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# **Factors influencing CO<sub>2</sub> Emission in China: A Nonlinear Autoregressive Distributed Lags Investigation**

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## **Abstract**

This paper investigates the environmental impact of economic growth, energy consumption, financial development and globalization in China over the period 1970Q1-2015Q4. In particular we consider four dimensions of globalization namely economic, social, political and overall globalization. The Nonlinear Autoregressive Distributed Lags (NARDL) model has been employed to capture the potential asymmetric impact of the determinants of dioxide carbon emissions in China. Interestingly, findings show that: (1) In the short-run: economic growth and financial development have a significant symmetric impact on CO<sub>2</sub> emissions. Energy consumption has a nonlinear and asymmetric influence on CO<sub>2</sub> emissions. However, economic globalization does not impact CO<sub>2</sub> emissions. (2) In the long-run: economic growth, financial development and economic globalization exhibit an asymmetric influence on carbon emissions in model including the economic dimension of globalization. Economic growth has a positive symmetric impact on CO<sub>2</sub> emissions in model including social globalization, however, it does not influence CO<sub>2</sub> emissions in case of political or overall globalization. In addition, energy consumption is positively linked to CO<sub>2</sub> emissions. Moreover, financial development does not influence carbon emission in models including respectively social, political and overall globalization. Social and overall globalization have a significant influence on CO<sub>2</sub> emissions. The results of this paper are important for policies that would promote sustainable development and environment protection.

**Keywords:** Energy, Growth, Financial Development, Globalization, CO<sub>2</sub> Emissions

## 1. Introduction

One of the most severe problems of the modern world is the climate change and its important negative consequences on environment. The human activity, particularly carbon dioxide emissions, has been considered among the main factors contributing to the changing of climate in the last decades (IPCC, 2007). It is thus crucial to determine factors that foster carbon dioxide emissions. In fact, heavy emission of carbon dioxide is nowadays harmful for environment with a high level of air pollution. Environmental policymakers should give a closer look at the latter issue because of the greenhouse effect and global warming. Renewable energy resources constantly renew themselves and have less negative effect on environment than fossil energy technologies, for instance. Renewable energy resources are largely preferred to their nonrenewable counterparts because they allow reducing carbon dioxide emissions (Dogan and Seker, 2016). Consequently, they enable to protect the environment. Recall that renewable energy is considered as a viable option to enhance access to energy and mitigate climate change (Moomaw et al. 2011). Kim and Park (2016) examine the relationship between financial development and renewable energy deployment and show that financial development is favorable for renewable sectors basically for renewable sectors characterized by high dependence on external finance. They conclude that financial development leads to a reduction in CO<sub>2</sub> emissions.

This paper highlights the factors impacting the CO<sub>2</sub> emissions in China. According to Lin et al. (2016), China's impressive economic performance in the past decades has resulted in increased carbon intensity. Over the past few years, China's economy has grown at an average of over 7%, exceeding that of the United States and the European Union combined. However, this impressive economic performance has led to an increase in carbon (CO<sub>2</sub>) emissions. The United States Energy Information Administration mentioned that China's primary energy consumption increased from 17.29 Quad BTU in 1980 to 103.72 Quad BTU in 2011. Similarly, her electricity net consumption increased from 261.49 billion kilowatthours in 1980 to 4207.70 billion kilowatthours in 2011, an increase of over 1500%. With the increase in total primary energy and coal-dominated electricity consumption, carbon emission also increased significantly. Carbon emission associated with electricity production and consumption in China is high because coal is the dominant fuel for electricity production in the country. As at 2012, China's energy consumption-related CO<sub>2</sub> emission stands at 8547.74 million metric tons compared to 1448.46 million metric tons in 1980, which makes it the largest CO<sub>2</sub> emitter in the world (Lin et al., 2016). Consequently, analyzing the drivers of this prominent increase of CO<sub>2</sub> emissions in China is crucial. Indeed, we study the impact of the gross domestic product growth, energy

consumption, financial development and globalization on carbon dioxide emission in China by using the Nonlinear Autoregressive Distributed Lags (NARDL) model.

This paper contributes to existing literature by four folds: First, we apply the multivariate nonlinear ARDL (NARDL) of Shin et al. (2014) to test for the possible nonlinear and asymmetric cointegration between carbon emissions and its determinants. The NARDL model is convenient for our setting as it jointly incorporates the long- and short-run asymmetric relationships among variables. Asymmetric and nonlinear linkages between economic variables could occur due to the complexity of economic systems and mechanisms leading to carbon emissions and its determinants. For instance, economic time series may be influenced by structural reforms, policy shifts, real and financial shocks, and regional and global imbalances. Contrary to economic and financial variables that depend on general macroeconomic instruments such as business cycle conditions, monetary policy adjustments and product market regulations, CO<sub>2</sub> emissions is rather more responsive to specific domestic and global energy market conditions.

Second, previous studies employed annual data and surprisingly none of them used higher frequency data such as quarterly or monthly. In this paper, we use quarterly data as the previous literature has pointed out the influence of data-frequency on empirical results (Narayan and Sharma, 2015).

Third, our study differs from previous studies by examining the effects of Gross Domestic Product (GDP), Energy Consumption (EC), Financial Development (FD) and Globalization (G) on carbon dioxide (CO<sub>2</sub>) emissions in China from 1970Q1 to 2015Q4. Although there have been several studies on CO<sub>2</sub> emissions in the field of energy economics, majority of these studies focused on emerging countries<sup>1</sup>. In contrast, this paper analyses the determinants of CO<sub>2</sub> emissions in the largest emerging economy “China” and the largest CO<sub>2</sub> emitter in the world. In addition, most of the empirical studies analyze the drivers of carbon emissions using panel data techniques, and do not adequately investigate countries individually. In contrast, we suspect that individual country characteristics such as GDP, energy consumption, financial development and globalization are important drivers of CO<sub>2</sub> emissions. The individual country characteristics and path ways are necessary given the significant differences in income level, financial development, energy consumption level and structure, technology advancement,

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<sup>1</sup>See for instance, Tamazian et al. (2009), Jalil and Feridun (2011), Ozturk and Acaravi (2013)

energy market structure, mitigation and adaptation capabilities, development policy goals across countries, and globalization (Lin et al. 2016).

Finally, little attention was given to financial variables as a key determinant of CO<sub>2</sub> emissions. Previous studies explained CO<sub>2</sub> emissions mostly by the gross domestic product and energy variables such as renewable and non-renewable energy (Dogan and Seker, 2016) and energy consumption (Alam et al. 2011). Recent studies enlarged the set of potential determinants of CO<sub>2</sub> emissions considering new variables that account for financial development such as financial capital market, quality of firms governance and financial development as well as variables accounting for globalization such as trade openness and foreign direct investment (Talukdar and Meisner 2001, Classens and Feijen 2007). In this paper, in addition to the GDP variable and energy consumption, we investigate the impact of financial development and globalization on CO<sub>2</sub> emissions in China. Particularly, we consider four dimensions of globalization namely economic globalization (EG), social globalization (SG), political globalization (PG) and overall globalization (OG). Disentangling the individual effect of each type of globalization on CO<sub>2</sub> emissions in China is important because it will help authorities taking right decision in order to reduce CO<sub>2</sub> emissions domestically.

Therefore, our study empirically investigates the drivers of carbon dioxide emissions in the largest emerging economy “China”. The results of this study have important implications for the implementation of future policies on financial policies in combination with macroeconomic policies in order to reduce CO<sub>2</sub> emissions. The importance of this paper lies with the fact that it takes into account the prominence of the effects of changes in economic growth, energy consumption, financial development and level of globalization on CO<sub>2</sub> emissions. In fact, the major findings are the following: (1) In the short-run, gross domestic product, energy consumption, financial development and globalization affect significantly CO<sub>2</sub> emissions; (2) In the long-run, energy consumption significantly deteriorates environment. However, in the short-run, more energy consumption increases CO<sub>2</sub> emissions while less energy consumption preserves environment; (3) In the long-run, economic globalization influences carbon dioxide emissions in asymmetric and nonlinear manners while social and overall globalizations reduce CO<sub>2</sub> emissions in a linear. Thus, higher level of economic globalization raises CO<sub>2</sub> emissions while lower level reduces CO<sub>2</sub> emissions. However, political globalization does not impact CO<sub>2</sub> emissions; (4) In the short-run, economic globalization does not impact CO<sub>2</sub> emissions. However, social, political and overall globalizations reduce CO<sub>2</sub> emissions and; (5) Financial development asymmetrically impacts dioxide carbon emissions only when it coupled with

economic globalization. It does not exert any influence on dioxide carbon emissions if coupled with social, political or overall globalization.

The paper is structured as follows. Section-2 discusses literature dealing with the factors influencing carbon dioxide emissions. The following section introduces the methodology and data used while the next section discusses the empirical findings. Finally, the last section concludes and provides policy implications.

## **2. Literature review**

### **2.1 Financial Development and CO<sub>2</sub> Emissions**

Financial development may pass-through to energy consumption through two possible channels. A first stream of literature argues that financial development rises energy consumption because it fosters economic growth, which in turn, requires more demand for energy and may influence CO<sub>2</sub> emissions (Sadorsky 2010). Sadorsky (2011) decomposed the global effect of financial development on energy consumption into three different effects: direct effect through the purchase of energy-consuming goods, business effect through energy demand and wealth effect through an increase of energy demand resulting from higher economic confidence. This shows that all these channels affect environmental quality by influencing carbon emissions. A second stream of literature believes that financial development reduces energy consumption which in resulting, decline carbon emissions. Indeed, financial development leads to using less-energy consuming technologies which results in a fall of energy consumption and CO<sub>2</sub> emissions as well. The latter transmission channel is known as the technological effect (Tamazian et al. 2009, Shahbaz et al. 2013, Mahalik and Mallick 2014).

There has been substantial research attention on CO<sub>2</sub> emissions in recent years since, CO<sub>2</sub> emissions is considered as one of the major air pollutant and hence environment degradation (Astöm et al. 2013, Henneman et al. 2016). This is the main reason for which we focus on carbon emissions and its long- and short-run determinants. A number of studies have been conducted showing the relationship between gross domestic product and CO<sub>2</sub> emissions. Financial development is now-a-days considered as a key factor driving the level of CO<sub>2</sub> emissions. Various recent studies investigate the impact of financial development on CO<sub>2</sub> emissions. For instance, Katircioglu and Taspinar (2017) examined the moderating role of financial development in environmental quality by using comprehensive index of financial development along with economic growth and energy consumption for Turkish economy<sup>2</sup>.

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<sup>2</sup> They used domestic credit by banking sector as share of GDP, domestic credit to private sector as share of GDP, broad money supply as share of GDP, liquid liabilities as share of GDP and the ratio of commercial bank assets to total assets to generate financial development index by using Principal Component Analysis (PCA).

Their empirical analysis reveals that financial development reduces stimulates carbon emissions but energy consumption stimulates it. Furthermore, financial development stimulates economic growth which in resulting, adds in CO<sub>2</sub> emissions. In case of South Asia<sup>3</sup>, Nasreen et al. (2017) noted that stable financial sector improves environmental quality by decreasing CO<sub>2</sub> emissions but energy consumption, economic growth and population density are major contributors of environmental degradation.

Abbasi and Riaz (2016) resort to the ARDL model to investigate the long-run relationship between CO<sub>2</sub> emissions and a set of economic and financial variables including financial development. They also employ an augmented version of the VAR model to capture the short-run linkage between carbon emissions and its determinants by splitting the sample period into two sub-samples in order to isolate the period 1988-2011 characterized by greater degree of liberalization and financial sector development. Their empirical results indicate that financial variables played a significant role in mitigating CO<sub>2</sub> emissions during the period 1988-2011 only. Additionally, the level of carbon dioxide emissions attributable to financial development is relatively small compared to the level of emission raising due to rising of per capita income. Economic growth increases energy consumption that, in turn, increases carbon dioxide emissions. The latter findings lead environmental deciders to adopt different mitigation policies able to attenuate the level of rising of carbon emission. Dogan and Seker (2016) employ the environmental Kuznets curve framework to study the influence of real output, renewable and non-renewable energy, trade and financial development on CO<sub>2</sub> emissions in a panel of 23 countries. They find that financial development and trade openness reduce carbon emission. For Portuguese economy, Shahbaz et al. (2016a) applied multivariate framework of carbon emissions function by financial development is major determinant of environmental degradation. Their results show that economic growth and energy intensity increase CO<sub>2</sub> emissions but financial development improves environmental quality by lowering CO<sub>2</sub>emissions. Javed and Sharif (2016) investigated the determinants of carbon emissions using Pakistani data for the period of 1972-2103. They reported that energy consumption and financial development have positive and significant impact on CO<sub>2</sub> emissions but trade openness declines carbon emissions. In UAE, Charfeddine and Khediri (2016) applied regime-switching cointegration approach for the investigation of carbon emissions function. Their results indicate the presence of EKC association between financial development and CO<sub>2</sub> emission. They also noted that electricity consumption, trade openness and urbanization

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<sup>3</sup> Pakistan, India, Bangladesh and Sri Lanka

improve environmental quality by lowering CO<sub>2</sub> emissions. Farhani and Ozturk (2016) applied bounds testing approach for examining the relationship between carbon emissions for Tunisian economy.

For Indian economic, Sehwat et al. (2015) reinvestigated the association between financial development and carbon emissions by applying ARDL cointegration approach. They show that financial development degrades environmental quality by adding CO<sub>2</sub> emissions<sup>4</sup>. Salahuddin et al. (2015) employ panel data econometric methodology to investigate the relationship between economic growth, electricity consumption, financial development and CO<sub>2</sub> emissions. Their findings show absence of short-run relationship between the variables. In addition, electricity consumption and economic growth have a positive relationship with CO<sub>2</sub> emissions while the relationship between financial development and CO<sub>2</sub> emissions is found to be negative. Their results also indicate that economic growth and electricity consumption stimulate dioxide emission while financial development reduces it. Boutabba (2014) uses the autoregressive distributed lag model (ARDL) to detect the long-run impact of economic growth, financial development, energy consumption and trade openness on carbon emissions in India. The empirical findings show evidence of long-run causality running from financial development to CO<sub>2</sub> emissions. The long-run impact of financial development on CO<sub>2</sub> emissions is positive indicating that higher level of financial development contributes more to environment degradation. Energy consumption is also a major determinant of CO<sub>2</sub> emissions in the long-run in India. Using the bounds test within the ARDL framework, Shahbaz et al. (2013) show evidence of long-run influence of financial development on carbon emissions in Malaysia. Indeed, financial development is found to reduce CO<sub>2</sub> emissions in Malaysia preserving then the environment from being degraded. The previous approach is employed by Shahbaz et al. (2013a) to investigate the effects of financial development, economic growth, coal consumption and trade openness on CO<sub>2</sub> emissions in South Africa. Their results indicate that economic growth raises dioxide emissions while financial development and trade openness reduce it. Also, coal consumption contributes to deteriorate the environment by increasing CO<sub>2</sub> emissions. Shahbaz et al. (2013b) examine the linkages between economic growth, energy consumption, financial development, trade openness and CO<sub>2</sub> emissions in Indonesia over the period 1975Q1-2011Q4. Their results mitigate those of previous studies. They show that economic growth and energy consumption increase CO<sub>2</sub> emissions while financial development

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<sup>4</sup> Their analysis also indicates that energy consumption, trade openness and urbanization are contributing factors to environmental degradation.



and trade openness reduces it. Similarly, Ozturk and Acaravci (2013) investigated carbon emissions function for Turkish economy. Their analysis indicated the insignificant effect of financial development on CO<sub>2</sub> emissions.

In case of China, Xiong and Qi (2017) applied the STIRPAT augmented model to analyze the relationship between financial development and carbon emissions using 30 provinces panel data for the period of 1997-2011. They noted that financial development improve environmental quality by reducing carbon emissions. Zhang (2011) investigates the impact of financial development on carbon emissions in China using a set of econometric techniques namely cointegration, Granger causality test and variance decomposition. Findings show that China's financial development acts as an important driver for increase of carbon emissions. Moreover, among the financial development indicators foreign direct investments influence the least carbon emissions in China. Jalil and Feridun (2011) reinvestigated the association between financial development and carbon emissions by adding energy consumption, trade openness and income as additional determinants in carbon emissions function. They noted that financial development declines CO<sub>2</sub> emissions but income, energy consumption and trade openness increase it. Once again, the results of these studies may be ambiguous due to use of different methods for different countries. Furthermore, the use linear econometric methodologies that impose similar size effect of positive and negative shocks of exogenous variables on the dependent variable. Ignorance of asymmetries occurred in time series data, may make previous empirical evidence meaningless. We found one study by Shahbaz et al. (2016b) filling this gap by using the recently developed nonlinear Autoregressive distributed lag model (NARDL) by considering Pakistani data. Their results show evidence of strong asymmetric among financial development, energy consumption, economic growth and CO<sub>2</sub> emission. They conclude that positive shocks of economic growth rise CO<sub>2</sub> emission revealing that economic growth and financial development contribute to the degradation of the environment in Pakistan.

### **2.3 Globalization and CO<sub>2</sub> Emissions**

However, globalization is regarded as economic tool of improving economic growth and welfare by removing restrictions on trade and investment inflows. Given that globalization also affects CO<sub>2</sub> emissions and economic activity via various channels. Having a country engaged in trade and investment activities, a higher amount of energy usage is required in production and consumption activities, which eventually releases more carbon dioxide to the environment. Globalization also improves environmental quality by reducing CO<sub>2</sub> emissions through technology and knowledge transfers (Shahbaz et al. 2015a). For example, the use of cleaner technology by firms requires lesser energy and also helps them to have higher economic growth

without hampering environmental quality. In addition, these trade, investment and technology channels have many implications for the environment and economic activity. First, if the country continues to release greater amounts of CO<sub>2</sub> emissions due to rising production and consumption activity, it surely hampers environmental quality via increasing carbon emissions. Second, environmental degradation further depends on the nature of technology used in production and consumption activities. In this case, if firms use dirty energy-intensive production techniques, no doubt, economic growth is increased but it also degrades environmental quality along with making climate change and global warming to the extent of worse situation. Third, if an economy continues to be a net importer of energy, then energy imports can again aggravate its balance of payments, which in turn further affects its economic growth process in the long-run.

Another stream of literature paid more attention to the influence of globalization on CO<sub>2</sub> emissions. For instance, Shahbaz et al. (2017) determined the factors contributing to CO<sub>2</sub> emissions for Chinese economy. Considering the role of globalization in carbon emissions function, their empirical analysis indicated that globalization (economic, social and political) contributes to environmental quality in the presence of environmental Kuznets curve. Jin (2015) considers a multi-region global model to simulate the effects of different dimensions of globalization on carbon emissions in China. The Simulation results show that traditional economic globalization policies such as trade and foreign direct investment liberalization lead dioxide emission to rise through boosting production output. However, technology globalization – such as removal of technology transfer barriers - reduces energy emissions in China. Consequently, authorities in China should adopt the right globalization-based policies in order the effectively preserve good domestic environment conditions. Fernandez-Amador et al. (2016) assess CO<sub>2</sub> emissions embodied in international trade and found that international trade increases global CO<sub>2</sub> emissions following the trade openness rise of developing countries. Shahbaz et al. (2016c) incorporate globalization in the Environmental Kuznets curve in 19 African countries to determine the link between globalization and CO<sub>2</sub> emission in African countries. Their findings show that globalization reduces CO<sub>2</sub> emissions 8 out of 19 countries while it increases it in 5 out of 19 countries. Ertugrul et al. (2016) analyze the relationship between carbon dioxide emissions and trade openness in top ten emitter countries. Employing three different econometric techniques namely Andrews-Zivot unit root test with structural breaks, the bounds testing for cointegration in presence of structural breaks and VECM Granger causality test, the authors find evidence that real income, energy consumption and trade openness are the main drivers of carbon emissions in the sample countries. They also

recommend that emerging countries should continue increase their real GDP as they found an inverted U-shaped relationship between economic growth and pollution in these countries. The authors conclude that trade openness coupled with economic growth contribute to environmental degradation in countries including China, Indonesia and Thailand. Doythch and Uctum (2016) analyze the impact of globalization, measured by capital inflows, on the environment using a large sample of countries over a long time span from 1984 to 2011. They argue that results are different according the industry receiving the foreign capital. Indeed, capital flows into manufacturing increase pollution while those into services preserve the environment from pollution. In addition, a different impact of capital flows on the environment is detected according the income level of countries in the sample. In fact, capital flows into low- and middle-income countries lead to the degradation of their environment while capital flowing into high-income countries benefits the environment. For Indian economy, Shahbaz et al. (2015) used index of globalization (economic globalization, social globalization and political globalization) developed by Dreher (2006) to examine the linkages between globalization and CO<sub>2</sub> emissions. Their results indicate that globalization detrimental effect on environmental quality. Ling et al. (2015) decomposed the trade-environment nexus using Malaysian data for the period of 1970-2011. They found that trade affects environmental quality via income, scale, composition and comparative effects. Their results show that income and comparative effects have positive effect on environmental degradation but scale and composition effects improve environmental quality. In case of Turkey, Shahbaz et al. (2013d) investigated the validation of environmental Kuznets curve in the presence of globalization. Their results show that globalization improves the environmental quality in the presence of environmental Kuznets curve.

Our previous review of the literature on how financial development and globalization clearly influence CO<sub>2</sub> emissions indicates mixed findings on the issue. So far, no definite answer to this question was formulated. As for the methods, we have found none of studies in case of China considered the role of asymmetries while investigating the association between carbon emissions and its determinant. We contribute to the existing literature by considering a new measure of the financial development index and four dimensions of globalization. We also employ a nonlinear ARDL model that more flexible than linear models in that, first, it enables to disentangle effects of positive and negative shocks of financial development and globalization on CO<sub>2</sub> emissions and second, it allows simultaneously variables having different orders of integration i.e. I(0) and I(1) variables.

### **3. Methodology and Data**

Few studies in existing energy literature investigated the contributory factors to CO<sub>2</sub> emissions by applying different methods which provide conflicting empirical results for case of China. Following Jalil and Feridun (2011), Zhang (2011) and, Xiong and Qi (2017) and Shahbaz et al. (2017), we apply augmented carbon emissions function by incorporating financial development and globalization as additional determinants of CO<sub>2</sub> emissions for Chinese economy. Therefore findings of existing studies in literature are ambiguous due to ignorance of relevant variables such as financial development and globalization. It is indicated by Sardosky (2011) that financial development affects carbon emissions by business, consumer and technological effects. Similarly, globalization affects CO<sub>2</sub> emissions by scale, technological and composition effects (Shahbaz et al. 2015). This shows the importance of financial development and globalization to be incorporated in carbon emissions function to determine the major determinants of CO<sub>2</sub> emissions.

In doing so, we employ the NARDL model developed by Shin et al. (2014) to assess the asymmetric influence of economic growth, financial development and globalization on carbon emissions in China. A general version of the NARDL model that includes both long- and short-run asymmetries is specified as follows:

$$\begin{aligned} \Delta CO2_t = & \mu + \rho_{CO2} CO2_{t-1} + \rho_{GDP}^+ GDP_{t-1}^+ + \rho_{GDP}^- GDP_{t-1}^- + \rho_{EC}^+ EC_{t-1}^+ + \rho_{EC}^- EC_{t-1}^- + \\ & \rho_{FD}^+ FD_{t-1}^+ + \rho_{FD}^- FD_{t-1}^- + \rho_{GLOB}^+ GLOB_{t-1}^+ + \rho_{GLOB}^- GLOB_{t-1}^- + \sum_{i=1}^{p-1} \alpha_i \Delta CO2_{t-i} + \\ & \sum_{i=1}^{q-1} (\beta_i^+ \Delta GDP_{t-i}^+ + \beta_i^- \Delta GDP_{t-i}^-) + \sum_{i=1}^{q-1} (\delta_i^+ \Delta EC_{t-i}^+ + \delta_i^- \Delta EC_{t-i}^-) + \sum_{i=1}^{q-1} (\gamma_i^+ \Delta FD_{t-i}^+ + \\ & \gamma_i^- \Delta FD_{t-i}^-) + \sum_{i=1}^{q-1} (\pi_i^+ \Delta GLOB_{t-i}^+ + \pi_i^- \Delta GLOB_{t-i}^-) + \varepsilon_t(1) \end{aligned}$$

Where CO<sub>2</sub> stands for carbon dioxide emissions, GDP for gross domestic product, EC for energy consumption, FD for financial development and GLOB for globalization. We run the previous model in equation-1 separately for each dimension of globalization. The subscripts + and – designate respectively the partial sum processes of positive and negative changes of the variables. For the globalization variable  $GLOB_t^+$  and  $GLOB_t^-$  are defined as follow:

$$\begin{aligned} GLOB_t^+ &= \sum_{i=1}^t \Delta GLOB_j^+ = \sum_{i=1}^t \max(\Delta GLOB_j, 0) \quad \text{and} \\ GLOB_t^- &= \sum_{i=1}^t \Delta GLOB_j^- = \sum_{i=1}^t \min(\Delta GLOB_j, 0) \end{aligned}$$

The long-run positive and negative coefficients can be calculated as  $\theta_{GDP}^+ = -\rho_{GDP}^+ / \rho_{CO2}$  and  $\theta_{GDP}^- = -\rho_{GDP}^- / \rho_{CO2}$ , respectively for GDP;  $\theta_{EC}^+ = -\rho_{EC}^+ / \rho_{CO2}$  and  $\theta_{EC}^- = -\rho_{EC}^- / \rho_{CO2}$ ,

respectively for EC;  $\theta_{FD}^+ = -\rho_{FD}^+/\rho_{CO_2}$  and  $\theta_{FD}^- = -\rho_{FD}^-/\rho_{CO_2}$ , respectively for FD and  $\theta_{GLOB}^+ = -\rho_{GLOB}^+/\rho_{CO_2}$  and  $\theta_{GLOB}^- = -\rho_{GLOB}^-/\rho_{CO_2}$ , respectively for globalization. The long-run symmetry of the influence of GDP, energy consumption, energy consumption, financial development and globalization are tested using the Wald test of the respective null hypotheses  $\rho_{GDP}^+ = \rho_{GDP}^-$ ,  $\rho_{EC}^+ = \rho_{EC}^-$ ,  $\rho_{FD}^+ = \rho_{FD}^-$  and  $\rho_{GLOB}^+ = \rho_{GLOB}^-$ . Likewise, short-run symmetry of the respective impacts of GDP, financial development and globalization on dioxide carbon emissions is tested using the Wald test of the respective null hypotheses  $\beta_i^+ = \beta_i^-$ ,  $\delta_i^+ = \delta_i^-$ ,  $\gamma_i^+ = \gamma_i^-$  and  $\pi_i^+ = \pi_i^-$  for  $i = 1, 2, \dots, q - 1$ .

Once asymmetry is detected – in the long-, short-run or both – it is possible to compute dynamic multipliers that give the predicted trajectory of CO<sub>2</sub> emissions following unit positive and negative unit changes of  $GDP^+$ ,  $GDP^-$ ,  $EC^+$ ,  $EC^-$ ,  $FD^+$ ,  $FD^-$ ,  $GLOB^+$  and  $GLOB^-$ , respectively. For example, the dynamic multipliers following a unit change of  $GLOB^+$  and  $GLOB^-$  are computed as follows:

$$m_{GDP,h}^+ = \sum_{j=0}^h \frac{\partial IP_{t+j}}{\partial GDP_t^+} \text{ and } m_{GDP,h}^- = \sum_{j=0}^h \frac{\partial IP_{t+j}}{\partial GDP_t^-}, \text{ respectively.}$$

Shin et al. (2014) show that  $m_{GDP,h}^+ \rightarrow \theta_{GDP}^+$  and  $m_{GDP,h}^- \rightarrow \theta_{GDP}^-$  when  $h \rightarrow \infty$ . Figures (1) to (4) depict the computed dynamic multipliers following a unit shock of each of the regressands considered in our study.

Unavailability of data on globalization has restricted us to consider the period of 1970-2015 for empirical analysis. We have utilized World Development Indicators (CD-ROM, 2015) to collect data for CO<sub>2</sub> emissions (metric tons), real GDP (in constant 2010 local currency) measure for economic growth, energy use (kg of oil equivalent) proxy of energy consumption and real domestic credit to private sector (in constant 2010 local currency) measure for financial development. This study uses globalization index developed by Dreher (2006). Fundamentally, Dreher (2006) uses three sub-indices: economic globalization, social globalization and political globalization to generate index for globalization. Economic globalization is understood in the form of trade and capital inflows. Social globalization is also understood in the form of personal contacts, telephone contacts, tourism, and migration of people among countries, information flows i.e. internet usage, televisions per 1000 people, trade in newspapers, and data on cultural

proximity (e.g., number of McDonald's restaurants, number of IKEA stores, trade in books). Finally, political globalization is described by the number of embassies in a country, membership in international organizations, participation in the UN secretary council membership and involvement in international treaties to generate an index of political globalization. From the point view of significance, the relative share in the overall globalization index contributed by economic globalization is 36% and followed by social globalization (38%) and by political globalization (26%). The weighted average overall globalization index and its sub-indices are created and maintained by ETH Zurich<sup>5</sup>. The total population series is used to transform all the variables into per capita units.

All studies investigated the determinants of CO<sub>2</sub> emissions by using annual frequency data but none of study used quarter frequency data while investigating the association between carbon emission and its determinants in case of China. It is clearly argued by Narayan and Sharma (2015), Phan et al. (2015a, b) and Hoang et al. (2015, 2016) that data frequency matters for empirical results. In doing so, we transform annual frequency data into quarter frequency data by employing quadratic match-sum method following Sbia et al. (2013). The quadratic match-sum method helps in avoiding the degree of freedom problem for small sample data. Moreover, Mack and Martinez-Garcia (2013) recommend for using the quadratic match-sum method while converting series from low to high frequency. This method seems to correct seasonal variations in the data.

#### **4. Findings and analysis**

Table-1 presents the Wald statistics and their corresponding p-values for the test that checks for the long-run ( $W_{LR}$ ) and short-run ( $W_{SR}$ ) asymmetry in the Nonlinear Autoregressive Distributed Lags (NARDL) model for the four considered dimensions of globalization namely economic globalization, social globalization, political globalization and overall globalization. The results show that in the long-run, the impact of gross domestic product growth (GDP) on CO<sub>2</sub> emissions is linear and symmetric regardless of the type of globalization introduced in the model except for economic globalization where the impact of GDP on CO<sub>2</sub> emissions is found to be nonlinear and asymmetric. Similarly, in the short-run, this relationship is linear and symmetric. As for energy consumption as explanatory variable, it impacts CO<sub>2</sub> emissions in a linear and symmetric fashion in the long-run. In contrast, energy consumption influences CO<sub>2</sub> emissions in an asymmetric and nonlinear manner in the short-run.

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<sup>5</sup> For more details, see <http://globalization.kof.ethz.ch/>

Financial development, globally, it has a nonlinear and asymmetric impact on CO<sub>2</sub> emissions in the long-run when economic globalization is considered in the model. However, in models including social, political or overall globalization financial development impacts CO<sub>2</sub> emissions in a symmetric and linear manner in the long-run. In the short-run, financial development has a linear and symmetric impact on CO<sub>2</sub> emissions for all globalization dimensions.

Turning to the analysis of the fourth explanatory variable, individual dimensions of globalization, the computed Wald statistics show that, in the long-run, the relationship between economic globalization and CO<sub>2</sub> emissions is asymmetric and nonlinear while between social, political and overall globalization is found to be linear and symmetric. However, in the short-run, globalization variable impacts CO<sub>2</sub> emissions in a linear and symmetric fashion.

**Table-1: Long-run and Short-run Asymmetry Test**

	$W_{LR}$	$W_{SR}$
CO2		
GDP	11.600*** [0.001]	0.003 [0.957]
EC	2.028 [0.156]	9.109*** [0.003]
FD	13.180*** [0.000]	0.6001 [0.439]
EG	8.659*** [0.004]	0.002 [0.964]
CO2		
GDP	0.355 [0.552]	0.028 [0.867]
EC	0.106 [0.745]	6.888** [0.010]
FD	1.447 [0.231]	0.538 [0.464]
SG	0.539 [0.464]	0.233 [0.630]
CO2		
GDP	1.099 [0.296]	0.080 [0.777]

EC	0.260 [0.611]	4.798** [0.030]
FD	1.917 [0.168]	0.001 [0.977]
PG	0.874 [0.351]	0.833 [0.363]
CO2		
GDP	1.011 [0.316]	0.091 [0.763]
EC	0.098 [0.755]	7.803*** [0.006]
FD	2.055 [0.154]	0.076 [0.783]
OG	0.654 [0.420]	0.260 [0.611]
Notes: WSR and WLR refer to the Wald statistics for the short- and long-run symmetry null hypotheses. The numbers in the brackets are the p-values. *** Indicates rejection of the null of symmetry at 1%. ** Indicates rejection of the null of symmetry at 5%. * Indicates rejection of the null of symmetry at 10%.		

Table-2 exhibits results of the estimated NARDL models by globalization dimension i.e. economic, social, political and overall globalization. Findings show that the coefficients related to the lagged CO<sub>2</sub> emissions, of all the four selected globalization types, are negative and statistically significant at 1% significance level for all models indicating the stability of the estimated NARDL models. However, the one-period lagged changes in CO<sub>2</sub> emissions have strong significant positive impact on current CO<sub>2</sub> emissions in all models regardless of the globalization dimension. The current and one-period lagged GDP growths have significant respective positive and negative short-run impact on CO<sub>2</sub> emissions in the four considered models. The overall short-run effect of GDP variations, calculated as the sum of all short-run coefficients of GDP, on current CO<sub>2</sub> emissions is positive i.e. 0.141, 0.152, 0.152 and 0.157 in the respective models of economic, social, political and overall globalization. In the long-run, results show that increases of GDP rise CO<sub>2</sub> emissions in the presence of economic globalization while decreases of GDP reduce CO<sub>2</sub> emissions. The latter asymmetric long-run impact of GDP on CO<sub>2</sub> emissions clearly reveals that decreases of GDP have stronger effect on CO<sub>2</sub> emissions than increases of GDP. In model including social globalization increases and decreases of GDP significantly rise and reduce CO<sub>2</sub> emissions by the same magnitude<sup>6</sup>. This reveals that overall GDP has positive and significant effect on CO<sub>2</sub> emissions. This empirical

<sup>6</sup> However, in presence of political and overall globalizations GDP does not influence CO<sub>2</sub> emission at all.



findings is similar with the studies of Jalil and Feridun (2011), Zhang (2011) and, Xiong and Qi (2017), which mention the contributory role of economic growth to environmental degradation.

In the long-run, energy consumption deteriorates the environment as more energy consumption leads more CO<sub>2</sub> emissions in China. Nevertheless, in the short-run consuming more energy entails more CO<sub>2</sub> emissions while reducing energy consumption leads to lower CO<sub>2</sub> emission of less magnitude in the case of social, political and overall globalization. Shahbaz et al. (2016), Ling et al. (2015), Shahbaz et al. (2013d) and Jalil and Feridun (2011) reported that energy consumption deteriorates environmental quality by adding in CO<sub>2</sub> emissions for Portugal, Malaysia, Turkey and China. In the economic globalization case, reducing energy consumption in the short-run does not reduce nor increase CO<sub>2</sub> emissions.

Financial development has increasing and decreasing impacts on long-run CO<sub>2</sub> emission via positive and negative shocks occurring in financial development when economic globalization is accounted for in the model. This indicates that an increase of financial development will lead CO<sub>2</sub> emissions to move down but opposite effect is noted in case of less financial development. Indeed, less financial development leads CO<sub>2</sub> emissions to move up with a stronger effect in the latter cases. Financial development does not impact CO<sub>2</sub> emissions in the long-run coupled with social, political or overall globalization. This shows that financial development affects carbon emissions insignificantly. This empirical evidence is consistent with Ozturk and Acaravci (2013) who noted that financial development does not affect carbon emissions for Turkish economy. This contrary is similar with Jalil and Feridun (2011), Zhang (2011) and, Xiong and Qi (2017) who reported that financial development improves environmental quality via lowering carbon emissions. Our findings show that the overall effect of financial development on CO<sub>2</sub> emission in the short-is significantly positive regardless of the type of globalization in the model.

Coming to the analysis of the effect of each of the considered globalization dimensions on CO<sub>2</sub> emission, our results show evidence of a nonlinear and asymmetric effect of economic globalization when considered in the model. Indeed, higher level of economic globalization will deteriorate environmental quality in the long-run while lower level of economic globalization will reduce CO<sub>2</sub> emissions with more pronounced effect in case of decrease the extend of economic globalization. Social and overall globalizations have significant negative long-run effects on CO<sub>2</sub> emissions meaning that higher levels of social and overall

globalizations would preserve the environment from being deteriorated<sup>7</sup>. The negative effect of globalization (social and overall globalization) on carbon emissions is consistent with Shahbaz et al. (2017) who noted that globalization improves environmental quality by decreasing carbon emissions without affecting the pattern of EKC effect. Higher level of economic globalization increases carbon emissions is contrary with Shahbaz et al. (2017) who reported that economic globalization is negative linked with carbon emissions i.e. economic globalization contributes to environmental quality for Chinese economy<sup>8</sup>. However, in the short-run social, political and overall globalizations are found to reduce CO<sub>2</sub> emissions while economic globalization has no short-run influence on CO<sub>2</sub> emissions.

**Table-2: Long-run and Short-run Asymmetric and Symmetric Analysis**

CO2-EG		GDP-SG		GDP-PG		GDP-OG	
GDP, FD and EG LR asymmetry & EC SR asymmetry		EC SR asymmetry		EC SR asymmetry		ECSR asymmetry	
$CO2_{t-1}$	-0.058*** (0.014)	$CO2_{t-1}$	-0.046*** (0.012)	$CO2_{t-1}$	-0.027** (0.012)	$CO2_{t-1}$	-0.035*** (0.011)
$GDP_{t-1}^+$	0.029** (0.013)	$GDP_{t-1}$	0.016** (0.007)	$GDP_{t-1}$	0.004 (0.005)	$GDP_{t-1}$	0.010 (0.007)
$GDP_{t-1}^-$	-0.137*** (0.050)	$EC_{t-1}$	0.029** (0.011)	$EC_{t-1}$	0.023* (0.012)	$EC_{t-1}$	0.023* (0.012)
$EC_{t-1}$	0.037* (0.019)	$FD_{t-1}$	0.002 (0.003)	$FD_{t-1}$	0.002 (0.004)	$FD_{t-1}$	0.003 (0.003)
$FD_{t-1}^+$	-0.020** (0.008)	$SG_{t-1}$	-0.009*** (0.003)	$PG_{t-1}$	-0.006 (0.005)	$OG_{t-1}$	-0.014** (0.006)
$FD_{t-1}^-$	0.072*** (0.019)	$\Delta CO2_{t-1}$	0.568*** (0.062)	$\Delta CO2_{t-1}$	0.557*** (0.065)	$\Delta CO2_{t-1}$	0.565*** (0.064)
$EG_{t-1}^+$	0.045** (0.017)	$\Delta GDP_t$	0.434*** (0.079)	$\Delta CO2_{t-4}$	-0.099** (0.049)	$\Delta GDP_t$	0.373*** (0.084)
$EG_{t-1}^-$	-0.060*** (0.021)	$\Delta GDP_{t-1}$	-0.282*** (0.083)	$\Delta GDP_t$	0.339*** (0.093)	$\Delta GDP_{t-1}$	-0.216** (0.087)
$\Delta CO2_{t-1}$	0.488*** (0.068)	$\Delta EC_t^+$	1.026*** (0.088)	$\Delta GDP_{t-1}$	-0.214** (0.095)	$\Delta EC_t^+$	1.089*** (0.093)
$\Delta GDP_t$	0.443***	$\Delta EC_{t-1}^+$	-0.542***	$\Delta EC_t^+$	1.047***	$\Delta EC_{t-1}^+$	-0.587***

<sup>7</sup> Being insignificant, the effect of political globalization on CO<sub>2</sub> emissions shows a similar pattern, i.e. political globalization reduces CO<sub>2</sub> emissions but it is insignificant.

<sup>8</sup> This difference in empirical findings shows the importance of asymmetries occurred in times series data to be considered while investigating the determinants of carbon emissions.

	(0.084)		(0.106)		(0.103)		(0.114)
$\Delta GDP_{t-1}$	-0.302*** (0.087)	$\Delta EC_t^-$	0.242* (0.143)	$\Delta EC_{t-1}^+$	-0.569*** (0.120)	$\Delta EC_{t-1}^-$	-0.370** (0.144)
$\Delta EC_t^+$	0.866*** (0.089)	$\Delta EC_{t-1}^-$	-0.373** (0.145)	$\Delta EC_{t-1}^+$	-0.272* (0.151)	$\Delta FD_t$	0.079*** (0.021)
$\Delta EC_{t-1}^+$	-0.329*** (0.110)	$\Delta FD_t$	0.063*** (0.020)	$\Delta FD_t$	0.087*** (0.022)	$\Delta FD_{t-1}$	-0.048** (0.021)
$\Delta FD_t$	0.072*** (0.022)	$\Delta SG_t$	-0.073*** (0.015)	$\Delta FD_{t-1}$	-0.055** (0.022)	$\Delta OG_t$	-0.195*** (0.038)
constant	-0.081*** (0.021)	$\Delta SG_{t-1}$	0.056*** (0.015)	$\Delta PG_t$	-0.134*** (0.036)	$\Delta OG_{t-1}$	0.153*** (0.040)
		Constant	-0.065*** (0.017)	$\Delta PG_{t-1}$	0.097** (0.038)	constant	-0.048*** (0.016)
				constant	-0.035** (0.016)		
$L_{GDP}^+$	0.510* [0.054]	$L_{GDP}$	0.359** [0.019]	$L_{GDP}$	0.169 [0.428]	$L_{GDP}$	0.307 [0.156]
$L_{GDP}^-$	-2.360*** [0.003]	$L_{EC}$	0.635*** [0.000]	$L_{EC}$	0.859*** [0.000]	$L_{EC}$	0.666*** [0.002]
$L_{EC}$	0.646*** [0.003]	$L_{FD}$	0.046 [0.510]	$L_{FD}$	0.070 [0.598]	$L_{FD}$	0.095 [0.303]
$L_{FD}^+$	-0.348** [0.035]	$L_{SG}$	-0.192*** [0.001]	$L_{PG}$	-0.244 [0.270]	$L_{OG}$	-0.404** [0.048]
$L_{FD}^-$	1.248*** [0.000]						
$L_{EG}^+$	0.788*** [0.005]						
$L_{EG}^-$	-1.038** [0.031]						
AIC	-1340.491	AIC	-1370.411	AIC	-1336.594	AIC	-1358.486
SIC	-1270.246	SIC	-1312.838	SIC	-1272.847	SIC	-1297.820
Q(40)	36.630 [0.622]	Q(40)	49.000 [0.155]	Q(40)	43.550 [0.323]	Q(40)	48.100 [0.177]

Notes: Only significant short-run coefficients are reported in this Table 2. Standard errors of the estimated coefficients are in parenthesis. The p-values of statistical tests are in brackets. LX<sup>+</sup> and LX<sup>-</sup> indicate the positive and negative long-run coefficients, respectively. Lag orders of the three NARDL models are selected according to the Akaike and Schwarz Information criteria. p = 5, q = 2 for CO2-PG ; p = 4, q = 2 for CO2-EG and CO2-OG; p = 3, q = 2 for CO2-SG. Q(40) refers to the Ljung-Box test of the null of independent residuals up to lag 40.  
\*\*\* Indicates significance at 1%.  
\*\* Indicates significance at 5%.  
\* Indicates significance at 10%.

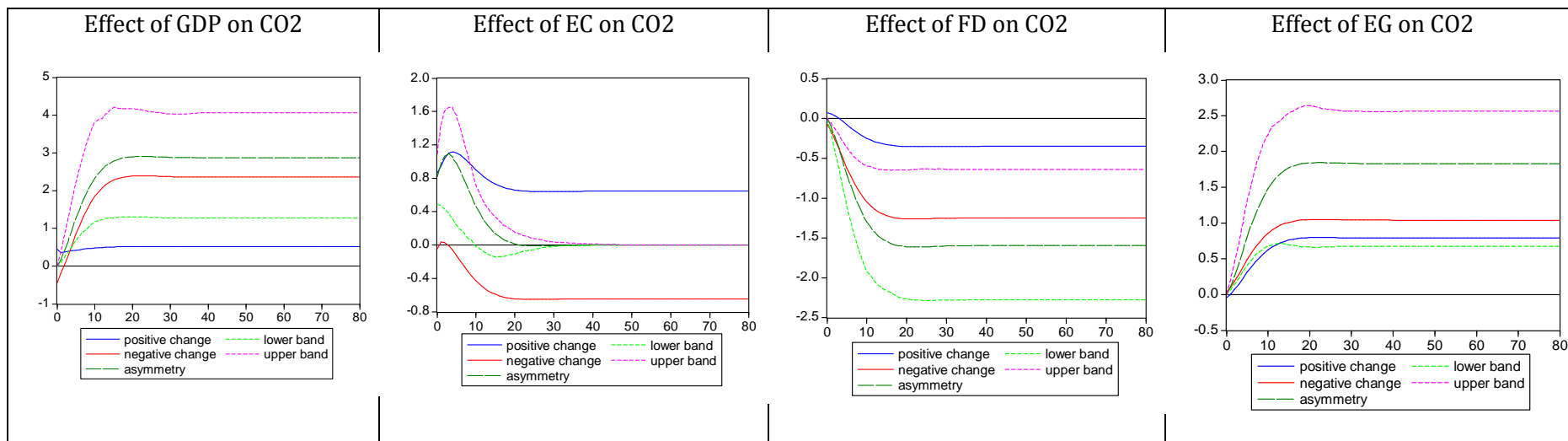
Figures 1 to 4 depict the dynamic multipliers of the four models. Each model includes gross domestic product, energy consumption, financial development and one of the four globalization dimensions considered namely economic globalization, social globalization, political globalization and overall globalization. The asymmetry curve shows a linear combination of the dynamic multipliers associated with positive and negative shocks. The positive change and

negative change curves indicate the adjustment paths after a positive and a negative change at a given forecasting horizon, respectively. The lower and upper bands indicate a 95% confidence interval for asymmetry.

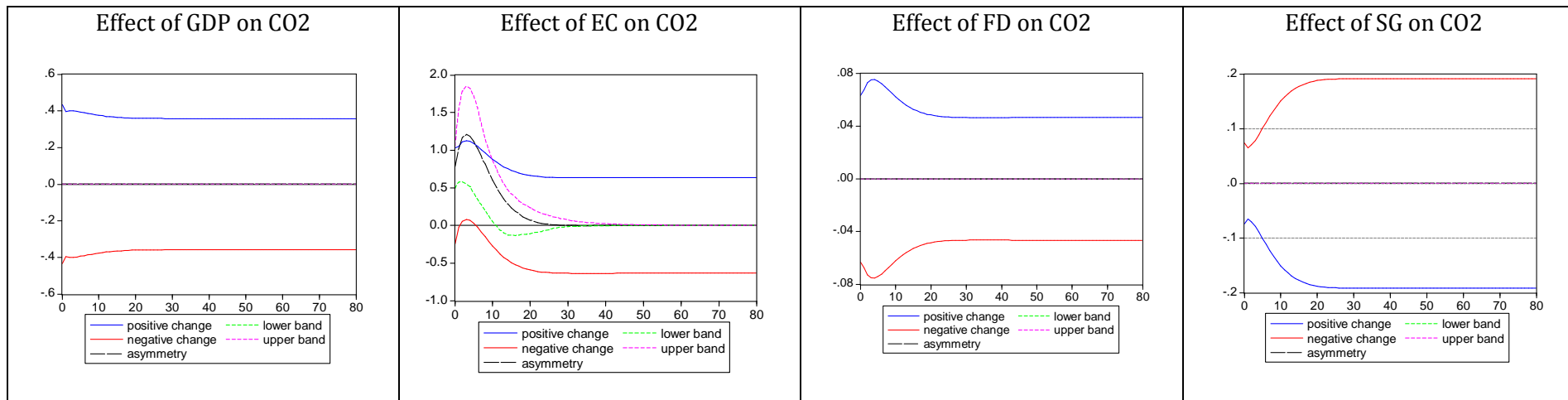
The shape of the dynamic multipliers of the effect of GDP on CO<sub>2</sub> emission changes across models. Indeed, in all four models it reveals that unitary shock to GDP implies a long-run response of CO<sub>2</sub> emissions that lasts approximately 20 quarters, and then the level of CO<sub>2</sub> emissions stabilizes with a lower reaction of CO<sub>2</sub> emissions in model including political globalization. The asymmetry curve showing the reaction of CO<sub>2</sub> emissions to a unit shock of energy consumption shows a similar pattern in the four estimated models. Indeed, after a quick increase of CO<sub>2</sub> emissions following a unit shock of energy consumption that lasts around two quarters CO<sub>2</sub> emissions level starts decreasing and vanishes after ten quarters.

Similarly, CO<sub>2</sub> emissions reaches a new equilibrium after roughly 20 quarters after a positive or a negative unitary shock to financial development occurs in models with social, political or overall globalization. However, when economic globalization is included in the model both unitary positive and negative shocks of financial development have negative impacts on CO<sub>2</sub> emissions leading to a negative asymmetry curve that stabilizes after around 18 quarters ahead. Turning to the effect of globalization on CO<sub>2</sub> emissions, the dynamic multipliers graphs show that the reaction of CO<sub>2</sub> emissions to positive and negative unitary changes in economic globalization is positive intensifying thus the asymmetry curve that is positive and stabilizes at the 1.7 level after around 18 quarters. Reactions of CO<sub>2</sub> emissions to social, political and overall globalization show similar patterns. Indeed, a unitary positive shock of economic (social) globalization shows a low positive effect on CO<sub>2</sub> emissions has a negative effect on CO<sub>2</sub> emissions while a negative unitary shock produces a positive effect of similar size on CO<sub>2</sub> emissions. Consequently, asymmetry path is null in the three previous cases.

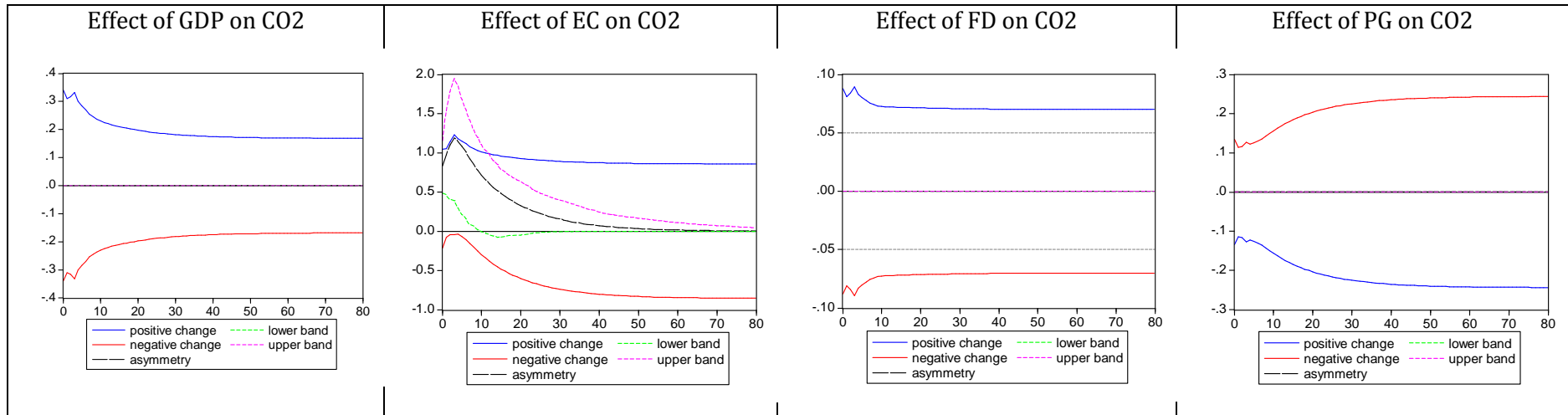
**Figure-1. Dynamic Multipliers for CO2-EG**



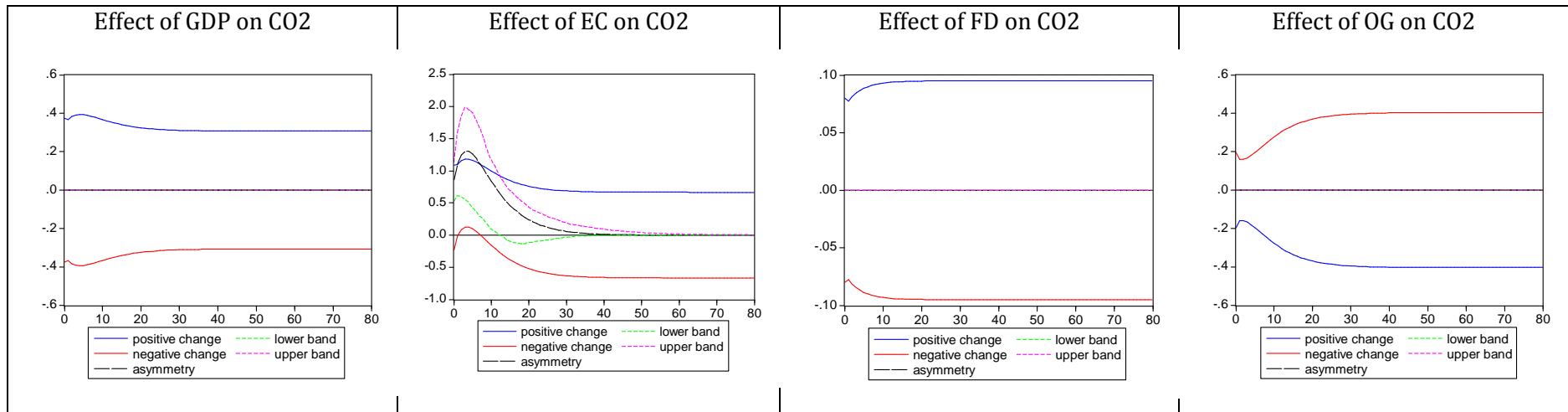
**Figure-2. Dynamic Multipliers for CO2-SG**



**Figure-3: Dynamic Multipliers for CO2-PG**



**Figure-4: Dynamic Multipliers for CO2-OG**



## 5. Conclusion and Policy Implications

The environment has experienced a substantial increase of dioxide carbon emissions which have nocuous effects on it in general and on households living conditions and health in particular. This enormous risk of environment degradation has led researchers and environment policy makers to concentrate on this topic in order to find appropriate solutions and adopt convenient policies. Nevertheless, reducing dioxide carbon emissions is still a challenge for several countries as the empirical findings regarding its determinants and their relationships with the level of carbon dioxide issued in the environment provide mixed results. Hence no clear definite answer on the influence of CO<sub>2</sub> emissions drivers on environment is set. We contribute to the existing literature on determinants of CO<sub>2</sub> emissions and their influence on it by analyzing the influence of four main drivers that we believe are extremely important within a nonlinear framework. To do so, we employ the NARDL model that is shown in previous studies to beat its linear counterpart i.e. the linear ARDL model. Indeed, even though the simple ARDL model accounts for the cointegrating relationships that may exist between CO<sub>2</sub> emissions and its determinants it fails to disentangle the effects of positive and negative shocks on CO<sub>2</sub> emissions as it assumes similar size to positive and negative shocks. Thus, the ARDL model allows for symmetric effects of positive and negative shocks. In contrast, the NARDL model permits to detect the asymmetric effects of positive and negative shocks of exogenous variables on CO<sub>2</sub> emissions both in the long-run and short-run. In addition, the NARDL model allows predicting the future effects of positive and negative shocks hitting the system equations by computing dynamic multipliers. Interestingly, findings show, (1) In the short-run, gross domestic product, energy consumption, financial development and globalization affect significantly CO<sub>2</sub> emissions; (2) In the long-run, energy consumption significantly deteriorates environment. However, in the short-run, more energy consumption increases CO<sub>2</sub> emissions while less energy consumption preserves environment; (3) In the long-run, economic globalization influences carbon dioxide emissions in asymmetric and nonlinear manners while social and overall globalizations reduce CO<sub>2</sub> emissions in a linear. Thus, higher level of economic globalization raises CO<sub>2</sub> emissions while lower level reduces CO<sub>2</sub> emissions. However, political globalization does not impact CO<sub>2</sub> emissions; (4) In the short-run, economic globalization does not impact CO<sub>2</sub> emissions. However, social, political and overall globalizations reduce CO<sub>2</sub> emissions and; (5) Financial development asymmetrically impacts dioxide carbon emissions only when it coupled with economic globalization. It does not exert any influence on dioxide carbon emissions if coupled with social, political or overall globalization. These empirical findings are important for forecasters, investors and



environmental policy makers who should consider the findings of the present research when formulating their investment decisions based on forecasts of the future evolution of environmental conditions and the related regulatory laws taken by the environmental deciders. This paper examines CO<sub>2</sub> emissions in China at the aggregate level. Future studies should aim at investigating the drivers of CO<sub>2</sub> emissions at the provincial and sectoral level in China, given the differences in energy consumption, financial development, globalization and economic growth level across provinces in China.

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