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Abstract

This study develops a Schumpeterian growth model with endogenous entry of heterogeneous firms to analyze the effects of monetary policy on economic growth via a cash-in-advance constraint on R&D investment. Our results can be summarized as follows. In the special case of a zero entry cost, an increase in the nominal interest rate decreases R&D, the arrival rate of innovations and economic growth as in previous studies. However, in the general case of a positive entry cost, an increase in the nominal interest rate affects the distribution of innovations that are implemented and would have an inverted-U effect on economic growth if the entry cost is sufficiently large. We also calibrate the model to aggregate data of the US economy and find that the growth-maximizing inflation rate is about 3%, which is consistent with recent empirical estimates.

JEL classification: O30, O40, E41
Keywords: monetary policy, inflation, economic growth, heterogeneous firms

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1 Introduction

This study develops a Schumpeterian growth model with endogenous entry of heterogeneous firms to analyze the effects of monetary policy on economic growth. The canonical Schumpeterian growth model in seminal studies such as Segerstrom et al. (1990), Grossman and Helpman (1991) and Aghion and Howitt (1992) features an identical step size of quality improvements across firms. In this study, we consider a Schumpeterian model with random quality improvements as in Minniti et al. (2013) but with the addition of a fixed entry cost to generate endogenous entry of firms with heterogeneous step sizes of quality improvements. To incorporate money demand into this growth-theoretic framework, we impose a cash-in-advance (CIA) constraint on R&D investment. Berentsen et al. (2012), Chu and Cozzi (2014) and Chu et al. (2015) provide extensive discussion on evidence for the presence of cash requirements on R&D expenditures. We capture these cash requirements using a CIA constraint on R&D.

In this monetary growth-theoretic framework, we derive the following results. In the special case of a zero entry cost, an increase in the nominal interest rate decreases R&D, the arrival rate of innovations and economic growth as in previous studies, such as Chu and Cozzi (2014) who consider a monetary Schumpeterian growth model with an identical step size of quality improvements, because the distribution of innovations that are implemented is exogenous under a zero entry cost despite random quality improvements. However, in the general case of a positive entry cost, monetary policy affects the distribution of innovations that are implemented. Specifically, an increase in the nominal interest rate decreases R&D and the arrival rate of innovations, which increases the present value of future profits. The resulting higher value of inventions leads to a lower threshold of quality improvements above which an innovation is implemented generating a positive effect on economic growth due to more entries. Together with the negative effect on the arrival rate of innovations, an increase in the nominal interest rate would have an inverted-U effect on economic growth if the entry cost is sufficiently large. Because the Fisher equation gives rise to a positive long-run relationship between the nominal interest rate and the inflation rate that is supported by empirical studies such as Mishkin (1992) and Booth and Ciner (2001), our result also implies an inverted-U relationship between inflation and economic growth. This theoretical prediction on an inverted-U relationship between inflation and economic growth is supported by empirical studies such as Bick (2010) and López-Villavicencio and Mignon (2011). Finally, we calibrate the model to aggregate data of the US economy to provide a quantitative analysis and find that the growth-maximizing inflation rate is 2.9%, which is close to the empirical estimate in López-Villavicencio and Mignon (2011) who identify a threshold inflation rate of 2.7% for industrialized countries.

For example, early empirical studies such as Hall (1992) and Opler et al. (1999) find a positive and significant relationship between R&D and cash flows in US firms. More recently, Bates et al. (2009) document that the average cash-to-assets ratio in US firms increased substantially from 1980 to 2006 and argue that this is partly driven by their rising R&D expenditures. Brown and Petersen (2011) provide evidence that firms smooth R&D expenditures by maintaining a buffer stock of liquidity in the form of cash reserves. Falato and Sim (2014) use firm-level data in the US to show that firms’ cash holdings increase (decrease) significantly in response to a rise (cut) in R&D tax credits. These results suggest that due to financial frictions, firms need to use cash to finance their R&D investment.
This study relates to the literature on innovation and economic growth. The R&D-based growth model originates from Romer (1990), who develops a variety-expanding growth model in which economic growth is driven by the development of new products. Then, Segerstrom et al. (1990), Grossman and Helpman (1991) and Aghion and Howitt (1992) develop the Schumpeterian quality-ladder growth model in which economic growth is driven by the quality improvement of existing products. For simplicity, these studies assume an identical step size for all quality improvements. A recent study by Minniti et al. (2013) generalizes the Schumpeterian model by allowing for heterogeneous step sizes of quality improvements that are randomly drawn from a distribution. Our study extends the elegant framework of Minniti et al. (2013) by introducing a fixed entry cost of implementing a developed invention in order to generate endogenous entries of heterogeneous firms, which turn out to have important implications on the effects of monetary policy.

This study also relates to the literature on inflation and innovation. In this literature, Marquis and Reflett (1994) is the seminal study that analyzes the effects of inflation on innovation in the Romer variety-expanding growth model. In contrast, we analyze the effects of inflation in a Schumpeterian quality-ladder model as in Chu and Lai (2013), Chu and Cozzi (2014), Chu et al. (2015) and He and Zou (2016), whose models however feature an identical step size of quality improvements across firms. Subsequent studies, such as Chu and Ji (2016) and Huang et al. (2015), consider monetary policy in a Schumpeterian growth model with both variety expansion and (identical) quality accumulation across firms. As in Marquis and Reflett (1994), these studies predict a monotonic relationship between inflation and economic growth. The present study contributes to this literature by allowing for the endogenous entry of firms with heterogeneous step sizes of quality improvements, which gives rise to a novel channel through which monetary policy affects innovation and growth. As a result, the model generates an inverted-U relationship between inflation and economic growth, which is supported by recent empirical studies.

The rest of this study is organized as follows. Section 2 presents and solves the model. Section 3 analyzes the effects of monetary policy. The final section concludes.

2 A Schumpeterian model with heterogeneous firms

The Schumpeterian quality-ladder growth model is based on Grossman and Helpman (1991). We extend their model by (a) introducing money demand via a CIA constraint on R&D to analyze monetary policy, (b) considering lab-equipment innovation and entry processes that use final goods (instead of labor) as the input, (c) allowing for random quality improvements as in Minniti et al. (2013), and (d) incorporating a fixed entry cost to generate endogenous entry of heterogeneous firms as in Melitz (2003). In summary, when a firm invents a higher quality product, the step size of the quality increment is randomly drawn from a Pareto distribution. See also Baldwin and Robert-Nicoud (2008), Haruyama and Zhao (2008) and Gustafsson and Segerstrom (2010) who adapt this fixed entry cost into the R&D-based growth model, but they do not consider random increments on the quality ladder.

3 The relationship between the two variables is usually found to be monotonically negative, but some of these studies also find that the relationship can be monotonically positive under some conditions.
distribution. If and only if the quality increment is sufficiently large, then the firm would pay the fixed entry cost to implement the invention and enter the market.

2.1 Household

In the economy, there is a representative household which has the following lifetime utility function:

\[ U = \int_0^\infty e^{-\rho t} \ln c_t dt, \tag{1} \]

where the parameter \( \rho > 0 \) is the subjective discount rate and \( c_t \) denotes consumption of final goods (numeraire) at time \( t \). The household maximizes utility subject to an asset-accumulation equation (expressed in real terms) given by

\[ \dot{a}_t + \dot{m}_t = r_t a_t - \pi_t m_t + i_t b_t + w_t + \tau_t - c_t. \tag{2} \]

\( a_t \) is the real value of financial assets (in the form of equity shares in monopolistic intermediate goods firms) owned by the household. \( r_t \) is the real interest rate. \( \pi_t \) is the inflation rate. \( m_t \) is the real money balance accumulated by the household. \( b_t \) is the amount of money borrowed by R&D entrepreneurs subject to the following constraint: \( b_t \leq m_t \). \( i_t \) is the interest rate on money \( b_t \) borrowed by R&D entrepreneurs, and it can be shown as a no-arbitrage condition that \( i_t \) must be equal to the nominal interest rate such that \( i_t = r_t + \pi_t \) from the Fisher equation. To earn the wage rate \( w_t \), the household inelastically supplies one unit of labor.\(^4\) \( \tau_t \) is a lump-sum transfer from the government to the household. From standard dynamic optimization, the familiar Euler equation is

\[ \frac{\dot{c}_t}{c_t} = r_t - \rho. \tag{3} \]

2.2 Final goods

Final goods are produced by perfectly competitive firms that employ labor and a composite of intermediate goods as inputs. The production function of final goods is \( Y_t = L_t^\theta K_t^{1-\theta} \), where \( L_t = 1 \) is labor input. \( K_t \) is a composite of intermediate goods produced with the following Cobb-Douglas aggregator:

\[ K_t = \exp \left\{ \int_0^1 \ln \left[ \sum_j q_t(\omega, j) y_t(\omega, j) \right] d\omega \right\}, \tag{4} \]

where the integer \( j \) in \( q_t(\omega, j) \) denotes the quality vintage of intermediate goods \( \omega \). Let \( j_\omega \) denotes the highest-quality vintage in industry \( \omega \). Firms are indifferent between the highest-quality vintage and the second-highest-quality vintage if their relative price is

\[ \frac{p_t(\omega, j_\omega)}{p_t(\omega, j_\omega - 1)} = \frac{q_t(\omega, j_\omega)}{q_t(\omega, j_\omega - 1)} \equiv \lambda_t(\omega), \tag{5} \]

\(^4\)Given that our model is already quite complex, we normalize the aggregate supply of labor to unity in order to sidestep the issue of scale effects; see for example, Peretto (1998, 2007) and Segerstrom (1998) for important ways of removing the strong scale effect in the Schumpeterian growth model.
where $\lambda_t(\omega) > 1$ is the quality increment between the two consecutive vintages of intermediate goods $\omega$ at time $t$. As usual, whenever this equality holds, we focus on the case in which firms buy the highest-quality intermediate goods only. In equilibrium, only the highest quality intermediate goods are traded. From profit maximization, the conditional demand function for intermediate goods $\omega \in [0, 1]$ is given by

$$y_t(\omega, j_\omega) = \frac{(1 - \theta) Y_t}{p_t(\omega, j_\omega)} = \frac{(1 - \theta) K_t^{1-\theta}}{p_t(\omega, j_\omega)}.$$  

(6)

Multiplying $q_t(\omega, j_\omega)$ to both sides of (6) and then aggregating the natural log of the resulting equation with respect to $\omega$, we derive

$$K_t = \left[(1 - \theta)Q_t/P_t\right]^{1/\theta},$$

(7)

where $Q_t \equiv \exp\left[\int_0^1 \ln q_t(\omega, j_\omega) d\omega\right]$ and $P_t \equiv \exp\left[\int_0^1 \ln p_t(\omega, j_\omega) d\omega\right]$ denote respectively the aggregate quality index and the aggregate price index of intermediate goods.

2.3 Intermediate goods

There is a unit continuum of industries $\omega \in [0, 1]$ producing differentiated intermediate goods. Each industry is temporarily dominated by a quality leader until the arrival and implementation of the next higher-quality product. The owner of the new innovation becomes the next quality leader.\(^5\) The current quality leader in industry $\omega$ uses one unit of final goods to produce one unit of intermediate goods $y_t(\omega, j_\omega)$, so that the marginal cost of production is one. From Bertrand competition,\(^6\) limit pricing yields the equilibrium price given by

$$p_t(\omega, j_\omega) = \lambda_t(\omega).$$

(8)

Therefore, the amount of monopolistic profit in industry $\omega$ is

$$\Pi_t(\omega, j_\omega) = [\lambda_t(\omega) - 1] y_t(\omega, j_\omega) = \left[\frac{\lambda_t(\omega) - 1}{\lambda_t(\omega)}\right] (1 - \theta) Y_t,$$

(9)

where the second equality uses (6) and (8).

2.4 R&D

R&D is performed by a continuum of competitive entrepreneurs. If an R&D entrepreneur employs $R_t(\omega)$ units of final goods to engage in innovation in industry $\omega$, then she is successful in inventing the next higher-quality product in the industry with an instantaneous probability given by

$$\phi_t(\omega) = R_t(\omega)/\alpha_t,$$

(10)

\(^5\)This is known as the Arrow replacement effect; see Cozzi (2007) for a discussion of the Arrow effect.

\(^6\)See Denicolò and Zanchettin (2010) for an analysis of Cournot competition in the Schumpeterian model.
where $\alpha_t \equiv \alpha Q_t^{(1-\theta)/\theta}$ inversely measures R&D productivity and is proportional to $Q_t^{(1-\theta)/\theta}$ to ensure balanced growth. To facilitate the payment of $R_t(\omega)$, the entrepreneur needs to borrow cash from the household, and the cost of borrowing is determined by the nominal interest rate $i_t$. Therefore, the cost of R&D is $(1 + i_t) R_t(\omega)$. Let $v_t^c(\omega, j_\omega + 1)$ denotes the expected value of an innovation before the realization of its quality increment. Then, the R&D free-entry condition is given by

$$v_t^c(\omega, j_\omega + 1) \phi_t(\omega) = (1 + i_t) R_t(\omega) \iff v_t^c(\omega, j_\omega + 1) / \alpha_t = (1 + i_t). \quad (11)$$

### 2.5 Random quality improvements

As in Minniti et al. (2013), when an R&D entrepreneur invents a higher-quality product in industry $\omega$, the quality increment $\lambda_t(\omega) > 1$ is drawn from a stationary Pareto distribution with the following probability density function:

$$f(\lambda) = \frac{1}{\kappa} \lambda^{-\frac{1+\kappa}{\kappa}}, \quad (12)$$

where the parameter $\kappa \in (0, 1)$ determines the shape of the Pareto distribution. Given that the expected value of $\lambda_t(\omega)$ is equal across industries, (9) implies that the expected value of $\Pi_t(\omega, j_\omega)$ is also the same across industries. Therefore, we will follow the standard treatment in the literature to focus on the symmetric equilibrium in which the arrival rate of innovations is equal across industries, such that $v_t(\omega) = v_t(\omega)$. Given the entry cost, a firm enters the market if and only if $v_t(\omega) \geq \phi_t(\omega)$, where $v_t(\omega)$ denotes the ex post value of an innovation (i.e., after the realization of the quality increment $\lambda$).\footnote{Cozzi et al. (2007) provide a theoretical justification for the symmetric equilibrium to be the unique rational-expectation equilibrium in the Schumpeterian model.} $v_t(\lambda)$ is monotonically increasing in $\lambda$ because $\Pi_t(\lambda) = (1 - \theta) Y_t(\lambda - 1) / \lambda$ is increasing in $\lambda$. Given that $v_t(1) = 0$ and $v_t(\lambda) / Q_t^{(1-\theta)/\theta}$ is stationary in equilibrium, it can be shown that there exists a stationary threshold value of $\lambda$, denoted as $\check{\lambda}$, above which firms implement their innovations and enter the market generating endogenous entry of firms with heterogeneous quality improvements.

### 2.6 Endogenous firm entry

To generate an endogenous distribution of heterogeneous firms, we follow Melitz (2003) and others to consider a fixed entry cost. The entry cost is given by $\beta_t \equiv \beta Q_t^{(1-\theta)/\theta}$, which is proportional to $Q_t^{(1-\theta)/\theta}$ to ensure balanced growth. Given the entry cost, a firm enters the market if and only if $v_t(\lambda) \geq \beta_t$, where $v_t(\lambda)$ denotes the ex post value of an innovation (i.e., after the realization of the quality increment $\lambda$).\footnote{We do not impose a CIA constraint on entry for the following reasons. Unlike R&D investment that is subject to uncertainty in innovation success, the entry cost is incurred after an innovation is already developed and patented. Therefore, banks should be available to extend credits to the firm, which can use the patent as a collateral.} $v_t(\lambda)$ is monotonically increasing in $\lambda$ because $\Pi_t(\lambda) = (1 - \theta) Y_t(\lambda - 1) / \lambda$ is increasing in $\lambda$. Given that $v_t(1) = 0$ and $v_t(\lambda) / Q_t^{(1-\theta)/\theta}$ is stationary in equilibrium, it can be shown that there exists a stationary threshold value of $\lambda$, denoted as $\check{\lambda}$, above which firms implement their innovations and enter the market generating endogenous entry of firms with heterogeneous quality improvements.\footnote{In a symmetric equilibrium with $\phi_t(\omega) = \phi_t$, the value of innovations does not depend on $\omega$.}
2.7 Asset prices

The ex-ante value of an innovation (i.e., before the realization of $\lambda$) is formally defined as

$$v^e_t(\omega, j\omega + 1) = \int_1^\lambda 0 \cdot f(\lambda)d\lambda + \int_1^\lambda [v_t(\lambda) - \beta_t]f(\lambda)d\lambda = \int_1^\lambda v_t(\lambda)f(\lambda)d\lambda - \Pr(\lambda \geq \bar{\lambda})\beta_t,$$

where $\Pr(\lambda \geq \bar{\lambda})$ denotes the probability of the innovation being implementable. In the symmetric equilibrium with $v^e_t(\omega, j\omega + 1) \equiv v^e_t$, the no-arbitrage condition for the ex-ante value of innovation can be derived as\(^{11}\)

$$r_t = \frac{\Pi^e_t + \dot{v}^e_t + \Pr(\lambda \geq \bar{\lambda})\beta_t - \Pr(\lambda \geq \bar{\lambda})\phi_t\left[v^e_t + \Pr(\lambda \geq \bar{\lambda})\beta_t\right]}{v^e_t + \Pr(\lambda \geq \bar{\lambda})\beta_t},$$

(13)

where $\phi_t$ is the arrival rate of innovation. $\Pr(\lambda \geq \bar{\lambda})\phi_t$ is the instantaneous probability that an innovation is created and implemented in an industry. The Pareto probability density function implies that

$$\Pr(\lambda \geq \bar{\lambda}) = \int_\bar{\lambda}^\infty f(\lambda)d\lambda = \bar{\lambda}^{-1/\kappa}.$$

(14)

Substituting (14) into (13) and rearranging terms yield

$$\frac{\Pi^e_t}{v^e_t + \bar{\lambda}^{-1/\kappa}\beta_t} = r_t + \bar{\lambda}^{-1/\kappa}\phi_t - \frac{\dot{v}^e_t + \bar{\lambda}^{-1/\kappa}\beta_t}{v^e_t + \bar{\lambda}^{-1/\kappa}\beta_t},$$

(15)

where the ex-ante value of monopolistic profits can be shown to be

$$\Pi^e_t = \left[\int_\bar{\lambda}^\infty \left(\frac{\lambda - 1}{\lambda}\right) f(\lambda)d\lambda\right] (1 - \theta)Y_t = \left[\frac{\bar{\lambda} - 1/(1 + \kappa)}{\bar{\lambda}^{1 + \kappa}/\lambda^{1 + \kappa}}\right] (1 - \theta)Y_t.$$

(16)

Similarly, the no-arbitrage condition for the ex-post value of an innovation with $\lambda \geq \bar{\lambda}$ is

$$\frac{\Pi_t(\lambda)}{v_t(\lambda)} = r_t + \bar{\lambda}^{-1/\kappa}\phi_t - \frac{\dot{v}_t(\lambda)}{v_t(\lambda)},$$

(17)

where the ex-post value of monopolistic profits with $\lambda \geq \bar{\lambda}$ is given by

$$\Pi_t(\lambda) = \left(\frac{\lambda - 1}{\lambda}\right) (1 - \theta)Y_t.$$

(18)

\(^{11}\)See Appendix A for the proof.
2.8 Monetary authority

The monetary policy instrument that we consider is the nominal interest rate \( i_t \), which is exogenously set by the monetary authority. Given \( i_t \), the inflation rate \( \pi_t \) is endogenously determined according to the Fisher equation such that \( \pi_t = i_t - r_t \), where \( r_t \) is the real interest rate and determined from the Euler equation in (3). Then, the growth rate of the nominal money supply is given by \( \mu_t = \pi_t + \dot{m}_t/m_t \), which becomes \( \mu = i - \rho \) on the balanced growth path.\(^{12}\) Finally, the monetary authority returns the seigniorage revenue as a lump-sum transfer \( \tau_t = \dot{m}_t + \pi_t m_t \) to the household.

2.9 Dynamics

In this section, we characterize the dynamics of the model. Lemma 1 shows that given a constant nominal interest rate \( i \), the economy immediately jumps to a balanced growth path. On this balanced growth path, each variable grows a constant (possibly zero) growth rate.

**Lemma 1** The economy jumps to a unique and saddle-point stable balanced growth path.

**Proof.** See Appendix B. ■

2.10 Economic growth

Recall that the (log of) aggregate quality index is \( \ln Q_t \equiv \int_0^1 \ln q_t(\omega, j_\omega) d\omega \). In industry \( \omega \), the quality \( q_t(\omega, j_\omega) \) jumps to \( q_t(\omega, j_\omega + 1) = \lambda(\omega) q_t(\omega, j_\omega) \) with probability \( \Pr(\lambda \geq \bar{\lambda}) \phi = \bar{\lambda}^{-1/\kappa} \phi \). The continuum of industries shares this random process of quality improvements. Therefore, the time derivative of \( \ln Q_t \) is given by

\[
\frac{\dot{Q}_t}{Q_t} = \left\{ \int_0^1 [\ln q_t(\omega, j_\omega + 1) - \ln q_t(\omega, j_\omega)] d\omega \right\} \bar{\lambda}^{-1/\kappa} \phi = \left[ \int_0^1 \ln \lambda(\omega) d\omega \right] \bar{\lambda}^{-1/\kappa} \phi. \tag{19}
\]

Using the law of large numbers, we obtain\(^{13}\)

\[
\frac{\dot{Q}_t}{Q_t} = \left[ \int_\lambda^\infty (\ln \lambda) \bar{f}(\lambda) d\lambda \right] \bar{\lambda}^{-1/\kappa} \phi = (\ln \bar{\lambda} + \kappa) \bar{\lambda}^{-1/\kappa} \phi, \tag{20}
\]

where \( \ln \bar{\lambda} + \kappa \) captures the average step size of implemented quality improvements and \( \bar{f}(\lambda) \) is defined as

\[
\bar{f}(\lambda) \equiv \frac{f(\lambda)}{\int_\lambda^\infty f(\lambda) d\lambda} = \bar{\lambda}^{\frac{\kappa}{2}} f(\lambda).
\]

\(^{12}\)It is useful to note that in this model, it is the growth rate of the money supply that affects the real economy in the long run, and a one-time change in the level of money supply has no long-run effect on the real economy. This is the well-known distinction between the neutrality and superneutrality of money. Empirical evidence generally favors neutrality and rejects superneutrality, consistent with our model; see Fisher and Seater (1993) for a discussion on the neutrality and superneutrality of money.

\(^{13}\)Derivations are available in an unpublished appendix; see Appendix C.
Finally, the growth rate of output $Y_t$ and consumption $c_t$ is equal to

$$g = \frac{1 - \theta}{\theta} \dot{Q}_t = \frac{1 - \theta}{\theta} (\ln \bar{\lambda} + \kappa) \bar{\lambda}^{-1/\kappa} \phi. \tag{21}$$

Equation (21) shows that the equilibrium growth rate depends on two endogenous variables, the arrival rate $\phi$ of innovations and the threshold step size $\bar{\lambda}$. We can determine $\phi$ using the R&D condition $v_t^e = (1 + i) \alpha Q_t^{(1-\theta)/\theta}$, where the balanced-growth value of $v_t^e$ is given by $v_t^e = \Pi_t^e / (\rho + \bar{\lambda}^{-1/\kappa} \phi) - \bar{\lambda}^{-1/\kappa} \beta Q_t^{(1-\theta)/\theta}$ using (15) and the Euler equation. Then, substituting (16) into the R&D condition, we obtain

$$(1 - \theta) \left[ \frac{\bar{\lambda} - 1/(1 + \kappa)}{\bar{\lambda}^{1+\kappa}} \right] \frac{Y_t}{Q_t^{(1-\theta)/\theta}} = \left[ (1 + i) \alpha + \bar{\lambda}^{-1/\kappa} \beta \right] (\rho + \bar{\lambda}^{-1/\kappa} \phi). \tag{22}$$

In Appendix B, we show that the production function of final goods can be expressed as

$$Y_t = \left( \frac{1 - \theta}{\bar{\lambda} e^\kappa} \right)^{(1-\theta)/\theta} Q_t^{(1-\theta)/\theta}. \tag{23}$$

Similarly, we can determine $\bar{\lambda}$ using the entry condition $v_t(\bar{\lambda}) = \beta Q_t^{(1-\theta)/\theta}$, where the balanced-growth value of $v_t(\bar{\lambda})$ is given by $v_t(\bar{\lambda}) = \Pi_t(\bar{\lambda}) / (\rho + \bar{\lambda}^{-1/\kappa} \phi)$ using (17) and the Euler equation. Then, substituting (18) into the entry condition, we obtain

$$(1 - \theta) \left( \frac{\bar{\lambda} - 1}{\bar{\lambda}} \right) \frac{Y_t}{Q_t^{(1-\theta)/\theta}} = \beta (\rho + \bar{\lambda}^{-1/\kappa} \phi). \tag{24}$$

Combining (22) and (24), we have the $\bar{\lambda}$ condition given by

$$(\bar{\lambda} - 1) \bar{\lambda}^{1/\kappa} = \frac{1}{1 + i} \frac{\beta}{\alpha} \frac{\kappa}{1 + \kappa}, \tag{25}$$

where the left-hand side is monotonically increasing in $\bar{\lambda}$. Therefore, (25) implicitly determines the unique equilibrium value of $\bar{\lambda}$. Using (23)-(25), we obtain the $\phi$ condition given by

$$\phi = \frac{\bar{\lambda}^{-1/\theta}}{(1 + i)} \frac{\kappa}{1 + \kappa} \frac{(1 - \theta)^{1/\theta}}{\alpha e^{\kappa(1-\theta)/\theta}} - \rho \bar{\lambda}^{-1/\kappa}. \tag{26}$$

Given the equilibrium value of $\bar{\lambda}$ from (25), (26) determines the unique equilibrium value of $\phi$.

### 3 Monetary policy and economic growth

In this section, we explore the effects of monetary policy on economic growth. In Section 3.1, we analytically derive the effects of the nominal interest rate. In Section 3.2, we calibrate the model to quantify the relationship between inflation and economic growth.
3.1 Qualitative analysis

Here we first derive the effects of increasing the nominal interest rate $i$ on the innovation-arrival rate $\phi$ and the threshold step size $\lambda$. Lemma 2 shows that $\phi$ is decreasing in $i$ for a given $\lambda$. Lemma 3 shows that $\lambda$ is decreasing in $i$. The intuition can be explained as follows. An increase in the nominal interest rate $i$ increases the cost of R&D and reduces the incentives for innovation; as a result, the innovation rate $\phi$ decreases for a given $\lambda$. Lemma 2 shows that $\lambda$ is decreasing in $i$. From the balanced-growth version of (15), we have $v_t^e = \Pi_t^e/(\rho + \lambda^{-1/\kappa} \phi) - \lambda^{-1/\kappa} \beta Q_t^{(1-\theta)/\theta}$, which shows that the decrease in $\phi$, by reducing creative destruction, increases the present value of the profit stream generated by implementing an innovation. This induces the implementation of innovations associated with smaller profit margins, thereby reducing the threshold mark-up $\lambda$ for entry.

**Lemma 2** For a given $\lambda$, the innovation rate $\phi$ is decreasing in the nominal interest rate $i$.

**Proof.** Use (26). ■

**Lemma 3** The threshold step size $\lambda$ is decreasing in the nominal interest rate $i$.

**Proof.** Use (25). ■

When the entry cost $\beta_t$ is zero, the nominal interest rate has no effect on the distribution of innovations that are implemented because all firms enter the market regardless of the size of quality increments. In this case, $\lambda = 1$, and $g = 1 - 2\kappa \phi$ is monotonically decreasing in $i$ via $\phi$. This result is the same as in Chu and Cozzi (2014), who consider a Schumpeterian growth model with an identical step size of quality improvements across firms. However, when the entry cost $\beta_t$ is positive, the nominal interest rate $i$ affects both $\lambda$ and $\phi$. In this case, $\Pr(\lambda \geq \lambda) = \lambda^{-1/\kappa}$ is increasing in $i$. In other words, an increase in the nominal interest rate reduces the threshold value $\lambda$ for entry and leads to more innovations being implemented for a given $\phi$. When the entry cost $\beta_t$ is sufficiently large, the overall effects of $i$ on the composite innovation rate $\lambda^{-1/\kappa} \phi$ and the equilibrium growth rate $g = 1 - 2\kappa \phi (\ln \lambda + \kappa) \lambda^{-1/\kappa} \phi$ become non-monotonic. Specifically, we find that when the nominal interest rate $i$ increases, $\lambda^{-1/\kappa} \phi$ and $g$ first increase and eventually decrease. We summarize these results in Proposition 1.

**Proposition 1** If the entry cost is sufficiently large (small), an increase in the nominal interest rate has an inverted-U (negative) effect on the composite innovation rate $\lambda^{-1/\kappa} \phi$ and the equilibrium growth rate $g$.

**Proof.** See the Appendix B. ■

Before we conclude this section, we explore the relationship between inflation and economic growth. The Fisher equation gives rise to a positive long-run relationship between the inflation rate and the nominal interest rate that is supported by empirical studies such
as Mishkin (1992) and Booth and Ciner (2001). In our model, the inflation rate is given by the Fisher equation \( \pi = i - r = i - g(i) - \rho \), where the second equality follows from the Euler equation. Therefore, so long as \( \partial g(i)/\partial i < 1 \), we have \( \partial \pi/\partial i = 1 - \partial g(i)/\partial i > 0 \).\textsuperscript{14} Given this positive relationship, inflation and economic growth would also exhibit an inverted-U relationship. Recent empirical studies such as Bick (2010) and López-Villavicencio and Mignon (2011) provide evidence that supports an inverted-U relationship between inflation and economic growth.

3.2 Quantitative analysis

In this section, we calibrate the model to aggregate data of the US economy to provide a quantitative illustration on the effects of monetary policy on economic growth. The model features the following structural parameters \( \{\rho, \theta, \alpha, \beta, \kappa\} \) and policy variable \( i \). For the discount rate, we set \( \rho \) to a standard value of 0.05. For the labor share, we set \( \theta \) to a value of 0.59; see Elsby et al. (2013) who document that the labor share in the US has fallen to less than 0.60 recently. According to the Conference Board Total Economy Database, the average growth rates of total factor productivity (TFP) in the US is about 0.6% from 1990 to 2014. We calibrate the R&D cost parameter \( \alpha \) by targeting the scenario in which domestic innovation drives half of the TFP growth in the US (i.e., \( g = 0.3\% \)).\textsuperscript{15} For the cost of entry, we calibrate \( \beta \) by setting the time between arrivals of innovation \( 1/\phi \) to about 3 years as in Acemoglu and Akcigit (2012). For the Pareto distribution parameter, we follow Minniti et al. (2013) to consider \( \kappa = 0.21 \) as our benchmark, but we also explore another value \( \kappa = 0.16 \) that has interesting implications. Finally, we calibrate the value of \( i \) by targeting the average inflation rate \( \pi \) in the US, which is about 2.5% in the past two decades. The parameter and variable values are summarized in Table 1.

<table>
<thead>
<tr>
<th>Table 1: Calibration</th>
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<tr>
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<tr>
<td>Parameters</td>
</tr>
<tr>
<td>0.210</td>
</tr>
<tr>
<td>0.160</td>
</tr>
</tbody>
</table>

Under our benchmark parameter values, we find that economic growth is an inverted-U function of the nominal interest rate. In Figures 1 and 2, we plot the equilibrium growth rate \( g \) against the inflation rate \( \pi \), which is monotonically increasing in the nominal interest rate \( i \). Figure 1 presents our benchmark result and shows that the relationship between economic growth and inflation follows an inverted-U shape. Furthermore, the growth-maximizing inflation rate is about 2.9%, which is close to the empirical estimate in López-Villavicencio and Mignon (2011) who find a threshold inflation rate of 2.7% for industrialized countries.

\textsuperscript{14}Under our calibrated parameter values, steady-state inflation is indeed increasing in the nominal interest rate.

\textsuperscript{15}See Chu (2010) who finds that domestic R&D drives less than half of the TFP growth in the US.
In the empirical literature, studies sometime find a monotonically negative effect of inflation on economic growth; see for example, Guerrero (2006) and Vaona (2012). Indeed, we find that our model is flexible enough to deliver a negative relationship between inflation and economic growth under reasonable parameter values. When we decrease the value of $\kappa$ to 0.16 and recalibrate the rest of the parameters, we find that the relationship between economic growth and inflation becomes monotonically negative. In this case, the smaller value of $\kappa$ implies a smaller ratio of $\beta/\alpha$, such that the negative effect of inflation dominates the positive effect.

Figure 1: Inflation and economic growth
$(\kappa = 0.21)$

Figure 2: Inflation and economic growth
$(\kappa = 0.16)$
4 Conclusion

In this study, we have developed a monetary Schumpeterian growth model with endogenous entry of firms with heterogeneous quality improvements. Given this monetary growth-theoretic framework, we explore the effects of monetary policy on economic growth and find that inflation could have an inverted-U effect on economic growth. Furthermore, we calibrate the model to aggregate data of the US economy to provide a quantitative investigation. Under our benchmark parameter values, we find that the growth-maximizing inflation rate is about 2.9%, which is consistent with recent empirical estimates. However, given that we have a stylized model, the quantitative analysis should be viewed as an illustrative exercise.

References


Appendix A: The stationary quality threshold

In the symmetric equilibrium $v^e_t(\omega, j_\omega + 1) \equiv v^e_t$, the \textit{ex-ante} value of an innovation is given by

$$v^e_t = \int_1^\bar{\lambda}_t 0 \cdot f(\lambda) d\lambda + \int_{\bar{\lambda}_t}^\infty [v_t(\lambda) - \beta_t] f(\lambda) d\lambda = \int_{\bar{\lambda}_t}^\infty v_t(\lambda) f(\lambda) d\lambda - \Pr(\lambda \geq \bar{\lambda}_t) \beta_t. \tag{A1}$$

Substituting the \textit{ex-post} no arbitrage condition $r_t v_t(\lambda) = \Pi_t(\lambda) + \dot{\pi}_t(\lambda) - \Pr(\lambda \geq \bar{\lambda}_t) \phi_t v_t(\lambda)$ into (A1) yields

$$r_t v^e_t = \Pi^e_t + \int_{\bar{\lambda}_t}^\infty \dot{\pi}_t(\lambda) f(\lambda) d\lambda - \Pr(\lambda \geq \bar{\lambda}_t) \phi_t \int_{\bar{\lambda}_t}^\infty [v_t(\lambda) - \beta_t] f(\lambda) d\lambda - \left[ \Pr(\lambda \geq \bar{\lambda}_t) \phi_t + r_t \right] \Pr(\lambda \geq \bar{\lambda}_t) \beta_t. \tag{A2}$$

Combining (A1) and the R&D condition (11) and also using (14), we obtain

$$\int_{\bar{\lambda}_t}^\infty v_t(\lambda) f(\lambda) d\lambda = (1 + \hat{i}_t) \alpha_t + \bar{\lambda}_t^{-1/\kappa} \beta_t, \tag{A3}$$

where $i_t$ is chosen exogenously by the monetary authority. Differentiating (A3) with respect to $t$, we use the Leibniz integral rule to derive

$$\int_{\bar{\lambda}_t}^\infty \dot{v}_t(\lambda) f(\lambda) d\lambda - \int_{\bar{\lambda}_t}^\infty v_t(\lambda) f(\lambda) d\lambda = (1 + i) \ddot{\alpha}_t + \bar{\lambda}_t^{-1/\kappa} \ddot{\beta}_t - \frac{1}{\kappa} \bar{\lambda}_t^{-1+\kappa} \dddot{\lambda}_t \beta_t. \tag{A4}$$

We substitute (12) and the entry condition $v_t(\bar{\lambda}_t) = \beta_t$ into (A4) to obtain

$$\int_{\bar{\lambda}_t}^\infty \dot{v}_t(\lambda) f(\lambda) d\lambda = (1 + i) \ddot{\alpha}_t + \bar{\lambda}_t^{-1/\kappa} \ddot{\beta}_t. \tag{A5}$$

Substituting (A5) into (A2), the \textit{ex-ante} no-arbitrage condition for an innovation can be expressed as

$$r_t = \frac{\Pi^e_t + \left[ \dot{v}^e_t + \bar{\lambda}_t^{-1/\kappa} \ddot{\beta}_t \right] - \bar{\lambda}_t^{-1/\kappa} \phi_t \left[ v^e_t + \bar{\lambda}_t^{-1/\kappa} \beta_t \right]}{v^e_t + \bar{\lambda}_t^{-1/\kappa} \beta_t}, \tag{A6}$$

which uses (14) and the R&D condition (11) again. Moreover, we make use of the R&D condition (11), $\alpha_t = \alpha Q_t^{(1-\theta)/\theta}$ and $\beta_t = \beta Q_t^{(1-\theta)/\theta}$ to derive

$$\frac{\dot{v}^e_t + \bar{\lambda}_t^{-1/\kappa} \ddot{\beta}_t}{v^e_t + \bar{\lambda}_t^{-1/\kappa} \beta_t} = \left( \frac{1 - \theta}{\theta} \right) \frac{\dot{Q}_t}{Q_t}.$$

With this expression, (A6) becomes

$$r_t = \frac{\Pi^e_t}{v^e_t + \bar{\lambda}_t^{-1/\kappa} \beta_t} + \left( \frac{1 - \theta}{\theta} \right) \frac{\dot{Q}_t}{Q_t} - \bar{\lambda}_t^{-1/\kappa} \phi_t. \tag{A7}$$
Meanwhile, the *ex-post* no-arbitrage condition for the threshold quality \( \lambda = \bar{\lambda}_t \) can be written as

\[
rt = \frac{\Pi_t(\bar{\lambda}_t)}{v_t(\bar{\lambda}_t)} + \left( \frac{1 - \theta}{\theta} \right) \frac{\dot{Q}_t}{Q_t} - \bar{\lambda}_t^{-1/\kappa} \phi_t. \tag{A8}
\]

By the R&D condition (11), the entry condition \( v_t(\bar{\lambda}_t) = \beta_t \), \( \alpha_t = \alpha Q_t^{(1-\theta)/\theta} \) and \( \beta_t = \beta Q_t^{(1-\theta)/\theta} \), (A7) and (A8) imply

\[
\frac{\Pi_t}{(1 + i) \alpha + \bar{\lambda}_t^{-1/\kappa} \beta} = \frac{\Pi_t(\bar{\lambda}_t)}{\beta}. \tag{A9}
\]

Given (16) and (18), (A9) can be rearranged as

\[
(\bar{\lambda}_t - 1)\bar{\lambda}_t^{1/\kappa} = \frac{1}{(1 + i) \alpha} \frac{\beta}{1 + \kappa}. \tag{A10}
\]

Equation (A10) shows that \( \bar{\lambda}_t \) is always stationary.
Appendix B: Proofs

**Proof of Lemma 1.** Using (8), we can express the aggregate price index of intermediate goods as

\[ P_t = \exp \left[ \int_0^1 \ln \lambda_t(\omega) d\omega \right] = \exp \left[ \int_{\tilde{\lambda}}^\infty (\ln \lambda) \tilde{f}(\lambda) d\lambda \right] = \tilde{\lambda} e^\kappa, \]  

(B1)

where \( \tilde{f}(\lambda) \) is defined as

\[ \tilde{f}(\lambda) = \frac{f(\lambda)}{\int_{\lambda}^\infty f(\lambda) d\lambda} = \tilde{\lambda}^\frac{1}{\kappa} f(\lambda). \]  

(B2)

Here we introduce a modified density function \( \tilde{f}(\lambda) \) in summing on \([\tilde{\lambda}, 1]\) because the distribution of \( \lambda \) in equilibrium is not on the original domain \([1, \infty)\), but instead on \([\tilde{\lambda}, \infty)\), due to endogenous entry. Note that \( \int_{\tilde{\lambda}}^\infty \tilde{f}(\lambda) d\lambda = 1 \). By (7) and (B1), we obtain

\[ K_t = h(1 - 1/\tilde{\lambda}) Q_t = (\tilde{\lambda} e^\kappa) \]

(B3)

noting \( L_t = 1 \). Recall that final goods are used for consumption, production of intermediate goods, R&D and entry. Consumption is given by \( c_t \). By (6) and (8), the amount of final goods used for the production of intermediate goods is

\[ X^m_t = \int_0^1 y_t(\omega, j_\omega) d\omega = \int_0^1 \frac{(1 - \theta) Y_t}{\lambda_t(\omega)} d\omega = (1 - \theta) Y_t \int_{\tilde{\lambda}}^\infty \frac{1}{\tilde{\lambda}} \tilde{f}(\lambda) d\lambda = \frac{(1 - \theta) Y_t}{(1 + \kappa) \tilde{\lambda}}. \]  

(B4)

Final goods for innovation and entry are given by

\[ X^r_t = \int_0^1 R_t(\omega) d\omega = \alpha_t \phi_t \text{ and } X^e_t = \int_{\omega \in \Omega_t} \beta_t d\omega = \beta_t \tilde{\lambda}^{-1/\kappa} \phi_t, \]  

(B5)

where \( \Omega_t \) is the set of industries in which innovations take place and are implemented at date \( t \). Finally, we substitute (B3), (B4) and (B5) into the market-clearing condition \( Y_t = c_t + X^m_t + X^r_t + X^e_t \) to derive

\[ \phi_t = \frac{1}{\alpha + \beta \tilde{\lambda}^{-1/\kappa}} \left[ \left( 1 - \theta \right) \frac{Y_t}{\tilde{\lambda} e^\kappa} \right]^{(1-\theta)/\theta} \left( 1 - \frac{1 - \theta}{(1 + \kappa) \tilde{\lambda}} \right) - C_t, \]  

(B6)

where \( C_t \equiv c_t/Q_t^{(1-\theta)/\theta} \) is a transformed variable that is stationary. We substitute (16) and the R&D condition (11) into (A7) to derive

\[ r_t = \frac{(1 - \theta) Y_t}{(1 + i) \alpha_t + \tilde{\lambda}^{-1/\kappa} \beta_t} \left[ \frac{\lambda - 1/(1 + \kappa)}{\lambda 1/\kappa} \right] - \tilde{\lambda}^{-1/\kappa} \phi_t + \frac{1 - \theta}{\theta} \dot{Q}_t. \]  

(B7)

\[ ^{16} \text{We achieve this by applying integration by parts to} \]

\[ \int_{\tilde{\lambda}}^\infty (\ln \lambda) \tilde{f}(\lambda) d\lambda = \tilde{\lambda}^{1/\kappa} \int_{\tilde{\lambda}}^\infty (\ln \lambda) \left( \frac{d}{d\lambda} \frac{\tilde{\lambda}^{1-1/\kappa}}{1 + \kappa} \right) d\lambda. \]
Finally, substituting (B3) and (B7) into (3) yields
\[
\frac{\dot{C}_t}{C_t} = r_t - \rho - \frac{1 - \theta}{\theta} \frac{\dot{Q}_t}{Q_t} = \frac{(1 - \theta)^{1/\theta}}{[(1 + i)\alpha + \bar{\lambda}^{-1/\kappa}\beta]e^{\kappa(1-\theta)/\theta}} \left[ \frac{\bar{\lambda} - 1/(1 + \kappa)}{\bar{\lambda}^{(1/\theta)+(1/\kappa)}} \right] - \bar{\lambda}^{-1/\kappa} \phi_t - \rho, \quad (B8)
\]
noting the definitions \(C_t \equiv c_t/Q_t^{(1-\theta)/\theta}\) and \(\alpha_t \equiv \alpha Q_t^{(1-\theta)/\theta}\). Substituting (B6) into (B8), we have an one-dimensional differential equation in \(C_t\). Given that \(\phi_t\) decreases with \(C_t\) in (B6), the right-hand side of (B8) is increasing in \(C_t\), so the dynamics of \(C_t\) is characterized by saddle-point stability, such that \(C_t\) must jump to its interior steady-state value. Given a stationary value of \(C_t\), (B6) implies that \(\phi_t\) is also stationary. ■

**Proof of Proposition 1.** In this proof, we first show that the relationship between \(i\) and \(\bar{\lambda}^{-1/\kappa} \phi\) is either inverted-U shaped or negative. Combining (25) and (26), we have
\[
\bar{\lambda}^{-1/\kappa} \phi = \frac{(1 - \theta)^{1/\theta}}{\beta e^{\kappa(1-\theta)/\theta}} \left( \frac{\bar{\lambda} - 1}{\bar{\lambda}^{1/\theta}} \right) - \rho. \quad (B9)
\]
By differentiating the right-hand side of (B9) with respect to \(\bar{\lambda}\), we can easily show that \(d(\bar{\lambda}^{-1/\kappa} \phi)/d\bar{\lambda} > (\ <) 0\) if \(\bar{\lambda} < (\ >) 1/(1 - \theta)\), implying an inverted-U relationship between \(\bar{\lambda}\) and \(\bar{\lambda}^{-1/\kappa} \phi\). In identifying the relationship with respect to \(i\), we naturally focus on a non-trivial range of \(\bar{\lambda}\), i.e., \((\bar{\Delta}, \bar{\lambda})\), where \(\bar{\lambda}^{-1/\kappa} \phi > 0\) holds. Given that \(\bar{\lambda}\) monotonically decreases with \(i\) (Lemma 3), \(i \geq 0\) provides another natural upper bound of \(\bar{\lambda}\), say \(\lambda_i\), which is defined by
\[
(\lambda_i - 1) \lambda_i^{1/\kappa} = \frac{\beta}{\alpha} \frac{\kappa}{1 + \kappa}. \quad (B10)
\]
When \(\lambda_i\) is large enough (exceeding \(1/(1 - \theta)\)), the relationship between \(i\) and \(\bar{\lambda}^{-1/\kappa} \phi\) is inverted-U-shaped on the non-trivial range \((\bar{\Delta}, \lambda_i)\); see Figure 3a. When \(\lambda_i\) is small enough (falling below \(1/(1 - \theta)\)), \(\bar{\lambda}^{-1/\kappa} \phi\) is monotonically decreasing in \(i\) on \((\bar{\Delta}, \lambda_i)\); see Figure 3b. Note that, by (B10), \(\lambda_i\) increases with \(\beta\) and, by (B9), \(\bar{\lambda}\) decreases with \(\beta\). This implies that for a larger (smaller) entry cost \(\beta\), accompanied by a larger (smaller) \(\lambda_i\), the relationship between \(i\) and \(\bar{\lambda}^{-1/\kappa} \phi\) becomes inverted-U (negative).

\[\text{17} \] Although \(\bar{\lambda}\) is an endogenous variable, it is stationary and a function of parameters as shown in (A10).

\[\text{18} \] The formal definition of \((\bar{\Delta}, \bar{\lambda})\) is given by incorporating \(\bar{\lambda}^{-1/\kappa} \phi = 0\) into (B9): \(\bar{\Delta}\) and \(\bar{\lambda}\) are equal to \(x\) such that \((x - 1)/x^{1/\theta} = \rho_\beta e^{\kappa(1-\theta)/\theta}/(1 - \theta)^{1/\theta}\). This has the two solutions such as \(x = \bar{\Delta}\) and \(\bar{\lambda}\) if and only if \(\rho_\beta < \theta (1 - \theta)^{2/\theta} e^{\kappa(1-\theta)/\theta}\). Otherwise, \(\bar{\lambda}^{-1/\kappa} \phi\) cannot be positive.
In the rest of this proof, we characterize the relationship between $i$ and $g$. For $\tilde{\lambda} < 1/(1 - \theta)$, it holds $d(\tilde{\lambda}^{-1/\kappa})/d\tilde{\lambda} > 0$ as shown above. Given that $(\ln \tilde{\lambda} + \kappa)$ is also increasing in $\tilde{\lambda}$, this implies $dg/d\tilde{\lambda} > 0$ for $\tilde{\lambda} < 1/(1 - \theta)$, by noting (21). To see the case where $1/(1 - \theta) < \tilde{\lambda}$, using (21) and (B9), we can obtain

$$\frac{dg}{d\tilde{\lambda}} = \frac{1 - \theta}{\theta \tilde{\lambda}^{1+1/\theta}} \left\{ \frac{(1 - \theta)^{1/\theta}}{\beta e^{\kappa(1-\theta)/\theta}} \left( \tilde{\lambda} - 1 \right) - \rho \tilde{\lambda}^{1/\theta} \right\} - \frac{(\ln \tilde{\lambda} + \kappa)}{\beta e^{\kappa(1-\theta)/\theta}} \frac{(1 - \theta)^{1/\theta}}{\theta} \frac{1 - \theta}{\left( \tilde{\lambda} - \frac{1}{1 - \theta} \right)}$$

Note the following properties: (a) $\zeta(1/(1 - \theta)) > 0$ and $\zeta(1/(1 - \theta)) = 0$; (b) $\zeta(\tilde{\lambda})$ is an unimodal function\(^{19}\) and $\xi(\tilde{\lambda})$ is a strictly increasing function; (c) $\zeta(\lambda) = \zeta(\tilde{\lambda}) = 0$; and (d) $\zeta(\tilde{\lambda})$ is strictly concave and $\xi(\tilde{\lambda})$ is strictly convex.

Using these properties, we can graphically show that $\xi(\tilde{\lambda})$ intersects $\zeta(\tilde{\lambda})$ from below only once at some point in $\tilde{\lambda} \in (1/(1 - \theta), \bar{\lambda})$, below (above) which $dg/d\tilde{\lambda} > (<) 0$. This implies an inverted-U relation between $\tilde{\lambda}$ and $g$ on $(\lambda, \bar{\lambda})$. The rest of Proposition 1 straightforwardly follows, noting that $\lambda_i$ is increasing in $\beta$. ■

\(^{19}\)It is useful to note that $\zeta(\tilde{\lambda})$ is upward sloping at $\tilde{\lambda} = 1/(1 - \theta)$.  

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Appendix C (not for publication)

**Equation (20):** Recall that the equilibrium distribution of $\lambda$ is given by $\tilde{f}(\lambda)$, which is defined by (B2) in Appendix B. Then we calculate

$$\int_{0}^{1} (\ln \lambda(\omega)) \, d\omega = \int_{\tilde{\lambda}}^{\infty} (\ln \lambda) \tilde{f}(\lambda) \, d\lambda = \frac{\lambda^{1/\kappa}}{\kappa} \int_{\tilde{\lambda}}^{\infty} (\ln \lambda) \lambda^{-\frac{1+\kappa}{\kappa}} \, d\lambda,$$

where the second equality uses (12) and (B2). Given that

$$\lambda^{-\frac{1+\kappa}{\kappa}} = \frac{d}{d\lambda} \frac{\lambda^{1 - \frac{1+\kappa}{\kappa}}}{1 - \frac{1+\kappa}{\kappa}},$$

we have

$$\int_{0}^{1} (\ln \lambda(\omega)) \, d\omega = \frac{\lambda^{1/\kappa}}{\kappa} \int_{\tilde{\lambda}}^{\infty} (\ln \lambda) \left[ \frac{d}{d\lambda} \left( \frac{\lambda^{1 - \frac{1+\kappa}{\kappa}}}{1 - \frac{1+\kappa}{\kappa}} \right) \right] \, d\lambda.$$ 

Applying integration by parts, we calculate

$$\frac{\lambda^{1/\kappa}}{\kappa} \int_{\tilde{\lambda}}^{\infty} (\ln \lambda) \left[ \frac{d}{d\lambda} \left( \frac{\lambda^{1 - \frac{1+\kappa}{\kappa}}}{1 - \frac{1+\kappa}{\kappa}} \right) \right] \, d\lambda = \frac{\lambda^{1/\kappa}}{\kappa} \left\{ \left| \frac{\lambda^{1 - \frac{1+\kappa}{\kappa}}}{1 - \frac{1+\kappa}{\kappa}} \ln \lambda \right|_{\tilde{\lambda}}^{\infty} - \int_{\tilde{\lambda}}^{\infty} \lambda^{-\frac{1+\kappa}{\kappa}} \, d\lambda \right\}.$$

From

$$\left. \frac{\lambda^{1 - \frac{1+\kappa}{\kappa}}}{1 - \frac{1+\kappa}{\kappa}} \ln \lambda \right|_{\tilde{\lambda}}^{\infty} = \kappa \tilde{\lambda}^{-\frac{1}{\kappa}} \ln \tilde{\lambda}$$

and

$$\int_{\tilde{\lambda}}^{\infty} \lambda^{-\frac{1+\kappa}{\kappa}} \, d\lambda = -\kappa^{2} \tilde{\lambda}^{-\frac{1}{\kappa}},$$

we have $\int_{0}^{1} (\ln \lambda(\omega)) \, d\omega = \ln \tilde{\lambda} + \kappa.$