative budget surplus at the same time.

Table A9: Perfect Storm Linear Probability Models

VARIABLES	DEFAULT	DEFAULT	DEFAULT	DEFAULT	DEFAULT	DEFAULT
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INTERACT1	0.0829 (0.132)					
INTERACT2		0.306** (0.130)				
INTERACT3			0.0557 (0.182)			
INTERACT4				0.0621 (0.099)		
INTERACT5					0.0403 (0.081)	
INTERACT6						0.172 (0.139)
Constant	0.459 (0.420)	0.347 (0.273)	0.5 (0.383)	0.5 (0.383)	0.5 (0.383)	0.5 (0.383)
Observations	299	299	299	299	299	299
R-squared	0.27	0.296	0.268	0.269	0.268	0.274
adjusted R-squared	0.147	0.177	0.144	0.145	0.144	0.152
State fixed effects?	Yes	Yes	Yes	Yes	Yes	Yes
Robust standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1						

The variable INTERACT2 is statistically significant and of the expected sign. However, the adjusted R² is only 0.18, not higher than the semi-final socioeconomic only model presented above. Part of this low explanatory power is due to the relatively crude way of categorizing variables as indicators. Nonetheless, although this variable performed as expected, it is not possible to say that this or any of the perfect storm variables contribute substantial explanatory power.

This completes our analysis of the fiscal ratio only models. The remaining discussion and analysis in this section is concerned with moving from semi-final to final models. In all the tables presented above, we found four estimated coefficients that were statistically significant, namely those on lnPOP, INT_BY_REV, SEI and INTERACT2.

Therefore, we collect these four variables together and include them in a linear probability model along with state fixed effects. The results are reported in Table A10 below.

In the first column of results in Table A10 we see that two of the coefficient estimates retain their statistical significance in the combined model, namely SEI and INT_BY_REV. Because, as emphasized above, we are interested here in presenting the most parsimonious model possible, we estimate one final model. This is presented in the second column of results in Table A10. We include SEI and INT_BY_REV only and we do not include state fixed effects. The coefficients of the two independent variables retain their statistical significance. Moreover, while their magnitude changes slightly, the change is less than 20%. Therefore, we feel comfortable in using the results presented in the second column of Table 10 as our *final model* for this Appendix.

Table A10: Final Linear Probability Models, Great Depression Era

VARIABLES	DEFAULT	DEFAULT
lnPOP	0.0404 (0.027)	
SEI	-0.0117** (0.005)	-0.00933** (0.004)
INT_BY_REV	2.537*** (0.645)	2.065*** (0.378)
INTERACT2	0.163 (0.134)	
Constant	-0.169 (0.446)	0.204 (0.132)
Observations	299	299
R-squared	0.38	0.15
adjusted R-squared	0.27	0.14
State fixed effects?	Yes	No

Robust standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1

We have used this final model to predict current-era defaults as an alternative to the hybrid model described in the main text. The equation below presents the model from the second column of Table 10 in equation format:

$$DEFAULT = 0.204 - 0.00933 * SEI + 2.065 * INT_BY_REV$$

We used this equation to predict current-period defaults, by plugging in the interest divided by revenue variable from the 2011 CAFR data set described in this report.

Although Duncan's SEI index is not available for these cities in 2010, we use median family income in its place. Data on median family income for cities is taken from the 2009-2011 three year estimates from American Community Survey. To make median family income more comparable to Depression-era SEI values, we divide by 2,000.

Using this equation to predict current-period defaults allows us to shed light on the extent to which we should have confidence in the ability of a historical model that only relies on statistical evidence to predict future defaults. To do this we make use of the 2012 defaults by San Bernardino and Stockton; if this model is able to predict defaults in these cities, or at least if our model could have indicated that these cities were at risk based on 2011 data, then one can have confidence that the model can be used as the basis for predicting future defaults.

Would this analysis have predicted defaults in Stockton and San Bernardino? Table A11 below describes the location of these two defaulting cities in the default likelihood ranking produced by following the procedure described above. Table A11 is divided into quintiles, which we have labeled Very safe, Safe, Moderately safe, At risk and Highly at risk. Our model did a fairly good job of placing San Bernardino, as our rankings put it in the "Highly at risk" category. Only 30 cities were predicted to be more likely to default than

San Bernardino in our rankings. On the other hand, our rankings placed Stockton only in the “At risk” category.

Table A11: Location of Defaulting Cities in Default Likelihood Ranking

Cites ranking	Risk level	Number of defaults
1-50	Very safe	0
51-100	Safe	0
101-150	Moderately safe	0
151-200	At risk	1
201-259	Highly at risk	1

Thus the ranking produced from the analysis of the Great Depression data only, which did not take into account any case study evidence or analysis of any current period data, correctly placed San Bernardino and Stockton in the bottom two quintiles. Of course, this is hardly a perfect ranking, and the hybrid model we described in the text did a far better job of ranking these cities. This illustrates the virtues of our hybrid approach, which incorporated case study evidence and analysis of current period data. However, the extensions we have presented in this Appendix illustrate how analysts could have predicted defaults after Vallejo but before Stockton, by leveraging historical data. Our hybrid model thus incorporates all relevant information, not just Depression-era historical evidence.

We also note that a socioeconomic variable reached our final specification. Thus, while the main text excludes socioeconomic variables and offers a case against their use, we understand that other researchers may wish to build upon our work by including SEI as an independent variable.

Appendix 2: Comparing data: CAFRs and the SCO's Cities Annual Report

The State Controller's Office (1912, 2012b) has been collecting fiscal data from all California cities on an annual basis since 1911 and publishes this data in a document entitled the Cities Annual Report. Since the report leverages a long established collection mechanism and cooperation from cities is legally mandated, it could theoretically serve as the data source for a default probability modeling tool. Unfortunately, some aspects of the SCO reporting mechanism limit its effectiveness for this purpose.

The report is published shortly after September 1 on a one year lag. Since the data pertain to the fiscal year ending June 30, the report becomes available at least 14 months after the year closes. For example, data as of June 30, 2012 will be available around September 1, 2013. By contrast, CAFRs for the last fiscal year began appearing in late 2012, and most were available by the end of April 2013.

On the other hand, some cities file audited financials on a very long delay. As of April 2013, the City of Adelanto had yet to publish any CAFR more recent than 2008, while Maywood's most recent audited financials were as of 2009. Both of these cities did provide 2011 fiscal data for the SCO report. At the same time, Cities Annual Report coverage for 2011 is also incomplete. The cities of Beaumont, Hawthorne, La Habra and Stockton did not provide SCO with 2011 financial data.

Another concern is that the Cities Annual Report data are collected through a process separate from that used to create the city's audited financials. City finance personnel enter data into a custom data collection instrument provided by SCO. Since the reporting process is divorced from the CAFR filing, the data may not be consistent.

To assess the applicability of the Cities Annual Report data to municipal default probability modeling, we compared data from the report to selected data obtained from four 2011 CAFRs. Since our proposed model is driven in part by interest expenses, pension contributions and total revenues, we investigated these data items. CAFR data presented in these comparisons is not the same data we used in our modeling. Our models are based on governmental fund totals; here we also included proprietary fund data for greater comparability with the SCO reporting scope.

Interest expenses and pension contributions do not appear on the printed reports issued by SCO. Expenditures are only presented by function, and apparently these two categories are distributed across a variety of functions. However, the Controller’s data collection instrument includes a form for reporting city expenditures by major object. This form provides fields for interest and retirement expenses. Data from these forms is aggregated into a Microsoft Access database available from SCO.

The following table compares CAFR and Cities Annual Report interest expense data (from the Access database) for four cities:

City	2011 Interest Expenditure in City’s CAFR	2011 Interest Expenditure in Cities Annual Report
Los Angeles	789,232,000	783,740,036
San Francisco	426,809,000	476,954,176
Twenty-Nine Palms	43	185,541
Walnut Creek	0	269,960

None foot precisely, but Los Angeles is quite close. Small discrepancies may be explained by the use of the modified accrual basis in CAFRs and budgetary basis in the SCO filings. Walnut Creek was one of many cities that did not report interest expense in the Cities Annual Report. According to Walnut Creek’s CAFR, its 2011 Interest Expense is

attributable to the Redevelopment Agency. It is possible that finance department personnel did not recognize the RDA as an entity of the city when it filed the SCO survey.

The following table compares CAFR and Cities Annual Report retirement expense data for four cities:

City	2011 Pension Contributions in City's CAFR	2011 Retirement Expenditure in Cities Annual Report
Los Angeles	909,831,000	796,818,827
San Francisco	308,823,000	343,717,123
Twenty-Nine Palms	462,025	481,867
Walnut Creek	5,174,324	7,398,862

In this case, orders of magnitude are consistent, but there are some significant differences. In the two instances in which the CAFR is significantly lower than the Cities Annual Report, it does not appear that OPEBs explain the difference. The survey includes a space for employee benefits, and it is likely that OPEBs are being included in this area.

Total annual revenues also show significant differences. The following table compares totals from the Statement of Revenues, Table 3 of the Cities Annual Report with aggregated revenues from CAFRs. The CAFR figures include governmental activities, business type activities and (in the case of Los Angeles and San Francisco) discretely reporting components. Had discretely reporting components been excluded the differences would have been larger.

City	2011 Total Revenue from CAFR	2011 Total Revenue in Cities Annual Report	Percentage Difference
Los Angeles	12,487,588,000	15,350,476,263	23%
San Francisco	7,194,037,000	8,527,657,110	19%
Twenty-Nine Palms	13,496,119	11,506,379	-15%
Walnut Creek	78,954,620	72,568,863	-8%

The Cities Annual Report divides revenue into two classifications: General and Functional. This appears to be similar to the distinction made in CAFRS between General and Program Revenues. The next two tables provide comparisons of these two revenue classifications.

City	2011 General Revenue from CAFR	2011 General Revenue in Cities Annual Report	Percentage Difference
Los Angeles	3,772,484,000	3,179,675,409	-16%
San Francisco	2,800,234,000	2,367,318,396	-15%
Twenty-Nine Palms	9,479,104	7,861,060	-17%
Walnut Creek	48,773,218	43,398,485	-11%

City	2011 Program Revenue from CAFR	2011 Functional Revenue in Cities Annual Report	Percentage Difference
Los Angeles	8,715,104,000	12,170,800,854	40%
San Francisco	4,393,803,000	6,160,338,714	40%
Twenty-Nine Palms	4,017,015	3,645,319	-9%
Walnut Creek	30,181,402	29,170,378	-3%

Given both the delayed appearance of the Cities Annual Report and significant differences from audited financials, we do not believe the SCO report provides a viable platform for a credit scoring system using the variables we have identified.

Appendix 3: Default Probability Scores for California Cities

In this appendix we present default probability estimates for 262 California cities with population over 25,000 that have filed Comprehensive Annual Financial Reports in fiscal year 2011 or 2012.

In 2012, there were 265 California cities with population in excess of 25,000. All of these cities had filed 2011 Comprehensive Annual Financial Reports by April 2013 with the exception of Adelanto, Bell and Maywood. We do not report results for these three cities, but note that their apparent failure to produce audited financials 22 months after the close of the 2011 fiscal year should be a cause for concern. Eastvale, a new city, did not have prior year revenues in 2011, so we exclude it from the 2011 reported population as well.

CITY	2011 DP	2012 DP
Alameda	0.88%	#N/A
Alhambra	1.28%	1.45%
Aliso Viejo	0.01%	0.01%
Anaheim	1.55%	1.61%
Antioch	1.39%	1.26%
Apple Valley	0.10%	0.18%
Arcadia	0.59%	0.65%
Atascadero	0.80%	0.81%
Atwater	2.44%	4.28%
Azusa	0.65%	#N/A
Bakersfield	0.87%	0.91%
Baldwin Park	0.35%	0.41%
Banning	0.23%	#N/A
Beaumont	0.50%	#N/A
Bell Gardens	0.02%	0.02%
Bellflower	0.09%	0.11%
Belmont	0.98%	0.96%
Benicia	1.32%	1.14%
Berkeley	1.02%	1.11%
Beverly Hills	0.40%	0.39%

CITY	2011 DP	2012 DP
Brawley	0.93%	#N/A
Brea	0.72%	0.58%
Brentwood	0.65%	0.69%
Buena Park	0.37%	0.51%
Burbank	0.49%	0.64%
Burlingame	0.84%	0.71%
Calexico	1.09%	0.87%
Camarillo	0.03%	0.03%
Campbell	0.41%	0.43%
Carlsbad	0.05%	0.05%
Carson	0.82%	0.68%
Cathedral City	0.90%	1.07%
Ceres	0.77%	1.06%
Cerritos	0.00%	0.03%
Chico	1.47%	1.68%
Chino	0.43%	0.49%
Chino Hills	0.37%	#N/A
Chula Vista	1.11%	1.33%
Citrus Heights	0.05%	0.07%
Claremont	0.48%	#N/A
Clovis	1.25%	1.28%
Coachella	0.46%	0.52%
Colton	1.71%	1.89%
Compton	6.52%	#N/A
Concord	1.18%	1.40%
Corona	0.40%	0.49%
Costa Mesa	0.58%	0.58%
Covina	0.77%	0.98%
Culver City	0.36%	0.49%
Cupertino	0.48%	0.32%
Cypress	0.15%	0.13%
Daly City	0.50%	1.08%
Dana Point	0.41%	0.37%
Danville	0.02%	0.03%
Davis	1.41%	1.54%
Delano	0.64%	0.80%
Desert Hot Springs	0.27%	0.65%
Diamond Bar	0.40%	0.35%

CITY	2011 DP	2012 DP
Downey	1.06%	1.15%
Dublin	0.06%	0.07%
East Palo Alto	0.10%	0.08%
Eastvale	#N/A	0.27%
El Cajon	0.81%	0.91%
El Centro	0.49%	0.33%
El Monte	0.60%	#N/A
El Paso de Robles (Paso Robles)	0.83%	0.54%
Elk Grove	0.57%	0.54%
Encinitas	0.17%	0.28%
Escondido	0.73%	0.67%
Eureka	1.62%	#N/A
Fairfield	1.55%	1.45%
Folsom	1.91%	1.84%
Fontana	0.26%	0.71%
Foster City	0.40%	0.41%
Fountain Valley	0.18%	0.17%
Fremont	1.20%	1.11%
Fresno	1.58%	#N/A
Fullerton	1.43%	1.53%
Garden Grove	0.54%	1.06%
Gardena	1.16%	1.16%
Gilroy	0.32%	0.36%
Glendale	0.32%	1.00%
Glendora	0.32%	0.48%
Goleta	0.01%	0.34%
Hanford	0.44%	#N/A
Hawthorne	0.37%	#N/A
Hayward	0.77%	1.01%
Hemet	1.57%	0.51%
Hesperia	0.40%	0.36%
Highland	0.41%	0.42%
Hollister	1.36%	1.74%
Huntington Beach	1.01%	#N/A
Huntington Park	0.15%	#N/A
Imperial Beach	0.05%	#N/A
Indio	1.90%	1.70%
Inglewood	1.33%	#N/A

CITY	2011 DP	2012 DP
Irvine	0.35%	0.38%
La Habra	0.70%	#N/A
La Mesa	0.64%	0.63%
La Mirada	0.02%	0.02%
La Puente	0.02%	0.33%
La Quinta	0.00%	0.00%
La Verne	0.96%	0.87%
Laguna Hills	0.70%	0.75%
Laguna Niguel	0.00%	0.00%
Lake Elsinore	0.49%	0.67%
Lake Forest	0.05%	0.29%
Lakewood	0.06%	0.05%
Lancaster	0.14%	0.41%
Lawndale	0.05%	#N/A
Lemon Grove	0.85%	0.81%
Lincoln	0.88%	#N/A
Livermore	0.92%	0.92%
Lodi	1.33%	1.35%
Lompoc	1.27%	1.39%
Long Beach	1.38%	#N/A
Los Altos	0.99%	1.02%
Los Angeles	1.51%	1.48%
Los Banos	0.58%	0.55%
Los Gatos	0.26%	0.30%
Lynwood	1.20%	1.35%
Madera	0.95%	#N/A
Manhattan Beach	0.80%	0.81%
Manteca	0.77%	0.69%
Martinez	0.91%	0.88%
Mendota	0.18%	0.22%
Menifee	0.62%	0.76%
Menlo Park	0.72%	0.53%
Merced	0.68%	0.74%
Milpitas	0.13%	0.60%
Mission Viejo	0.38%	0.59%
Modesto	1.36%	1.42%
Monrovia	1.82%	3.75%
Montclair	0.75%	#N/A

CITY	2011 DP	2012 DP
Montebello	1.75%	1.71%
Monterey	0.54%	0.55%
Monterey Park	0.84%	0.73%
Moorpark	1.13%	0.88%
Moreno Valley	0.51%	0.59%
Morgan Hill	1.08%	1.16%
Mountain View	0.25%	0.23%
Murrieta	0.38%	0.61%
Napa	1.04%	#N/A
National City	0.79%	0.63%
Newark	0.99%	0.91%
Newport Beach	0.60%	0.51%
Norco	1.65%	1.41%
Norwalk	0.35%	0.46%
Novato	0.21%	0.17%
Oakland	0.67%	0.61%
Oakley	0.19%	0.13%
Oceanside	1.04%	1.01%
Ontario	0.59%	0.59%
Orange	0.54%	#N/A
Oxnard	1.03%	1.10%
Pacifica	1.44%	1.41%
Palm Desert	0.02%	0.02%
Palm Springs	0.98%	1.03%
Palmdale	0.48%	#N/A
Palo Alto	0.91%	0.94%
Paradise	1.45%	#N/A
Paramount	0.24%	0.25%
Pasadena	1.20%	1.14%
Perris	0.16%	0.13%
Petaluma	2.01%	1.88%
Pico Rivera	0.08%	0.05%
Pittsburg	0.61%	0.54%
Placentia	1.03%	1.14%
Pleasant Hill	0.29%	0.59%
Pleasanton	1.00%	1.03%
Pomona	1.78%	2.04%
Porterville	0.18%	0.38%

CITY	2011 DP	2012 DP
Poway	0.05%	0.06%
Rancho Cordova	0.37%	0.37%
Rancho Cucamonga	0.20%	0.15%
Rancho Palos Verdes	0.14%	0.11%
Rancho Santa Margarita	0.15%	0.13%
Redding	1.55%	1.72%
Redlands	0.70%	0.67%
Redondo Beach	1.26%	1.27%
Redwood City	0.99%	1.15%
Rialto	0.42%	#N/A
Richmond	0.97%	#N/A
Ridgecrest	2.14%	4.58%
Riverside	0.50%	1.32%
Rocklin	0.18%	#N/A
Rohnert Park	0.93%	0.79%
Rosemead	0.27%	0.27%
Roseville	0.56%	0.65%
Sacramento	1.54%	1.35%
Salinas	1.50%	1.67%
San Bernardino	2.09%	#N/A
San Bruno	0.96%	0.76%
San Buenaventura	0.82%	0.98%
San Carlos	0.38%	0.28%
San Clemente	0.45%	0.51%
San Diego	1.30%	1.23%
San Dimas	0.03%	0.02%
San Francisco	1.38%	1.30%
San Gabriel	0.95%	1.34%
San Jacinto	0.03%	0.03%
San Jose	1.30%	1.20%
San Juan Capistrano	0.73%	0.63%
San Leandro	0.90%	1.24%
San Luis Obispo	1.03%	1.06%
San Marcos	0.29%	0.44%
San Mateo	1.35%	1.15%
San Pablo	0.09%	0.07%
San Rafael	1.60%	1.83%
San Ramon	1.34%	1.19%

CITY	2011 DP	2012 DP
Santa Ana	1.74%	1.55%
Santa Barbara	1.03%	1.09%
Santa Clara	0.64%	0.15%
Santa Clarita	0.09%	0.29%
Santa Cruz	1.17%	0.74%
Santa Maria	0.13%	0.29%
Santa Monica	0.15%	0.08%
Santa Paula	1.15%	#N/A
Santa Rosa	1.11%	1.00%
Santee	0.99%	1.08%
Saratoga	0.65%	0.47%
Seaside	1.13%	1.25%
Simi Valley	0.34%	0.39%
Soledad	1.68%	#N/A
South Gate	0.10%	0.11%
South Pasadena	0.39%	#N/A
South San Francisco	1.15%	1.14%
Stanton	0.02%	0.09%
Stockton	1.72%	#N/A
Suisun City	0.39%	#N/A
Sunnyvale	0.32%	0.25%
Temecula	0.42%	0.57%
Temple City	0.00%	0.00%
Thousand Oaks	0.10%	0.09%
Torrance	0.97%	0.99%
Tracy	0.56%	0.60%
Tulare	0.73%	0.85%
Turlock	0.43%	#N/A
Tustin	0.00%	0.00%
Twentynine Palms	0.03%	#N/A
Union City	1.03%	1.20%
Upland	1.51%	2.06%
Vacaville	1.78%	2.07%
Vallejo	1.12%	#N/A
Victorville	1.82%	1.73%
Visalia	0.55%	0.71%
Vista	0.64%	0.31%
Walnut	0.08%	0.11%

CITY	2011 DP	2012 DP
Walnut Creek	0.62%	0.75%
Wasco	1.39%	1.22%
Watsonville	2.02%	1.61%
West Covina	0.59%	0.69%
West Hollywood	0.07%	0.08%
West Sacramento	0.56%	0.43%
Westminster	0.49%	0.59%
Whittier	0.40%	0.50%
Wildomar	1.30%	1.25%
Windsor	0.19%	0.25%
Woodland	1.19%	#N/A
Yorba Linda	0.02%	#N/A
Yuba City	0.39%	0.66%
Yucaipa	0.00%	0.00%

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