



Munich Personal RePEc Archive

## **Welfare costs of reclassification risk in the health insurance market**

Pashchenko, Svetlana and Porapakkarm, Ponpoje

Uppsala University, University of Macau

19 October 2011

Online at <https://mpra.ub.uni-muenchen.de/48013/>  
MPRA Paper No. 48013, posted 05 Jul 2013 04:21 UTC

# Welfare costs of reclassification risk in the health insurance market <sup>\*</sup>

Svetlana Pashchenko<sup>†</sup>

Ponpoje Porapakkarm<sup>‡</sup>

Uppsala University

University of Macau

June 23, 2013

## Abstract

One of the major problems of the U.S. health insurance market is that it leaves individuals exposed to reclassification risk. Reclassification risk arises because the health conditions of individuals evolve over time, while a typical health insurance contract only lasts for one year. A change in the health status can lead to a significant change in the health insurance premium. We study how costly this reclassification risk is for the welfare of consumers. More specifically, we use a general equilibrium model to quantify the implications of introducing guaranteed renewable contracts into the economy calibrated to replicate the key features of the health insurance system in the U.S. Guaranteed renewable contracts are private insurance contracts that can provide protection against reclassification risk even in the absence of consumer commitment or government intervention. We find that the welfare effects from introducing this type of contracts are small implying that the presence of reclassification risk does not impose large welfare losses on consumers. This happens because two institutions in the current U.S. system substitute for the missing explicit contracts that insure reclassification risk: employer-sponsored health insurance and government means-tested transfers. If these two institutional features are removed, the average welfare gains from having access to guaranteed renewable contracts are large and can exceed 2% of the annual consumption.

Keywords: health insurance, reclassification risk, dynamic insurance, guaranteed renewable contracts, general equilibrium

JEL Classification Codes: D52, D58, D91, G22, I11

---

<sup>\*</sup>We thank all seminar participants at the University of Virginia, Federal Reserve Bank of Richmond, QSPS Summer Workshop, PET Meeting in Bloomington, SED Meeting in Ghent, GRIPS, Asian Meeting of the Econometric Society in Seoul, Uppsala University, and especially, Huberto Ennis, John Jones, and Eric Young, for their comments and suggestions. Porapakkarm acknowledges financial support from the Research and Development Administration Office at the University of Macau. All errors are our own.

<sup>†</sup>Email: sap9v@virginia.edu

<sup>‡</sup>Email: ponpojep@umac.mo

# 1 Introduction

An important feature of the health insurance market is that a typical insurance policy only lasts for one year while a disease can last for any period of time. This creates the problem of reclassification risk - a risk to face a drastic increase in health premiums when one's health status deteriorates. The fact that standard health insurance contracts leave individuals exposed to reclassification risk is considered an important market failure in the health insurance market (Hendel and Lizzeri, 2003; Diamond, 1992). Moreover, the observation that sick people face high health insurance premiums was an important argument for an additional regulation on health insurance market during 2009-2010 health reform debate.<sup>1</sup> The goal of this paper is to evaluate how important is the lack of protection against reclassification risk for the welfare of consumers.

One way to do this is to compare the current system with the first best solution to the problem of reclassification risk. The first best is to enroll everyone into a long-term health insurance contract. The price of such a contract depends on the average expected medical expenses of all participants. In other words, healthy people make transfers to the sick equalizing the insurance price for all risk categories. These contracts require consumer's commitment because healthy individuals will tend to drop out. As shown by Cochrane (1995), the lack of commitment can be overcome by introducing a special arrangement such as illiquid accounts<sup>2</sup>. Another problem with ensuring participation in these contracts is incomplete labor markets. Since premiums are based on average medical expenses but not on individual income, consumers experiencing a sequence of bad income shocks may be unable to pay the premium. This can be solved by introducing income-based transfers. However, since all income redistributive measures have a non-trivial effect on welfare, in the presence of these transfers it is hard to measure a pure welfare effect of reclassification risk.

---

<sup>1</sup>This debate resulted in the major health reform bill that was signed in 2010. Among other things, this bill forbids insurance companies to risk-adjust premiums in the individual market. This can be considered as a regulatory way to remove reclassification risk. We discuss implications of our results for the health reform in Section 7.

<sup>2</sup>More specifically, Cochrane's idea is to substitute long-term contracts with a sequence of short-term contracts that require consumers who turn out to be healthy to make transfers to insurance firms. Illiquid accounts are needed to enforce these transfers.

To overcome this problem, we consider a special type of contract that can provide insurance against reclassification risk but does not require consumers' commitment, income-based transfers or any other special arrangements. These are guaranteed renewable contracts discussed in details by Pauly et al (1995). These contracts are front-loaded: a consumer is required to prepay part of his future premiums and this prepayment locks him into the contract. In return, a consumer is guaranteed that i) he will be able to renew his health insurance contract in the future; ii) the prespecified renewal price will be independent of his future health realizations. A key feature of this contract is that reclassification risk is insured not by making healthy people pay for the sick but by allowing individuals to make state-contingent savings that pay off when their premiums increase.<sup>3</sup> To evaluate welfare costs of reclassification risk, we consider how much welfare improvement can be achieved from introducing guaranteed renewable contracts in the individual health insurance market.<sup>4</sup>

We construct a general equilibrium overlapping generations model where people face uninsurable labor income risk and medical expense risk that can be partially insured. Several types of health insurance are available. First, some individuals have access to employer-based insurance. Second, lowest-income individuals can get Medicaid. Finally, all individuals can buy insurance policy directly in the individual market. In the individual market premiums are risk-rated, i.e. depend on the current health conditions of individuals. All policies last for one year while medical expenditures are persistent, which creates the problem of reclassification risk.

Our model reflects two institutional features that are important when evaluating the importance of reclassification risk in the U.S. health insurance markets. First, a large

---

<sup>3</sup>It is important to point out the fundamental difference between these guaranteed renewable contracts and the regulatory guaranteed renewability provision that exists in some states of the US. The guaranteed renewability provision forbids insurance companies to deny coverage to individuals who already have an insurance contract and want to renew it. However, this provision does not require that the price at renewal is fixed at some prespecified level. In contrast, the key feature of guaranteed renewable contracts is that they allow to predetermine the renewal price in advance.

<sup>4</sup>More specifically, in order to measure welfare costs of reclassification risk we introduce a frictionless market offering guaranteed renewable contracts. An alternative modeling strategy is to allow this market to exist in the baseline economy but with frictions that prevent people from using it. Then we can quantify how large are these frictions. Our results are robust to this alternative modeling strategy. In Appendix D we show that relatively small fixed costs are enough to deter people from buying guaranteed renewable contracts, which possibly explain why we do not observe these contracts offered in reality.

fraction of non-elderly adults gets their insurance from employer-based market. This market is community rated, i.e. premiums are independent of the health conditions of individuals. People with permanent access to this market are protected from the risk of premium fluctuations. Also, lowest-income individuals can get public insurance from Medicaid for free. Second, for people who face high medical shock and/or bad labor income shock, the government provides protection in the form of the consumption minimum floor. This consumption floor can also mitigate the consequences of the lack of an explicit insurance against reclassification risk.

We calibrate the model using the Medical Expenditure Panel Survey dataset to match the key insurance statistics for the U.S. Using the calibrated model we study the quantitative implications of introducing frictionless guaranteed renewable contracts in the individual market.

We find that comparing to the situation when only standard short-term insurance contracts are available, introduction of guaranteed renewable contracts can noticeably decrease uninsurance rates - from 25.9% to 19.4% due to the higher participation in the individual insurance market. Also, if both standard and guaranteed renewable contracts are available, most of the consumers prefer to buy the later type of contract. Our results show that people who hold guaranteed renewable contracts face almost no fluctuations in their health insurance premiums even if their health deteriorates. This implies that these contracts provide a good protection against reclassification risk.

In terms of welfare, we find that introduction of guaranteed renewable contracts brings small welfare gains. This suggests that in the current U.S. health insurance system people are not very concerned about the absence of an explicit insurance against reclassification risk. This happens because two institutional features provide good implicit insurance against reclassification risk. First, employer-sponsored health insurance that protects mostly high-income people; and second, the consumption minimum floor that protects mostly people with low income. If these two institutional features are removed, the average welfare gains from having access to guaranteed renewable contracts are large and can exceed 2% of the annual consumption. Our results are robust to the alternative

design of guaranteed renewable contracts and the degree of actuarial unfairness in the health insurance market.

This paper is structured as follows. Section 2 reviews the related literature. Section 3 illustrates how a guaranteed renewable contract works using a simple example. Section 4 presents the model. Section 5 explains our calibration. Section 6 discusses the results. Section 7 considers implications of our results for the upcoming health insurance reform. Section 8 concludes.

## 2 Related literature

This paper belongs to two strands of literature. First is the literature studying how private markets can provide insurance against reclassification risk if buyers cannot commit to a contract. A seminal paper in this area is Cochrane (1995) who characterizes a set of contracts that can provide long-term health insurance in such an environment. His insight is to combine standard one-period insurance contracts with premium insurance, i.e. insurance against future premium fluctuations. One requirement for such premium insurance to work is that each consumer needs to open a special account that works as a clearing house between him and the insurance company. An important condition is that consumers cannot freely withdraw money from this account. One special case in this set of contracts that can work without a special account are front-loaded guaranteed renewable contracts. These contracts were studied in more details by Pauly et al (1995) who showed that guaranteed renewable contracts can provide a good degree of reclassification risk insurance without creating liquidity problem if consumers buy them while still young and healthy. Front-loaded contracts were also studied by Hendel and Lizzeri (2003) for the case of life insurance market. They showed that the structure of premiums in this market is consistent with front-loaded contracts that emerge in the absence of consumer commitment. However, Fang and Kung (2010) and Daily et al (2008) showed that the growing life settlement market can limit the degree of reclassification risk insurance that life insurers can provide. Finkelstein et al. (2005) studied front-loaded contracts in the

long-term care insurance market and showed that the amount of front-loading currently existing is not enough to lock consumers into the contracts. To our knowledge, our paper is the first one that studies welfare effects of guaranteed renewable contracts in the health insurance market in a general equilibrium framework and taking into account the existing institutions in the US.

The second strand of literature this paper belongs to studies quantitative heterogeneous agent models with incomplete markets augmented by (i) medical expense shocks and (ii) health insurance markets where individuals can partially insure these shocks. This branch of incomplete market literature has emerged recently and includes, among others, papers by Kitao and Jeske (2009) who study subsidies for employer-based insurance, Hansen et al (2012) who evaluate the consequences of expanding Medicare program, Hsu (2009) who studies the effect of private health insurance on savings, and Pashchenko and Porapakarm (2013) who study the current health reform in the U.S. These studies consider an environment when only one type of contract is available in the individual health insurance market. Our contribution to this literature is that we expand the contract space by allowing insurance firms to offer guaranteed renewable contracts. Since it is a long-term contract, this extension involves solving a dynamic contracting problem within a general equilibrium framework.

### 3 Simple illustration

This section constructs a simple example that illustrates how a guaranteed renewable contract works. Consider an individual whose health is good, and the price he pays for a standard one-period health insurance contract is  $p_L$ . With probability  $v$  an individual may still be in good health in the next period, in which case his health insurance premium will stay unchanged. However, with probability  $1 - v$  his health status may deteriorate. If this happens, his health insurance premium for the standard contract will raise to  $p_H$ , where  $p_H > p_L$ . If an individual buys the standard one-period contract, he is exposed to reclassification risk - the risk that his health premium will rise from  $p_L$  to  $p_H$ .

Suppose an individual has an option to buy a guaranteed renewable contract at the price  $p_1^{GR}$ . This contract insures his medical expenditure in the next period like the standard one-period contract. On top of that, it guarantees that in the next period an individual can buy health insurance at the prespecified price  $p_2^{GR}$  *that does not depend on his health status realization*. If his health status remains the same he can buy a standard contract at price  $p_L$ . However, if his health status deteriorates he can renew his guaranteed renewable contract at price  $p_2^{GR} < p_H$ . Under the assumption of perfect competition in the insurance market, the price of such a guaranteed renewable contract is determined in the following way:

$$p_1^{GR} = p_L + (1 - v)(p_H - p_2^{GR}). \quad (1)$$

Note that the guaranteed renewable contract is more expensive than the regular one-period contract because of the front-loading part  $(1 - v)(p_H - p_2^{GR})$ . This front-loading takes into account the fact that an individual can become unhealthy but the price of renewing his health insurance ( $p_2^{GR}$ ) cannot be readjusted.

## 4 Model

### 4.1 Households

#### Demographics and preferences

The economy is populated by two overlapping generations: young and old. A young individual stays young with probability  $\zeta^y$  and becomes old with a probability  $1 - \zeta^y$ . An old individual survives to the next period with probability  $\zeta^o$ .<sup>5</sup> The population is assumed to remain constant. Old agents who die are replaced by the entry of new young

---

<sup>5</sup>We assume a stochastic aging environment because a full life-cycle model is computationally impractical in our framework. The most time-consuming part of our computations is to find equilibrium prices of guaranteed renewable contracts. In a stochastic aging model this price depends only on health status. In the full life-cycle model the price will be a function of both age and health which makes our model computationally infeasible especially when it comes to transition calculations. In Section 5.3 we explain how we adjust our calibration strategy to approximate for the key life-cycle features in our model.



agents.

An individual discounts his future utility by the discount factor  $\beta$ . Preferences are described by the CRRA utility function with the risk aversion parameter  $\sigma$ :

$$u(c) = \frac{c^{1-\sigma}}{1-\sigma}.$$

## Health insurance

An individual's health status  $h$  is indexed by  $\{1, 2, \dots, H\}$ . An increasing number implies deteriorating health status. Health status evolves according to a  $H$ -state Markov process, where  $G^y(h'|h)$  stands for the young and  $G^o(h'|h)$  for the old. The current health status of an individual determines his current medical expenditures  $x(h)$ , where  $x$  is a deterministic and strictly monotone-increasing function, different between the young and the old.<sup>6</sup> Thus, in the following, we will refer to health status ( $h$ ) and medical expenditures ( $x$ ) interchangeably.

Each young individual can buy insurance against medical expenditures in the individual insurance market where two types of contracts are offered. The first type is a standard one-year contract that covers some fraction of the next period medical expenditures. The price of this contract depends on the current health status of an individual and is denoted by  $p^I(h)$ . The second type of contract is guaranteed renewable. This contract covers a fraction of the next period's medical expenditures like a standard one-year contract. In addition, a guaranteed renewable contract provides an option to renew insurance in the following period at the same price regardless of the new health status<sup>7</sup>. Guaranteed renewable contracts do not have a termination date, i.e. an individual can renew the same contract as long as he is still young. An important condition for an individual to be able to renew this contract is continuous participation. In other words, if an individual does not renew the contract once, he will lose the option to renew it in the

---

<sup>6</sup>We represent medical expenses as exogenous shocks as opposed to an outcome of people's optimal decision making. We discuss this assumption in details in Appendix B.

<sup>7</sup>There are several ways to design a guaranteed renewable contract by changing the price that an insurer guarantees at the renewal. In our main experiments we assume that the renewal price is the same as the price of the original contract. Later on we relax this assumption by letting the renewal price to differ from the original price. Detailed discussion of these experiments is provided in section 6.

future. The premium of a *newly issued* guaranteed renewable contract is a function of the *current* health status of an individual. The price of a guaranteed renewable contract that is already *in force* is fixed and determined by the health status of an individual *at the time of the contract initiation*.

In each period, with some probability, a young individual can get an offer to buy employer-sponsored health insurance (ESHI). This is denoted by  $g$ :  $g = 1$  if an individual gets an ESHI offer,  $g = 0$  if he does not.<sup>8</sup> The out-of-pocket premium of employer-based insurance is equal to

$$\bar{p} = (1 - \psi) p.$$

Here  $p$  is the premium charged to all participants of the employer-based pool, and  $\psi$  is the fraction of this premium paid by the employer.

Low-income individuals are eligible to enroll in Medicaid that provides health insurance for free. To become eligible for Medicaid, an individual's total resources net of out-of-pocket medical expenses must be below a certain level which is denoted by  $y^{pub}$ .

We use  $i$  to index the current health insurance status as follows:

$$i = \left\{ \begin{array}{ll} -2 & ; \text{ if uninsured} \\ -1 & ; \text{ if insured by Medicaid} \\ 0 & ; \text{ if holding a standard one-period insurance or ESHI} \\ 1, 2, \dots, H & ; \text{ if holding a guaranteed renewable contract originated when} \\ & \text{ his health status equals } i. \end{array} \right\}$$

If an individual holds a guaranteed renewable contract,  $i$  keeps track of the health status when the contract was initiated. For a newly purchased contract  $i$  is the current health status  $h$ . We denote the premium for a newly issued guaranteed renewable insurance as  $p^{GR}(h)$ , and the premium for a guaranteed renewable contract that is already in force as  $p^{GR}(i)$  for  $i = \{1, 2, \dots, H\}$ .

---

<sup>8</sup>We incorporate an important feature of the data that many individuals do not have a permanent access to ESHI but can lose this access several times over their working life. We do not explicitly model COBRA which provides a possibility for some individuals to extend their access to the employer-based pool for a limited time period after they lose their ESHI offer. Modeling COBRA requires us to add an additional state variable making our computation of the transition impractical. In addition, COBRA is not effective in eliminating the problem of fragmented access to ESHI and thus will not conceptually change our model.

If a young person is insured, the insurance will cover a fraction  $q(i, x)$  of his current medical expenses. This fraction depends on his medical expenditures ( $x$ ) and the type of insurance he has ( $i$ ).

All retired households are enrolled in Medicare. Medicare charges a premium of  $p^{med}$ . We denote the fraction of medical expenses covered by Medicare by  $q^{med}(x)$ .

## Labor income

A young individual supplies labor inelastically. We denote his earnings by  $\tilde{w}z$ , where  $\tilde{w}$  is the adjusted wage per effective labor unit and  $z$  is his idiosyncratic productivity. We model the productivity, an ESHI offer, and health status as a joint Markov process. The productivity of the old is set to zero.

## Taxation and social transfers

Each household has to pay income tax  $\mathcal{T}(y)$ . The taxable income  $y$  is based on both labor income and capital income. We incorporate two features of the current U.S. tax code related to the taxation of health-related expenses into our definition of  $y$ . First, households can tax-exempt their medical expenses in excess of 7.5% of their income. Second, households buying group insurance can subtract the out-of-pocket group premium  $\bar{p}$  from their taxable income.

We also assume a social welfare system,  $T^{SI}$ . The social welfare system guarantees that a household will have a minimum consumption level at  $\underline{c}$ . This reflects the U.S. public transfer programs such as Medically Needy part of Medicaid, food stamps, Supplemental Security Income (SSI), and transfers to finance uncompensated care.<sup>9</sup>

All old individuals are retired. They receive Social Security benefits in the amount  $ss$ .

---

<sup>9</sup>This structure reflects an important feature of the US social insurance system - the coexistence of ex-ante and ex-post insurance. Ex-ante insurance is represented in our model by Medicaid. Medicaid enrollees get into the program based on their current total resources and this program covers their *next period's* medical expenses. In contrast, ex-post insurance that is represented by the consumption floor is available to individuals impoverished by their out-of-pocket medical costs *this period*.

## Optimization problem

**Retired individuals** The state variables of an old individual include liquid capital ( $k \in \mathbb{K} = R^+ \cup \{0\}$ ) and health status ( $h \in \mathbb{H} = \{1, 2, \dots, H\}$ ). The value function of the old can be written as follows:

$$\mathbf{V}^o(k, h) = \max_{c, k'} u(c) + \beta \zeta^o E_t \mathbf{V}^o(k', h') \quad (2)$$

$$s.t. \quad k(1+r) + ss + T^{SI} = c + \zeta^o k' + x(1 - q^{med}(x)) + p^{med} + \mathcal{T}(y) \quad (3)$$

where

$$T^{SI} = \max(0, \underline{c} + x(1 - q^{med}(x)) + \mathcal{T}(y) + p^{med} - ss - k(1+r)) \quad (4)$$

$$y = \max(0, \tilde{y}) \quad (5)$$

$$\tilde{y} = rk + ss - \max(0, x(1 - q^{med}(x)) - 0.075(rk + ss)) \quad (6)$$

Equation (3) is the budget constraint. We assume that there is an actuarially-fair annuity market. Thus each retired individual needs to save only  $\zeta^o k'$  instead of  $k'^{10}$ . Equation (6) takes into account the tax-deductibility of medical expenses in excess of 7.5% of the total income.<sup>11</sup>

**Young individuals** The state variables for a young individual include liquid capital ( $k \in \mathbb{K} = R^+ \cup \{0\}$ ), health status ( $h \in \mathbb{H} = \{1, 2, \dots, H\}$ ), idiosyncratic labor productivity ( $z \in \mathbb{Z} = R^+$ ), ESHI offer status ( $g \in \mathbb{G} = \{0, 1\}$ ), and index of health insurance status ( $i \in \mathbb{I} = \{-2, -1, 0, 1, 2, \dots, H\}$ ).

Each period an individual chooses his consumption ( $c$ ), saving ( $k'$ ), and health insurance status for the next period ( $i^H$ ). Depending on one's Medicaid eligibility, ESHI offer and insurance status, he can choose not to buy any insurance ( $NB$ ), buy a guaranteed renewable contract ( $BGR$ ), renew the existing guaranteed renewable contract ( $RGR$ ), buy a standard individual policy ( $BI$ ), buy a group insurance ( $BG$ ), or enroll in Medicaid

<sup>10</sup>Alternatively, one can assume that the accidental bequests are evenly distributed to all young. Since the distributed amount is small, it will not affect our results. But the computational cost is higher since one needs to wait until the convergence of total bequests to get the invariant distribution.

<sup>11</sup>The problem of a newly retired household is slightly different from a retired household since he is still covered by his pre-retirement insurance. The difference lies in the state variables and the out-of-pocket medical expenditure. For the newly retired, the state variables are  $\{k, h, i\}$ ; and in the budget constraint  $x(1 - q^{med}(x))$  is replaced by  $x(1 - q(i, x))$ .

(*BM*). We summarize the insurance choices as follows.

- If a household currently has a guaranteed renewable contract,  $i = \{1, 2, 3, \dots, H\}$ ,<sup>12</sup>

$$i^H = \begin{cases} \{BGR, RGR, BI, BG, BM\} & \text{if } g = 1 \text{ and eligible for Medicaid} \\ \{BGR, RGR, BI, BM\} & \text{if } g = 0 \text{ and eligible for Medicaid} \\ \{NB, BGR, RGR, BI, BG\} & \text{if } g = 1 \text{ and not eligible for Medicaid} \\ \{NB, BGR, RGR, BI\} & \text{if } g = 0 \text{ and not eligible for Medicaid} \end{cases}$$
- If a household does not have a guaranteed renewable contract,  $i = \{-2, -1, 0\}$ ,
$$i^H = \begin{cases} \{BGR, BI, BG, BM\} & \text{if } g = 1 \text{ and eligible for Medicaid} \\ \{BGR, BI, BM\} & \text{if } g = 0 \text{ and eligible for Medicaid} \\ \{NB, BGR, BI, BG\} & \text{if } g = 1 \text{ and not eligible for Medicaid} \\ \{NB, BGR, BI\} & \text{if } g = 0 \text{ and not eligible for Medicaid} \end{cases}$$

The value function of a working-age household can be written as follows:

$$\mathbf{V}^y(k, h, z, g, i) = \max_{c, k', i^H} u(c) + \beta \zeta^y E \mathbf{V}^y(k', h', z', g', i') + \beta(1 - \zeta^y) E \mathbf{V}^o(k', h', i') \quad (7)$$

$$s.t. \quad k(1+r) + \tilde{w}z + T^{SI} = c + k' + x(1 - q(i, x)) + P(h, i, i^H) + \mathcal{T}(y) \quad (8)$$

where

$$\tilde{w} = \begin{cases} w & ; & \text{if } g = 0 \\ w - c_E & ; & \text{if } g = 1 \end{cases} \quad (9)$$

$$P(h, i, i^H) = \begin{cases} 0 & ; & \text{if } i^H = NB \text{ or } BM \\ p^I(h) & ; & \text{if } i^H = BI \\ p^{GR}(h) & ; & \text{if } i^H = BGR \\ p^{GR}(i) & ; & \text{if } i^H = RGR \\ \bar{p} & ; & \text{if } i^H = BG \end{cases} \quad (10)$$

$$y = \max(0, \tilde{y}) \quad (11)$$

$$\tilde{y} = \begin{cases} \tilde{w}z + rk - \max(0, x(1 - q(i, x)) - 0.075(\tilde{w}z + rk)) & ; & \text{if } i^H \neq BG \\ \tilde{w}z + rk - \max(0, x(1 - q(i, x)) - 0.075(\tilde{w}z + rk)) - \bar{p} & ; & \text{if } i^H = BG \end{cases} \quad (12)$$

$$T^{SI} = \max(0, \underline{c} + x(1 - q(i, x)) + \mathcal{T}(y) - \tilde{w}z - k(1+r)) \quad (13)$$

---

<sup>12</sup>Note, that if a household is eligible for Medicaid he cannot stay uninsured because Medicaid is free.

$$i' = \left\{ \begin{array}{ll} -2 & ; \text{ if } i^H = NB \\ -1 & ; \text{ if } i^H = BM \\ 0 & ; \text{ if } i^H = \{BI, BG\} \\ i & ; \text{ if } i^H = RGR \\ h & ; \text{ if } i^H = BGR \end{array} \right\} \quad (14)$$

The conditional expectation on the right-hand side of equation (7) is over  $\{h', z', g'\}$ . The second equation is the budget constraint. In equation (9),  $w$  is the wage per effective labor unit. If a household has an ESHI offer, then the employer partly pays for the premium. In order to break even, the employer deducts  $c_E$  from the wage per effective labor unit to get an adjusted wage  $\tilde{w}$ . Equation (12) reflects the tax deductibility of the ESHI premium and medical expenses exceeding 7.5% of the income. Equation (14) maps the current health insurance status and health insurance choices into the next period health insurance status. The income eligibility of Medicaid program requires that

$$k(1+r) + \tilde{w}z - x(1 - q(i, x)) \leq y^{pub}.$$

**Distribution of households** To simplify the notations, we denote the space of a household' state variables by  $\mathbb{S}$ :  $\mathbb{S} \equiv \mathbb{K} \times \mathbb{H} \times \mathbb{Z} \times \mathbb{G} \times \mathbb{I}$  for young individuals,  $\mathbb{S} \equiv \mathbb{K} \times \mathbb{H} \times \mathbb{I}$  for just-retired individuals, and  $\mathbb{S} \equiv \mathbb{K} \times \mathbb{H}$  for retirees. Let  $\mathbf{s} \in \mathbb{S}$  and denote by  $\Gamma^y(\mathbf{s})$  and  $\Gamma^o(\mathbf{s})$  the measure of young and retired people correspondingly.

## 4.2 Production sector

There are two stand-in firms that act competitively. Their production functions are Cobb-Douglas,  $AK^\alpha L^{1-\alpha}$ , where  $K$  and  $L$  are aggregate capital and aggregate labor and  $A$  is the total factor productivity. The first stand-in firm offers ESHI to its workers. The second stand-in firm does not<sup>13</sup>. Under the competitive market assumption, the second firm pays each employee his marginal product of labor. Because capital is freely allocated between the two firms, the Cobb-Douglas production function implies that the

<sup>13</sup>An alternative setup is that there are two islands, one offers ESHI and the other does not. Workers are stochastically allocated between the two islands but there are no frictions in the capital market. Inside each island, the labor market is competitive.

capital-labor ratios of both firms are the same. Consequently we have<sup>14</sup>

$$r = \alpha AK^{\alpha-1}L^{1-\alpha} - \delta, \quad (15)$$

$$w = (1 - \alpha)AK^\alpha L^{-\alpha} \quad (16)$$

where  $\delta$  is the depreciation rate.

The first firm has to partially finance health insurance premiums for its employees. These costs are fully passed on to the employees through a wage reduction. In specifying this wage reduction we follow Jeske and Kitao (2009). The first firm subtracts an amount of  $c_E$  from the marginal product per effective labor. The total wage reduction of each employee with an ESHI offer is  $c_E z$ .<sup>15</sup> The zero profit condition implies

$$c_E = \frac{\psi p \left( \int \mathbf{1}_{\{i^H=BG\}} \Gamma^y(\mathbf{s}) \right)}{\int \mathbf{1}_{\{g=1\}} z \Gamma^y(\mathbf{s})}. \quad (17)$$

where  $\mathbf{1}_{\{\cdot\}}$  is a function that is equal to one if its argument is true, otherwise the function is equal to zero.

### 4.3 Private health insurance sector

We model the health insurance sector under the following assumptions. First, both individual and group insurance markets are competitive implying zero expected profit for each insurance contract. Second, there are administrative costs associated with issuing an insurance policy and these costs are proportional to the total value of the contract. Third,

---

<sup>14</sup>Define  $\{K_1, L_1\}$  and  $\{K_2, L_2\}$  as aggregate capital and labor in firms 1 and 2. Since capital can move freely between firms, the Cobb-Douglas production implies  $r + \delta = \alpha A \left( \frac{K_1}{L_1} \right)^{\alpha-1} = \alpha A \left( \frac{K_2}{L_2} \right)^{\alpha-1}$ . Next we can write

$$\frac{K}{L} = \frac{K_1 + K_2}{L_1 + L_2} = \frac{\frac{K_1}{L_1} + \frac{K_2}{L_2} \frac{L_2}{L_1}}{1 + \frac{L_2}{L_1}} = \frac{K_1}{L_1}.$$

The last equality uses the fact that  $\frac{K_1}{L_1} = \frac{K_2}{L_2}$ .

<sup>15</sup>The assumed structure implies a proportional transfer from high-income to low-income people inside the employer-based pool. This assumption is not important for our results since all changes in our study happen in the individual insurance market. An alternative assumption is a lump-sum wage reduction. This alternative structure is difficult to implement in our setup since some workers will end up earning zero or negative wage.

health insurance companies can observe only the current health status of an individual.<sup>16</sup>

### Standard one-period insurance

The zero profit condition implies that the premium for a standard one-period insurance contract is equal to the expected discounted medical costs covered by an insurance company multiplied by administrative load ( $\gamma^I$ ):

$$p^I(h) = (1+r)^{-1} \gamma^I EM(h) \quad (18)$$

Here  $EM(h)$  is the expected medical expenses of an individual with health status  $h$  covered by an insurance company:

$$EM(h) = \sum_{h'} x(h') q(0, x(h')) G^y(h'|h)$$

### Guaranteed renewable insurance

The price of a newly issued guaranteed renewable contract depends on the current health status of an individual. To determine the premium, an insurer needs to assign a probability to an event that an individual will continue to renew the contract. Consider an individual with health status  $h_t$  who chooses to buy a new guaranteed renewable contract in period  $t$ . Denote by  $\pi_{t+j}(h_{t+j}|h_t)$  an insurer's belief that this individual will continue to renew the same insurance contract every period up to a period  $t+j$  when his health status becomes  $h_{t+j}$ . The zero profit condition allows us to write the premium of a new guaranteed renewable contract as follows:

$$p^{GR}(h_t) = p^I(h_t) + \sum_{j=1}^{\infty} \frac{1}{(1+r)^j} \sum_{h_{t+j}=1}^H \pi_{t+j}(h_{t+j}|h_t) (p^I(h_{t+j}) - p^{GR}(h_t)) \quad (19)$$

---

<sup>16</sup>For standard one-period insurance contracts only health status matters for pricing. For guaranteed renewable contracts an additional factor that affects pricing is the probability that the contract will be renewed in the future. This probability depends not only on health, but also on other state variables, in particular assets and labor income. We do not allow prices to be conditioned on assets or labor income because these variables are difficult for insurance companies to verify.



The first term on the right hand side is the premium for a standard insurance contract that covers medical expenses in the next period. The second term is the extra payment for the option to renew the contract in the future. It arises because an insurance company will not be able to readjust the price in the future even if an individual's health deteriorates.

The beliefs of the insurer  $\pi_{t+j}(h_{t+j}|h_t)$  should be consistent with households' optimal decisions in equilibrium. Denote the measure of young people with health status  $h_t$  who choose to buy a new guaranteed renewable contract in period  $t$  by  $\Gamma^y(h_t, i_t^H = BGR)$ . Denote by  $\mathcal{F}(h_{t+j}, i_{t+j}^H = RGR || h_t, i_t^H = BGR)$  the measure of those people in this group who have been renewing the same contract every period from period  $t$  to period  $t+j$  when their health become  $h_{t+j}$ . Thus  $\pi_{t+j}(h_{t+j}|h_t)$  can be defined as

$$\pi_{t+j}(h_{t+j}|h_t) = \frac{\mathcal{F}(h_{t+j}, i_{t+j}^H = RGR || h_t, i_t^H = BGR)}{\Gamma^y(h_t, i_t^H = BGR)} \quad (20)$$

### Employer-based group insurance

The premium in the group insurance market does not depend on the health status of individuals<sup>17</sup>. Using the zero profit condition, the premium can be written as a weighted average of the expected covered medical costs of participating employees multiplied by the administrative load ( $\gamma^G$ ).

$$p = (1+r)^{-1} \gamma^G \frac{\int \mathbf{1}_{\{i^H=BGR\}} \times EM(h) \Gamma^y(\mathbf{s})}{\int \mathbf{1}_{\{i^H=BGR\}} \Gamma^y(\mathbf{s})}, \quad (21)$$

## 4.4 Government constraint

We assume that the government runs a balanced budget. This implies:

$$\int \mathcal{T}(y) \Gamma^y(\mathbf{s}) + \int \mathcal{T}(y) \Gamma^o(\mathbf{s}) = \int (ss + xq^{med}(x) - p^{med}) \Gamma^o(\mathbf{s}) + \int T^{SI} \Gamma^y(\mathbf{s}) + \int T^{SI} \Gamma^o(\mathbf{s}) \quad (22)$$

---

<sup>17</sup>The U.S. regulation prohibits employers to charge employees with different health-related characteristics different insurance premiums.

The left-hand side is the total income tax. The first term on the right-hand side is the net expenditure on Social Security and Medicare systems for the old. The last two terms are the costs of running the means-tested transfer program, i.e. to keep households above the consumption minimum floor.

We define the competitive equilibrium of the economy in Appendix A.

## 5 Data and Calibration

### 5.1 Data

We calibrated the model using the Medical Expenditure Panel Survey (MEPS) dataset. The MEPS collects detailed records on demographics, income, medical costs and insurance for a nationally representative sample of households. It consists of two-year overlapping panels and covers the period of 1996-2006. We use eight waves of the MEPS, from 1999 to 2007<sup>18</sup>.

The MEPS links people into one household based on eligibility for coverage under a typical family insurance plan. This Health Insurance Eligibility Unit (HIEU) defined in the MEPS dataset corresponds to our definition of a household. All statistics we use were computed for the head of the HIEU, i.e. we use individual-level data where each individual is a head of a household. We define the head as the person who has the highest income in the HIEU. A different definition of the head (based on gender) does not give significantly different results. We use longitudinal weights provided in the MEPS to compute all the statistics. Given that all individuals are observed for at most two years, we pool together all eight waves of the MEPS. Since each wave is a representation of the population in each year, the weight of each individual was divided by eight in the pooled sample.

In our sample we include all non-student heads whose age is at least 20 and whose labor income (to be defined later) is non-negative. The sample size for each wave is

---

<sup>18</sup>We do not use the first two waves of the MEPS because they do not contain the variables we use for constructing a household unit.

presented in Table 1. We use 2003 as a base year. All level variables were normalized to the base year using the Consumer Price Index (CPI).

⟨ Insert Table 1 here ⟩

When measuring the insurance status in the data, we use the following approach. In the MEPS the question about the source of insurance coverage is asked retrospectively for each month of the year. We define a person as having employer-based insurance if he reports having ESHI for at least eight months during the year (variables PEGJA-PEGDE). The same criteria was used when defining public insurance (variables PUBJA-PUBDE) and individual insurance status (variables PRIJA-PRIDE)<sup>19</sup>. In addition, we assume that a person has an ESHI offer if he reports having an offer in at least two out of three interview rounds during a year (variables OFFER31x, OFFER42x, OFFER53x).

## 5.2 Demographics, preferences and technology

The period in the model is one year. Young agents are born at age 20 and stay young on average 45 years, so the probability to stay young,  $\zeta^y$ , is set to 44/45. The survival probability of an old individual  $\zeta^o$  is set to make the fraction of the old in the population equal to 20%; thus  $1 - \zeta^o = 4(1 - \zeta^y)$ . To keep the total measure of population equal to one, the measure of newborns in every period is set to  $\frac{(1 - \zeta^y)(1 - \zeta^o)}{2 - \zeta^y - \zeta^o}$ .

The risk aversion parameter  $\sigma$  is equal to 3 which is in the range commonly used in the macroeconomic literature. The discount factor  $\beta$  is calibrated to match the aggregate capital output ratio of 3.0.

The Cobb-Douglas function parameter  $\alpha$  is set to 0.33 which corresponds to the U.S.'s capital income share. The annual depreciation rate  $\delta$  is calibrated to achieve the interest

---

<sup>19</sup>For those few individuals who switch the source of coverage during the year, we define insurance status in the following way. If a person has both ESHI and individual insurance in one year, and each coverage lasted for less than eight months but with a total duration of coverage of more than eight months, we classify this person as individually insured. Likewise, when a person has a combination of individual and public coverage that altogether lasts for more than eight months, we define that individual as having public insurance. Our results do not change significantly if we change the cutoff point to 6 or 12 months.

rate of 4.0% in the baseline economy. The total factor productivity  $A$  is normalized to make the average labor income equal to one in the baseline model.

### 5.3 Joint process of health, labor income, and ESHI offer

#### Health status and Medical expenses

The medical expenses in our model correspond to the total amount paid for the health care services (variable: TOTEXP). This includes both out-of-pocket payments and payments made by insurance companies but it does not include over-the-counter drugs. In our model there is a one-to-one mapping between medical expenses and health status. We categorize medical expenses into five bins and each bin corresponds to a different health status (Table 2).

⟨ Insert Table 2 here ⟩

The average amount of medical expenses corresponding to each health status are (0.001,0.016,0.075, 0.318,1.483) for young households and (0.021,0.083,0.251,0.917,2.317) for retired households. These numbers are based on the medical expenses in 2003/2004 wave normalized by the average labor income (\$35,624).

To construct a transition matrix for health status, we compute the fraction of household moving from one bin to another.<sup>20</sup> The resulting transition matrix for young households,  $G^y(h'|h)$ , is

0.619	0.264	0.092	0.022	0.002
0.261	0.432	0.260	0.044	0.003
0.094	0.257	0.517	0.122	0.010
0.070	0.142	0.414	0.341	0.034
0.013	0.096	0.274	0.372	0.245

---

<sup>20</sup>We assume that medical shocks follow a Markov process since MEPS allows us to observe only two consecutive periods of medical expenses for each individual. The implied first and second autocorrelation of total medical expenses are 0.37 and 0.16. French and Jones (2004) provide a detailed examination of health costs in Health and Retirement (HRS) dataset and report first and second autocorrelation equal to 0.45 and 0.34 correspondingly. It is important to point out that these numbers should be compared with caution for the following reasons. First, French and Jones's estimates are based on out-of-pocket medical expenses while ours - on total medical expenses. Second, HRS includes only people above age 50 while our estimates include young people and medical shocks for young people are less persistent. Finally, data frequency for medical expenses in MEPS is one year while in HRS it is two years.

while the transition matrix for retired households,  $G^o(h'|h)$ , is

0.626	0.225	0.111	0.037	0.001
0.257	0.416	0.265	0.058	0.005
0.131	0.324	0.427	0.108	0.011
0.090	0.170	0.455	0.242	0.043
0.056	0.174	0.388	0.336	0.046

Here the first row corresponds to  $h = 1$  and the first column corresponds to  $h' = 1$ .

### Labor income

We define labor income as a sum of wages (variable WAGEP) and 75% of income from business (variable BUSNP). This definition is the same as used in the Panel Study of Income Dynamics Dataset (PSID) that has been commonly used for income calibration in the macroeconomic literature. We categorized labor income into five quintiles ( $5 \times 20\%$ ). The labor income level in each quintile is based on the value for 2003/2004 wave normalized by the average income. These numbers are 0.091, 0.477, 0.802, 1.226, and 2.417.

The dashed lines in Panels (a) and (b) in Figure 1 show the relationship between labor income and medical expenses/health observed in the data. The hump shape in Panel (a) can be explained by the life-cycle profile of labor income. Our model does not have age dimension so the age profile of labor income is partially captured by health status. In the data households in good health ( $h = 1$ ) are more likely to be young, while those in bad health ( $h = 4$  or  $h = 5$ ) are more likely to be near retirement. These two groups tend to have lower incomes than the middle-age households.

Panel (b) also shows that the average medical expenses of households in the first income quintile are two times higher than the average medical expenses of the high income group. This pattern is driven by two facts. First, the distribution of medical expenses is highly skewed: the medical expenses of people with  $h = 5$  is more than four times higher than the medical expenses of those with  $h = 4$ . Second, households with serious health problems,  $h = 5$ , are more likely to experience a very low income shock.

⟨ Insert Figure 1 here ⟩

When constructing a joint Markov process of labor income and health status, our goal is to capture the above pattern. To do this we divide our sample into four subsamples based on the health status in the *second* year of each wave. The first, second, and third subsamples include households whose health status in the second year equals 1, 2, and 3 respectively. The fourth subsample include households whose health status in the second year equals 4 or 5. Then we construct a transition matrix of labor income for each subsample by calculating the fraction of households who move from one quintile to another. The resulting four transition matrixes capture the dynamics of labor income conditional on health shock in the second period, and are denoted as  $Q(z'|z, h' = 1)$ ,  $Q(z'|z, h' = 2)$ ,  $Q(z'|z, h' = 3)$ , and  $Q(z'|z, h' = 4)$ . Due to the small sample size, we cannot get the transition matrix conditional on  $h' = 5$  directly. So we define

$$Q(z'|z, h' = 5) = a \times Q(z'|z, h' = 4) + (1 - a) \times D; \quad 0 \leq a \leq 1,$$

where  $D$  is a  $5 \times 5$  matrix with the first column equal to one and the remaining columns equal to zero. If  $a = 1$ ,  $Q(z'|z, h' = 5) = Q(z'|z, h' = 4)$ . But if  $a = 0$ ,  $Q(z'|z, h' = 5) = D$ , meaning that the income of those households who have serious health problems drops to the level of the lowest income quintile. In our calibration, we choose  $a$  to make the average labor income of those with  $h = 5$  match the data as shown in Panel (a) of Figure 1.

The joint transition matrix of health status and labor income is constructed by combining the transition matrix of health status,  $G^y(h'|h)$ , with the conditional transition matrix of labor income  $Q(z'|z, h')$ . The advantage of this approach is that the conditional expected medical expenses depend only on the current health status. This dramatically simplifies the computation since we can compute the premiums of standard one-period insurance directly from  $G^y(h'|h)$ <sup>21</sup>.

---

<sup>21</sup>If the conditional expected medical expense also depend on the current labor income, say  $E(x'|x, z = 1) \neq E(x'|x, z = 2)$ , and the insurance company does not observe  $z$ , then the premiums of standard one-period contracts will depend on households' insurance decision and the equilibrium distribution of households.

## ESHI Offer status

The dashed line in Panel (d) in Figure 1 shows that there is a strong correlation between the probability to get access to ESHI and labor income. We assume that the probability of getting an ESHI offer is a logistic function:

$$Prob_t = \frac{\exp(u_t)}{1 + \exp(u_t)},$$

where the variable  $u_t$  is an odds ratio that takes the following form:

$$u_t = \eta_0 + \eta_1 D_{gt-1} + \eta_h D_{ht} + \eta_z D_{zt} + \eta_{year} D_{year}, \quad (23)$$

where  $D_{gt-1}$  is a dummy variable for an ESHI offer in period  $t - 1$ ,  $D_{ht}$  and  $D_{zt}$  are the sets of dummy variables for health status and income quintile in period  $t$ , and  $D_{year}$  is a set of dummy variables for each year.

To calibrate the joint distribution  $\{h, z, g\}$  of newborns, we use the empirical joint distribution of households aged 20-35 from the data. This allows us to approximate the life-cycle features absent from our model. In particular, an important characteristic of a life-cycle model is an increasing age profile of medical expenses and labor income. In our calibration newborns are clustered around relatively low medical expenses and relatively low labor productivity. Given the persistence of the processes for medical shocks and productivity, young people in our model are slowly moving up the medical expenses and labor income ladders. The average medical expense as a fraction of average labor income in our model is 7.1% for young people and 20.0% for old people, comparing to 7.3% and 20.4% in MEPS. For newborns, the fraction of average medical expense in average labor income constitutes 4.6% comparing to 5.2% for people aged 20-35 in MEPS.

Figure 1 allows to compare our simulations of  $\{h, z, g\}$  with the data (simulations are plotted with the solid lines). Overall, we are able to match the key features of the data well. However, the simulated offer rate (59.1%) is slightly lower than in the data (64%)<sup>22</sup>.

---

<sup>22</sup>This mismatch mostly arises from the absence of educational heterogeneity in our model. As shown in Pashchenko and Porapakkarm (2011), people with low educational attainment have a significantly

## 5.4 Insurance policies

We use the MEPS to find the fraction of medical costs covered by an average insurance policy. We estimate the following equation

$$InsCov = \beta_0 + \beta_1 x + \beta_2 x^2 + \Theta D_{year}$$

separately for private insurance, Medicaid, and Medicare.  $InsCov$  is medical expenses paid by insurance (variables: TOTPRV, TOTMCD, TOTMCR). We include only people with positive medical expenses when estimating this regression. Then we use our estimates to compute the fraction of medical expenses covered by insurance for each health status and truncate it to be between 0 and 1. Table 3 reports the results for each type of insurance.

⟨ Insert Table 3 here ⟩

## 5.5 Government constraint

In calibrating the tax function  $\mathcal{T}(y)$  we use a nonlinear relationship specified and estimated by Gouveia and Strauss (1994):

$$\mathcal{T}(y) = a_0 [y - (y^{-a_1} + a_2)^{-1/a_1}]$$

Here  $a_0$  controls the marginal tax rate levied on people with the highest income,  $a_1$  determines the progressivity of the tax code, and  $a_2$  is a scaling parameter. We set  $a_0$  and  $a_1$  to the original estimates of Gouveia and Strauss (0.258 and 0.768 correspondingly). The parameter  $a_2$  is used to balance the government budget.

The consumption minimum floor  $\underline{c}$  in the baseline economy was calibrated so that the fraction of households with assets less than \$5,000 in the model is the same as in the data. Based on the 1989-2001 Survey of Consumer Finance (SCF) dataset this fraction is 

---

lower probability to get access to ESHI.



20.0% (Kennickell, 2003). To match this fraction,  $\underline{c}$  is set to 0.92 of the Federal Poverty Line (FPL), or \$8,807.

The Social Security replacement rate is set to 45% of the average labor income. This number is obtained by applying the Social Security benefit formula to the average labor earnings profile.

## 5.6 Medicaid and private insurance

The Medicaid eligibility rules differ from state to state. As of 2009, 14 states had an income eligibility threshold below 50% of FPL, 20 states had it between 50% and 99% of FPL, and 17 states had it higher than 100% of FPL (Kaiser Family Foundation, 2008). We set  $y^{pub}$  to 48.0% of FPL, or \$4,595, to match the fraction of people insured by Medicaid.

In our baseline model, we assume that only standard one-year contracts are offered in the individual market. To match the fraction of those buying individual insurance, we set the administrative load of an individual insurance policy  $\gamma^I$  to 1.208.

The administrative load for the group insurance  $\gamma^G$  is set to 1.11 (Kahn et al, 2005). We set the share of health insurance premium paid by the firm ( $\psi$ ) to 83.0%. This number is consistent with the data in which the premiums of group insurance paid by employers range from 77% to 89% (Sommers, 2002).

## 5.7 Performance of the baseline model

Tables 4 and 5 summarize the parameters used in our baseline model. Table 6 reports the fraction of non-elderly adults with different insurance statuses and the numerical results from the baseline model. The model slightly underestimates the fraction of people with ESHI because our calibrated offer rate is lower than that in the data. As a result the fraction of uninsured is slightly overestimated.

⟨ Insert Table 4, Table 5, and Table 6 here ⟩

To evaluate the performance of our baseline model, we use health insurance statistics not targeted by our calibration. Figures 2 and 3 show the decomposition of health insurance status along the dimension of labor income and health status. Our model is able to replicate the insurance statistics for people in different income and health categories.

⟨ Insert Figure 2 and Figure 3 here ⟩

## 6 Results and discussions

This section discusses how the baseline economy changes once guaranteed renewable contracts are introduced. We provide analysis based on the open economy case, i.e. we fix the interest rate and the wage but allow all insurance prices to adjust in equilibrium<sup>23</sup>.

### 6.1 Effects on premiums

Figure 4 compares the premium for a *newly issued* guaranteed renewable contract with that for a standard one in the new steady state. Guaranteed renewable contracts are more expensive due to the extra payment for the renewability. The difference in prices between the two types of contracts declines as health status deteriorates. For example, for the healthiest group the premium for a guaranteed renewable contract is almost three times higher than that for a standard contract. On the other extreme, for people in the worst health status, the premiums for guaranteed renewable and standard insurance are the same. For this group of people health status cannot deteriorate any further, so the price of a guaranteed renewable contract does not include the extra payment for renewability.

To understand how well guaranteed renewable contracts provide protection against reclassification risk, Figure 5 compares premiums for standard contracts with the average

---

<sup>23</sup>We do this to isolate the pure effect of providing insurance against reclassification risk from the effect of change in aggregate capital. For the closed economy case, the aggregate capital slightly decreases by 0.4%.

premiums for guaranteed renewable contracts including those that are already *in force for at least one period*. An important observation is that on average people who hold guaranteed renewable contracts face insurance premiums that are almost independent of their health status. This happens because most people initiate guaranteed renewable contracts when they are healthy and later they face low premiums even if their health becomes worse. In contrast, people who buy standard contracts face a steep increase in their premiums once their health status deteriorates. This implies that guaranteed renewable contract is a good means to eliminate the risk of premium fluctuations.

⟨ Insert Figure 4 and Figure 5 here ⟩

## 6.2 Effects on health insurance decisions

Table 7 shows how households' insurance purchasing decisions change after guaranteed renewable contracts are introduced. The fraction of uninsured in the new steady state noticeably decreases from 25.4% to 19.4%. The fraction of people with individual insurance increases from 8.2% to 14.2%, and most of this people (9.8%) hold guaranteed renewable contracts.

⟨ Insert Table 7 here ⟩

Table 8 shows how people move between different insurance statuses once guaranteed renewable contracts are available<sup>24</sup>. Around 19% of previously uninsured people start buying insurance once there is the option of guaranteed renewability. This suggests that guaranteed renewability makes the individual insurance market more attractive. Indeed, around 45% of people who were previously buying standard contracts switch to use guaranteed renewable ones.

⟨ Insert Table 8 here ⟩

---

<sup>24</sup>This table is constructed for the first period of transition to the new steady-state once guaranteed renewable contracts are available.

Figures 6 and 7 show the decomposition of health insurance decisions by income quintile and health status. Figure 6 shows that once guaranteed renewable contracts become available, the participation in the individual market increases for people both in good and bad health meaning that the risk-sharing increases. More specifically, the percentage of uninsured among people in the worst health status decreases from 12.7% to 9.3%, while for people in the best health status this number goes down from 24.3% to 22.4%. This can be explained by the fact that individuals buy guaranteed renewable insurance when they are still in good health and therefore are able to renew it at a relatively low premium once their health deteriorates. Table 9 illustrates this point further by showing that people buying guaranteed renewable contracts tend to have higher expected medical expenses than those buying standard contracts.

⟨ Insert Figure 6, Figure 7 and Table 9 here ⟩

Figure 7 shows that guaranteed renewable contracts crowd out standard contracts and reduce the fraction of uninsured individuals for all income quintiles. Interestingly, people in the two lowest income quintiles show the largest participation in the market for guaranteed renewable contracts. Table 9 shows that on average individuals buying guaranteed renewable contracts have lower income than those buying standard contracts. This seems surprising at first given that guaranteed renewable contracts are more expensive than standard ones. To investigate this issue further, Figure 8 plots the fraction of people buying guaranteed renewable contracts in each asset and income quintiles. One can see that the negative correlation between income and demand for guaranteed renewable contracts comes from the top two asset quintiles. In other words, individuals who buy guaranteed renewable contracts have accumulated enough assets to afford this type of contract but their income is low. These individuals are less likely to get access to ESHI, and as will be shown later, this is an important factor determining the demand for guaranteed renewable contracts.

⟨ Insert Figure 8 here ⟩

### 6.3 Welfare analysis

The first row of Table 10 illustrates the welfare gains when moving to an economy where guaranteed renewable contracts are available. Despite the fact that guaranteed renewable contracts provide good protection against reclassification risk, the resulting welfare gains are small. A newborn in the new economy needs a compensation equivalent to 0.0170% of his annual consumption if he is to live in the baseline economy. If we take transition periods into account, the average welfare gains among all young slightly increase to 0.0696%.

⟨ Insert Table 10 here ⟩

Figure 9 shows that the consumption equivalent variation in the first period where guaranteed renewable contracts become available differs substantially by income and asset quintiles. People with low income but high assets are the ones who value guaranteed renewable contracts most. This is the same group that have the highest demand for guaranteed renewable insurance as shown in Figure 8.

⟨ Insert Figure 9 here ⟩

The small welfare gains from having an explicit insurance against reclassification risk imply that the effect of reclassification risk on consumption smoothing is not large. To investigate why this is the case, we consider several factors which may affect how much individuals are concerned about reclassification risk and how much they value the new insurance contracts. In particular, we consider the following five factors: i) implicit insurance against reclassification risk provided by ESHI and Medicaid, ii) the consumption minimum floor, iii) different degree of front-loading, v) labor income risk, and vi) actuarial unfairness of premiums. The first two factors affect how well individuals are protected against reclassification risk in the baseline economy. The last three factors affect individuals' valuation of guaranteed renewable contracts as a means to provide reclassification risk insurance.

In all experiments, when computing welfare gains for all young we control for the distribution of the households. In general, the distribution of households can change significantly from one experiment to the other. To make sure our comparisons are valid, we always compute the average welfare gains for all young using the same distribution. More specifically, in all experiments we use the steady-state distribution of the baseline economy as an initial distribution of the transition period.

### **ESHI and Medicaid**

In the baseline economy there are two institutions that can provide an implicit insurance against reclassification risk. These institutions are Medicaid and employer-based insurance. Both Medicaid and ESHI provide health insurance at a risk-independent rate. Medicaid is free, and premiums for ESHI are community rated, i.e. they are the same for all participants in the employer-based pool. Thus, an agent with a high probability of getting access to these insurance schemes is less concerned about the risk that his premium will increase when his health deteriorates.

To understand how quantitatively important these effects are, we consider several counterfactual experiments. We remove ESHI, Medicaid or both of these programs from the baseline economy, and then reevaluate the welfare gains from introducing guaranteed renewable contracts. The results are presented in the third to fifth rows of Table 10. The corresponding changes in the individuals' insurance decisions are shown in the second and third rows of Table 11.

⟨ Insert Table 11 here ⟩

The welfare effects from introducing guaranteed renewable contracts do not change much once Medicaid is removed: the consumption equivalent variation goes up from 0.0696% to 0.0715%. People who rely on Medicaid are low-income people who cannot afford health insurance on their own. As observed from the second row of Table 11, most of the publicly insured people become uninsured once Medicaid is removed. So they are indifferent between having access to guaranteed renewable contracts or not. This

happens because people are exposed to reclassification risk when they buy insurance contracts that deviate from the first best. For people who never buy private insurance contracts, these deviations from the first best do not matter.

The situation is very different when ESHI is removed. As can be seen in the third row of Table 10, the removal of ESHI increases the consumption equivalent variation almost three times, from 0.0696% to 0.1774%. This implies that without ESHI individuals are more exposed to reclassification risk, thus guaranteed renewable contracts become more valuable.

Figures 10 and 11 illustrate this point further. The elimination of Medicaid has almost no effect on the demand for guaranteed renewable insurance for people in all income and asset quintiles. In contrast, if there is no ESHI, the take-up rates of guaranteed renewable insurance increase dramatically. The most noticeable changes are observed among high-income people in the top two asset quintiles. Previously this group had a very low demand for guaranteed renewable contracts. Once ESHI is removed, the majority of this group start buying new contracts. As a result, the negative relationship between the take-up rates of guaranteed renewable contracts and income observed in Figure 8 disappears.

Figure 11 shows how welfare effects from the new contracts differ by income and asset quintiles in the environment when either Medicaid or ESHI is not available. People who gain the most from having an explicit insurance against reclassification risk in the absence of ESHI are those in the high-income group. In the baseline economy most of these people have access to community rated insurance through their employers. For them ESHI is a good source of reclassification risk insurance. Once this institutional feature is removed, high-income people place much higher value on having access to guaranteed renewable contracts.

⟨ Insert Figure 10 and Figure 11 here ⟩

### **Minimum consumption floor**

A major problem with reclassification risk is that it decreases the insurability of health shocks. If premiums increase following a deterioration of the health status, insurance

may become unaffordable. Thus, people are more concerned about reclassification risk if it is very painful to be uninsured.

The consumption minimum floor provides support for people who depleted all resources. This includes uninsured people with high medical costs. Thus, the consumption floor mitigates the consequences of being uninsured and decreases the concern of lacking protection against reclassification risk.

To understand the quantitative significance of this effect, we reevaluate the welfare gains from guaranteed renewable contracts in an economy with a reduced consumption minimum floor. The seventh to tenth rows of Table 10 show the welfare effects when the consumption minimum floor is equal to 75%, 50%, 25% and 10% of its level in the baseline model. The resulting changes in welfare gains are substantial. When the consumption floor decreases to 10% of the baseline level, the average consumption equivalent variation increases more than 30 times - from 0.0696% to 2.3293%.

To illustrate the role of the minimum consumption floor in more details, Figures 12 and 13 show how the demand for guaranteed renewable contracts and welfare gains change in response to a decline in the consumption floor for people with different income and asset levels. In terms of the demand for new insurance contracts, most noticeable changes are observed among people in the bottom two asset quintiles. When the consumption floor is reduced to 25% of the baseline level, a lot of people in this group start buying guaranteed renewable contracts while previously their participation in this market was almost zero (Figure 8). We do not see a similar response from the high-asset group because these people buy guaranteed renewable contracts even when the consumption floor is high. Those high-asset individuals who do not buy guaranteed renewable contracts are insured by ESHI and a change in the consumption floor does not affect their insurance decisions.

In terms of welfare, the consumption equivalent variation increases substantially for all people except those in the very bottom of both income and asset distribution. The later group has no resources and always qualifies even for the least generous means-tested transfers. It is important to note that even people with high assets value guaranteed renewable contracts substantially more once the consumption minimum floor decreases.



This happens because these people may also face unaffordable health insurance premiums after a sequence of bad health shocks. Since there is less chance they can rely on the consumption floor in this situation, they value an explicit insurance against unaffordability of premiums more<sup>25</sup>.

⟨ Insert Figure 12 and Figure 13 here ⟩

### Different degree of front-loading

The welfare gains from the availability of an explicit insurance against reclassification risk may also be affected by the design of this insurance. Guaranteed renewable contracts are front-loaded and it may be the case that the amount of front-loading is in sharp contrast with what would be optimal from the point of view of intertemporal consumption smoothing. In general, if guaranteed renewable contracts are more front-loaded they provide more reclassification risk insurance because they lock more consumers into the contract, thus having better risk composition as time goes by. This comes at the cost of being more expensive and also being further away from the optimal intertemporal allocation of resources. In other words, there is a tradeoff between better insurance against reclassification risk and better intertemporal allocation. This tradeoff may become worse in the environment with uninsurable labor income risks. In such an environment consumers want to keep a buffer stock of savings against negative labor income shock and thus may be less interested in front-loaded contracts that require prepayments for risks that will be realized far into the future.

To understand whether the tradeoff between optimal consumption smoothing and reclassification risk insurance plays an important role in consumers' valuation of guaranteed renewable contracts, we consider two experiments. We reduce the degree of front-loading by increasing the price that renewable contracts guarantee, first, to 125% and then to 180% of the original price<sup>26</sup>. In other words, if previously an individual is guaranteed to

---

<sup>25</sup>This result is consistent with the findings of De Nardi et al. (2010) who showed that social insurance has a large effect even on people at the top end of income distribution.

<sup>26</sup>Using example from Section 2, this is equivalent to setting  $p_2^{GR} = 1.25 * p_1^{GR}$  and  $p_2^{GR} = 1.8 * p_1^{GR}$ . In all the previous experiments we have  $p_2^{GR} = p_1^{GR}$ .

be able to buy health insurance at the unchanged price, now he is guaranteed the price will not increase more than 25% or 80% of the original price.

Table 10 shows that lowering the degree of front-loading makes welfare gains smaller: the consumption equivalent variation decreases to 0.0645% and 0.0622% for the case of 125% and 180% contracts correspondingly. This suggests that design of guaranteed renewable contracts does not affect our evaluation of welfare costs of reclassification risk.

### Labor income risk

Another factor that can affect how much people value guaranteed renewable contracts is labor income risk. Uninsurable and persistent labor income shocks can affect both people’s attitude towards reclassification risk and their ability to participate in long-term insurance contracts.

Labor income risks can make it harder to participate in long-term insurance contracts. Guaranteed renewable contracts require periodic payments to stay in force. Individuals who experience a bad income shock may find their next payment unaffordable and thus have to terminate the contract. On the other hand, labor income risks make people more concerned about being uninsured because if a medical shock coincides with a negative labor income shock it will make their situation worse.

To understand whether labor income shocks significantly affect people’s valuation of insurance against reclassification risk, we conduct an experiment where we reduce labor income risk. Specifically, in this experiment we change the labor income distribution in such a way that the cross-sectional variance of labor income is equal to 15% of the baseline case <sup>27</sup>. As shown in Table 10, the welfare gain from having a protection against reclassification risks is still small; more specifically the consumption equivalence drops from 0.0696% to 0.0244%. Table 11 shows that less people buy guaranteed renewable contracts when facing lower labor income risks: the fraction of people with new contracts

---

<sup>27</sup>Technically, we keep the joint transition matrix of health, labor income, and ESHI offer the same as in the baseline model but assign a new labor income for each income grid. Denote  $z_j$  and  $\hat{z}_j$  as the original and new value for each income grid  $j$ . We define  $\hat{z}_j = 0.15z_j + 0.75\bar{z}$ , where  $\bar{z}$  is the cross-sectional average labor income in the baseline model. Since the invariant distribution over each income grid is the same, it is easy to show that the cross-sectional average of  $\hat{z}$  is  $\bar{z}$ , while its cross-sectional variance is 15% of that in the baseline case.

goes down from 9.8% to 6.7%. This suggests that labor income risk does not prevent people from buying guaranteed renewable contracts, on the contrary it makes the additional insurance more valuable.<sup>28</sup>

### **Actuarial unfairness of premiums**

Finally, we consider whether actuarial unfairness plays an important role in the valuation of guaranteed renewable contracts. Even if reclassification risk is costly in terms of welfare, people may not value insurance against this risk if it is actuarially unfair. We consider the case when administrative loads are entirely eliminated from both standard and guaranteed-renewable contracts<sup>29</sup>. The results of this experiment are presented in the last column of Table 10. The welfare gains change very little, going up from 0.0696% to 0.0905%, suggesting that actuarial unfairness does not significantly affect people's valuation of guaranteed renewable contracts.

## **7 Implication for the health insurance reform**

In March of 2010 President Obama signed the Patient Protection and Affordable Care Act that is going to introduce significant changes in the U.S. health insurance system. This reform has two key components. First, it introduces a wide range of income-based transfers, i.e. subsidies and expansion of public coverage. Second, it changes the rules under which the individual insurance market operates. In particular, the new law does not allow insurance companies to differentiate premiums by individual's health status. In other words, it introduces community rating in the individual insurance market. To prevent cream-skimming behavior of insurers, the reform also prohibits insurance companies to deny coverage to anyone. Finally, the new law mandates individuals to buy health insurance unless their income is very low.

Pashchenko and Porapakarm (2013) evaluate how different components of the reform

---

<sup>28</sup>As discussed in De Santis (2007), the welfare function is convex in the overall consumption risk. Labor income shocks augment overall risk; thus removing the labor income risk makes the welfare cost of any additional uncertainty smaller.

<sup>29</sup>In other words, we set  $\gamma^I = \gamma^G = 0$ .

contribute to its welfare outcome and find that the contribution of community rating is very small. This paper confirms their results and provides an insight for why this is the case. Community rating is a regulatory approach to eliminate reclassification risk<sup>30</sup>. In this paper we show that this risk is already to a large extent insured by such institutions as employer-based insurance and means-tested transfers. Thus, the introduction of community rating have a small effect on welfare.

Another implication of our findings is that even if reclassification risk is important for welfare, good protection against it can be obtained through private markets. Community rating accompanied by individual mandates is a large scale intervention in the insurance market. As such it has non-trivial distorting effects on both households' and insurance firms' decisions. In this light a private market approach to solving the problem of reclassification risk may be an alternative worth considering.

## 8 Conclusion

This paper studies how important reclassification risk is for the welfare of consumers. Reclassification risk is believed to be an important problem in the individual health insurance market. Premiums in this market are risk-rated while a typical contract lasts for only one year. Individuals whose health status deteriorates can see a drastic increase in their health insurance premiums, and this reduces their ability to obtain health insurance.

We constructed a general equilibrium model and calibrated it using the MEPS dataset to replicate the key features of the U.S. economy. To evaluate welfare costs of reclassification risk, we consider the effect of introducing into this economy guaranteed renewable health insurance contracts. Guaranteed renewable contracts are private insurance contracts that provide protection against reclassification risk without requiring consumer's commitment or income based transfers.

We find that the welfare gains from having access to the explicit insurance against reclassification risk through guaranteed renewable contracts are small. This is because two

---

<sup>30</sup>Kifman (2002) provides a detailed comparison between guaranteed renewable contracts and community rating as a means to insure reclassification risk.

institutional features of the current system - employer-based insurance and consumption minimum floor - provide a good implicit protection against reclassification risk. While employer-sponsored insurance mostly provides insurance to high-income people, low-income people are protected by the consumption minimum floor. If these two institutions are removed, welfare gains from having access to guaranteed renewable contracts are large and can exceed 2% of the annual consumption. Our results are robust to the alternative design of guaranteed renewable contracts and the degree of actuarial unfairness in the health insurance market.

## References

- [1] Brook, R., Ware, J., Rogers, W., Keeler, E., Davies, A., Donald, C., Goldberg, G., Lohr, K., Masthay, P., Newhouse, J., 1983. Does free care improve adults health? Results from a randomized trial. *New England Journal of Medicine*, 309(23), 1426-1434
- [2] Cochrane, J., 1995. Time-consistent Health Insurance. *The Journal of Political Economy*, 103(3), 445-473
- [3] Daily, G., Hendel, I., Lizzeri, A., 2008. Does the Secondary Life Insurance Market Threaten Dynamic Insurance? *American Economic Review*, 98(2), 151-156
- [4] De Nardi, M., French, E., Jones, J., 2010. Why Do the Elderly Save? *Journal of Political Economy*, 118(1), 39-75
- [5] De Santis, M., 2007. Individual Consumption Risk and the Welfare Cost of Business Cycles. *American Economic Review*, 97(4), 1488-1506
- [6] Diamond, P., 1992. Organizing the Health Insurance Market. *Econometrica*, 60(6), 1233-1244
- [7] Fang, H., Kung, E., 2010. How Does Life Settlement Market Affects the Primary Life Insurance Market?. Manuscript, University of Pennsylvania
- [8] Finkelstein, A., McGarry, K., Sufi, A., 2005. Dynamic Inefficiencies in Insurance Markets: Evidence from Long-term Care Insurance. *American Economic Review Papers and Proceedings*, 95(2), 224-228
- [9] Fonseca, R., Michaud, P., Galama, T., Kapteyn, A., 2009. On the Rise of Health Spending and Longevity. Working Papers 722, RAND Corporation
- [10] French, E., Jones, J., 2004. On the distribution and dynamics of health care costs. *Journal of Applied Econometrics* 19, 705-721.
- [11] French, E., Jones, J., 2011. The Effects of Health Insurance and Self-Insurance on Retirement Behavior. *Econometrica*, 79(3), 693-732
- [12] Gouveia, M., Strauss, R., 1994. Effective Federal Individual Income Tax Functions: An Exploratory Empirical Analysis. *National Tax Journal*, 47, 317-339
- [13] Hansen, G., Hsu, M., Lee, J., 2012. Health Insurance Reform: The impact of a Medicare Buy-In. Mimeo, GRIPS.
- [14] Hendel, I., Lizzeri, A., 2003. The Role of Commitment in Dynamic Contracts: Evidence from Life Insurance. *The Quarterly Journal of Economics*, 118(1), 299-327
- [15] Herring, B., Pauly, M., 2006. Incentive-Compatible Guaranteed Renewable Health Insurance Premiums. *Journal of Health Economics*, 25, 395-417
- [16] Hsu, Minchung, 2010. Health Insurance, the Social Welfare System and Household Saving. Manuscript, National Graduate Institute for Policy Studies (GRIPS)

- [17] Jeske, K., Kitao, S., 2009. U.S. Tax Policy and Health Insurance Demand: Can a Regressive Policy Improve Welfare? *Journal of Monetary Economics*, 56(2), 210-221
- [18] Kaiser Family Foundation, 2008. Tax Subsidies for Health Insurance: An Issue Brief, available at <http://www.kff.org/insurance/7779.cfm>
- [19] Kennickell, A., 2003. A Rolling Tide: Changes In the Distribution of Wealth in the U.S., 1989-2001. Working paper, Federal Reserve Board
- [20] Kahn, J., Kronick, R., Kreger, M., Gans, D. 2005. The Cost of Health Insurance Administration in California: Estimates for Insurers, Physicians, and Hospitals, *Health Affairs* 24
- [21] Kiffman, M., 2002. Insuring Premium Risk in Competitive Health Insurance Market. Ed. Mohr Siebek
- [22] Kopecky, K., Koreshkova, T., 2011. The Impact of Medical and Nursing Home Expenses and Social Insurance Policies on Savings and Inequality. Mimeo, Federal Reserve Bank of Atlanta
- [23] Ozkan, S., 2011. Income Differences and Health Care Expenditures over the Life Cycle. Mimeo, Federal Reserve Board
- [24] Pashchenko, S., Porapakkarm, P., 2013. Quantitative Analysis of Health Insurance Reform: Separating Regulation from Redistribution. *Review of Economic Dynamics*, 16(3), 383-404
- [25] Pauly, M., Kunreuther, H., Hirth, R., 1995. Guaranteed Renewability in Insurance. *Journal of Risk and Uncertainty*, 10, 143-156
- [26] Racine, A., Kaestner, T., Joyce, T., Colman, G., 2001. Differential Impact of Recent Medicaid Expansions by Race and Ethnicity. *Pediatrics* 108 (5), 1135-1142
- [27] Ross, C., Mirowsky, J., 2000. Does Medical Insurance Contribute to Socioeconomic Differentials in Health? *Milbank Quarterly*, 78(2), 291-321
- [28] Rust, J., Phelan, C., 1997. How Social Security and Medicare Affect Retirement Behavior in a World of Incomplete Markets. *Econometrica*, 65, 781-831
- [29] Scholz, J., Seshadri, A., 2010. Health and Wealth in a Lifecycle Model. Mimeo, University of Wisconsin Madison
- [30] Sommers, J., 2002. Estimation of Expenditures and Enrollments for Employer-Sponsored Health Insurance. Agency for Healthcare Research and Quality, MEPS Methodology Report 14
- [31] Young, E., 2010. Solving the Incomplete Markets Model with Aggregate Uncertainty Using the Krusell-Smith Algorithm and Non-Stochastic Simulations. *Journal of Economic Dynamics and Control*, 34(1), 36-41.

Panel	99/00	00/01	01/02	02/03	03/04	04/05	05/06	06/07	Total
Obs.	4,954	4,017	8,248	6,244	6,464	6,417	6,200	6,656	49,200

Table 1: Number of observations in eight waves of MEPS (1999-2007)

	$h = 1$	$h = 2$	$h = 3$	$h = 4$	$h = 5$
medical expenses (percentile)	$< 30^{th}$	$30^{th} - 60^{th}$	$60^{th} - 90^{th}$	$90^{th} - 99^{th}$	$> 99^{th}$

Table 2: Health status and medical expenses

	$h = 1$	$h = 2$	$h = 3$	$h = 4$	$h = 5$
Medicaid: $q(-1, x)$	1.00	1.00	0.70	0.52	0.50
Private insurance: $q(i, x)$ for $i = \{0, 1, \dots, 5\}$	0.00	0.40	0.71	0.78	0.81
Medicare: $q^{med}(x)$	0.00	0.35	0.56	0.64	0.65

Table 3: Fraction of medical expenses covered by insurance

Parameter name	Notation	Value	Source
Risk aversion	$\sigma$	3	-
Cobb-Douglas parameter	$\alpha$	0.33	Capital share in output
Tax function parameters	$a_0$	0.258	Gouveia and Strauss (1994)
	$a_1$	0.768	Gouveia and Strauss (1994)
Social Security replacement rates	$ss$	45%	-
Group insurance loads	$\gamma^G$	1.11	Kahn et all (2005)
Employer's contribution	$\psi$	0.83	Sommers (2002)
Medicare premium	$p^{med}$	\$1,071	Total premiums =2.11% of $Y$

Table 4: Parameters set outside the model

Parameter name	Notation	Value	Target
Discount factor	$\beta$	0.908	$\frac{K}{Y} = 3$
Depreciation rate	$\delta$	0.07	$r = 0.04$
Individual Insurance loads	$\gamma^I$	1.21	% of individual insurance=8.2%
Medicaid's income eligibility	$y^{pub}$	\$4,595	% of public insurance=9.1%
Consumption floor	$\underline{c}$	\$8,807	% with assets < \$5,000=20%

Table 5: Parameters used to match some targets



	uninsured	public ins	individual ins	ESHI
data	21.45%	9.10%	8.20%	61.30%
model	25.4%	9.10%	8.20%	57.30%

Table 6: Percentage of non-elderly adults with different insurance status (2003/2004)

	Baseline	+GR contracts
Uninsured (%)	25.4	19.4
Individually insured (%)	8.2	14.2
- by standard contracts	8.2	4.4
- by GR contracts	—	9.8
Publicly insured (%)	9.1	9.1
Insured by ESHI (%)	57.3	57.3

Table 7: Insurance statistics before and after introduction of GR contracts (steady-state)

		Insurance decisions if GR insurance is available				
		Uninsured	Medicaid	ESHI	Std ins	GR ins
Original decisions	Uninsured	80.7%	0.00%	0.00%	0.02%	19.26%
	Medicaid	0.00%	100.0%	0.00%	0.00%	0.00%
	ESHI	0.00%	0.00%	100.0%	0.00%	0.00%
	Std ins	0.60%	0.00%	0.00%	54.63%	44.76%

Table 8: Changes in insurance decisions if GR contracts are available

	Insurance	Average $E(x)$	Average labor inc	Average total inc
Baseline	Std ins	0.057	1.107	1.246
New steady-state with GR ins	Std ins	0.038	1.326	1.433
	GR ins	0.084	0.628	0.828

Table 9: Average income and medical expenses for people choosing different types of contracts

Experiments	average CEV	
	newborn	all young
Benchmark	0.0170%	0.0696%
<i>Effect of Medicaid and ESHI</i>		
- No Medicaid program	0.0171%	0.0715%
- No ESHI program	0.0537%	0.1774%
- No Medicaid and ESHI program	0.0542%	0.1812%
<i>Effect of consumption floor</i>		
- 0.75¢ (\$6,605)	0.0269%	0.1862%
- 0.50¢ (\$4,403)	0.0571%	0.4134%
- 0.25¢ (\$2,201)	0.2136%	1.0319%
- 0.10¢ (\$880)	0.8575%	2.3293%
<i>Effect of front-loading</i>		
- 125% of $p^{GR}$	0.0151%	0.0645%
- 180% of $p^{GR}$	0.0149%	0.0622%
<i>Effect of labor income risk</i>		
- reduced labor income risk	0.0303%	0.0244%
<i>Effect of actuarial unfairness</i>		
- No administrative load ( $\gamma^{GR} = \gamma^I = 0$ )	0.0150%	0.0905%

Table 10: Consumption equivalent variation after introducing GR contracts<sup>a</sup>

<sup>a</sup>The above welfare changes are computed by comparing two economies: an economy with a setup corresponding to each experiment and an economy with the same setup except having guaranteed renewable contracts. The CEV of newborns corresponds to the comparative statics between the two steady-states, while the CEV of all young takes into account the steady-state distribution in the baseline model and the transition periods.

	uninsured	Std ins	GR ins	Pub ins	ESHI
Benchmark	19.4	4.4	9.8	9.1	57.3
No Medicaid	28.0	4.6	10.7	—	57.3
No ESHI	33.1	7.5	50.1	9.3	—
0.75¢ (\$6,605)	12.0	6.2	17.4	6.6	57.8
0.50¢ (\$4,403)	7.3	5.9	24.2	4.2	58.3
0.25¢ (\$2,201)	3.3	4.7	30.9	2.9	58.2
0.10¢ (\$880)	2.9	3.7	33.7	2.4	57.4
125% of $p^{GR}$	18.9	4.6	10.1	9.1	57.3
180% of $p^{GR}$	18.3	1.5	13.8	9.1	57.3
Reduced labor income risk	21.2	12.9	6.7	0.2	59.1
$\gamma^{GR} = \gamma^I = 0$	7.8	12.4	13.5	9.0	57.4

Table 11: Insurance statistics for model with GR contracts for different experiments (steady-state)

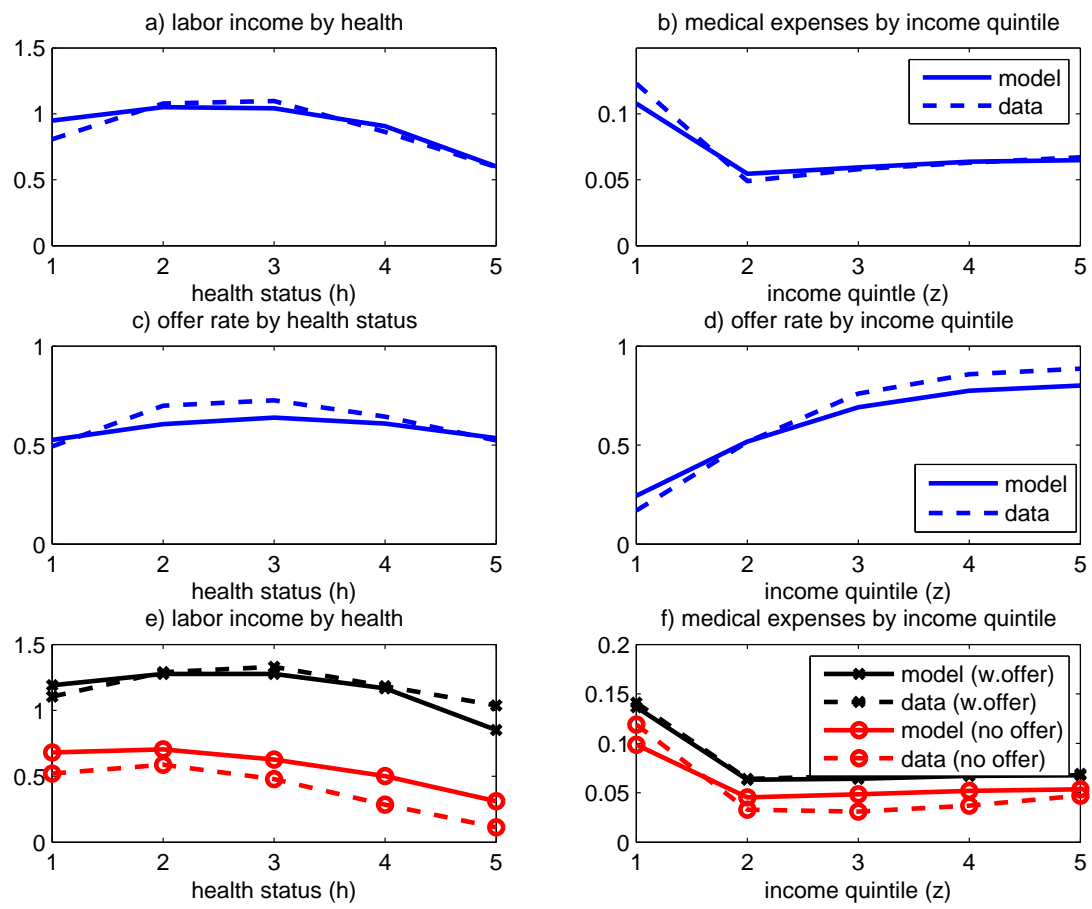


Figure 1: Relationship between ESHI offer, labor income, and medical expenses

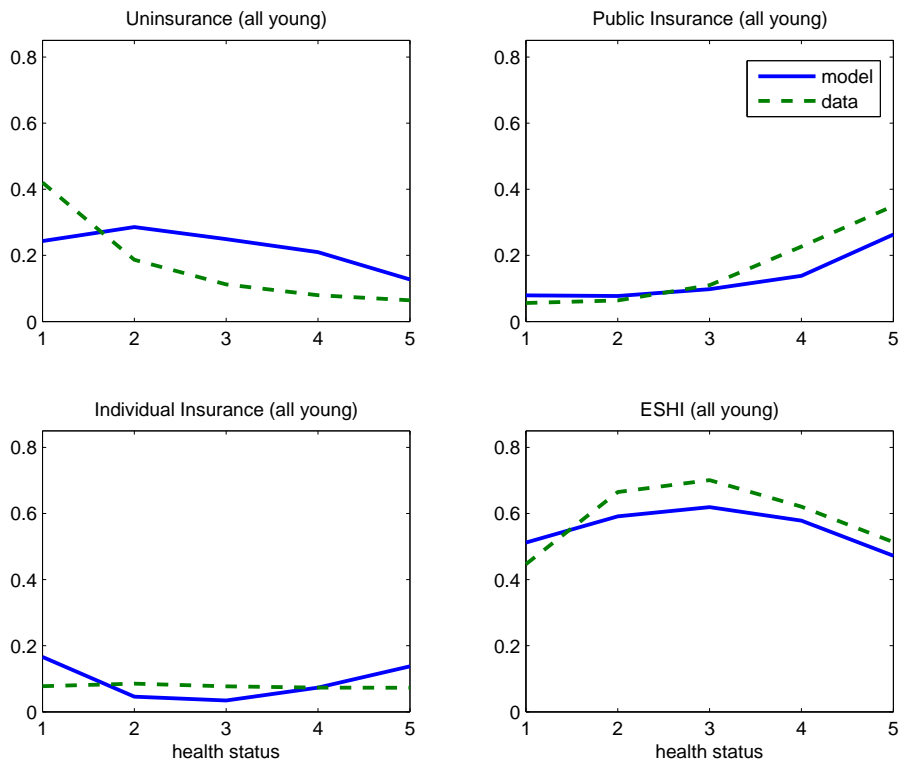


Figure 2: Insurance decision by health status (baseline model)

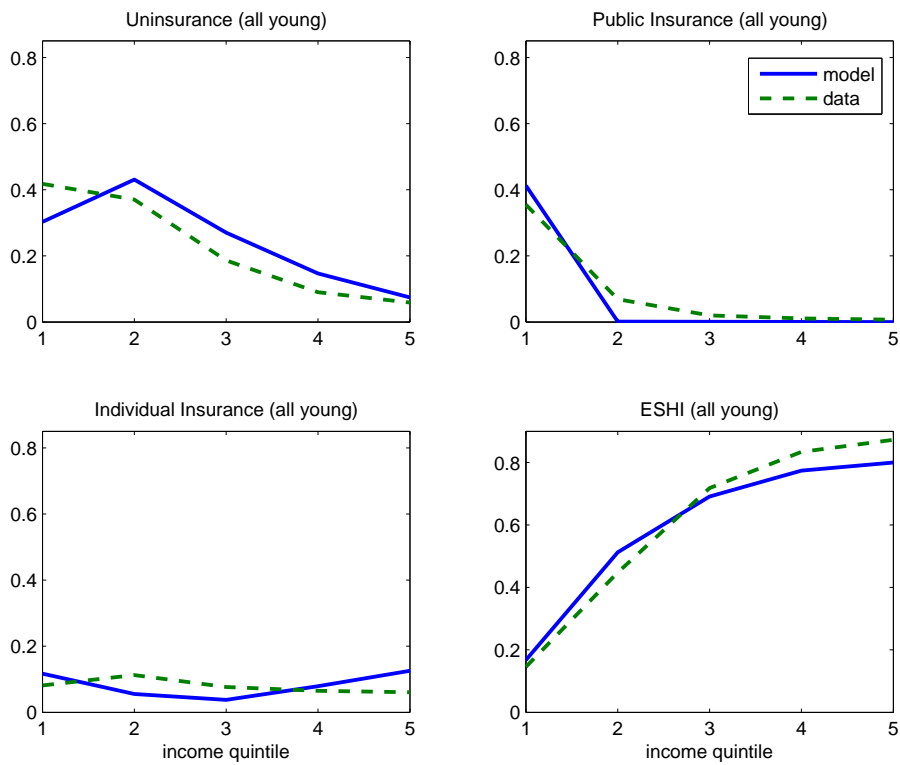


Figure 3: Insurance decision by labor income (baseline model)

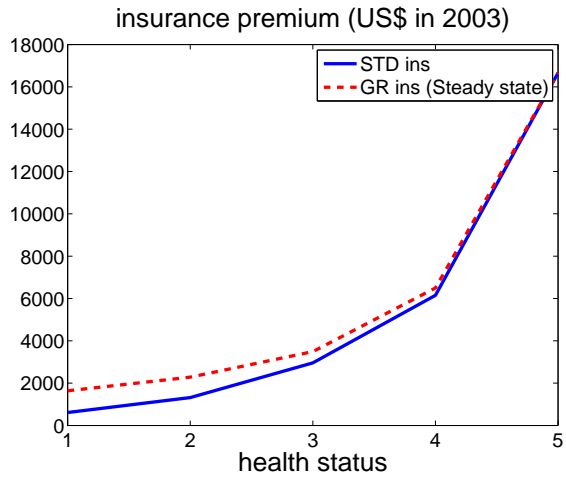


Figure 4: Premiums for new contracts

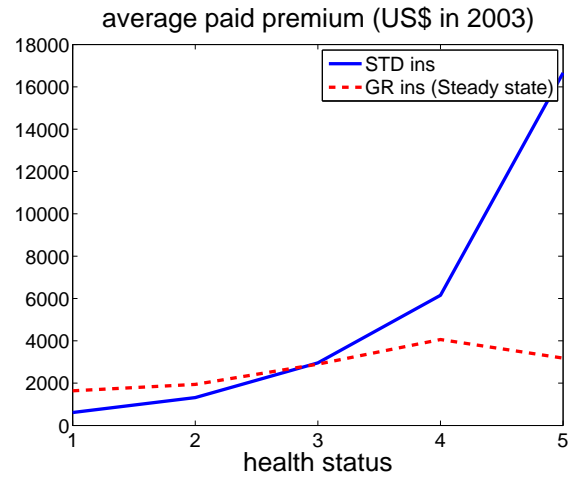


Figure 5: Average premiums for existing contracts

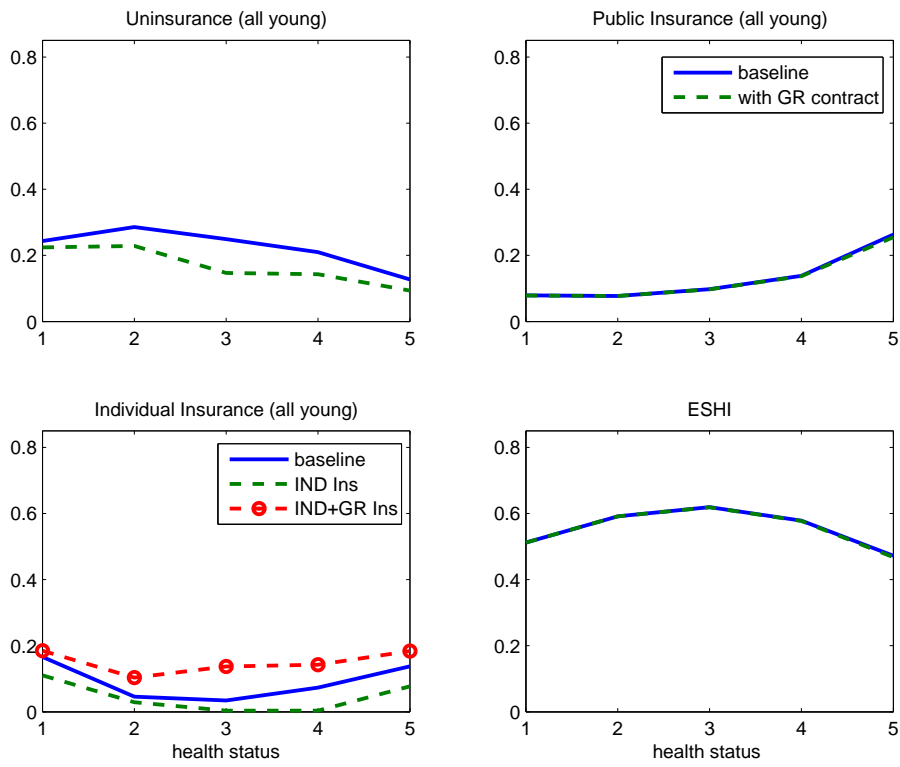


Figure 6: Insurance decisions by health status in the steady-state (+GR contract)

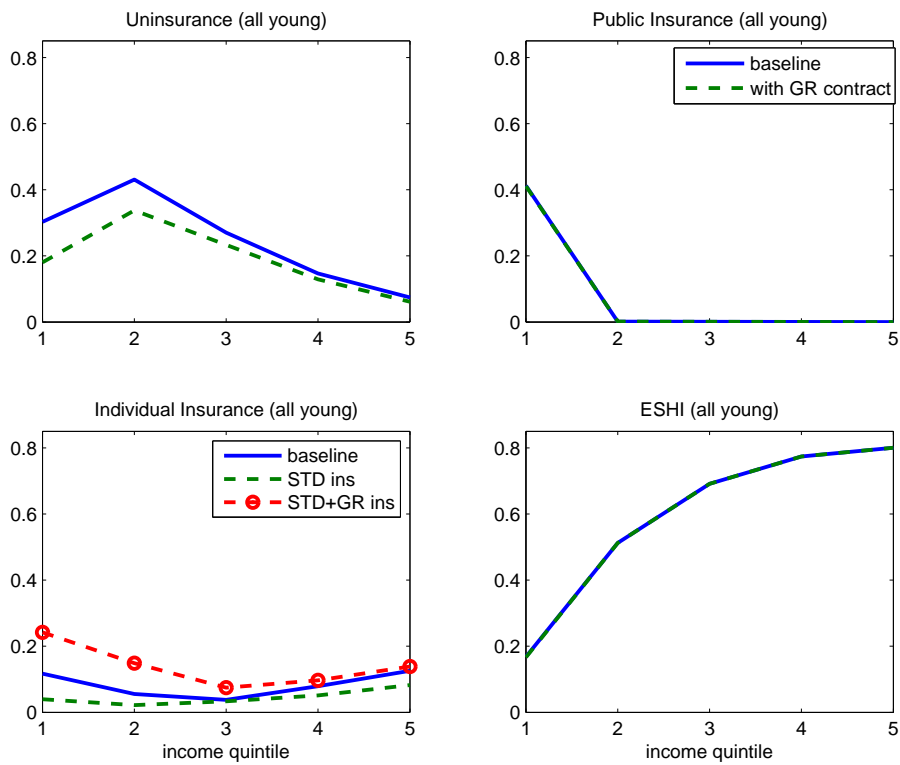


Figure 7: Insurance decisions by labor income in the steady-state (+GR contract)

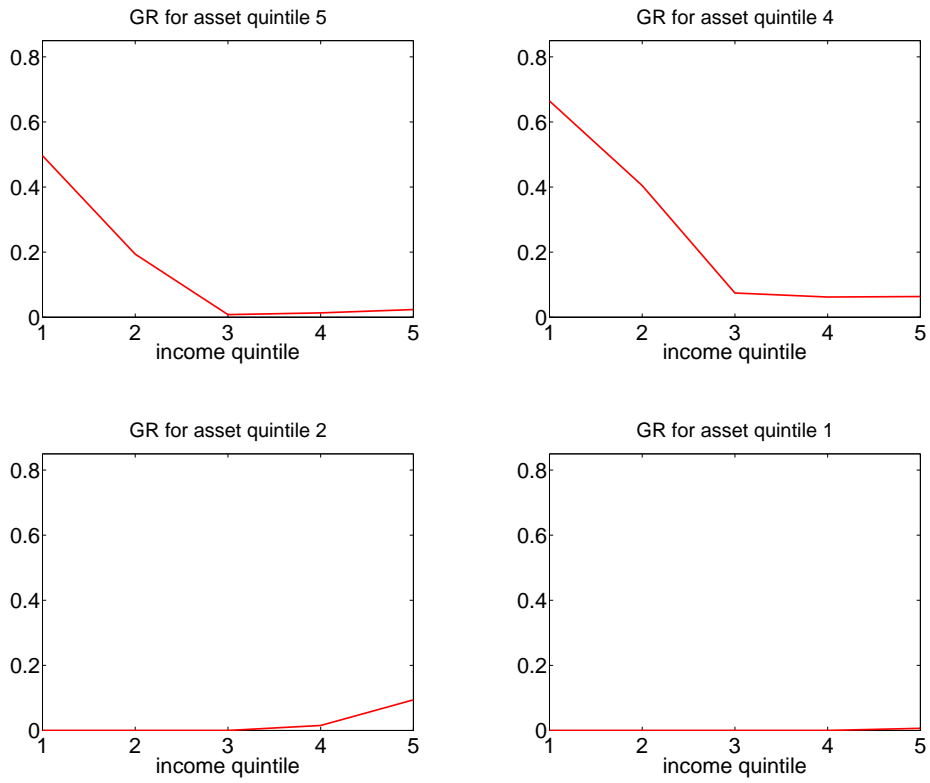


Figure 8: Fraction of people buying GR contracts by income and asset quintile

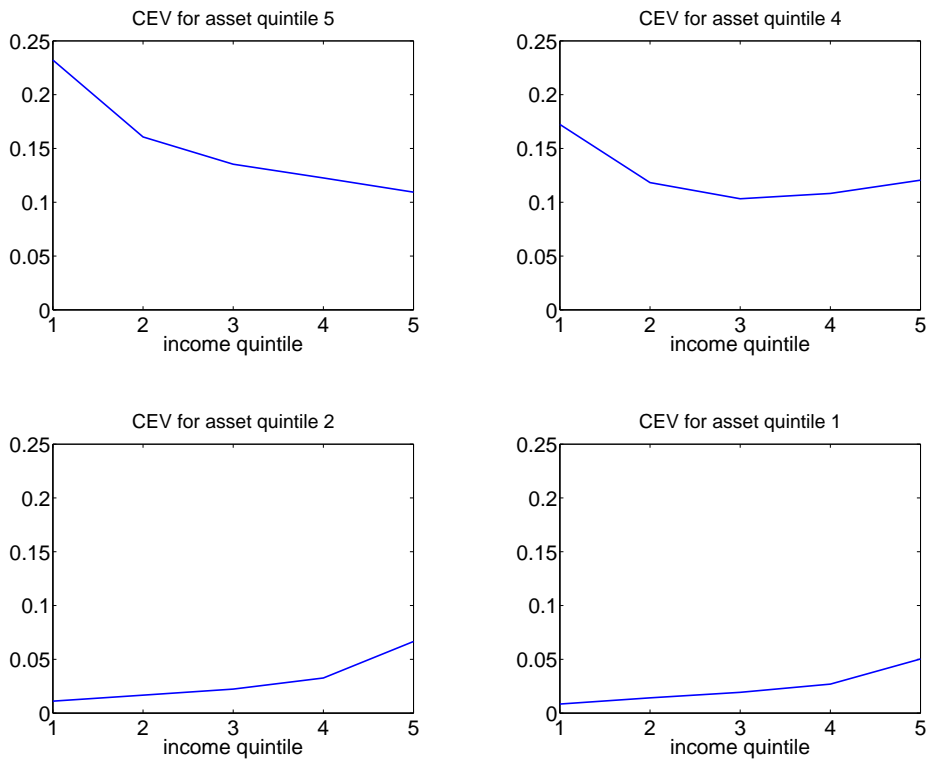


Figure 9: Consumption Equivalence by income and asset quintile (benchmark)



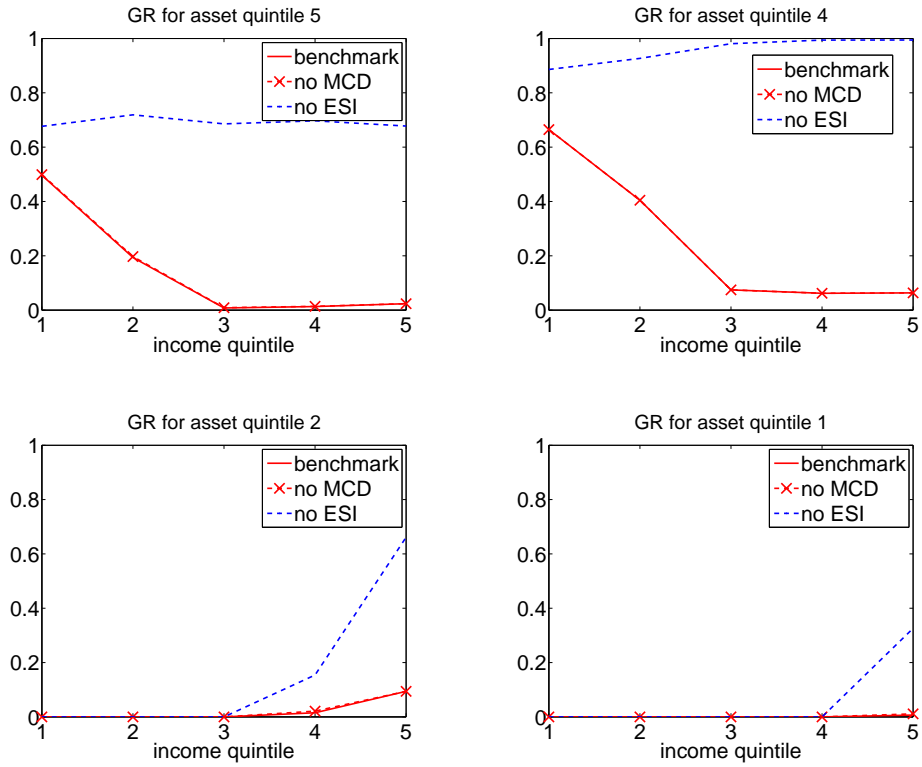


Figure 10: Fraction of people buying GR contracts by income and asset quintile (effect of ESHI/MCD)

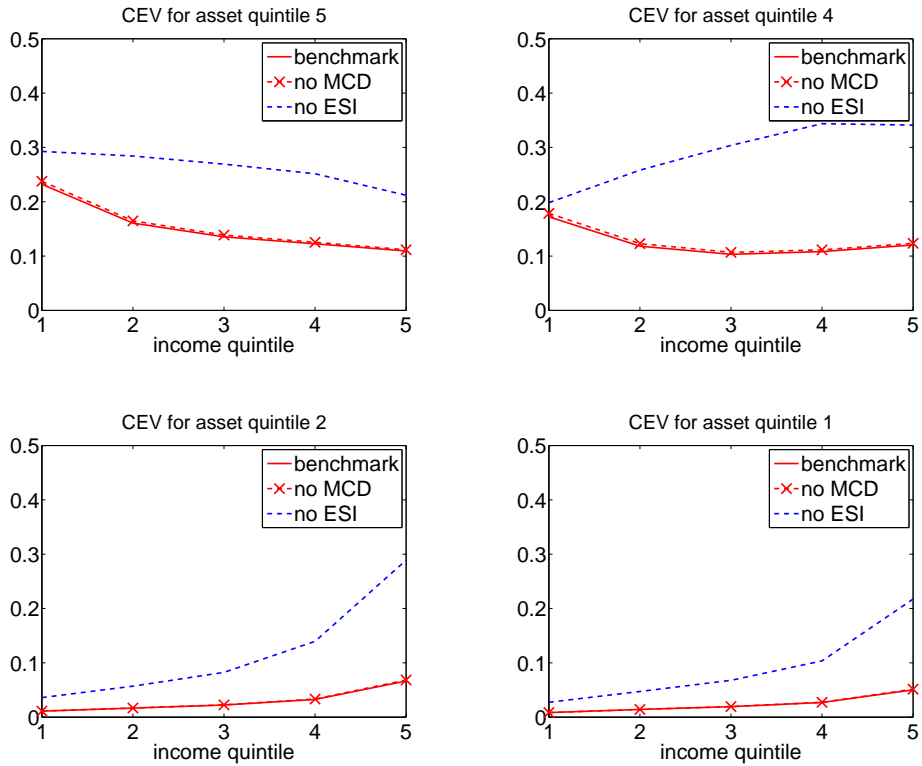


Figure 11: Consumption Equivalence by income and asset quintile (effect of ESHI/MCD)

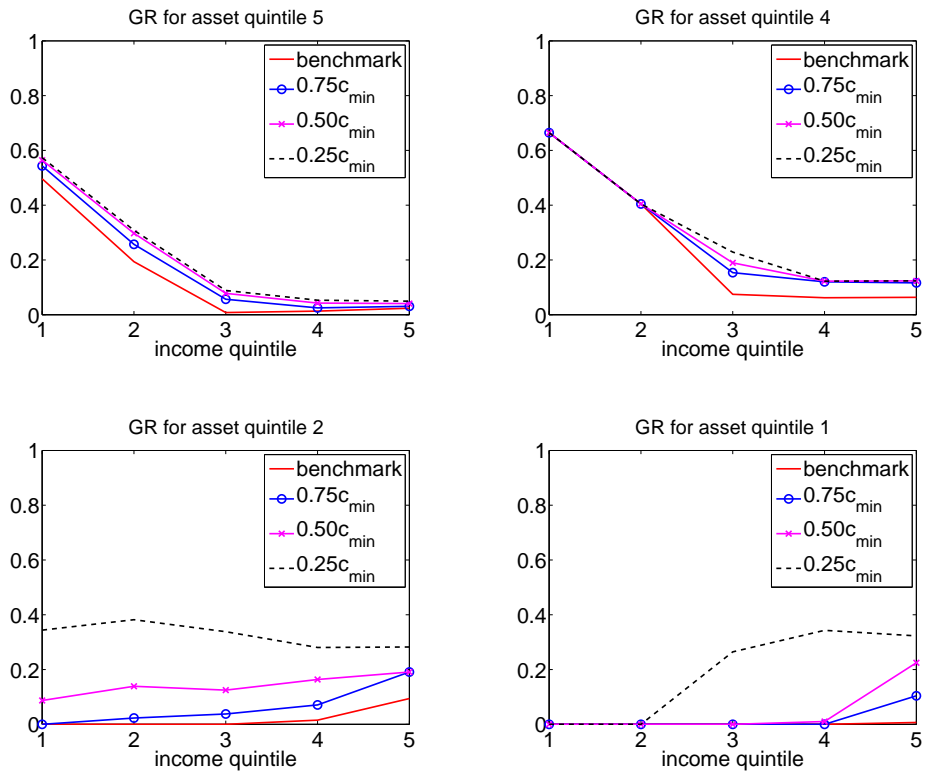


Figure 12: Fraction of people buying GR contracts by income and asset quintile (effect of  $c$ )

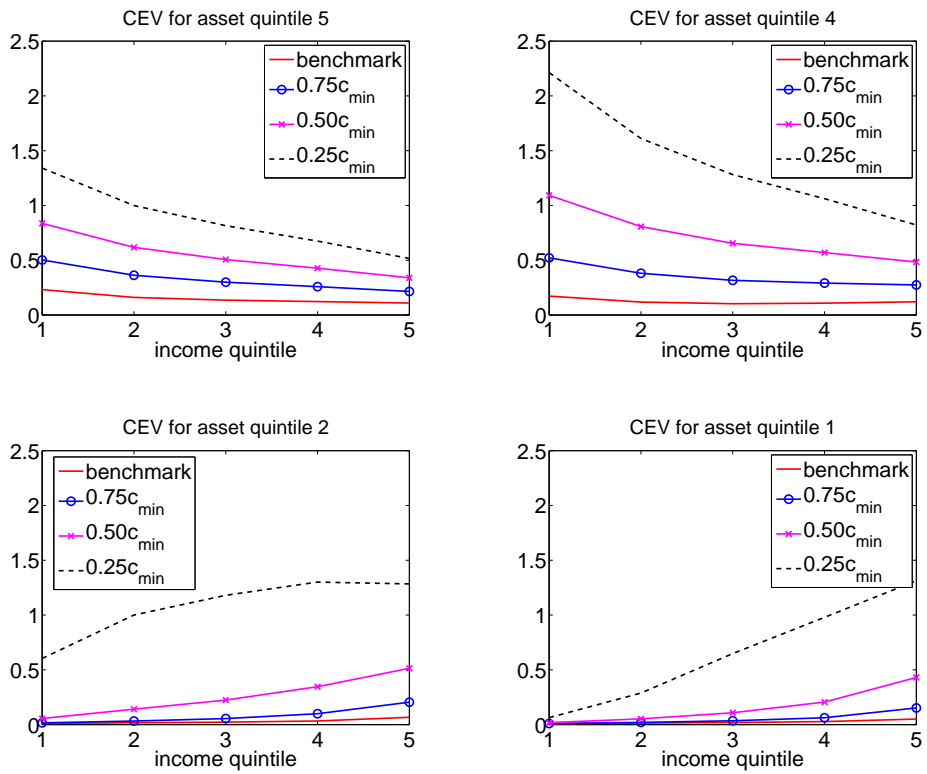


Figure 13: Consumption Equivalence by income and asset quintile (effect of  $c$ )

## Appendix

### A Competitive equilibrium with asymmetric information <sup>31</sup>

Given the government programs  $\{\underline{c}, ss, q^{med}(x), p^{med}\}$ , the insurance coverage  $\{q(i, x)\}$ , and the fraction of the group premium contributed by the employer ( $\psi$ ), the competitive equilibrium with asymmetric information consists of the set of equilibrium prices  $\{w, r, p, p^I(h), p^{GR}(i)\}$ , wage reduction  $\{c_E\}$ , households' value functions  $\{\mathbf{V}^y(\mathbf{s}), \mathbf{V}^o(\mathbf{s})\}$ , decision rules for the young  $\{c(\mathbf{s}), k'(\mathbf{s}), i^H(\mathbf{s})\}$  and for the old  $\{c(\mathbf{s}), k'(\mathbf{s})\}$ , the tax function  $\{\mathcal{T}(y)\}$ , time-invariant distributions  $\{\Gamma^y(\mathbf{s}), \Gamma^o(\mathbf{s})\}$ , and the set of insurers' beliefs  $\{\pi_{t+j}(h_{t+j}|h_t); j > 0, \forall t\}$  such that the following conditions are satisfied:

1. Given the set of prices and the tax function, decision rules and value functions solve the individuals' optimization problems (2) and (7).
2. Wage ( $w$ ) and rent ( $r$ ) satisfy Equations (15) and (16).
3. Labor market clears:  $L = \int z\Gamma^y(\mathbf{s})$
4. Capital market clears. Since guaranteed renewable contracts are front-loaded, there will be a balance carrying over time for each contract. We need to take this balance into account when computing the aggregate capital. Denote by  $\theta_{t+j}^t(h_t)$  an ex-post balance at time  $t+j$  of a unit of contract sold at time  $t$  to an individual with health status  $h_t$ . One period after a contract is originated this balance takes the following form:

$$\begin{aligned} \theta_{t+1}^t(h_t) &= p^{GR}(h_t)(1+r) - \gamma^I EM(h_t) + \\ & p^{GR}(h_t) \int_{h_{t+1}} \frac{\mathcal{F}(h_{t+1}, i_{t+1}^H = RGR || h_t, i_t^H = BGR)}{\Gamma^y(h_t, i_t^H = BGR)}. \end{aligned}$$

---

<sup>31</sup>We refer to this equilibrium as asymmetric information equilibrium because insurance companies observe only one state variable - health status. For guaranteed renewable contracts health is not the only variable relevant for pricing which creates an asymmetric information environment.

The first term on the right-hand side is the premium collected at the initiation of the contract and carried on to the next period. The second term is the cost of medical claims in period  $t + 1$ . The last term is the revenue from the contract renewal. We can define recursively the ex-post balance  $j$  periods after the contract is originated as follows<sup>32</sup>:

$$\begin{aligned} \theta_{t+j}^t(h_t) &= \theta_{t+j-1}^t(h_t)(1+r) - \\ &\gamma^I \int_{h_{t+j-1}} EM(h_{t+j-1}) \frac{\mathcal{F}(h_{t+j-1}, i_{t+j-1}^H = RGR || h_t, i_t^H = BGR)}{\Gamma^y(h_t, i_t^H = BGR)} + \\ &p^{GR}(h_t) \int_{h_{t+j}} \frac{F(h_{t+j}, i_{t+j}^H = RGR || h_t, i_t^H = BGR)}{\Gamma^y(h_t, i_t^H = BGR)}. \end{aligned}$$

Thus the capital market clearing condition in period  $t$  can be written as

$$\begin{aligned} K &= \int k'(\mathbf{s}) \Gamma^y(\mathbf{s}) + \int k'(\mathbf{s}) \Gamma^o(\mathbf{s}) + \\ &p \int \mathbf{1}_{\{i^H(\mathbf{s})=BG\}} \Gamma^y(\mathbf{s}) + \int \mathbf{1}_{\{i^H(\mathbf{s})=BI\}} p^I(h) \Gamma^y(\mathbf{s}) + \\ &\int \mathbf{1}_{\{i^H(\mathbf{s})=BGR\}} p^{GR}(h) \Gamma^y(\mathbf{s}) + \sum_{j=1}^{\infty} \int \theta_{t-j}^{t-j}(h_{t-j}) \Gamma^y(h_{t-j}, i_t^H(\mathbf{s}) = BGR) \end{aligned}$$

5.  $c_E$  satisfies Equation (17); thus the firm offering ESHI earns zero profit.
6. The tax function  $\{\mathcal{T}(y)\}$  satisfies the government budget balance in Equation (22).
7. Standard one-period insurance premiums,  $p^I(h)$ , satisfy Equation (18), guaranteed renewable premiums  $p^{GR}(i)$ ,  $i = 1, \dots, H$ , satisfy Equation (19), and the group insurance premium ( $p$ ) satisfies Equation (21). Thus health insurance companies earn zero expected profit on each contract.
8. Insurance companies' beliefs  $\{\pi_{t+j}(h_{t+j,t}|h_t); j > 0, \forall t\}$  satisfy Equation (20) if

---

<sup>32</sup>By recursively substituting  $\theta_{t+j-1}^t$ , this equation is equivalent to Equation (19).

$\Gamma^y(h_t, i_t^H = BGR) \neq 0$ . Otherwise,

$$\pi_{t+j}(h_{t+j}|h_t) = 0 \quad ; \quad j > 0, \forall t. \quad (24)$$

The last equation is the off-equilibrium belief of insurers. When no one with health status  $h_t$  buys a guaranteed renewable contract, insurers believe that if one with health  $h_t$  buys a guaranteed renewable contract, he will not renew the contract in the next period<sup>33</sup>.

## B Discussion of the assumption of exogenous medical expenses

Currently, two approaches exist for the modeling of medical expenses in macroeconomic and structural studies. The first approach takes a stand that medical expenses are exogenous shocks that result in monetary losses (see for example, Jeske and Kitao (2009), Hansen et al (2012), French and Jones (2010), Kopecky and Koreshkova (2011)). The second approach assumes that people can choose the amount of their medical spending (Fonseca et al (2010), Ozkan (2011), Scholz and Seshadri (2010)). It is well known that in reality medical spending has both discretionary and non-discretionary part. However, to the best of our knowledge the literature lacks a model that can unite the two approaches described above and reproduce the empirical patterns of discretionary vs. non-discretionary spending.

Our choice of the model of exogenous medical spending is determined by the focus of our study. We evaluate the value of different risk-sharing mechanisms available through the current health insurance system in the U.S. As Rust and Phelan (1997) emphasize, the value of health insurance to a large extent depends on the variance and skewness of medical expenses. In addition, the value of health insurance depends on the correlation of medical and labor income shocks. To adequately measure welfare effects of risk-sharing in health insurance markets, we need to carefully represent the joint distribution of medical

---

<sup>33</sup>Our results are robust to an alternative specification of the off-equilibrium beliefs.

expenses and labor income. To the best of our knowledge, none of the existing models of endogenous medical expenses can simultaneously reproduce the empirical variance and skewness of medical spending and its correlation with labor income.

If we were able to incorporate adjustments in medical spending resulting from people's optimal decision-making in our model, we can expect that better insurance options provided by guaranteed renewable contracts can improve health outcomes and consequently increase our estimated welfare effects. However, we expect these effects to be small based on the numerous empirical studies that suggest health insurance has little effect on subsequent health outcomes. A seminal contribution in this area is RAND Health Insurance Experiment, where a random set of individuals were given co-payment-free health insurance over 3- to 5-year period, while a control group faced standard co-payments. Brook et al (1983) found that the group with free health insurance had very similar health outcomes than the control group, even though the former group had substantially higher utilization rate of medical services. Several other studies find similar results using different methods and data. Racine et al (2001) find that expanded Medicaid eligibility for children between 1989 and 1995 reduced uninsurance but had very limited effect on health. Ross and Mirowsky (2000) using longitudinal data find no difference in health status between privately insured and uninsured.

## C Computational algorithm

We solved for the steady state equilibrium of the baseline model as follows.

1. Guess an initial interest rate  $r$ , price in the group insurance market  $p$ , the amount the firm offering ESHI subtracts from the wage of their workers  $c_E$ , prices of guaranteed renewable contracts  $p^{GR}(h), h = 1..H$ , and the tax parameter  $a_2$ <sup>34</sup>.
2. Guess value functions for young and old. Solve the problems for young and old. We optimize with respect to savings and insurance decisions and evaluate the value function for points outside the state space grid using a Piecewise Cubic Hermite Interpolating

---

<sup>34</sup>We cannot prove the uniqueness of the equilibrium in the health insurance market, however our results are robust to alternative initial guesses of insurance prices ( $p$  and  $p^{GR}(h), h = 1..H$ ).

Polynomial (PCHIP). Update the value functions and continue iterating until both value functions converge. Use convergent value functions to find policy functions.

3. Given the policy functions, simulate the households distribution using a non-stochastic method as in Young (2010).

4. Use the distribution of households and policy functions to compute government budget deficit/surplus. Gradually update the tax function parameter  $a_2$ , the interest rate  $r$ , insurance prices  $p^{GR}(h)$ ,  $h = 1..H$ ,  $p$ , and the subtraction from wage  $c_E$ . Repeat steps 2-3 until all these variables converge.

## D Guaranteed renewable insurance market with frictions

Our results show that if individuals gain access to a frictionless market offering guaranteed renewable contracts this results in only small welfare gains.<sup>35</sup> An alternative approach to understand the value of guaranteed renewable insurance is to introduce these contracts in the baseline economy and quantify the size of frictions needed to deter people from buying them (the outcome observed in reality). In this section we take this second approach and introduce frictions in the market for guaranteed renewable contracts as an extra cost embedded in the premium. One can think of the extra cost arising from two sources. First, to offer guaranteed renewable contracts, insurance companies need to commit to the long-term contract and thus require a compensation for future uncertainties, which are potentially large. The growth rate of aggregate medical expenses is difficult to predict, particularly in the far future.<sup>36</sup> In addition, the invention of new treatments, and changes in clinical guidelines or standards can significantly alter the dynamics of reclassification

---

<sup>35</sup>The premium for guaranteed renewable contract consists of two parts. First is the premium for the regular one-period health insurance. We assume this premium is the same as in the market offering standard one period contracts meaning it includes an administrative load. Second is the premium for the right to renew the contract at the prespecified price. This part of the premium has no administrative costs or any other loads. We refer to the guaranteed renewable market introduced in the benchmark experiments of our paper as frictionless to stress the fact that the option to renew the contract is priced actuarially fair.

<sup>36</sup>Cutler(1996) argues that the lack of long-term health insurance can be explained by the uncertainty in aggregate medical costs.

risks, specifically the transition probability of medical expenses in our model, which is the basis for pricing guarantee renewable contracts. Second, guaranteed renewable contracts are more complicated than regular insurance contracts, and consumers may need to put more efforts to understand them.

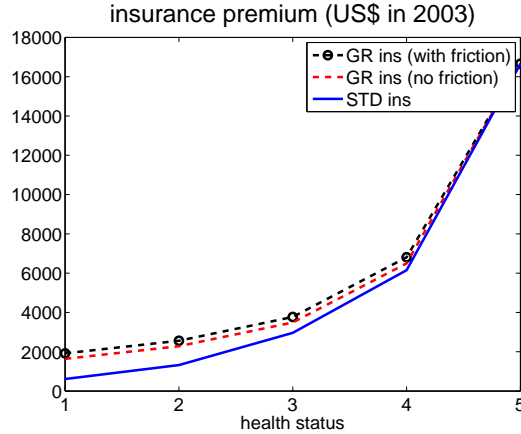


Figure 14: The difference in premiums between frictionless guaranteed renewable contracts and guaranteed renewable contracts with an extra cost

To illustrate how large the extra cost that eliminates the demand for guaranteed renewable contracts should be, we use the following approach. We start with the benchmark economy with frictionless guaranteed renewable contracts (as described in Section 6) and gradually increase its premiums,  $p^{GR}(h)$ , till the take-up rate of each GR contract is less than 0.25%.<sup>37</sup> Figure (14) compares the premiums of guaranteed renewable contracts in this experiment with the steady-state premiums reported in Section 6. The differences between the two premiums are \$283, \$276, \$280, and \$310 for the contract initiated by a person in medical grid 1, 2, 3, and 4 respectively. In terms of the percentage difference, the premiums in the market with frictions are higher than frictionless premiums by 17%, 12%, 8%, and 5% for people in medical grid 1, 2, 3, and 4 respectively. Thus, relatively small fixed costs can eliminate demand for guaranteed renewable contracts, which is consistent with our result that these contracts bring small welfare gains when employers' sponsored health insurance and the minimum consumption guarantee coexist.

<sup>37</sup>Instead of completely eliminating the demand for guaranteed renewable contracts, we change the premiums until the size of the market becomes very small. We do this to avoid the result when our assumption about off-equilibrium beliefs influences the equilibrium. One can argue that firms would stop offering insurance contracts if there is too few participants in the market.



It is worth noting that the additional costs needed to wipe out the demand for guarantee renewable contracts need not to be large partly because of the adverse selection effect. A small increase in premiums for guaranteed renewable contracts will repel people with low probability to renew the contract in later periods. This increases the costs for insurers to provide the option to renew the contract. Consequently, the premium increases which decreases the demand of guaranteed renewable contracts even further.