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Exchange Arrangements and Currency Crises: What's the matter with the exchange rate classification?

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Abstract

The purpose of this paper is to empirically investigate whether certain exchange rate arrangements are more prone to currency crises using a probit model. We define a currency crisis as a period characterised by the presence of intense foreign exchange market pressure. The definition is based on a foreign exchange market pressure index (MPI). If the value of the MPI is above a certain threshold, we define that period as a crisis state; otherwise the period is defined as a tranquil state. The definition of currency crises used in this paper focuses on discrete events.

Keywords: Exchange rate regimes, currency crises, speculative attacks.

JEL classification: F31, F33.

1 Introduction

Since the financial crises of the 1990s in emerging markets, the issue of the suitability of exchange rates regimes has returned to the international finance research agenda. More precisely, the debate over fixed and floating exchange arrangements has once again taken centre stage in academic circles. Some economists maintain that the first round of this debate

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was won by those advocating floating arrangements because all crisis episodes took place in countries which had adopted a variety of mechanisms for pegging their exchange regimes. Conversely, the advocates of fixed exchange regimes suggest that there are bad fixes and good fixes (like official dollarization) and good or truly fixed arrangements that allow countries to achieve credibility.

An important recent development in the debate over optimal exchange rate regimes is the recognition that the choice of an exchange rate arrangement is different between particular groups of countries. The choice of an exchange rate regime for developed countries is different from the one of developing countries or emerging economy countries. Developing countries are often beset by a lack of credibility and limited access to international capital markets. Hence, fixed exchange rate regimes play a useful role by providing policymakers with a nominal anchor for monetary policy and by helping to establish a degree of policy credibility. In contrast, emerging market economies are more integrated with global financial markets but they have encountered more currency crises under pegged exchange rate arrangements (Husain et al., 2005). Developed countries have obtained more benefits from flexible exchange regimes because they are more developed economically and institutionally, and more integrated in global financial markets (Rogoff et al., 2003).

Contrary to a large number of theoretical studies in the literature, relatively few studies attempt to empirically investigate the impact of an exchange rate regime on currency crises in developed, emerging and developing countries, separately. This is perhaps, because such an empirical investigation is fraught with difficulties, including the problem concerning the classification of exchange arrangement. This article addresses the issue of measurement errors in the classification of exchange rate regimes by using four different classification schemes. Three *de facto* and one *de jure* classifications are used. Consequently, the sensitivity of these results to alternative exchange rate classifications is also tested. The principal conclusion emerging from this study is the following: emerging and developing countries adopting fixed exchange rate arrangements have a lower probability of currency crises.

The remainder of this article is organised in the following way: Section 2 shows a brief review on exchange arrangement classifications. Section 3 presents a brief literature review focusing on the link between exchange rate regimes and currency crises. Section 4 discusses the issues of exchange market pressure indicators and currency crises. Section 5 describes the empirical framework. A preliminary analysis of the data is presented in Section 6. Section 7 reports empirical findings. Section 8 concludes the findings of this article.

2 Regime Classification

A common problem in the empirical analysis of exchange rate systems is regime classification. The literature identifies two approaches to this problem: the *de jure* classification and the *de facto* classification. The former classifies countries by what they say they do (*de jure*). However, countries often act differently to what they declare they do. In particular, a self-declared independent floating regime, in reality, often operates a managed peg regime. This phenomenon of operating a disguised peg is referred to as "fear of floating" (Calvo and Reinhart, 2002). Classifying countries by what they actually do is a *de facto* classification. Some authors develop *de facto* classifications using various methods (Ghosh et al., 1997; Bailliu et al., 2001; Poirson, 2002; Bubula and Otker-Rober, 2002; Reinhart and Rogoff, 2004; Shambaugh, 2004; Dubas et al., 2005; Levy-Yeyati and Sturzenegger, 2005; Bérnassy-Quééré et al., 2006; Frankel and Wei, 2008; Ilzetski et al., 2010), but these are fundamentally based

on data that presents the behaviour of nominal exchange rates, international reserves and interest rates².

Some empirical studies simply employ the *de facto* classification because the *de jure* classification may reach incorrect results³, particularly about floating regimes. On the other hand, some research employs the *de jure* classification arguing that it suffers from less drawbacks than the *de facto* classification⁴.

In this article we employ a combination of three *de facto* and one *de jure* classifications. Firstly, we use the *de facto* classification developed by Levy-Yeyati and Sturzenegger (2005), henceforth known as the "LYS classification". These authors apply a cluster analysis to a data set with three variables: changes in the nominal exchange rate, the volatility of these changes, and the volatility of international reserves from all IMF reporting countries in the period 1974-2000. Secondly, the "natural classification" developed by Reinhart and Rogoff (2004) is employed. Reinhart and Rogoff (2004) reclassified exchange rate regimes based on market determined dual and parallel exchange rates, and use official rates only if the exchange rates are unified⁵. These authors examine the chronologies of the exchange rate history for 153 countries in the period 1946-2001. They are able to distinguish among floating by high inflation countries (freely falling) from floating by others. They define the category of "freely falling" rates when the 12-month rate of inflation exceeds 40% and when, during these periods of high inflation there is no official announcement of the regime by the authorities⁶. In addition, they define hyperfloats as those episodes of macroeconomic instability that are characterised by hyperinflation where the monthly inflation rate is 50% or more. Thirdly, an alternative classification scheme developed by Bailliu et al. (2001) is used. These authors develop a Hybrid Mechanical Rule (HMR) classification. This system classifies exchange rate regimes in terms of their observed flexibility and takes into account external shocks and revaluations. Their analysis is based on a sample of 60 countries for the period 1973-1998. Finally, the *de jure* classification from the IMF is used⁷.

In our analysis all the different classifications are grouped into three broader regimes: fixed, intermediate and floating exchange rate regimes (see Table 1). Managed floating is classified under the floating category, with respect to the term managed, in the context of the Reinhart-Rogoff classification, does not necessarily imply active or frequent foreign exchange market intervention.

² To a literature reviews on why many countries follow *de facto* regimes different from their *de jure* regimes see Cruz-Rodríguez (2013).

³ This could be the results of measurement error in the classification of exchange rate arrangements.

⁴ The *de facto* classification has the advantage of being based on observable behaviour, but it does not capture the distinction between stable nominal exchange rates resulting from the absence of shocks, and stability that stems from policy actions offsetting shocks. More importantly, it fails to reflect the commitment of the central bank to intervene in the foreign exchange market. Although the *de jure* classification captures this formal commitment, it falls short of capturing policies inconsistent with the commitment, which lead to a collapse or frequent adjustments of the parity.

⁵ In case where there are no dual or multiples rates or parallel markets are not active.

⁶ In situations where the currency crisis marks a sudden transition from a fixed or quasi-fixed regime to a managed or independently floating regime, they label an exchange rate as freely falling during the six months immediately following a currency crisis.

⁷ The data on the *de jure* classification of exchange rate regimes is taken from Ghosh et al. (2002) and from the IMF's Annual Report on Exchange Arrangements and Exchange Restrictions.

Table 1: Classification of Exchange Rate Regime

| Fixed | Intermediate | Floating |
|---|--|--|
| <i>De facto</i> Classification by Levy-Yeyati and Sturzenegger | | |
| (1) Fixed | (2) Crawling peg (3) Dirty floats | (4) Float |
| <i>De facto</i> Classification by Reinhart and Rogoff | | |
| (1) No separate legal tender (2) Pre-announced peg or currency board arrangement (3) Pre-announced horizontal band that is narrower than or equal to $\pm 2\%$ (4) <i>De facto</i> peg | (5) Pre-announced crawling peg (6) Pre-announced crawling band that is narrower than or equal to $\pm 2\%$ (7) <i>De facto</i> crawling peg (8) <i>De facto</i> crawling band that is narrower than or equal to $\pm 2\%$ (9) Pre-announced crawling band that is wide than or equal $\pm 2\%$ (10) <i>De facto</i> crawling band that is narrower than or equal to $\pm 5\%$ (11) Moving band that is narrower than or equal to $\pm 2\%$ | (12) Managed floating (13) Freely floating (14) Freely falling (15) Hyperfloating |
| <i>De facto</i> Classification by Bailliu, Lafrance and Perrault | | |
| (1) Currency boards (2) Single currency peg (3) Basket pegs (4) Crawling pegs with narrow bands | (5) Flexibility index ≤ 1 | (6) Flexibility index ≥ 1 |
| <i>De jure</i> Classification by Ghosh, Gulde and Wolf | | |
| (1) Pegged regimes | (2) Intermediate regimes | (4) Floating regimes |

Note: Inconclusive classifications from Levy-Yeyati and Sturzenegger are not considered in our analysis.

Sources: Bailliu et al. (2001); Bailliu et al. (2003); Ghosh et al. (2002); Reinhart and Rogoff (2004); and Levy-Yeyati and Sturzenegger (2005).

3 Exchange Rate Regimes and Currency Crises: A Survey of the Literature

Earlier contributions to the theoretical literature on currency crises pointed almost exclusively to deteriorating economic fundamentals as the trigger for currency crises. However, few studies have made an attempt to investigate empirically whether a particular exchange rate regime is more prone to a currency crisis. Some empirical research suggests that currency crises are more likely to occur under fixed or intermediate exchange regimes. However, a study developed by the IMF (1997), based on the IMF's *de jure* classifications, finds that close to half of the currency crashes (sharp changes in the exchange rate) occur under floating regimes, implying that crises can arise under both pegged and floating regimes⁸. Similarly, Ghosh et al. (2002) find that *de jure* pegged regimes have the lowest probability of a

⁸ An important observation is that many exchange rate regimes are improperly classified as flexible when they are in fact, pegged regimes.

currency crisis⁹. Likewise, Falcetti and Tudela (2006) show that currency crises in developing and emerging markets are less frequent under *de jure* fixed exchange rates than under *de jure* flexible regimes in the period 1970-1997. On the other hand, Rogoff et al. (2003) find that currency crises tend to occur more frequently in *de facto* intermediate regimes especially in emerging markets. Similar conclusions are drawn by Peltonen (2006) who finds, using the *de facto* classification from Reinhart and Rogoff (2004), that emerging markets with more rigid exchange rate regimes were less prone to currency crises during the last two decades. Empirical case studies conducted by Jakubiak (2001) demonstrate that a floating exchange rate regime does not guarantee an emerging country avoiding a currency crisis. Haile and Pozo (2006), using the IMF's *de jure* and the LYS *de facto* classifications analyse the incidence of currency crises in emerging markets according to the exchange regime in place between 1974 and 1998. Their results suggest that the *de facto* exchange regime plays no role in determining currency crisis period. As a consequence, fixed exchange regimes that are not truly fixed appear to invite speculation against the currency, increasing the likelihood of currency crisis.

In the same way, Bubula and Otker-Rober (2003), using their own *de facto* classification¹⁰, find that pegged regimes, as a whole, are more prone to currency crises compared with floating regimes, particularly for developed and emerging market economies that are integrated with international capital markets, in the period 1990-2001¹¹. On the contrary, Coulibaly (2009), using panel data of 192 countries from 1970 through 1999, and 195 currency crisis episodes, examines the effect of membership in a currency union on the probability of experiencing a currency crisis. Both parametric and non-parametric estimates suggest that membership in a currency union reduces the likelihood of a currency crash. Angkinand et al. (2009), using a logit model and a panel of 90 countries observed annually from 1990 to 2001, show that results from using Reinhart and Rogoff (2004) regime are that middle regimes such as adjustable parities, crawls, and moving bands are relatively prone to crises, while managed floats have the lowest probability of crises among intermediate regimes. However, when authors turn to LYS classification, they do not find any significant result in explaining the correlation between exchange rate regimes and currency crises.

Esaka (2010a) examines the link between *de facto* exchange rate regimes and the incidence of currency crises in 84 countries from 1980 to 2001 using probit models. The author employs the *de facto* classification of Reinhart and Rogoff (2004) and finds no evidence that intermediate regimes have a significantly higher probability of currency crises than both hard pegs and free floats. Similarly, Esaka (2010b) examines whether *de facto* exchange rate regimes affect the occurrence of currency crises in 84 countries over the 1980–2001 period by using the probit model and the *de facto* classification of Reinhart and Rogoff (2004). His results show that pegged regimes significantly decrease the likelihood of currency crises compared with floating regimes. On the other hand, Asici (2011) applied a multinomial logit framework to 163 developed and developing countries over the period from 1990 to 2007. His regression results suggest that countries experiencing currency crisis are those that have chosen regimes inconsistent with their individual features.

Karimi and Voia (2014) analyze the effect of exchange rate regimes and capital account liberalization policies on the occurrence of currency crises for 21 countries over the period of 1970-1998. The authors examine changes of the likelihood of currency crises under *de jure* IMF classification and two *de facto* exchange rate regimes (Reinhart and Rogoff and LYS). Their results show that the likelihood of currency crises changes significantly under *de*

⁹ However, the impact of a currency crisis is more severe under pegged and intermediate regimes than under floating regimes.

¹⁰ For details on this classification, see Bubula and Otker-Rober (2002).

¹¹ They define currency crises as episodes of severe market pressures, reflected by sharp movements in both exchange and interest rates.

facto regimes. While Reinhart and Rogoff based models show that fixed exchange rate arrangements are least susceptible to speculative attacks, LYS based models point to the intermediate exchange rate regimes as the least crisis prone. However, Esaka (2014), using data on currency crises and exchange rate regimes from 84 countries for the period of 1980–1998 and the *de jure* IMF classification to identify official announced exchange rate regimes and the *de facto* Reinhart and Rogoff (2004) classification, evaluates the treatment effect of consistent pegs on the occurrence of currency crises to examine whether consistent pegs are indeed more prone to currency crises than other regimes. Using matching estimators as a control for the self-selection problem of regime adoption, the author finds that countries with consistent pegs have a significantly lower probability of currency crises than countries with other exchange rate policies. On the other hand, Ghosh et al. (2015) using the IMF *de facto* classification¹² and a sample of 50 emerging economies over the 1980-2011 period, show that macroeconomic and financial vulnerabilities are significantly greater under less flexible intermediate regimes, including hard pegs, as compared to floats. Conversely, Combes et al. (2016) revisit the link between crises and exchange rate regimes. Using a panel of 90 developed and developing countries over the period 1980-2009, and two *de facto* classifications (the IMF *de facto* classification and the Ilzetski et al., 2010, classification). Their results reject that intermediate regimes are more vulnerable to crises compared to the hard peg and the fully floating regimes.

4 The Exchange Market Pressure Indicator and Currency Crisis Periods

In any empirical analysis of currency crises, the first issue is to define the nature of a crisis. A currency crisis can be understood as a sudden decline in the confidence to an individual currency usually leading to a speculative attack against it. Since, in a currency crisis situation, a speculative attack may lead to sharp currency depreciation, an increase of interest rates and/or a substantial reserve loss, the most straightforward approach is to employ an index of speculative pressure¹³. This technique is common in the empirical literature on currency crises. The exchange market pressure indicator was originally developed by Girton and Roper (1977) to describe the composite behaviour of nominal exchange rates and international reserves, and later modified by Eichengreen et al. (1996). In the interest of measuring currency crises Eichengreen et al. (1996) add a third term: changes in the nominal interest rate. The idea behind this is that an excess demand for foreign exchange can be met through several channels. Depreciation or devaluation occurs if the speculative attack is successful, but monetary authorities may instead accommodate the pressure by running down their international reserves or deter the attack by raising interest rates. This methodology, which identifies currency crises using an exchange market pressure indicator, has been followed, in principle, by Sachs et al. (1996); Kaminsky et al. (1998); Tudela (2004); Peltonen (2006); Haile and Pozo (2006); Falcetti and Tudela (2006), among others.

¹² Critics constantly moved away from the official International Monetary Fund (IMF) classification to construct a *de facto* classification system in 1999. The new IMF classification combines the available information on exchange rates and monetary policy frameworks, and the formal or informal policy intentions of authorities, with data on actual exchange rates and reserve movements to reach an assessment of the actual exchange rate regime (Habermeier et al., 2009, provide information on revisions to this classification system in early 2009). However, it can be argued that the new IMF classification system is still one of the *de jure* regimes, since it still relies heavily on official information and looks mainly at the behaviour of official exchange rates (Reinhart and Rogoff, 2004).

¹³ In theoretical literature, a currency crisis is mostly defined only in the case of fixed exchange rate regimes, usually as the official devaluation or abandonment of the fixed exchange rate regime. However, this definition is not flexible enough to serve a use in empirical research, since many currencies are not formally pegged to a specific currency and many countries use various forms of floating exchange rate regimes.

In this article, the exchange Market Pressure Indicator (MPI) is calculated as the weighted average of percentage changes in the exchange rate (e), percentage changes in the interest rate (i), and percentage changes in international reserves (r)¹⁴, using the United States as the country of reference¹⁵. The exchange market pressure index is defined as follows:

$$MPI = w_1 \Delta e + w_2 \Delta i - w_3 \Delta r \quad (1)$$

where e represents the price of US\$1 in domestic currency, i the interest rate, and r international reserves. Since the volatilities of foreign reserves, exchange rates and interest rates are very different, the weights w_1 , w_2 and w_3 , attached to each component are used to equalise the volatilities of each of the three MPI components, thereby preventing any one of them from dominating the index, and are defined as the inverse of the standard deviation of each of the individual series. Formally:

$$w_j = \frac{\frac{1}{StDev_j}}{\frac{1}{StDev_e} + \frac{1}{StDev_i} + \frac{1}{StDev_r}} \quad (2)$$

where j stands for any of the three variables and $StDev$ stands for the standard deviation. According to equation (1), if a country has a fixed exchange rate regime, a speculative attack may lead to sharp currency devaluation, an increase of interest rates and/or a substantial foreign reserve loss. On the other hand, if a country has a flexible exchange rate regime, a speculative attack may lead to sharp currency depreciation, and then to an increase of interest rates and/or a substantial international reserve loss, but only if monetary authorities want to deter the attack.

A crisis period is defined to occur when the value of the MPI exceeds an arbitrary threshold¹⁶. Following Eichengreen et al. (1996) we define crisis periods as MPI values that are greater than 1.5 standard deviations over the mean of the series. Formally:

$$Crisis = \begin{cases} 1 & \text{if } MPI > \mu_{MPI} + 1.5\sigma_{MPI} \\ 0 & \text{otherwise} \end{cases} \quad (3)$$

where μ_{MPI} and σ_{MPI} denote the mean and the standard deviation of the sample of the MPI. Hence, a crisis takes place for an individual country when its MPI variable takes an "extreme"

¹⁴ A decrease rather than increase in international reserves is used, since an increase in speculative pressure tends to increase the exchange rate and the interest rate, but tends to reduce foreign reserves.

¹⁵ Variables in logarithms.

¹⁶ Unsuccessful speculative attacks are also included in our definition of a currency crisis since they point to the vulnerability of the system reflected or that can be seen in a fall in international reserves and a rise in interest rates.

positive value¹⁷. The total number of crises identified is 227 (43 in advanced, 52 in emerging and 132 in developing countries) and about 88% of the countries experienced at least one currency crisis over the sample period¹⁸. The MPI is a continuous variable, while our currency crisis definition is a discrete binary variable. Also, the last definition is sensitive to the threshold used.

5 Empirical Methodology

The analysis of the relation between exchange arrangements and currency crises will be based on the discrete choice model method (probit model). Given our indicators, the model estimates the probability of a currency crisis. The estimated model takes the form:

$$Prob. (y_{it} = 1|x_{it}, \beta_t) = F(x_{it}, \beta_t) \quad (4)$$

where x_t corresponds to our set of indicators and β_t is a vector of unknown parameters. The observed variable y_{it} assumes a value of 0 or 1 depending on whether a currency crisis has occurred or not. With a probit model, the right hand side of the model is constrained between 0 and 1, and is compared to the observed value y_{it} . The probit model assumes that the probability distribution function (y_{it} conditional on x_{it}) corresponds to normal distribution. The model with a success probability $F(x_{it}, \beta_t)$ and independent observations leads to the joint probability.

6 The Data

The sample consists of panel data for 125 countries classified by the World Bank according to their income. Advanced or developed countries are those economies classified as upper income countries. Emerging markets countries are defined according to the Morgan Stanley Capital International (MSCI) index¹⁹ at that moment. The rest of the countries are designated as developing. Table 2 provides a list of countries classified in each group.

The data set is annual, spanning from 1974 through to 1999. Data availability differs across countries. Particularly, the data for East-European countries which starts from the 1990s.

¹⁷ We use country specific thresholds because the standard deviations are computed within each-country and not for the whole sample.

¹⁸ For an analysis of sensitivity to different threshold we use the MPI greater in value than 1.3, 1.4, 1.6, 1.7, 2.0, 2.5 and 3.0 standard deviations over the country's own mean value. The total of number of crises according to the conventional criterion of 2.5 or 3.0 standard deviations are very small (50 and 25 respectively). Similarly, the total of number of crises when we use high threshold (mean plus 1.7 or 2.0 standard deviations) are 164 and 101, respectively. On the contrary, when we use low threshold (mean plus 1.3 standard deviations) we obtained 295 crises. However, when we use the threshold of mean plus 1.4 or 1.6 standard deviation we obtained 256 and 195 crises, these results are closer to those obtained when we use 1.5 standard deviations.

¹⁹ The MSCI index classifies a country into an emerging market in line with a number of factors relating to international capital market access.

Table 2: List of Countries

| Advanced Countries | Emerging Markets | Developing Countries | | |
|---------------------------|-------------------------|-----------------------------|---------------|-------------------------|
| Australia | Argentina | Algeria | Haiti | Niger |
| Austria | Brazil | Antigua & Barbuda | Honduras | Nigeria |
| Belgium | Chile | Benin | Ivory Coast | Panama |
| Canada | China | Bolivia | Jamaica | Paraguay |
| Cyprus | Colombia | Botswana | Kazakhstan | Romania |
| Denmark | Czech Republic | Burkina Faso | Kenya | Saudi Arabia |
| Finland | Egypt | Burundi | Kyrgyz Rep. | Senegal |
| France | Hungary | Cameron | Lao Dem. Rep. | Slovak Rep. |
| Germany | India | Chad | Latvia | Sri Lanka |
| Greece | Indonesia | Congo, Rep. of | Lebanon | St. Lucia |
| Iceland | Israel | Costa Rica | Lesotho | St. Kitt & Nevis |
| Ireland | Jordan | Croatia | Liberia | St. Vicent & Grenadines |
| Italy | Korea, Rep. | Dominica | Libya | Suriname |
| Japan | Malaysia | Dominican Rep. | Lithuania | Swaziland |
| Kuwait | Mexico | Ecuador | Macedonia | Tanzania |
| Luxembourg | Morocco | El Salvador | Madagascar | Togo |
| Netherlands | Pakistan | Equatorial Guinea | Malawi | Tunisia |
| Norway | Peru | Estonia | Mali | Uganda |
| Portugal | Philippines | Gabon | Malta | Ukraine |
| Singapore | Poland | Gambia, the | Mauritius | Uruguay |
| Slovenia | Rusia | Georgia | Moldova | Zambia |
| Spain | South Africa | Ghana | Mongolia | Zimbabwe |
| Sweden | Thailand | Grenada | Myanmar | |
| Switzerland | Turkey | Guatemala | Nepal | |
| United Kingdom | Venezuela | Guinea-Bissau | New Zealand | |
| United States | | Guyana | Nicaragua | |

Note: Emerging market economies are those that are included in the Morgan Stanley Capital International (MSCI) index. Advanced economies are those that are classified as upper income economies by the World Bank, with the exception of Israel, which is in an emerging market. The remaining countries were designated as developing countries.

Most of the macroeconomic and financial variables used in our analysis are taken from the World Bank's World Development Indicators and the IMF's World Economic Outlook databases. A few series are taken from the CD-ROM version of the International Monetary Fund's International Financial Statistic (IFS). The data from the *de jure* IMF classification can be obtained from the IMF's Annual Report on Exchange Arrangements and Exchange Restrictions and Ghosh et al. (2002). For the Market Pressure Index (MPI) calculations, we employ total non-gold international reserves, average period exchange rates and short-term interest rates. Money market rates were used for all the countries where available, and t-bill rates, bank lending or deposit rates otherwise; in a number of cases, discount rates were used, when no other interest rate data were available (see Table 3).

Table 3: Interest Rate Used for the Corresponding Countries

| Money Market | T-bill | Bank Lending | Bank Deposit | Discount |
|----------------|-------------------|--------------------------|----------------|----------------|
| Argentina | Belgium | Antigua & Barbuda | Algeria | Benin |
| Australia | France | Dominica | Bolivia | Botswana |
| Austria | Guyana | El Salvador | Chile | Burkina Faso |
| Brazil | Jamaica | Er. Guinea | Dominican Rep. | Burundi |
| Canada | Kazakhstan | Estonia | Greece | Cameroon |
| Croatia | Kenya | Gabon | Guatemala | Chad |
| Czech Rep. | Kyrgyz Rep. | Grenada | Guinea-Bissau | China |
| Denmark | Lesotho | Honduras | Haiti | Colombia |
| Finland | Moldova | Israel | Hungary | Congo, Rep. of |
| Georgia | Romania | Liberia | Iceland | Costa Rica |
| Germany | St. Kitts & Nevis | Macedonia | Indonesia | Cyprus |
| Ireland | | Nigeria | Korea | Ecuador |
| Italy | | Panama | Lao Dem. Rep. | Egypt |
| Japan | | Poland | Libya | Gambia, The |
| Latvia | | Slovak Rep. | Lithuania | Ghana |
| Luxembourg | | Slovenia | Madagascar | India |
| Malaysia | | St. Lucia | Mexico | Ivory coast |
| Mauritius | | St. Vincent & Grenadines | Mongolia | Jordan |
| Netherlands | | Suriname | Morocco | Kuwait |
| Norway | | Swaziland | Myanmar | Lebanon |
| Pakistan | | Ukraine | Nicaragua | Malawi |
| Paraguay | | Uruguay | Saudi Arabia | Mali |
| Philippines | | Zambia | Turkey | Malta |
| Russia | | | Uganda | Nepal |
| Singapore | | | | New Zealand |
| South Africa | | | | Niger |
| Spain | | | | Peru |
| Sri Lanka | | | | Portugal |
| Sweden | | | | Senegal |
| Switzerland | | | | Tanzania |
| Thailand | | | | Togo |
| Tunisia | | | | Venezuela |
| United Kingdom | | | | |
| United States | | | | |
| Zimbabwe | | | | |

Notes: Money Market is the rate on short-term lending between financial institutions. Treasury bill rate is the rate at which short-term securities are issued or traded in the market. Lending rate is the bank rate that usually meets the short- and medium-term financing needs of the private sector. This rate is normally differentiated according to the creditworthiness of borrowers and objectives of financing. Deposit rate usually refers to rates offered to resident customers for demand, time or saving deposits. Discount rate is the rate at which the central banks lend or discount eligible paper for deposit money banks, typically shown on an end-of-period basis.

The variables used in this analysis and their descriptions are listed in Table 4. These variables were selected on the basis of previous theoretical and empirical literature. Government balance is defined as current and capital revenue and official grants received, less total expenditure and lending minus repayments. This variable considers central governments only. Short-term debt is defined as debt that has an original maturity of one year or less. Available data does not permit a distinction between public and private non-guaranteed short-term debt. The ratio of bank liquid reserves to bank assets is the ratio of domestic currency holding and deposits with the monetary authorities to claims on other governments, nonfinancial public enterprises, the private sector, and other banking institutions. Money and quasi money are defined as the sum of currency outside banks, demand deposits other than those of the central government, and the time, savings, and foreign currency deposits of resident sectors other than the central government. This definition of money supply is frequently called M2. Foreign direct investment is the sum of equity capital, reinvestment of

earnings, other long-term capital, and short-term capital as shown in the balance of payments. Current account balance is the sum of the credits less the debits arising from international transactions in goods, service, income, and current transfers. Unemployment refers to the share of the labour force that is without work but available for and seeking employment. International reserves are the sum of a country's monetary authorities' holdings of special drawing rights, its reserve position in the IMF, its holdings of foreign exchange, and its holdings of gold. Variables expressed in US dollar were converted to the natural logarithmic scale. The rest of variables were expressed in percentage. Finally, floating and intermediate exchange rate regimes are identified with a dummy variable that received the value of one in which these regimes prevail in a country in a particular year.

Table 4: List of variables used in the estimations

| Variable | Description |
|-----------------|--|
| Gov. Balance | Central government balance (% of GDP) |
| Stdebratio | Short-term debt/Total debt (%) |
| Debt | Total debt/GNI (%) |
| Domfin | Domestic financing, total (% of GDP) |
| Debtsx | Debt service/Exports of goods and services (%) |
| Bnkres | Ratio of bank liquid reserves to bank assets (%) |
| Dcrep | Domestic credit to private sector (% GDP) |
| M2gdp | Money and quasi money (% GDP) |
| M2res | Money and quasi money (% Reserves) |
| Resdebt | Reserves/Total debt (%) |
| Resimp | Reserves/Imports of goods and services (%) |
| Fdigni | Foreign direct investment (% of GNI) |
| Cagni | Current account balance (% GDP) |
| Inflation | The consumer price index (%) |
| Unempl | Unemployment, total (% of total labour force) |
| Usirate | USA short-term interest rate (%) |
| Reserves | International reserves (US\$) |
| Per capita GDP | Per capita real GDP growth (%) |
| Real GDP | Real GDP growth (%) |
| Openness | Exports plus imports of goods and services (% GDP) |
| Floating | Dummy variable capturing float exchange rate regimes |
| Intermediate | Dummy variable capturing intermediate arrangements |

Notes: The table does not include the dependent variables, which are explained in the text. Variables expressed in US dollars were converted to the natural logarithmic scale for the purpose of estimation.

7 Estimation Results

In order to examine which exchange arrangements are more prone to a currency crisis we use a probit model where the dependent variable is the probability of a currency crisis and the independent variables are all the above-mentioned variables simultaneously (not reported), but insignificant variables were gradually eliminated, until the most parsimonious representation of the data was achieved²⁰.

The impact of exchange regimes on the probability of currency crises is shown in Tables 5 and 6. The signs of independent variables are mostly as expected. Also, the statistical characteristics of the models are favourable. Most variables are significant to the level of 10%.

²⁰ However, in some cases the dummy variables of exchange rates were statistically not significant but they are not excluded.

The LR statistic shows the general statistical significance of the models (zero hypothesis of no significance of all the coefficients in the models was rejected with a significance of 1%)²¹. However, McFadden R² indicates relatively low goodness-of-fit in the models (between 5% and 28%).

Table 5: The Impact of Exchange Arrangements on Currency Crises in All Countries and Advanced Economies

| | All Countries | | | | Advanced Economies | | | |
|-------------------------|--------------------|-------------------|--------------------|--------------------|--------------------|-------------------|--------------------|-------------------|
| | Natural | LYS | HMR | <i>De jure</i> | Natural | LYS | HMR | <i>De jure</i> |
| Constant | -1.26 (-6.80)* | -1.69 (-8.31)* | -1.08 (-4.99)* | -1.20 (-7.08)* | -1.58 (-4.94)* | -1.15 (-3.23)* | -0.86 (-2.46)** | -1.33 (-4.59)* |
| Per cap. GDP | -0.01 (-1.00) | -0.02 (-1.28) | -0.04 (-2.85)* | -0.02 (-1.84)^ | -0.05 (-2.31)# | -0.09 (-2.02)# | -0.10 (-2.46)# | -0.06 (-2.41)# |
| Gov. balance | -0.03 (-2.61)* | -0.04 (-3.01)* | -0.04 (-2.23)# | -0.03 (-3.04)* | | | | |
| Dcrep | 0.01 (2.32)# | 0.002 (0.60) | 0.01 (2.92)* | 0.01 (2.05)# | 0.01 (1.75)^ | 0.004 (1.15) | 0.01 (1.47) | 0.004 (1.42) |
| Resimp | -0.02 (-4.84)* | -0.02 (-4.11)* | -0.02 (-2.92)* | -0.02 (-4.97)* | -0.03 (-3.29)* | -0.03 (-2.83)* | -0.04 (-3.41)* | -0.02 (-3.08)* |
| Resdebt | 0.004 (2.357)# | 0.001 (0.31) | -0.003 (-0.44) | 0.004 (2.16)# | | | | |
| Inflation | 0.0001 (2.12)# | 0.0001 (1.45) | 0.0003 (1.64) | 0.0002 (2.58)* | 0.02 (1.55) | 0.01 (1.28) | 0.02 (1.66)^ | 0.02 (1.91)^ |
| Openness | -0.004 (-2.34)# | -0.000 (-0.18) | -0.004 (-1.68)^ | -0.004 (-2.28)# | | | | |
| Floating | 0.53 (3.81)* | 0.74 (4.56)* | -0.03 (-0.09) | 0.39 (2.63)* | 0.19 (0.65) | 0.16 (0.68) | -0.09 (-0.38) | 0.10 (0.45) |
| Intermediate | 0.09 (0.65) | 1.01 (6.32)* | 0.18 (1.07) | 0.37 (2.83)* | 0.34 (1.39) | -0.43 (-1.33) | -0.45 (-1.88)^ | -0.06 (-0.26) |
| Observations | 1370 | 1168 | 706 | 1345 | 581 | 418 | 472 | 581 |
| Obs.= 0 | 1260 | 1079 | 644 | 1236 | 540 | 388 | 436 | 540 |
| Obs.= 1 | 110 | 89 | 62 | 109 | 41 | 30 | 36 | 41 |
| LR Stat. | 85.558 | 95.553 | 58.586 | 79.946 | 25.644 | 23.243 | 31.088 | 23.987 |
| Prob. LR | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| McFadden R ² | 0.11 | 0.15 | 0.14 | 0.11 | 0.07 | 0.11 | 0.12 | 0.08 |

Notes: The dependent variable is currency crises. z -statistics are displayed in brackets. (*) denote significance at the 1 per cent level, (#) at the 5 per cent and (^) at the 10 per cent level.

Source: Author's estimates.

In order to evaluate the goodness-of-fit in the models, we carry out two goodness-of-fit tests: the power of the models in predicting a currency crisis in the sample and the Hosmer-Lemeshow test. The standard method of a probit model to evaluate its predictive power is to compare the estimated probabilities of a crisis with actual occurrences. For this purpose, a probability threshold was set to serve as a criterion for the decision whether a model signals a crisis or not. In case the probability of a crisis exceeds the threshold or cut-off level, the model is considered to send a signal and vice versa. Using a cut-off level for the probability of a crisis as 50%, the models issue hardly any wrong signals, but they missed most of the crises in the sample²². As shown in Table 7, the probability threshold, as the value separating the crisis period from the tranquil period, was set at 15%.

²¹ LR statistic is analogous to the F statistic in the models estimated OLS.

²² Similar to the results surveyed by Berg and Pattillo (1999) our diagnostic statistics reveal that the models rarely generate a predicted probability of crises above 50%.

Table 6: The Impact of Exchange Arrangements on Currency Crises in Emerging and Developing Countries

| | Emerging Economies | | | | Developing Countries | | | |
|-------------------------|-------------------------------|-------------------|--------------------|-------------------------------|----------------------|-------------------|-------------------|--------------------|
| | Natural | LYS | HMR | <i>De jure</i> | Natural | LYS | HMR | <i>De jure</i> |
| Constant | -0.44 (-0.89) | -1.73 (-2.47)# | -0.35 (-0.59) | -0.32 (-0.66) | -1.52 (-10.95)* | -1.52 (-9.76)* | -1.52 (-6.16)* | -1.46 (-11.09)* |
| Per cap. GDP | 0.087 (-3.59)* | -0.10 (-3.32)* | -0.10 (-2.61)* | -0.06 (-2.56)* | | | | |
| Gov. balance | -0.09 (-3.59)* | 0.02 (0.38) | 0.10 (1.06) | 0.06 (1.14) | -0.03 (-3.42)* | -0.03 (-2.53)# | -0.07 (-3.99)* | -0.03 (-3.58)* |
| Dcrep | 0.01 (2.36)# | 0.004 (0.91) | 0.01 (1.57) | 0.01 (1.42) | | | | |
| Domfin | 0.05 (0.71) | 0.06 (0.84) | 0.10 (0.99) | 0.04 (0.54) | | | | |
| Resimp | -0.06 (-4.31)* | -0.04 (-3.07)* | -0.03 (-2.40)# | -0.05 (-4.25)* | -0.01 (-2.47)# | -0.02 (-2.97)* | -0.01 (-0.64) | -0.01 (-2.61)* |
| Resdebt | 0.01 (1.27) | 0.01 (0.74) | 0.003 (0.21) | 0.01 (1.41) | -0.001 (-0.23) | -0.004 (-0.89) | -0.02 (-1.45) | -0.0004 (-0.17) |
| Debt | 0.01 (0.29) | -0.03 (-0.65) | -0.04 (-0.73) | 0.02 (0.62) | | | | |
| Debtsx | | | | | 0.001 (0.37) | 0.0002 (0.04) | 0.004 (0.49) | 0.003 (0.89) |
| Inflation | 4.70e ⁰⁵ (0.22) | -0.001 (-0.78) | 0.004 (1.04) | 8.83e ⁰⁶ (0.04) | | | | |
| Openness | -0.002 (-0.54) | 0.01 (0.89) | -0.0002 (-0.03) | -0.002 (-0.50)* | | | | |
| Floating | -0.19 (-0.52) | 1.05 (2.10)# | -9.99 (-0.88) | 0.26 (0.71) | 0.35 (2.62)* | 0.55 (3.09)* | 0.89 (1.29) | 0.22 (1.39) |
| Intermediate | 0.41 (1.59) | 1.27 (2.41)# | -0.34 (-1.01) | -0.18 (-0.59) | 0.13 (0.88) | 0.90 (5.79)* | 0.49 (2.39)# | 0.17 (1.15) |
| Observations | 377 | 318 | 261 | 388 | 1210 | 1021 | 437 | 1191 |
| Obs.= 0 | 345 | 294 | 241 | 355 | 1114 | 943 | 396 | 1095 |
| Obs.= 1 | 32 | 24 | 20 | 33 | 96 | 78 | 41 | 96 |
| LR Stat. | 58.506 | 49.226 | 39.708 | 54.244 | 37.567 | 61.257 | 28.836 | 34.472 |
| Prob. LR | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| McFadden R ² | 0.27 | 0.24 | 0.28 | 0.24 | 0.06 | 0.11 | 0.11 | 0.05 |

Notes: The dependent variable is currency crises. z -statistics are displayed in brackets. (*) denote significance at the 1 per cent level, (#) at the 5 per cent and (^) at the 10 per cent level.

Source: Author's estimates.

Lowering the cut-off level to 15% leads to a strong improvement in the models' ability to predict currency crises in the sample, while the number of wrong signals rises only moderately. Also, the majority of correct predictions are for tranquil periods²³. Given the cut-off probability of 15% the models correctly call between 15.5% and 70.8% of the crises and between 84% and 96.7% of the tranquil periods. Also, the count R² indicates a relatively good goodness-of-fit in the models (see Table 8)²⁴.

²³ We also used a cut-off level for the probability of a crisis as 25%, but the models accurately predicted crises in emerging and developing countries particularly using LYS and *de Jure* classifications.

²⁴ The count R² is another comparatively simple measure of goodness-of-fit. It is defined as:

$$\text{Count } R^2 = \frac{\text{number of correct predictions}}{\text{total number of observations}}$$

Table 7: Goodness-of-Fit of the Probit Models

| Classification | Predicts | All countries | | Advanced | | Emerging | | Developing | |
|----------------|----------|---------------|--------|----------|--------|----------|--------|------------|--------|
| | | Tranquil | Crises | Tranquil | Crises | Tranquil | Crises | Tranquil | Crises |
| Natural | Tranquil | 1129 | 70 | 511 | 29 | 307 | 12 | 1029 | 79 |
| | Crises | 131 | 40 | 29 | 12 | 48 | 21 | 85 | 17 |
| LYS | Tranquil | 938 | 49 | 350 | 22 | 253 | 7 | 837 | 48 |
| | Crises | 141 | 40 | 38 | 8 | 41 | 17 | 106 | 30 |
| HMR | Tranquil | 567 | 37 | 385 | 22 | 213 | 10 | 343 | 27 |
| | Crises | 77 | 25 | 51 | 14 | 10 | 10 | 53 | 14 |
| De Jure | Tranquil | 1129 | 76 | 522 | 32 | 293 | 12 | 1028 | 81 |
| | Crises | 107 | 33 | 18 | 9 | 52 | 20 | 67 | 15 |

Source: Author's calculations.

Table 8: Measure of Goodness-of-Fit: The Count R²

| Classification | All Countries | Advanced | Emerging | Developing |
|----------------|-----------------------|-----------------------|-----------------------|-----------------------|
| Natural | R ² = 0.85 | R ² = 0.90 | R ² = 0.85 | R ² = 0.86 |
| LYS | R ² = 0.84 | R ² = 0.86 | R ² = 0.85 | R ² = 0.85 |
| HMR | R ² = 0.84 | R ² = 0.85 | R ² = 0.85 | R ² = 0.82 |
| De Jure | R ² = 0.86 | R ² = 0.91 | R ² = 0.83 | R ² = 0.88 |

Source: Author's calculations.

Alternatively, if the average of the predicted values approaches the average of the observed outcomes successfully, a model is considered to be well fitted. The Hosmer-Lemeshow test statistics are commonly used to assess these properties. In order to calculate these test statistics, the data set is sorted in ascending order by the predicted probability of a currency crisis. The data set is then split into the subsets by grouping the first quantile of observations into the first set, and so forth. For each subset, the difference between the observed and predicted number of currency crises is determined on which the Hosmer-Lemeshow test statistics are based. Most Hosmer-Lemeshow test statistics lead to not rejecting the null hypothesis of no difference between observed and predicted values at 10% level (see Table 9).

Table 9: Hosmer-Lemeshow Goodness-of-Fit Test

| Classification | All Countries | Advanced | Emerging | Developing |
|----------------|--------------------------|--------------------------|--------------------------|--------------------------|
| Natural | $\chi^2(8) = 4.33(0.83)$ | $\chi^2(8) = 8.69(0.37)$ | $\chi^2(8) = 9.28(0.32)$ | $\chi^2(8) = 7.29(0.51)$ |
| LYS | $\chi^2(8) = 2.79(0.95)$ | $\chi^2(8) = 2.52(0.96)$ | $\chi^2(8) = 3.76(0.88)$ | $\chi^2(8) = 3.82(0.87)$ |
| HMR | $\chi^2(8) = 5.91(0.65)$ | $\chi^2(8) = 8.08(0.43)$ | $\chi^2(8) = 4.49(0.81)$ | $\chi^2(8) = 4.72(0.79)$ |
| De Jure | $\chi^2(8) = 9.97(0.27)$ | $\chi^2(8) = 7.96(0.44)$ | $\chi^2(8) = 9.25(0.32)$ | $\chi^2(8) = 5.71(0.68)$ |

Source: Author's calculations.

Considering the results shown in Tables 5 and 6, can be pointed out that the probability of currency crises increases along with a low ratio of foreign reserves to import of goods and services, high inflation, increases in the ratio of domestic financing to GDP and the ratio of

domestic credit to private sector to GDP, low ratio of international reserves to total debt and increases in the ratio of total debt to Gross National Income. As expected, increases in per capita GDP growth rate, among others, reduce the probability of currency crises.

Table 10: Exchange Arrangements Performance on Currency Crises

| | Natural | LYS | HMR | De Jure |
|--|-------------------------------------|-------------------------------------|-------------------------------------|-------------------------------------|
| Ranking from the best to the worst performance | <i>All Countries</i> | | | |
| | Fixed Intermediate* Floating | Fixed Floating Intermediate | Floating* Fixed Intermediate* | Fixed Intermediate Floating |
| | <i>Advanced Economies</i> | | | |
| | Fixed Floating* Intermediate* | Intermediate* Fixed Floating* | Intermediate Floating* Fixed | Intermediate* Fixed Floating* |
| | <i>Emerging Economies</i> | | | |
| | Floating Fixed Intermediate* | Fixed Floating Intermediate | Floating* Intermediate* Fixed | Intermediate* Fixed Floating* |
| | <i>Developing Countries</i> | | | |
| Fixed Intermediate* Floating | Fixed Floating Intermediate | Fixed Intermediate Floating* | Fixed Intermediate* Floating* | |

Note: (*) insignificant variables.

Source: Author's calculations.

In addition, our results suggest that floating and intermediate exchange regimes are associated with a higher probability of currency crises than fixed regimes (see Table 10). Developing countries using fixed arrangements have a lower likelihood of currency crises relative to similar countries using floating or intermediate regimes. An explanation is that countries with underdeveloped or weak financial systems are also likely to have problems accommodating large exchange rate movements under flexible regimes.

For advanced and emerging economies, our results are not clear. They are sensitive to regime classification and different classifications can lead to very different results. Notwithstanding this, when the HMR classification is used in emerging countries, floating regimes show the best performance (not statistically significant) while emerging countries using fixed regimes increase the probability of currency crises. Conversely, when our model is applied to all the samples the results suggested that fixed arrangements are less prone to currency crises. These result are similar to findings by Ghosh et al. (2002); Falcetti and Tudela (2006), Haile and Pozo (2006) and Esaka (2010b), and contrary to Bubula and Otker-Rober (2003). However, when we use the LYS classification is found that intermediate exchange regimes are more prone to currency crises in all countries, as well as, in emerging and developing economies. These results could show a lower popularity of intermediate regimes, according to Rogoff et al. (2003). As a consequence, our results do not clarify whether floating or intermediate exchange arrangements are more prone to currency crises.

To summarise, currency crises tend to occur more frequently in countries using floating or intermediate regimes than those countries using fixed regimes, but it is not clear which exchange rate regime is more prone to currency crises. These results tend to suggest that the

affirmations on emerging and developing countries should allow for more exchange rate flexibility as a means to reduce the probability of currency and financial crises is not well founded.

8 Concluding Remarks

The academic debate on the most appropriate exchange rate regime for a country or group of countries has been one of the most controversial topics in theoretical and empirical literature. Notwithstanding its increasing relevance to policy, the literature offers relatively few empirical studies about the impact of the exchange rate regime on a currency crisis in developed, emerging and developing countries, separately. This article has provided an empirical analysis of the impact of different exchange rate regimes on currency crises in advanced, emerging and developing countries. To this end, we have attempted to make two contributions. To begin with, we distinguish between the *de jure* and the three *de facto* classifications system. We have used the IMF *de jure* classification and checked the robustness of our results with three different *de facto* classifications: the LYS classification based on a clustered analysis, the natural classification based mainly on market determined dual and parallel exchange rates, and the HMR classification based on exchange rate regimes and taking into account external shocks and revaluations. The most complete *de facto* exchange rate classifications are made by Reinhart and Rogoff (2004).

Secondly, our results also suggest that currency crises tend to occur more frequently in developing countries with floating or intermediate regimes and, also, with higher ratio of domestic financing to GDP, higher ratio of domestic credit to private sector to GDP and lower ratio of international reserves to imports, among other indicators, than developing countries adopting fixed arrangements. However, the results are not clear in terms of which exchange rate regime is more prone to currency crises. Conversely, our results on which exchange rate regimes increase the probability of currency crises in emerging and advanced economies that already have well developed financial sectors are not clear. They are sensitive to regime classification because different classifications can lead to very different results.

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