

# Understanding inflation trends in Finland: A univariate approach

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#### **Understanding Inflation Trends in Finland: A Univariate Approach**

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#### ABSTRACT

This research uses annual time series data on inflation rates in Finland from 1960 to 2017, to model and forecast inflation using ARIMA models. Diagnostic tests indicate that F is I(1). The study presents the ARIMA (1, 1, 3) model. The diagnostic tests further imply that the presented optimal ARIMA (1, 1, 3) model is stable and acceptable in predicting Finnish inflation. The results of the study apparently show that F will be hovering around 1% over the next 10 years. Policy makers and the business community in Finland are expected to take advantage of the anticipated stable inflation rates over the next decade.

Key Words: Forecasting, Inflation

**JEL Codes:** C53, E31, E37, E47

#### **INTRODUCTION**

Inflation is the sustained increase in the general level of prices and services over time (Blanchard, 2000). The negative effects of inflation are widely recognized (Fenira, 2014). Inflation is one of the central terms in macroeconomics (Enke & Mehdiyev, 2014) as it harms the stability of the acquisition power of the national currency, affects economic growth because investment projects become riskier, distorts consuming and saving decisions, causes unequal income distribution and also results in difficulties in financial intervention (Hurtado *et al*, 2013).

As the prediction of accurate inflation rates is a key component for setting the country's monetary policy, it is especially important for central banks to obtain precise values (Mcnelis & Mcadam, 2004). To prevent the aforementioned undesirable outcomes of price instability, central banks require proper understanding of the future path of inflation to anchor expectations and ensure policy credibility; the key aspects of an effective monetary policy transmission mechanism (King, 2005). Inflation forecasts and projections are also often at the heart of economic policy decision-making, as is the case for monetary policy, which in most industrialized economies is mandated to maintain price stability over the medium term (Buelens, 2012). Economic agents, private and public alike; monitor closely the evolution of prices in the economy, in order to make decisions that allow them to optimize the use of their resources (Hector & Valle, 2002). Decision-makers hence need to have a view of the likely future path of inflation when taking measures that are necessary to reach their objective (Buelens, 2012). To avoid adjusting policy and models by not using an inflation rate prediction can result in imprecise investment and saving decisions, potentially leading to economic instability (Enke &

Mehdiyev, 2014). In this study, we seek to model and forecast inflation in Finland using ARIMA models.

#### LITERATURE REVIEW

Kock & Terasvirta (2013) forecasted Finnish consumer price inflation using Artificial Neural Network models with a data set ranging over the period March 1960 – December 2009 and established that direct forecasts are more accurate then their recursive counterparts. Nyoni (2018) studied inflation in Zimbabwe using GARCH models with a data set ranging over the period July 2009 to July 2018 and established that there is evidence of volatility persistence for Zimbabwe's monthly inflation data. Nyoni (2018) modeled inflation in Kenya using ARIMA and GARCH models and relied on annual time series data over the period 1960 – 2017 and found out that the ARIMA (2, 2, 1) model, the ARIMA (1, 2, 0) model and the AR (1) – GARCH (1, 1) model are good models that can be used to forecast inflation in Kenya. Sarangi *et al* (2018) analyzed the consumer price index using Neural Network models with 159 data points and revealed that ANNs are better methods of forecasting CPI in India. Nyoni & Nathaniel (2019), based on annulation rates from 1960 to 2016 and found out that the ARMA (1, 0, 2) model is the best model for forecasting inflation rates in Nigeria.

#### **MATERIALS & METHODS**

#### **Box – Jenkins ARIMA Models**

One of the methods that are commonly used for forecasting time series data is the Autoregressive Integrated Moving Average (ARIMA) (Box & Jenkins, 1976; Brocwell & Davis, 2002; Chatfield, 2004; Wei, 2006; Cryer & Chan, 2008). For the purpose of forecasting inflation rate in Finland, ARIMA models were specified and estimated. If the sequence  $\Delta^{d}F_{t}$  satisfies an ARMA (p, q) process; then the sequence of  $F_{t}$  also satisfies the ARIMA (p, d, q) process such that:

which we can also re – write as:

where  $\Delta$  is the difference operator, vector  $\beta \in \mathbb{R}^p$  and  $\alpha \in \mathbb{R}^q$ .

#### The Box – Jenkins Methodology

The first step towards model selection is to difference the series in order to achieve stationarity. Once this process is over, the researcher will then examine the correlogram in order to decide on the appropriate orders of the AR and MA components. It is important to highlight the fact that this procedure (of choosing the AR and MA components) is biased towards the use of personal judgement because there are no clear – cut rules on how to decide on the appropriate AR and MA components. Therefore, experience plays a pivotal role in this regard. The next step is the

estimation of the tentative model, after which diagnostic testing shall follow. Diagnostic checking is usually done by generating the set of residuals and testing whether they satisfy the characteristics of a white noise process. If not, there would be need for model re – specification and repetition of the same process; this time from the second stage. The process may go on and on until an appropriate model is identified (Nyoni, 2018).

#### **Data Collection**

This study is based on a data set of annual rates of inflation in Finland (INF or simply F) ranging over the period 1960 - 2017. All the data was taken from the World Bank.

#### **Diagnostic Tests & Model Evaluation**

#### **Stationarity Tests: Graphical Analysis**



Figure 1

#### The Correlogram in Levels

Autocorrelation function for INF \*\*\*, \*\*, \* indicate significance at the 1%, 5%, 10% levels.

LA	G AG	CF	PACF	Q-stat. [p-value]
1	0.8373	***	0.8373 ***	42.7976 [0.000]
2	0.6609	***	-0.1340	69.9438 [0.000]

3	0.5459	***	0.1041	88.7997 [0.000]
4	0.4743	***	0.0520	103.2977 [0.000]
5	0.4167	***	0.0124	114.6999 [0.000]
6	0.4438	***	0.2832 **	127.8837 [0.000
7	0.4420	***	-0.0973	141.2144 [0.000]
8	0.3925	***	-0.0455	151.9385 [0.000]
9	0.3352	**	0.0044	159.9170 [0.000]
10	0.2634	**	-0.1314	164.9489 [0.000]
11	0.1763		-0.0376	167.2500 [0.000]

## The ADF Test in Levels

Table 2: Levels-intercept

Variable	ADF Statistic	Probability	Critical Values	5	Conclusion	
F	-1.467281	0.5423	-3.560019	@1%	Non-stationary	
			-2.917650	@5%	Non-stationary	
			-2.596689	@10%	Non-stationary	

Table 3: Levels-trend & intercept

Variable	ADF Statistic	Probability	Critical Values	5	Conclusion	
F	-3.465043	0.0532	-4.130526	@1%	Non-stationary	
			-3.492149	@5%	Non-stationary	
			-3.174802	@10%	Stationary	

Table 4: without intercept and trend & intercept

Variable	ADF Statistic	Probability	Critical Values		Conclusion
F	-1.392548	0.1504	-2.609324	@1%	Non-stationary
			-1.947119	@5%	Non-stationary
			-1.612867	@10%	Non-stationary

Figure 1 and tables 1 - 4 show that F is non-stationary in levels.

# The Correlogram (at 1<sup>st</sup> Differences)

Autocorrelation function for d\_INF \*\*\*, \*\*, \* indicate significance at the 1%, 5%, 10% levels.

LAG	ACF	PACF	Q-stat. [p-value]

- 1 0.0318 0.0318 0.0609 [0.805]
- 2 -0.1954 -0.1966 2.3951 [0.302]

3	-0.1233	-0.1141	3.3424	[0.342]
4	-0.0194	-0.0542	3.3662	[0.499]
5	-0.2808	** -0.3449 ***	8.46	40 [0.132]
6	0.0972	0.0821	9.0867	[0.169]
7	0.1321	-0.0162	10.2610	[0.174]
8	0.0253	-0.0340	10.3049	[0.244]
9	0.0685	0.1306	0.6341	[0.302]
10	0.0327	-0.0546	10.7108	8 [0.380]
11	-0.0827	0.0249	11.2112	2 [0.426]

# ADF Test in 1<sup>st</sup> Differences

Table 6: 1<sup>st</sup> Difference-intercept

Variable	ADF Statistic	Probability	Critical Values		Conclusion
F	-4.710223	0.0003	-3.560019	@1%	Stationary
			-2.917650	@5%	Stationary
			-2.596689	@10%	Stationary

Table 7: 1<sup>st</sup> Difference-trend & intercept

Variable	ADF Statistic	Probability	Critical Values	5	Conclusion		
F	-4.608372	0.0027	-4.140858	@1%	Stationary		
			-3.496960	@5%	Stationary		
			-3.177579	@10%	Stationary		

 Table 8: 1<sup>st</sup> Difference-without intercept and trend & intercept

Variable	ADF Statistic	Probability	Critical Values		Conclusion
F	-4.711857	0.0000	-2.609324	@1%	Stationary
			-1.947119	@5%	Stationary
			-1.612867	@10%	Stationary

Tables 5-8 indicate that the F series became stationary after taking first differences.

# Evaluation of ARIMA models (without a constant)

Model	AIC	ME	MAE	RMSE
ARIMA (1, 1, 1)	265.8227	-0.044202	1.7068	2.3636
ARIMA (1, 1, 0)	264.1306	-0.044966	1.7284	2.37
ARIMA (0, 1, 1)	264.0954	-0.044121	1.7218	2.3693
ARIMA (2, 1, 1)	263.4818	-0.088189	1.636	2.2773
ARIMA (1, 1, 2)	263.7031	-0.093457	1.6373	2.2826

ARIMA (2, 1, 0)	263.8906	-0.057138	1.6667	2.3238
ARIMA (2, 1, 2)	264.8469	-0.072912	1.6116	2.268
ARIMA (3, 1, 1)	265.4751	-0.087669	1.6357	2.2771
ARIMA (1, 1, 3)	261.5062	-0.07683	1.6365	2.209
ARIMA (2, 1, 3)	263.5026	-0.076251	1.6365	2.2091
ARIMA (3, 1, 0)	265.1973	-0.062971	1.6614	2.3095
ARIMA (0, 1, 3)	263.0044	-0.088872	1.6388	2.2677
ARIMA (1, 1, 4)	263.5045	-0.076554	1.6365	2.2091
ARIMA (0, 1, 2)	263.7017	-0.068374	1.6622	2.3201
ARIMA (3, 1, 2)	266.6212	-0.081179	1.6501	2.261

A model with a lower AIC value is better than the one with a higher AIC value (Nyoni, 2018). The study will only consider the AIC as the criteria for choosing the best model for predicting inflation in Finland. Hence, the ARIMA (1, 1, 3) model is selected finally.

#### **Residual & Stability Tests**

#### ADF Tests of the Residuals of the ARIMA (1, 1, 3) Model

#### Table 10: Levels-intercept

Variable	ADF Statistic	Probability	Critical Values		Conclusion
R <sub>t</sub>	-4.697559	0.0003	-3.557472	@1%	Stationary
			-2.916566	@5%	Stationary
			-2.596116	@10%	Stationary

#### Table 11: Levels-trend & intercept

Variable	ADF Statistic	Probability	Critical Values		Conclusion
R <sub>t</sub>	-4.646931	0.0024	-4.137279 @1%		Stationary
			-3.495295	@5%	Stationary
			-3.176618	@10%	Stationary

#### Table 12: without intercept and trend & intercept

Variable	ADF Statistic	Probability	Critical Values		Conclusion
R <sub>t</sub>	-4.732690	0.0000	-2.608490 @1%		Stationary
			-1.946996	@5%	Stationary
			-1.612934	@10%	Stationary

Tables 10, 11 and 12 show that the residuals of the ARIMA (1, 1, 3) model are stationary and hence the ARIMA (1, 1, 3) model is suitable for forecasting inflation in Finland.

#### Stability Test of the ARIMA (1, 1, 3) Model

#### Figure 2



Since the corresponding inverse roots of the characteristic polynomial lie in the unit circle, it illustrates that the chosen ARIMA (1, 1, 3) model is stable and suitable for predicting inflation in Finland over the period under study.

### FINDINGS

#### **Descriptive Statistics**

Table 13

Description	Statistic
Mean	4.7353
Median	3.415
Minimum	-0.21
Maximum	17.81
Standard deviation	4.2692
Skewness	1.2295
Excess kurtosis	1.0207

As shown above, the mean is positive, i.e. 4.7353%. The minimum is -0.21% and the maximum is 17.81%. The skewness is 1.2295 and the most striking characteristic is that it is positive, indicating that the inflation series is positively skewed and non-symmetric. Excess kurtosis was found to be 1.0207; implying that the inflation series is not normally distributed.

### **Results Presentation**<sup>1</sup>

<sup>&</sup>lt;sup>1</sup> The \*, \*\* and \*\*\* means significant at 10%, 5% and 1% levels of significance; respectively.

$\Delta F_{t-1} = -0.67$ P: (0.0 S. E: (0.1	77853 $\Delta F_{t-1} + 0$ 005) (0 941) (0	ARIMA (1, 1, 3) $0.742466\mu_{t-1} - 0.2$ $0.0000$ )(0.09)(0.1950)	<b>3) Model:</b> $52819\mu_{t-2} - 0.46$ 989) (0.000 532) (0.124	58618µ <sub>t-3</sub> [3] 12) 11)
Variable	Coefficient	Standard Error	Z	p-value
AR (1)	-0.677853	0.194148	-3.491	0.0005***
MA (1)	0.742466	0.195027	3.807	0.0001***
MA (2)	-0.252819	0.153219	-1.65	0.0989*
MA (3)	-0.468618	0.124056	-3.777	0.0002***

# Forecast Graph



Figure 3



8

Year	Prediction	Std. Erro	or 95%	Confidence Interval
2018	0.59	2.182	-3.69 -	4.86
2019	0.67	3.187	-5.57 -	6.92
2020	0.77	3.600	-6.29 -	7.82
2021	0.70	3.762	-6.67 -	8.08
2022	0.75	4.046	-7.18 -	8.67
2023	0.72	4.225	-7.56 -	9.00
2024	0.74	4.452	-7.99 -	9.46
2025	0.72	4.631	-8.35 -	9.80
2026	0.73	4.827	-8.73 -	10.19
2027	0.73	5.001	-9.07 -	10.53

Table 15

Figure 3 (with a forecast range from 2018 - 2027) and table 15, clearly show that inflation in Finland is projected to be hovering closer to 1% in the next 10 years. This is clear testimony to the fact that there is price stability in Finland and this is indeed predicted to exist over the next decade, ceteris paribus. The current and projected price stability in Finland could be attributed to prudent macroeconomic policy formulation and implementation.

#### CONCLUSION

The ARIMA model was employed to investigate annual inflation rates in Finland from 1960 to 2017. The study planned to forecast inflation in Finland for the upcoming period from 2018 to 2027 and the best fitting model was carefully selected. The ARIMA (1, 1, 3) model is stable and most suitable model to forecast inflation in Finland for the next ten years. Based on the results, policy makers in Finland should continue to engage proper economic policies in order to maintain macroeconomic stability in Finland.

#### REFERENCES

- [1] Blanchard, O (2000). Macroeconomics, 2<sup>nd</sup> Edition, *Prentice Hall*, New York.
- [2] Box, G. E. P & Jenkins, G. M (1976). Time Series Analysis: Forecasting and Control, *Holden Day*, San Francisco.
- [3] Brocwell, P. J & Davis, R. A (2002). Introduction to Time Series and Forecasting, *Springer*, New York.

- [4] Buelens, C (2012). Inflation modeling and the crisis: assessing the impact on the performance of different forecasting models and methods, *European Commission*, Economic Paper No. 451.
- [5] Chatfield, C (2004). The Analysis of Time Series: An Introduction, 6<sup>th</sup> Edition, *Chapman & Hall*, New York.
- [6] Cryer, J. D & Chan, K. S (2008). Time Series Analysis with Application in R, Springer, New York.
- [7] Enke, D & Mehdiyev, N (2014). A Hybrid Neuro-Fuzzy Model to Forecast Inflation, Procedia Computer Science, 36 (2014): 254 – 260.
- [8] Fenira, M (2014). Democracy: a determinant factor in reducing inflation, *International Journal of Economics and Financial Issues*, 4 (2): 363 375.
- [9] Hector, A & Valle, S (2002). Inflation forecasts with ARIMA and Vector Autoregressive models in Guatemala, *Economic Research Department*, Banco de Guatemala.
- [10] Hurtado, C., Luis, J., Fregoso, C & Hector, J (2013). Forecasting Mexican Inflation Using Neural Networks, *International Conference on Electronics*, *Communications and Computing*, 2013: 32 – 35.
- [11] Kharimah, F., Usman, M., Elfaki, W & Elfaki, F. A. M (2015). Time Series Modelling and Forecasting of the Consumer Price Bandar Lampung, *Sci. Int (Lahore).*, 27 (5): 4119 – 4624.
- [12] King, M (2005). Monetary Policy: Practice Ahead of Theory, *Bank of England*.
- [13] Kock, A. B & Terasvirta, T (2013). Forecasting the Finnish Consumer Price Inflation using Artificial Network Models and Three Automated Model Section Techniques, *Finnish Economic Papers*, 26 (1): 13 – 24.
- [14] Mcnelis, P. D & Mcadam, P (2004). Forecasting Inflation with Think Models and Neural Networks, *Working Paper Series*, European Central Bank.
- [15] Nyoni, T & Nathaniel, S. P (2019). Modeling Rates of Inflation in Nigeria: An Application of ARMA, ARIMA and GARCH models, *Munich University Library – Munich Personal RePEc Archive (MPRA)*, Paper No. 91351.
- [16] Nyoni, T (2018). Modeling and Forecasting Inflation in Zimbabwe: a Generalized Autoregressive Conditionally Heteroskedastic (GARCH) approach, *Munich University Library Munich Personal RePEc Archive (MPRA)*, Paper No. 88132.
- [17] Nyoni, T (2018). Modeling and Forecasting Inflation in Kenya: Recent Insights from ARIMA and GARCH analysis, *Dimorian Review*, 5 (6): 16–40.

- [18] Nyoni, T. (2018). Box Jenkins ARIMA Approach to Predicting net FDI inflows in Zimbabwe, *Munich University Library – Munich Personal RePEc Archive (MPRA)*, Paper No. 87737.
- [19] Sarangi, P. K., Sinha, D., Sinha, S & Sharma, M (2018). Forecasting Consumer Price Index using Neural Networks models, *Innovative Practices in Operations Management and Information Technology* – Apeejay School of Management, pp: 84 – 93.
- [20] Wei, W. S (2006). Time Series Analysis: Univariate and Multivariate Methods, 2<sup>nd</sup> Edition, Pearson Education Inc, Canada.